

1 The genetic architecture and evolutionary consequences of the human pelvic 2 form

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21 Abstract

22 Human pelvic shape has undergone significant evolutionary change since the divergence
23 from the chimpanzee lineage. This transformation, involving the reduction of the pelvic canal
24 size to support bipedal locomotion, is thought to give rise to the obstetrical dilemma, a
25 hypothesis highlighting the mismatch between the large brain size of infants and the narrowed
26 birth canal in females. Empirical evidence for this classic hypothesis has been equivocal, largely
27 due to a lack of sample size and appropriate types of data. To elucidate the genetic underpinnings
28 of pelvic morphology, we applied a deep learning model to 31,115 dual-energy X-ray
29 absorptiometry (DXA) from the UK Biobank, extracting a set of seven pelvic proportion (PP)
30 phenotypes, including measures of the birth canal. All PPs were found to be highly heritable
31 (~25-40%) and a genome-wide association study of these traits identified 179 independent loci.
32 Unlike other skeletal proportions including long bone lengths, the subpubic angle associated with
33 the birth canal exhibits a genetic correlation between sexes significantly less than 1, in line with
34 sex-specific reproductive function. PPs were also left-right asymmetric but not heritable and
35 instead associated with handedness. We conducted phenotypic and genetic association analyses
36 to link PPs to 3 facets of the dilemma: locomotion, pelvic floor function and childbirth. Larger
37 birth canal phenotypes were associated with reduced walking pace, decreased risk of back pain,
38 and increased risk of hip osteoarthritis - phenotypes linked to locomotor efficiency. We also
39 observed that a narrower birth canal width was associated with a reduced risk of pelvic floor
40 disorders. When examining childbirth-related outcomes, narrower birth canal phenotypes were
41 associated with increased risk of emergency cesarean sections and obstructed labor due to
42 insufficient dilation, but not obstructed labor due to positioning of the fetus. Finally, we
43 examined whether the dilemma might have been alleviated through evolution. We found no
44 association between any PPs and gestational duration, contrary to the initial prediction by
45 Washburn in 1960. However, we found that the birth weight of the child, a proxy for skull and
brain size, was genetically correlated with birth canal width but not with other PPs. Collectively,
our study offers fresh insight on a 60-year-old debate in human evolutionary studies. Our results
support the idea that the obstetrical dilemma has played a central role in the co-evolution of the
human brain and pelvis, while also highlighting the potential role of associated factors such as
pelvic floor health.

46 Introduction

47

48 The human skeleton has undergone significant morphological change associated with the
49 transition to bipedalism. Some of the most significant changes occurred in the pelvis