

# The landscape of regional missense mutational intolerance quantified from 125,748 exomes

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1 **Abstract**

2 Missense variants can have a range of functional impacts depending on factors such as the  
3 specific amino acid substitution and location within the gene. To interpret their deleteriousness,  
4 studies have sought to identify regions within genes that are specifically intolerant of missense  
5 variation<sup>1–12</sup>. Here, we leverage the patterns of rare missense variation in 125,748 individuals in  
6 the Genome Aggregation Database (gnomAD)<sup>13</sup> against a null mutational model to identify  
7 transcripts that display regional differences in missense constraint. Missense-depleted regions  
8 are enriched for ClinVar<sup>14</sup> pathogenic variants, *de novo* missense variants from individuals with  
9 neurodevelopmental disorders (NDDs)<sup>15,16</sup>, and complex trait heritability. Following ClinGen  
10 calibration recommendations for the ACMG/AMP guidelines, we establish that regions with less  
11 than 20% of their expected missense variation achieve moderate support for pathogenicity. We  
12 create a missense deleteriousness metric (MPC) that incorporates regional constraint and  
13 outperforms other deleteriousness scores at stratifying case and control *de novo* missense  
14 variation, with a strong enrichment in NDDs. These results provide additional tools to aid in  
15 missense variant interpretation.

16

17 **Main text**

18 Over the last decade, exome and genome sequencing have enabled variant discovery across  
19 hundreds of thousands of individuals<sup>13,17–21</sup>. These large reference databases have provided the  
20 opportunity to study selective forces acting on the human genome and to identify genomic  
21 regions under selective constraint by, for example, identifying regions with fewer variants than  
22 expected based on mutational models<sup>13,18,22–25</sup>. Gene-level metrics of predicted loss-of-function  
23 (pLoF) variant depletion have proven to be valuable in variant classification and identification of  
24 novel disease genes<sup>15,16,26–28</sup>. The functional impact and selective pressures relevant to  
25 missense variation, by contrast, remain challenging to predict, as the effect of a missense  
26 variant is governed by the gene housing the variant, the position of the variant in the gene, and  
27 the specific amino acid substitution caused by the variant. To address this, prior work has  
28 sought to identify regions within coding genes that are specifically intolerant of missense  
29 variation as a way to improve interpretation<sup>1–12</sup>. Here, we expand upon previous work<sup>1</sup> and show  
30 a sub-genic measure of missense intolerance leveraging population-level variation facilitates  
31 variant classification and risk stratification for association studies with *de novo*, rare, and  
32 common variants.

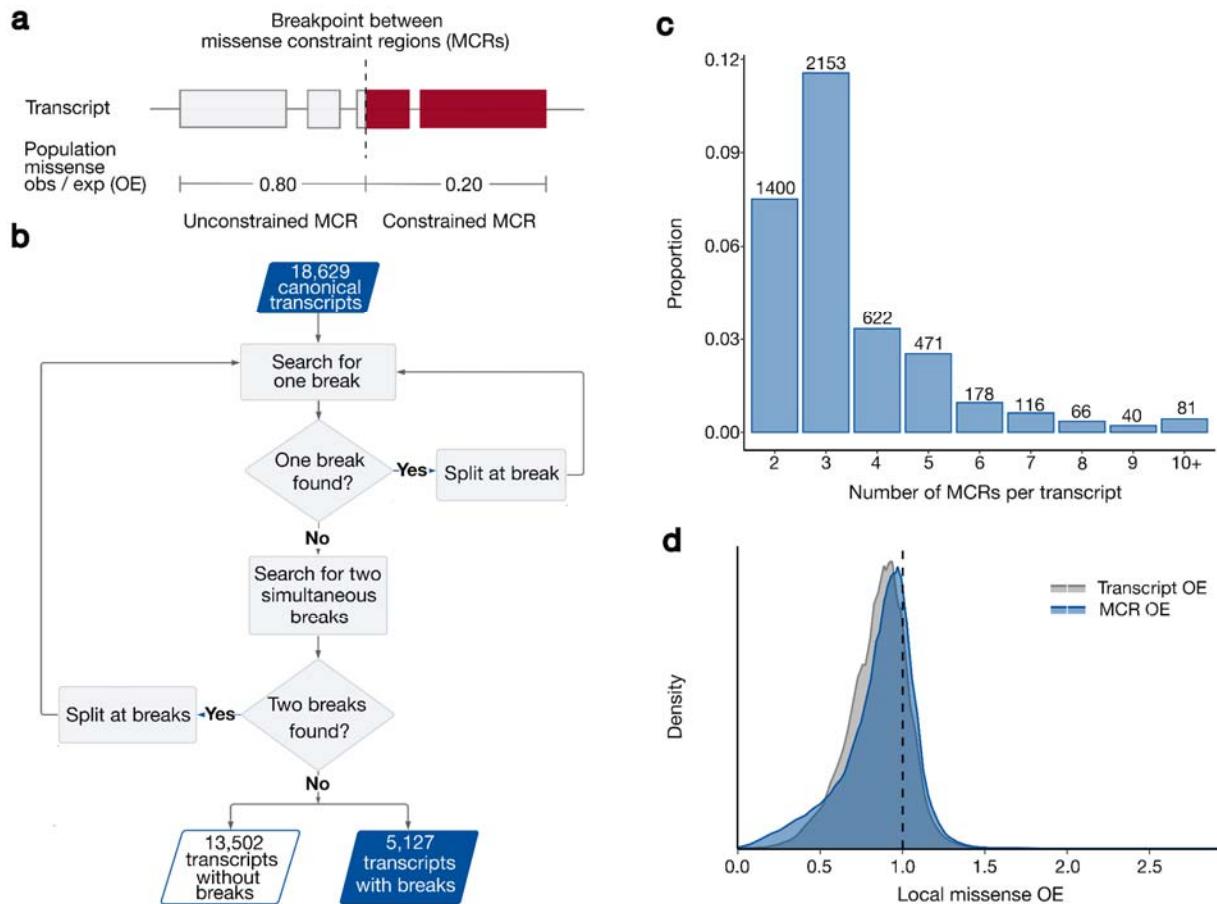
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34 We explored the patterns of rare missense variant presence or absence in 125,748 exomes in  
35 the Genome Aggregation Database (gnomAD) v2.1.1 on GRCh37 to quantify missense  
36 depletion at the sub-genic level. We searched 18,629 canonical protein-coding transcripts for  
37 variability in missense constraint, quantified as the number of rare (allele frequency [AF] <  
38 0.1%) missense variants observed in gnomAD divided by the number expected under neutral  
39 evolution as estimated from previously described mutational models<sup>13</sup>(observed/expected [OE]).  
40 For each transcript, we applied a recursive search based on likelihood ratio tests over all  
41 potential rare missense sites looking for breaks that divide the transcript coding sequence  
42 (CDS) into distinct missense constraint regions (MCRs; **Fig. 1a, b**). We discover 5,127  
43 transcripts (28%) harbor regional variability in missense constraint (**Fig. 1c**), i.e., have two or  
44 more MCRs (minimum coding length 49bp, median 461bp; **Supplementary Fig. 1**). We thus  
45 refine the resolution of missense constraint for 42% of coding sites (coding space in the 5,127  
46 transcripts vs. 18,629 total assessed). After recalibrating the missense OE distribution over all  
47 potential sites of missense variants using MCR-wide rather than transcript-wide missense OE  
48 measurements, we discover widespread signatures of negative and neutral selection that are  
49 obscured when quantifying over the unit of whole transcripts (**Fig. 1d**). We find a larger  
50 proportion of the exome lies within strongly constrained sequences (5.6% vs. 1.7% at OE < 0.4;

51 see **Supplementary Note** for OE threshold selection), and the mode of the distribution shifts  
52 toward an OE indicative of evolutionary neutrality at approximately 1 (40.6% vs. 36.5% at  $0.9 <$   
53  $OE \leq 1.1$ ).

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57 **Fig. 1:** 28% of protein-coding genes in the human genome are discovered to harbor regional  
58 variation in population-level missense depletion.

59 **a**, An example transcript that has two missense constraint regions (MCRs) with significantly  
60 different levels of population-wide missense depletion, defined as the number of missense  
61 variants observed in gnomAD at rare frequency ( $AF < 0.1\%$ ) divided by the number of rare  
62 missense variants expected under neutral evolution (observed/expected or OE). Lower OE  
63 values correspond to greater variant depletion in the population and suggest stronger constraint.

64 **b**, Flow chart describing the process of searching for breakpoints that divide a transcript into  
65 multiple MCRs. Searching for breakpoints is recursive and leverages likelihood ratio tests at a  
66 significance threshold of  $p = 0.001$ . **c**, The number of MCRs within the 5,127 transcripts  
67 discovered to harbor regional differences in missense constraint. The other 13,502 transcripts  
68 are deemed to have a single MCR (that is, a constant level of constraint across their entirety)  
69 and are not shown. **d**, The distribution of local missense OE at all coding sites in canonical  
70 transcripts. Local missense OE is defined as the OE calculated over the whole transcript (for

71 “transcript OE”) or over the MCR (for “MCR OE”) where the site is located. Transcript OE and  
72 MCR OE are equivalent for transcripts with one MCR.

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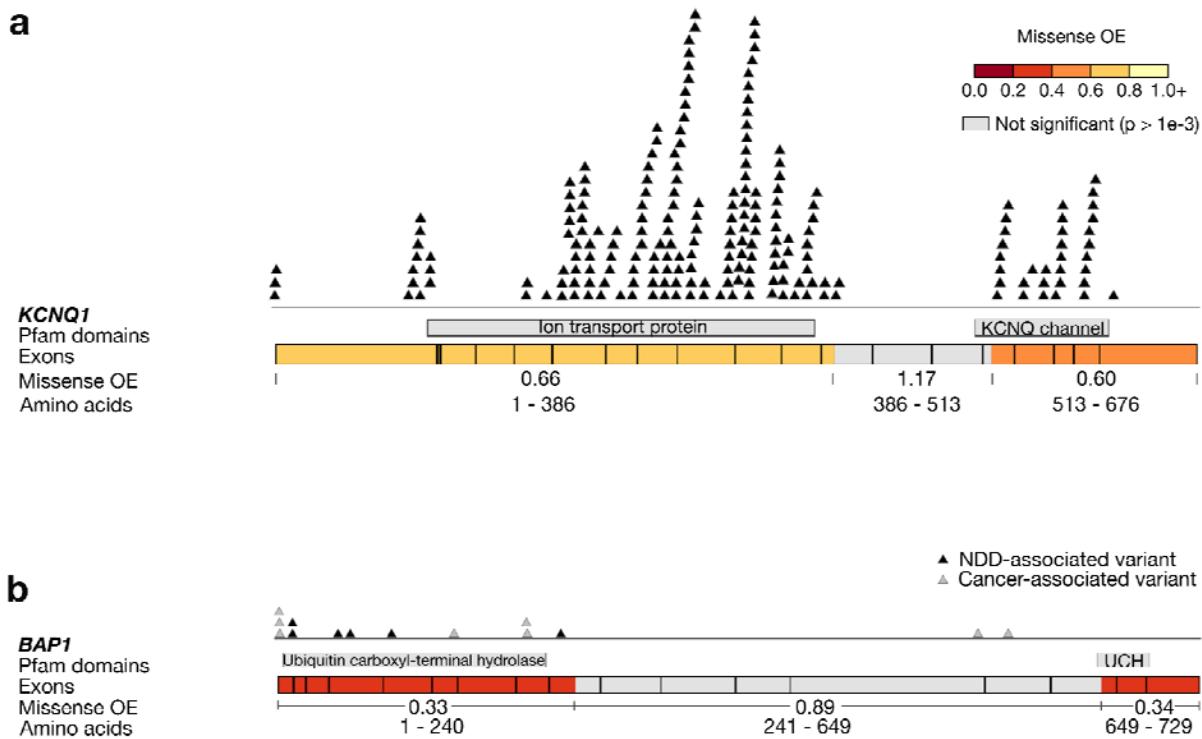
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75 Furthermore, we find that constrained MCRs overlap established disease-associated mutational  
76 hotspots, including critical protein domains. One example is in the well-characterized *KCNQ1*, a  
77 voltage-gated potassium channel gene, in which pathogenic variants cause cardiac disorders  
78 such as long QT syndrome. We discover one moderately constrained MCR (missense OE =  
79 0.60) overlapping the highly conserved C-terminus<sup>29</sup> and another (missense OE = 0.66)  
80 encompassing the voltage-sensing and pore domains (**Fig. 2a**). Both the C-terminus of *KCNQ1*  
81 and its voltage-sensing domain are established “hotspot” regions (specific missense-  
82 constrained regions with ACMG/AMP hotspot/functional domain moderate support [PM1] for  
83 pathogenicity)<sup>29–31</sup>. All but two ClinVar pathogenic/likely pathogenic (P/LP) missense variants in  
84 this gene fall within these two missense-constrained MCRs.

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86 We also find that missense constraint within MCRs is able to identify regions associated with  
87 severe, early-onset disease. One example is in *BAP1*, which plays a key role in chromatin  
88 modeling by mediating histone deubiquitination. Disease-causing variants in this gene are linked  
89 to cancer or, as recently discovered, Kury-Isidor syndrome<sup>32</sup>. The first highly missense-  
90 constrained MCR (missense OE = 0.33) in *BAP1* encompasses the ubiquitin C-terminal  
91 hydrolase domain connected to Kury-Isidor syndrome<sup>32</sup> (**Fig. 2b**), and all 11 variants reported to  
92 be causal for Kury-Isidor fall within this MCR. The only ClinVar P/LP variants that do not fall  
93 within any missense-constrained MCRs in *BAP1* are associated with cancer phenotypes, which  
94 may be under weaker selection than neurodevelopmental disorders (NDDs).

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97 **Fig. 2:** Missense constraint regions (MCRs) and the distribution of ClinVar pathogenic/likely  
98 pathogenic (P/LP) missense variants in two genes associated with early-onset developmental  
99 disorders. Exons are delineated with black outlines and MCRs are delineated by color. MCRs  
100 are colored based on their missense observed/expected (OE) ratio, and MCRs with missense  
101 OEs not significantly different from 1 ( $p > 0.001$ ) are shaded gray. **a**, *KCNQ1*. Only two of the  
102 210 P/LP missense variants in *KCNQ1* do not fall within either constrained MCR. The first  
103 constrained MCR encompasses the voltage-sensing and pore domains of this gene, and the  
104 other constrained MCR overlaps the C-terminus. Both domains contain previously reported  
105 hotspot regions, with some regions reaching moderate level (PM1) support for pathogenicity<sup>31</sup>.  
106 Ion transport protein: domain that contains both the transmembrane voltage-sensing and pore  
107 domains. KCNQ channel: C-terminal cytoplasmic domain that overlaps four helices (A-D). **b**,  
108 *BAP1*. Variants in this gene can lead to cancer-predisposition syndromes, increased risk of  
109 certain cancers, or the neurodevelopmental disorder Kury-Isidor syndrome<sup>32</sup>. All of the ClinVar  
110 P/LP variants associated with Kury-Isidor fall within the first MCR with a highly depleted  
111 missense OE of 0.33. An additional five variants reported in Kury *et al.*<sup>32</sup> but not ClinVar fall  
112 within either highly constrained MCR in this gene. P/LP variants associated with Kury-Isidor are  
113 colored in black, and all other cancer-associated P/LP variants are colored in gray. UCH:  
114 Ubiquitin carboxyl-terminal hydrolase isozyme L5 domain. ClinVar data are from the October 15,  
115 2023 release.  
116  
117  
118 Next, we sought to determine if the signatures of selection revealed by MCRs recapitulated  
119 biological and disease relevance of coding sequences. Overall, most transcripts that are

120 intolerant to pLoF variation (as measured by the loss-of-function observed/expected upper  
121 bound fraction [LOEUF] score<sup>13</sup>) also tend to be intolerant to missense variation. This trend is  
122 markedly more prominent when measuring missense constraint at the sub-genic level vs. the  
123 transcript-level (**Supplementary Note; Supplementary Fig. 2**). We also discovered that 64%  
124 (1697/2659) of genes that are both LOEUF- and MCR missense-constrained do not have  
125 disease associations in OMIM<sup>33</sup>, suggesting the existence of many undocumented genes  
126 containing variants of significant consequence for disease (**Supplementary Fig. 3**). In a set of  
127 730 strongly mutationally intolerant genes, defined here as exhibiting both population depletion  
128 of pLoF variants (first three LOEUF deciles) and association with a developmental phenotype  
129 (high-confidence membership in any non-cancer Gene2Phenotype [G2P]<sup>34</sup> gene list with  
130 dominant inheritance), we observed strong transcript-wide missense depletion that was even  
131 stronger for genes with multiple MCRs (**Fig. 3a** and **Supplementary Fig. 4**; Wilcoxon  $p < 10^{-50}$ ).  
132 Given that we have greater power to detect missense constraint variability over longer  
133 sequences (**Supplementary Fig. 5**), we controlled for transcript length but still found that  
134 intolerant transcripts are eight times more likely to harbor multiple MCRs ( $p < 10^{-50}$ ). These  
135 strongly intolerant transcripts are highly enriched for severely depleted regions (three times  
136 more likely to have minimum MCR OE  $< 0.4$  after regressing out transcript length,  $p < 10^{-18}$ ),  
137 whereas the most constrained MCRs in not strongly intolerant transcripts are less depleted and  
138 more evenly distributed across the OE spectrum. Finally, we observe a group of genes with  
139 strong overall missense depletion in which we did not detect multiple MCRs ( $n = 459$  with  
140 missense OE  $< 0.4$ ; **Supplementary Table 1**), suggesting these genes are robustly intolerant to  
141 missense variants across their length. When comparing missense constraint to selection over  
142 longer timescales (measured by evolutionary conservation in placental mammals, phyloP<sup>35</sup>), we  
143 found that genes with more conserved coding sequences also tended to be more overall  
144 depleted of human missense variation (Spearman  $p = 0.56$ ,  $p < 10^{-50}$ ). However, a substantial  
145 number of strongly constrained MCRs appear widely unconserved across mammals, potentially  
146 pointing to human-specific negative selection pressures that are obscured at the whole-  
147 transcript level (**Supplementary Fig. 6**).  
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149 We next aggregated *de novo* missense variants from 31,058 individuals with a severe  
150 developmental disorder<sup>15</sup> (DD), 15,036 autistic individuals (AUT), and 5,492 siblings not  
151 diagnosed with a DD<sup>16</sup> (**Fig. 3b**). The distribution of *de novo* missense variants across the  
152 missense OE spectrum in unaffected siblings largely mirrored the exome-wide missense OE  
153 distribution. In contrast, *de novo* missense variants in autistic individuals are enriched in

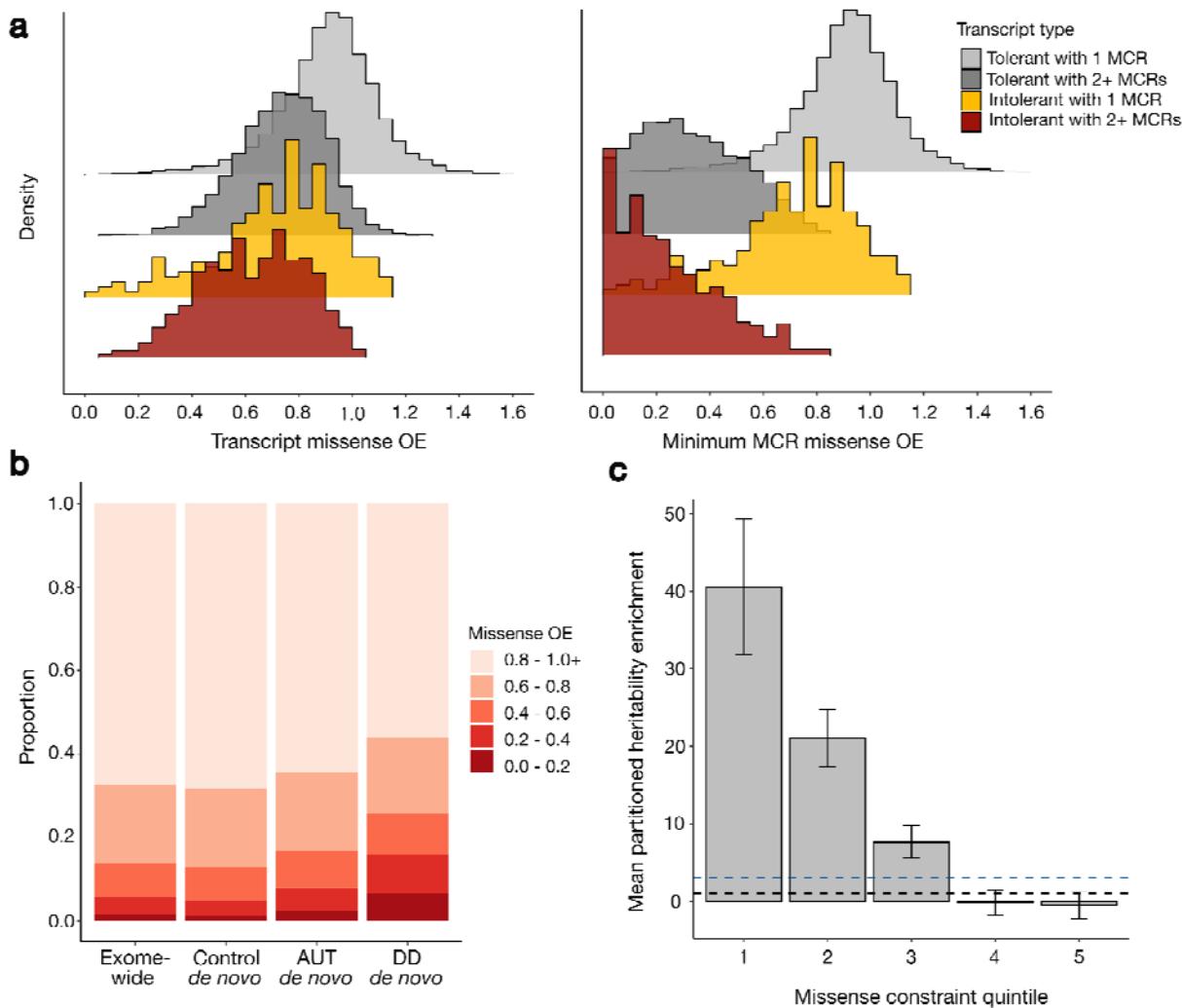
154 missense-constrained sequences, and this pattern is more striking in individuals with DDs. For  
155 example, relative to unaffected siblings, the rate of *de novo* missense variants in MCRs with OE  
156 < 0.2 is 2-fold higher in autistic individuals ( $p < 10^{-23}$ ) and 6.6-fold higher in individuals with DDs  
157 ( $p < 10^{-50}$ ) (**Supplementary Fig. 7**; see **Supplementary Note** for OE threshold selection). This  
158 is consistent with the expectations that a small subset of *de novo* missense variants in  
159 individuals with developmental phenotypes are causal for those traits and that variants causal  
160 for DD are generally more selectively deleterious than those for autism.

161

162 Beyond large-effect rare and *de novo* variation in traits under strong negative selection, we  
163 additionally investigated whether our MCR metric, which was calculated using rare variants,  
164 correlates with functional effects of common variants. Prior work found that pLoF-constrained  
165 genes and their flanking 100kb sequences are enriched for SNP heritability across hundreds of  
166 independent traits in the UK Biobank (UKBB) and other large genome-wide association studies  
167 (GWAS)<sup>13</sup>. We partitioned common (AF > 5%) variant heritability of the same 268 independent  
168 traits across MCRs to investigate relative enrichment. To establish a baseline, we computed the  
169 heritability enrichment over all coding sequences comprising MCRs (3-fold). The most  
170 constrained MCRs have the strongest heritability enrichment; the first quintile of MCR missense  
171 OE harbors a 41-fold enrichment (**Fig. 3c**). Coding SNPs in missense-unconstrained MCRs  
172 (e.g., in the two least constrained quintiles of MCR missense OE) harbor no detectable  
173 heritability enrichment relative to the average genome-wide SNP. These findings suggest that:  
174 1) regions depleted of rare missense variation can help prioritize common coding variants  
175 important for complex traits (i.e., improve GWAS fine-mapping variant prioritization), and 2)  
176 there exists a subset of coding sequence with no appreciable heritability enrichment, which rare  
177 variant depletion can help identify.

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**Fig. 3:** Regional missense depletion reveals constraint obscured by gene-level measures. **a**, Left: The distribution of transcript-wide missense observed/expected (OE) across 18,629 transcripts stratified by the combination of two factors: whether the transcript is strongly mutationally intolerant (within first three LOEUF deciles and association with a developmental phenotype in Gene2Phenotype [G2P]<sup>34</sup>) and whether we detect multiple missense constraint regions (MCRs). Number of transcripts in each category are: strongly intolerant with multiple MCRs (n=581; red), strongly intolerant with one MCR (n=149; yellow), not strongly intolerant with multiple MCRs (n=4,546; dark gray), not strongly intolerant with one MCR (n=13,353; light gray). X-axis is cut off at 1.6 for visibility. Right: Minimum MCR missense OE using the same groupings. Minimum MCR missense OE is the same as transcript missense OE for transcripts with a single MCR. **b**, MCR missense OE at all sites of possible exome-wide missense variants vs. sites of *de novo* missense variants in controls, autistic individuals (AUT), or individuals with DD. *De novo* variants from individuals with developmental phenotypes are enriched in more constrained sequences, with a more pronounced enrichment in DD than autism. **c**, Enrichment in per-variant heritability explained by common (AF > 5%) protein-coding SNPs stratified by MCR missense OE quintile, relative to the average SNP genome-wide. Enrichment is estimated by linkage disequilibrium score regression, accounting for number of SNPs in each quintile, and

198 is averaged across 268 independent traits in UKBB and other large genome-wide association  
199 studies. Black dashed line at 1 indicates no enrichment. Blue dashed line at 3 indicates average  
200 coding enrichment. Error bars represent 95% confidence intervals.

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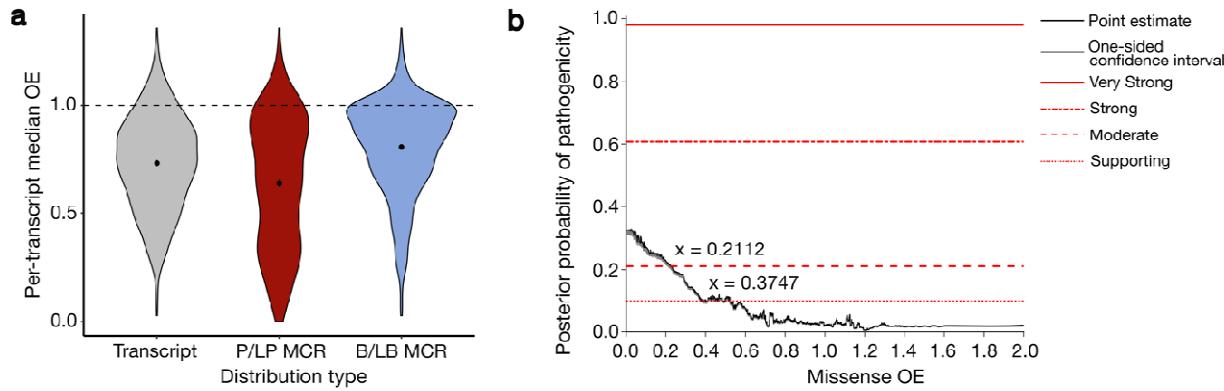
203 We examined the localization of high-quality ClinVar<sup>36</sup> missense variants classified as P/LP  
204 within genes with both unconstrained (missense OE > 0.9) and constrained (missense OE <  
205 0.2) MCRs and found that P/LP variants occur much more frequently in missense constrained  
206 MCRs (odds ratio [OR] = 15.2;  $p < 10^{-50}$ ). We also examined the localization of P/LP and  
207 benign/likely benign (B/LB) variants within MCRs in autosomal dominant disease-associated  
208 genes and found that P/LP variants tend to localize to regions that are more strongly missense-  
209 constrained than the overall transcript (Wilcoxon  $p = 3.5 \times 10^{-10}$ ), while B/LB variants show the  
210 opposite effect and tend to occur in regions with OEs closer to 1 (Wilcoxon  $p < 10^{-18}$ ; **Fig. 4a**).  
211 While more subtle, these same patterns are also significant in autosomal recessive disease-  
212 associated genes (**Supplementary Fig. 8**).

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214 To enable use of our missense constraint metric in ACMG/AMP clinical variant classification, we  
215 applied previously established probabilistic frameworks<sup>37</sup> to determine the MCR missense OE  
216 thresholds that met different levels of clinical evidence strengths evaluated under the  
217 hotspot/functional domain (PM1) and benign *in silico* prediction (BP4) criteria codes<sup>30</sup>. MCR  
218 missense OE  $\leq 0.37$  met supporting (PM1\_Supporting) and OE  $\leq 0.21$  met moderate (PM1)  
219 levels of evidence for pathogenicity (**Fig. 4b**), but no MCR missense OE threshold met any  
220 levels of evidence to support benignity. However, separate calibration specifically in transcripts  
221 with multiple MCRs found that MCR missense OE  $\geq 1.56$  met moderate and OE  $\geq 0.97$  met  
222 supporting evidence for BP4, indicating that in transcripts where we are powered to characterize  
223 regional constraint, MCRs with OEs close to one harbor an indication of benignity  
224 (**Supplementary Fig. 9**). Calibration of two additional regional constraint metrics, Constrained  
225 Coding Regions (CCRs<sup>9</sup>) and COnact Set MIssense tolerance (COSMIS<sup>12</sup>), which incorporates  
226 predicted 3D structure information, revealed that these metrics also reach moderate support for  
227 pathogenicity (PM1), and COSMIS only reaches supporting levels for benignity  
228 (**Supplementary Fig. 9; Supplementary Table 2**).

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232 **Fig. 4:** ACMG/AMP calibration of missense constraint. **a**, The distribution within genes with  
233 autosomal dominant disease associations of transcript-wide missense observed/expected (OE;  
234 gray) and missense constraint region (MCR) OE for ClinVar pathogenic/likely pathogenic  
235 missense variants (P/LP; red) and benign/likely benign missense variants (B/LB; blue). We  
236 filtered to 1,007 transcripts with at least one P/LP and one B/LB missense variant. For the P/LP  
237 and B/LB distributions, we annotated each variant with the missense OE across the MCR they  
238 fell in and collapsed these values within each transcript by taking the respective medians. **b**,  
239 Local posterior probabilities of pathogenicity given MCR missense OE in all transcripts. Gray  
240 shading indicates the one-sided 95% confidence interval on the more stringent side. Horizontal  
241 lines indicate thresholds required to meet ACMG/AMP evidence levels. From bottom to top:  
242 supporting, moderate, strong, very strong. MCR missense OE reaches supporting (OE ≤ 0.37)  
243 and moderate (OE ≤ 0.21) level evidence for PM1 (hotspot/functional domain).

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246 We transformed our regional missense constraint measure into a variant-level predictor of  
247 missense deleteriousness named MPC (Missense deleteriousness Prediction by Constraint)  
248 that additionally incorporates information about amino acid substitution type and local context.  
249 The logistic regression-based model integrates regional missense constraint-derived metrics  
250 together with BLOSUM<sup>38</sup>, Grantham<sup>39</sup>, and PolyPhen-2<sup>40</sup> and is trained on ClinVar pathogenic  
251 and gnomAD common (AF > 0.1%) variants in 2,987 genes defined as haploinsufficient in  
252 Collins *et al.*<sup>41</sup> and 366 genes with DD associations in G2P through non-LoF mechanisms.

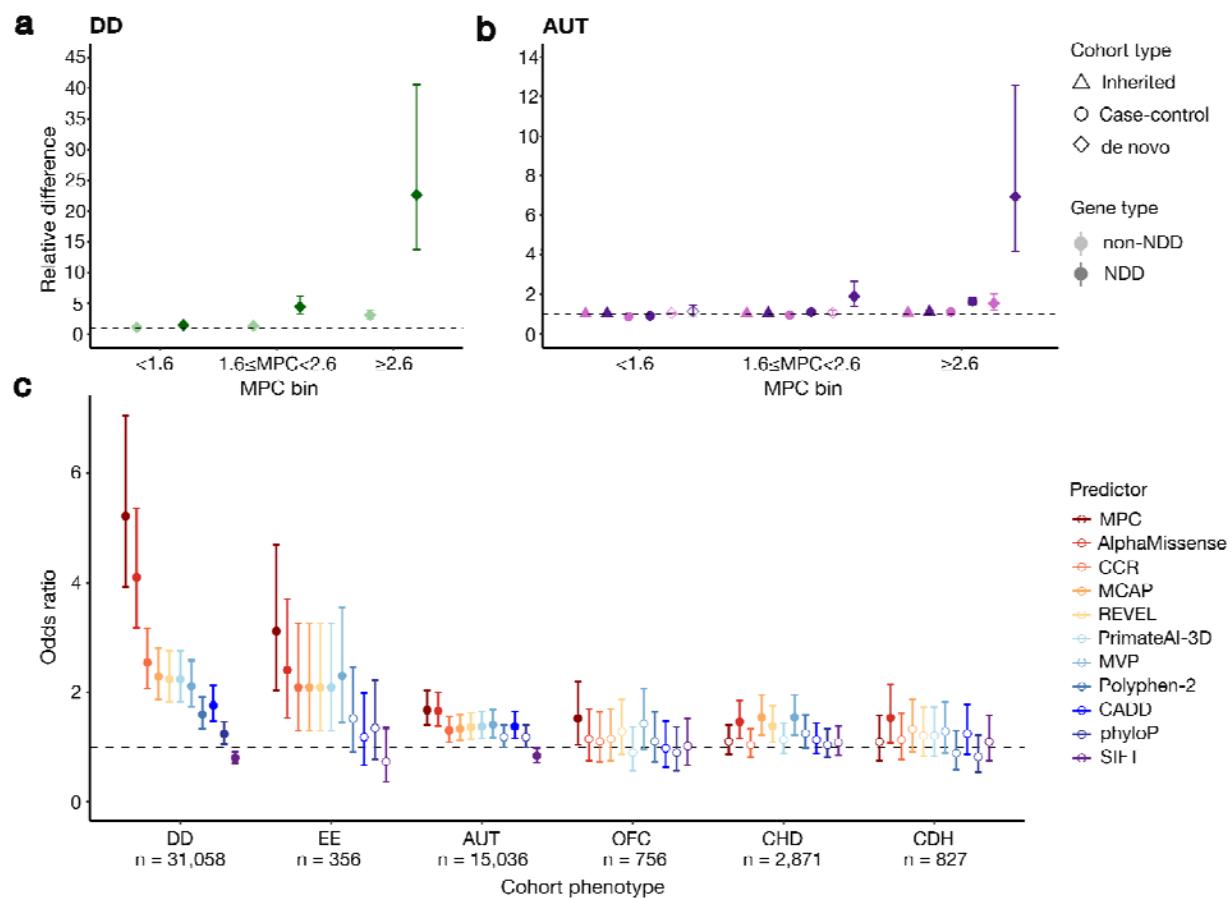
253 Higher scores predict greater deleteriousness (**Supplementary Fig. 10, 11**). We assessed the  
254 utility of MPC in prioritizing potentially disease-causing variation by evaluating its ability to  
255 stratify case and control rare and *de novo* missense variation. Consistent with the regional  
256 constraint results, the *de novo* missense variants from DD and AUT cases are enriched for high  
257 MPC scores compared to controls (**Supplementary Fig. 12**). We further stratified by presence  
258 in 373 genes previously associated with NDD<sup>16</sup> and three bins of MPC scores (< 1.6, 1.6-2.6, ≥  
259 2.6; see **Supplementary Note** for calibration of these bins), and found a very strong enrichment  
260 of *de novo* missense variants in the two most deleterious bins among both the DD (**Fig. 5a**) and

261 AUT cases (**Fig. 5b**) compared to unaffected individuals. However, while the enrichment in the  
262 373 NDD-associated genes was significant for missense variants with MPC  $\geq 2.6$  (RR in DD  
263 cases =  $22.7$ ,  $p < 10^{-50}$ ; RR in AUT cases =  $6.9$ ,  $p < 10^{-21}$ ), as well as missense variants with  
264 MPC between 1.6-2.6 (RR in DD cases =  $4.5$ ,  $p < 10^{-35}$ ; RR in AUT cases =  $1.9$ ,  $p = 3.0 \times 10^{-5}$ ), it  
265 was only significant in NDD-unassociated genes for missense variants with MPC  $\geq 2.6$  (RR in  
266 DD cases =  $3.1$ ,  $p < 10^{-28}$ ; RR in AUT cases =  $1.5$ ,  $p = 5.9 \times 10^{-4}$ ). This suggests that while there  
267 is a sizable reservoir of potentially causal variants in genes yet to be associated with NDDs,  
268 they will be more difficult to find as they must reach stricter deleteriousness criteria. For autism,  
269 we additionally assessed inheritance rates of rare missense variants (AF < 0.1%) from parents  
270 to 13,384 probands and case-control rates for an additional 5,591 cases and 8,597 controls  
271 without *de novo* information. While we did not find substantial enrichment in inheritance rates in  
272 any missense category, we discovered substantial enrichment in the case-control analysis for  
273 variants in the 373 NDD-associated genes with MPC  $\geq 2.6$  (RR =  $1.6$ ,  $p < 10^{-12}$ ), which we infer  
274 is from *de novo* variants that are not recognizable as such due to lack of parental information.

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276 We extended our assessment of case-control *de novo* stratification for a comparison of our  
277 model against several other missense deleteriousness predictors: AlphaMissense<sup>42</sup>, CCRs<sup>9</sup>, M-  
278 CAP<sup>43</sup>, REVEL<sup>44</sup>, PrimateAI-3D<sup>45</sup>, MVP<sup>46</sup>, Polyphen-2<sup>40</sup>, CADD<sup>47,48</sup>, mammalian conservation  
279 phyloP<sup>35</sup>, and SIFT<sup>49</sup>. For this assessment, we evaluated four additional early-onset  
280 development-related phenotypes: epileptic encephalopathy (EE), orofacial cleft (OFC),  
281 congenital heart disease (CHD), and congenital diaphragmatic hernia (CDH). To compare  
282 across predictors with different score distributions, we used a ranking-based performance  
283 assessment. For each predictor, we ranked the *de novo* missense variants from each case  
284 cohort against those in the 5,492 controls and computed the OR of case vs. control variants in  
285 the top percentiles of these rankings (**Fig. 5c**). At the top 10% of variants, MPC displays the  
286 highest OR for DD (OR =  $5.2$ , Fisher's exact  $p < 10^{-48}$ ), EE (OR =  $3.1$ ,  $p = 2.2 \times 10^{-7}$ ), AUT (OR =  
287  $1.7$ ,  $p = 8.9 \times 10^{-9}$ ), and OFC (OR =  $1.5$ ,  $p = 0.025$ ), although there is substantial confidence  
288 interval overlap with other predictors. This indicates that MPC effectively ranks high-impact *de*  
289 *novo* variants in the most deleterious prediction regimes. Of the other predictors,  
290 AlphaMissense also performs consistently well across all phenotypes. In particular, in CHD and  
291 CDH, which have the least *de novo* enrichment across predictors, we observe MPC lagging in  
292 performance, while AlphaMissense is one of the top performers. This may suggest that causal  
293 *de novo* variants in these phenotypes may occur at a narrow set of sites where 3D structure is  
294 important, which AlphaMissense can more deftly capture through integration of protein structure

295 prediction. These observations are more or less consistent over a range of thresholds used to  
296 define the ranking top percentiles (**Supplementary Fig. 13**).  
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299  
300 **Fig. 5:** MPC effectively stratifies case and control variation.  
301 **a**, The difference relative to controls of missense variants stratified by MPC score and  
302 localization to genes associated with neurodevelopmental disorders (NDDs) for **a**, individuals  
303 with DD and **b**, autistic individuals (AUT). Relative difference is calculated as: for *de novo*  
304 variants, the average rate of variants in probands divided by that in sibling controls; for case-  
305 control, the average rate of variants in cases divided by that in controls from case-control data;  
306 for inherited, the average rate in probands of transmitted variants divided by that of  
307 untransmitted variants. Error bars represent 95% confidence intervals calculated from a  
308 binomial test. **c**, The odds ratio of case to control *de novo* missense variants in the top 10% vs.  
309 bottom 90% of respective rankings. *De novo* missense variants from each case cohort are  
310 ranked against those in the 5,492 controls for each predictor. DD: developmental disorders, EE:  
311 epileptic encephalopathy, AUT: autism, OFC: orofacial cleft, CHD: congenital heart disease,  
312 CDH: congenital diaphragmatic hernia. Error bars represent 95% confidence intervals. Only  
313 variants scored by all predictors are included. Points are solid colored if the difference from 1 is  
314 statistically significant (binomial or Fisher exact  $p < 0.05$ ).  
315

316

317 We have developed a method to identify sub-genic regions with differential intolerance to  
318 missense variation at base-level resolution. We demonstrate that coding regions depleted for  
319 missense variation in the general population are enriched for established disease-associated  
320 variation, *de novo* variants from individuals with NDDs, and heritability for 268 complex traits  
321 from the UK Biobank and other large GWAS. Additionally, we have calibrated these constraint  
322 scores to establish that regions with less than 20% of their expected variation can achieve  
323 moderate evidence for association to disease following ACMG/AMP guidelines. Finally, we  
324 incorporated regional missense intolerance information into the missense deleteriousness  
325 metric, MPC, and show that MPC effectively separates potentially risk-carrying variants  
326 identified in various developmental disorder cases from those seen in controls.

327

328 At current sample sizes, we are unable to characterize constraint at single amino acid  
329 resolution. Furthermore, because our approach relies on variant presence or absence in a large  
330 reference dataset, many of the constrained regions we find are linked to variants that cause  
331 severe, early-onset disease. However, the true nature of the variation we capture is more  
332 accurately linked to reproductive fitness and the strength of selection acting on heterozygotes<sup>50</sup>.  
333 Our methodology specifically searches for linear sub-genic regions in canonical transcripts that  
334 are depleted of missense variants compared to a null mutational model. This means that our  
335 model is unable to find depleted sequences that are clustered specifically in 3D space and is  
336 also currently ignorant of coding sequences not present in the Ensembl canonical transcript.  
337 However, we note that our linear metric achieves similar evidence for both pathogenicity and  
338 benignity as the structural constraint-based COSMIS model<sup>12</sup> (**Supplementary Fig. 9**).

339

340 In summary, we identify 28% of canonical transcripts with variable levels of missense constraint  
341 and demonstrate that coding regions specifically depleted of missense variation in the general  
342 population are enriched for disease-associated variation. Additionally, we show that this  
343 depletion of missense variation can be used as moderate evidence when classifying variants  
344 according to ACMG/AMP guidelines and that incorporation of regional missense constraint into  
345 an *in silico* predictor effectively prioritizes a subset of *de novo* missense variation in individuals  
346 with developmental phenotypes for association testing. We have publicly released these data  
347 for use in both research and clinical settings. We anticipate refined resolution of these metrics  
348 as datasets grow, both in size and in ancestral diversity, and with the incorporation of  
349 complementary structural or functional data.

350 **Methods**

351 **Transcripts**

352 This study analyzed only canonical, coding transcripts as defined by GENCODE v19/Ensembl  
353 v74. We excluded the same set of transcripts from this analysis that were excluded in the  
354 previous gnomAD v2.1.1 genic constraint estimates<sup>13</sup>. Briefly, we excluded transcripts that had  
355 outlier counts of variants expected under neutrality (zero expected pLoF, missense, or  
356 synonymous variants; too many observed pLoF, missense, or synonymous variants compared  
357 to expectation; or too few observed synonymous variants compared to expectation). In total, this  
358 study analyzed 18,629 transcripts.

359

360 **gnomAD variants**

361 All analyses in this paper were conducted using the 125,748 gnomAD v2.1.1 exomes<sup>13</sup> on  
362 GRCh37. Median coverage was calculated on a random subset of the gnomAD exomes as  
363 described previously<sup>13</sup>. We defined the set of sites with possible missense variants using a  
364 synthetic Hail Table (HT) containing all possible single nucleotide variants in the exome. We  
365 annotated this HT with the Variant Effect Predictor (VEP, version 85) against GENCODE  
366 version 19, and filtered to variants with the consequence "missense\_variant" in the canonical,  
367 coding transcripts as defined in *Transcripts*. We then further filtered to variants that fit one of the  
368 following criteria: (1) allele count (AC) > 0 and AF < 0.001, variant QC PASS, and median  
369 coverage > 0 in gnomAD v2.1.1 exomes; or (2) AC = 0, i.e. variants not seen in gnomAD v2.1.1  
370 exomes.

371

372 **ClinVar variants**

373 We annotated functional consequences for ClinVar<sup>14</sup> (v.20230305) variants using the VEP table  
374 described in *gnomAD variants*. Missense ClinVar variants with non-conflicting P, LP, B, LB  
375 classification and a review status of at least one star were selected for analysis.

376

377 **Rare and *de novo* variants from developmental cohorts**

378 Case *de novo* mutations for association analyses were obtained from studies of developmental  
379 disorders<sup>15</sup> (DD), autism<sup>16</sup> (AUT), congenital heart disease<sup>51</sup> (CHD), orofacial cleft<sup>52</sup> (OFC),  
380 congenital diaphragmatic hernia<sup>53</sup> (CDH), and epileptic encephalopathy<sup>54</sup> (EE). Control *de novo*  
381 mutations were obtained from neurotypical siblings of the autistic probands<sup>16</sup>. Variants from the  
382 autism study were lifted over from GRCh38 to GRCh37 using the "liftover" function in Hail.  
383 Variant functional consequences were re-annotated using the VEP table described in *gnomAD*  
384 *variants*. Variants transmitted and not transmitted from parents to autistic probands were  
385 procured from previously published ASC-SSC and SPARK cohorts, and case-control variants  
386 for autism were procured from previously published iPSYCH and Swedish cohorts<sup>16</sup>. Both the  
387 inherited/uninherited and case-control variant sets were filtered to AF < 0.1%.

388

389 Training, validation, and test datasets

390 To generate independent training and test sets, we selected 80% (14,894 transcripts) of the  
391 18,629 canonical coding transcripts to comprise the training set and the remaining 20% (3,735  
392 transcripts) to the test set. To ensure the training and test transcripts have similar distributions  
393 of features that may impact constraint estimates, we used stratified randomization to match the  
394 training and test transcripts on  $s_{het}$  coefficients (as a measure of selection) and number of  
395 potential missense sites (as a measure of power to detect transcript-wide constraint changes).  
396 The training set was used for MPC model training and MCR model selection, and the test set  
397 was used for MPC model evaluation. No similar hold-outs of data were performed for training of  
398 the mutational model used to compute expected variant counts (see *Modeling of mutation rates*  
399 and *expected neutral missense variation*).

400

401 Modeling of mutation rates and expected neutral missense variation

402 Expected missense variant counts were determined as described previously<sup>13</sup>. Briefly, we  
403 created a model using the 15,708 gnomAD v2.1.1 genomes that estimated the mutation rate for  
404 each single nucleotide substitution with one base of context (e.g., ACT > AGT) in non-coding  
405 regions of the genome. We then calibrated this mutation rate against the proportion observed of  
406 each context at synonymous sites to adjust for the larger size of the gnomAD v2 exomes,  
407 adjusting for low coverage regions (median coverage < 40x) and methylation levels at CpG sites  
408 using methylation data from the Roadmap Epigenomics Consortium<sup>55</sup>. We created three  
409 separate models (referred to as "plateau" models moving forwards): one for autosomal and  
410 pseudoautosomal sites, one for chromosome X sites, and one for chromosome Y sites. Each of  
411 these models contains mutation rate estimates for each substitution, context, and methylation  
412 level. We then applied the plateau models to the proportion observed of each substitution and  
413 its context, exome coverage, and methylation level. We counted all possible variants in our  
414 synthetic Hail Table (HT) that passed the following criteria: (1) Median coverage > 0; (2) no low-  
415 quality variant observed in gnomAD v2 exomes; (3) no variants above 0.1% AF observed in  
416 gnomAD v2 exomes. We then correlated this proportion observed value with the mutation rate  
417 calculated using the appropriate model above. For low coverage sites (median coverage below  
418 40x), we calculated a scaling factor as described previously<sup>13</sup>: briefly, we computed the total  
419 number of observed synonymous variants in the gnomAD v2 exomes divided by the total  
420 number of possible synonymous variants in the synthetic HT multiplied by the mutation rate  
421 aggregated across all possible substitutions and their contexts and methylation levels. We used  
422 this scaling factor to create a model to adjust the proportion of expected variation for low  
423 coverage sites (coverage model).

424

425 Identifying breakpoints within transcripts of regional missense constraint

426 Observed missense variant counts were calculated using sites from the 125,748 gnomAD  
427 exomes that passed all the following criteria: (1) Allele count (AC) > 0; (2) allele frequency (AF)  
428 < 0.001; (3) variant QC PASS (passed gnomAD variant QC filters, including random forest  
429 filters); (4) median coverage > 0. We filtered the gnomAD v2 exomes Hail Table (HT) to the

430 sites that matched the above criteria and then annotated the synthetic HT with whether that  
431 variant (chromosome/locus plus reference and alternate alleles) was observed in the gnomAD  
432 exomes. We then aggregated the total number of observed variant counts per locus by  
433 summing the number of observed variants for each possible substitution (reference and  
434 alternate allele) at each locus. Finally, we grouped the synthetic HT by transcript annotation to  
435 sum the total number of observed missense variants per transcript.

436

437 As previously described<sup>13</sup>, we applied the two models (plateau and coverage) described in  
438 *Modeling of mutation rates and expected neutral missense variation* to calculate the total  
439 proportion of expected missense variation. Briefly, we summed the mutation rate (mu\_agg) for  
440 each substitution, context, and methylation level across the exome. We then applied the  
441 appropriate plateau model (autosomal/pseudoautosomal, chromosome X, chromosome Y) and  
442 adjusted CpG vs. non-CpG sites separately. After applying the appropriate plateau model, we  
443 applied the coverage model to low coverage (median coverage < 40x) sites to create the final  
444 adjusted mutation rate (mu\_adj). We then aggregated the raw mutation rate sum (mu\_agg) and  
445 adjusted mutation rate (mu\_adj) per transcript to get the total mutation rate sum and proportion  
446 of expected missense variation per transcript.

447

448 We implemented a minimum number of expected missense variants to prevent finding  
449 breakpoint positions that would create very small (i.e., a handful of base pairs in size) transcript  
450 subsections (see **Supplementary Note**).

451

452 We applied a likelihood ratio test to determine whether the missense observed/expected (OE)  
453 ratio was uniform along a transcript or whether a transcript had evidence of distinct sections of  
454 missense constraint. We used the observed and expected missense counts to search for  
455 positions that would divide a transcript into two or more regions with varying levels of missense  
456 depletion. For our analyses, we assume that the observed missense counts should follow a  
457 Poisson distribution around the expected missense counts. We defined our null model as  
458 transcripts not having any evidence of regional variability in missense depletion (where the  
459 expectation, the OE ratio, is consistent across the length of the transcript). Our alternative  
460 model was that transcripts exhibited evidence of distinct sections of missense depletion (OE  
461 ratio calculated per transcript subsection). Because the alternative model should always have a  
462 better fit than the null model, we require a chi square value above a given threshold ( $p = 0.001$ )  
463 to establish significance. We used the following formulas to determine the significance of a  
464 breakpoint that would split a transcript into two sections, A and B:  
465

- 466 •  $p_0 = Pois(obs_A, exp_A * OE) * Pois(obs_B, exp_B * OE)$
- 467 •  $p_1 = Pois(obs_A, exp_A * OE_A) * Pois(obs_B, exp_B * OE_B)$
- 468 •  $\chi^2 = 2(\ln(p_1) - \ln(p_0))$

469

470 where OE is the missense observed/expected ratio across the entire transcript,  $obs_A$  is the  
471 number of observed missense variants in transcript section A,  $exp_A$  is the number of expected  
472 missense variants in transcript section A,  $OE_A$  is the OE ratio across transcript section A,  $obs_B$  is  
473 the number of observed missense variants in transcript section B,  $exp_B$  is the number of

474 expected missense variants in transcript section A,  $OE_B$  is the OE ratio across transcript section  
475 B, and Pois is the Poisson likelihood.

476

477 We used the following formulas to determine the significance of a breakpoint that would split a  
478 transcript into three sections, A, B, and C:

479

$$p_0 = \text{Pois}(obs_A, exp_A * OE) * \text{Pois}(obs_B, exp_B * OE) * \text{Pois}(obs_C, exp_C * OE)$$

- $p_1 = \text{Pois}(obs_A, exp_A * OE_A) * \text{Pois}(obs_B, exp_B * OE_B) * \text{Pois}(obs_C, exp_C * OE_C)$
- $\chi^2 = 2(\ln(p_1) - \ln(p_0))$

483

484 where OE is the missense observed/expected ratio across the entire transcript,  $obs_A$  is the  
485 number of observed missense variants in transcript section A,  $exp_A$  is the number of expected  
486 missense variants in transcript section A,  $OE_A$  is the OE ratio across transcript section A,  $obs_B$  is  
487 the number of observed missense variants in transcript section B,  $exp_B$  is the number of  
488 expected missense variants in transcript section B,  $OE_B$  is the OE ratio across transcript section  
489 B,  $obs_C$  is the number of observed missense variants in transcript section C,  $exp_C$  is the number  
490 of expected missense variants in section C, and Pois is the Poisson likelihood.

491

492 For the purposes of our analyses, all transcript subsections with more observed variants than  
493 expected were capped at an OE of 1, as we were looking for areas of missense depletion and  
494 not missense enrichment. We also converted the expected counts for transcript subsections  
495 with zero expected variants from 0 to  $10^{-9}$  to prevent nonfinite OE values.

496

497 To search for a single breakpoint that would divide a transcript into two subsections, we  
498 calculated chi square statistics (as discussed above) to conduct likelihood ratio tests  
499 simultaneously for every eligible position within a transcript. The positions we considered were  
500 positions with a possible missense variant substitution that had at least 16 expected missense  
501 counts in either direction (i.e., both transcript subsections created by dividing the transcript at  
502 this point would have at least 16 expected missense variants). We then aggregated chi square  
503 values across each transcript to find the maximum value per transcript, and we marked any  
504 positions as breakpoints if the chi square calculated at that position was equal to the maximum  
505 chi square value over all sites in the transcript and significant at  $p = 0.001$ .

506

507 Any transcripts that did not have a single significant breakpoint moved forwards into our two  
508 simultaneous breaks search flow. In this search flow, we again calculated chi square statistics to  
509 conduct likelihood ratio tests for every eligible position pair. For every position with a possible  
510 missense, we calculated the chi square statistic of that position paired with each possible  
511 position downstream as long as the two positions created transcript subsections with at least 16  
512 expected missense variants (i.e., all three of the transcript subsections created would have at  
513 least 16 expected missense variants). Because of the large number of pairwise computations,  
514 this step is the most computationally intensive portion of our algorithm. After completing the

515 single and two simultaneous break search workflows, we merged the results from both search  
516 types.

517  
518 Our breakpoint search flow is recursive, and the steps are as follows: Search for a single  
519 significant breakpoint dividing a transcript into two subsections. If no single significant  
520 breakpoint was found in the transcript, search for two simultaneous breakpoints. Merge the  
521 results from the single and two simultaneous breakpoint searches. Repeat the steps above,  
522 treating each separate transcript subsection as if it were an independent transcript, until no  
523 more significant breakpoints are found.

524

525 Modeling deleteriousness of missense substitution classes with missense constraint  
526 We incorporated two MCR OE-based metrics to measure the increased deleteriousness of  
527 amino acid substitution classes (e.g., Met to Tyr) in functionally important areas of proteins: the  
528 overall OE for each substitution and the second derivative of this OE value per OE bin of  
529 missense constraint (**Supplementary Fig. 14**). To calculate the first metric, the substitution  
530 overall OE, we divided the total number of rare, high quality variants (see *gnomAD variants*)  
531 causing that substitution by the total number of expected variants (see *Modeling of mutation*  
532 *rates and expected neutral missense variation*). To calculate the second metric, the substitution  
533 OE second derivative, we aggregated the OEs of each substitution by MCR OE bin in 10 bins  
534 from 0 to 1.0+ (i.e., for the 0-0.1 OE bin, we calculated all of the observed substitutions that  
535 occurred within regions with a OE between 0 and 0.1 and divided that number by the total  
536 number of expected substitutions occurring in those regions).

537

538 Modeling deleteriousness of individual missense variants

539 We designed a missense variant deleteriousness predictor to explicitly incorporate information  
540 on amino acid substitution class and position-specific variant effects. A logistic regression model  
541 was first trained to differentiate pathogenic from benign missense variants. The pathogenic  
542 training set consisted of high-quality ClinVar variants (see *ClinVar variants*) labeled as  
543 pathogenic or likely pathogenic in 2,987 likely-haploinsufficient genes, defined as having  
544 probability of haploinsufficiency (pHaplo)  $\geq 0.86^{41}$ , or in 366 genes with DD associations in G2P  
545 through non-LoF mechanisms. The latter gene set was created by filtering on the G2P DD panel  
546 (accessed October 6, 2023) to select genes where: 1. confidence\_category is either definitive or  
547 strong evidence, 2. allelic\_requirement was monoallelic, and 3. mutation\_consequence included  
548 altered gene product structure or increased gene product level. The benign training set  
549 consisted of high-quality common variants as described in *gnomAD variants*. Variants matching  
550 criteria for both the benign and pathogenic sets were removed from the training data. We  
551 evaluated models with all possible combinations of the following complementary features: amino  
552 acid substitution overall OE and OE second derivative (see *Modeling deleteriousness of*  
553 *missense substitution classes*); BLOSUM<sup>38</sup> and Grantham<sup>39</sup> scores of amino acid substitution  
554 class severity; the local missense constraint level of a variant (missense OE across the MCR if  
555 applicable, else across the transcript); and PolyPhen-2<sup>40</sup>. We selected BLOSUM, Grantham,

556 and PolyPhen-2 because of the orthogonal information added on top of our OE-based (and  
557 therefore population allele frequency-dependent) metrics. For each model, only variants with all  
558 relevant annotations were used in training the regression model and the subsequent  
559 calculations to produce deleteriousness scores. The deleteriousness score prediction for any  
560 missense variant  $i$  is given as:

561

$$d_i = -\log_{10}(m_i/M)$$
$$m_i = \max(0.83, f_i)$$

564  
565 where  $d_i$  is the deleteriousness score prediction,  $f_i$  is the number of common missense variants  
566 with a fitted value from the regression that is less than the fitted value for variant  $i$ , and  $M$  is the  
567 number of common missense variants (equivalent to the number of benign training variants for  
568 the regression).  $m_i$  is set to have a minimum value of 0.83 to avoid a mathematical error in the  
569 log when the fitted value for a given variant is less than those of all common variants. Larger  
570 values of  $d_i$  indicate stronger predicted-deleteriousness. The best model was chosen to be the  
571 model featurized with all six possible features. The training set for this model consisted of  
572 64,023 benign variants and 12,955 pathogenic variants. This model was then applied to  
573 produce MPC scores for the 68,576,965 possible exome-wide missense variants with all  
574 features, and the distribution of these MPC scores is given in **Supplementary Figs. 10, 11, and**  
575 **12**.

576 Comparison of MPC to other predictors

577 We compared our model to the following missense deleteriousness predictors:  
578 AlphaMissense<sup>42</sup>, Constrained Coding Regions (CCRs)<sup>9</sup>, MVP<sup>46</sup>, M-CAP<sup>43</sup>, PrimateAI-3D<sup>45</sup>,  
579 REVEL<sup>44</sup>, CADD<sup>47,48</sup>, PolyPhen-2<sup>40</sup>, and SIFT<sup>49</sup>. We annotated the case and control *de novo*  
580 missense variants described in *Rare and de novo variants from developmental cohorts* and  
581 ranked the variants based on their annotated scores. To assess each predictor's ability to  
582 stratify case and control variation, we assessed the proportion of case to control variants among  
583 the variants with the top 10% for each score and compared this number to the overall proportion  
584 of case to control variation using a Fisher exact test.

## 585 Data availability

586 The missense constraint regions (MCRs) are displayed on the gnomAD v2 browser  
587 (<https://gnomad.broadinstitute.org>) and available for download on the gnomAD website  
588 (<https://gnomad.broadinstitute.org/downloads#v2>) and in the gnomAD v2 public datasets on  
589 Google, Amazon, and Microsoft clouds. MPC scores for all possible variants in canonical  
590 transcripts is available in the gnomAD v2 public datasets on Google (gs://gcp-public-data-  
591 gnomad/release/2.1.1/regional\_missense\_constraint/gnomad\_v2.1.1\_mpc.ht). gnomAD v2  
592 exome, genome, and coverage data and the table of all possible single nucleotide  
593 polymorphisms used to calculate mutational models and search for MCRs are also available in  
594 the gnomAD public buckets and are easily accessed using code in the gnomAD Hail utilities  
595 GitHub repository  
596 ([https://github.com/broadinstitute/gnomad\\_methods/blob/7c0c994883f321492a48962674d5cae](https://github.com/broadinstitute/gnomad_methods/blob/7c0c994883f321492a48962674d5cae)

597 b289df4c7/gnomad/resources/grch37/gnomad.py#L107 and  
598 [https://github.com/broadinstitute/gnomad\\_methods/blob/7c0c994883f321492a48962674d5caeb](https://github.com/broadinstitute/gnomad_methods/blob/7c0c994883f321492a48962674d5caeb)  
599 289df4c7/gnomad/utils/vep.py#L161).  
600  
601 ClinVar data were downloaded from ClinVar's FTP server  
602 (<https://ftp.ncbi.nlm.nih.gov/pub/clinvar/>). *De novo* variants were extracted from the  
603 supplemental files of the cited studies.  
604  
605 AlphaMissense scores were downloaded from <https://github.com/google-deepmind/alphamissense>. CCRs were downloaded from <https://github.com/quinlan-lab/CCRHTML>.  
606 MCAP scores were downloaded from <http://bejerano.stanford.edu/MCAP/>. REVEL scores were  
607 downloaded from <https://sites.google.com/site/revelgenomics/>. PrimateAI-3D scores were  
608 downloaded from <https://primad.basespace.illumina.com/download>. MVP scores were  
609 downloaded from  
610 [https://figshare.com/articles/dataset/Predicting\\_pathogenicity\\_of\\_missense\\_variants\\_by\\_deep\\_Learning/13204118](https://figshare.com/articles/dataset/Predicting_pathogenicity_of_missense_variants_by_deep_Learning/13204118). PolyPhen-2 and SIFT scores were obtained from VEP<sup>56</sup>. CADD scores were  
611 downloaded from the CADD website (<https://cadd.gs.washington.edu/download>). phyloP scores  
612 were downloaded from the UCSC browser (<https://genome.ucsc.edu/cgi-bin/hgTrackUi?db=hg38&g=cons241way>).  
613  
614

## 616 **Code availability**

617 Code to determine missense constraint regions (MCRs) and calculate MPC is available at  
618 [https://github.com/broadinstitute/regional\\_missense\\_constraint](https://github.com/broadinstitute/regional_missense_constraint). Code used to generate the  
619 mutational models is available at [https://github.com/broadinstitute/gnomad\\_lof](https://github.com/broadinstitute/gnomad_lof) and  
620 <https://github.com/broadinstitute/gnomad-constraint>. The Hail library is available at  
621 <https://hail.is/>.

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636

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639 **Contributions**

640 K.R.C., L.W., K.E.S., B.M.N, and M.J.D conceived and designed experiments. K.R.C, L.W., and  
641 K.E.S. performed primary writing of the manuscript. K.R.C., L.W., R.P., C.L., H.A., and R.Y.  
642 performed the analyses and generated figures. P.S. and J.C. were instrumental to developing  
643 methods. R.H.G., N.A.W., and M.S. developed visualizations for the web browser. B.W., W.P.,  
644 M.W.W., K.M.L., J.K.G, K.J.K, and G.T. completed code review for methods. D.G., J.I.G., C.V.,  
645 and T.P. helped debug runtime compute. J.A.K. provided data and analysis suggestions. S.B.  
646 contributed analysis suggestions. H.L.R., B.M.N., M.E.T, D.G.M, A.O.D.L., K.J.K., P.R., M.J.D.,  
647 and K.E.S. supervised the research. All authors listed under the Genome Aggregation Database  
648 Consortium contributed to the generation of the primary data incorporated into the gnomAD  
649 resource. All authors reviewed the manuscript.

650 **Ethics declarations**

651 Competing interests/Declaration of interests

652 J.A.K is a current employee of Regeneron Genetics Center. T.P. and G.T. are founders of E9  
653 Genomics, Inc.. H.L.R. has received support from Illumina and Microsoft to support rare disease  
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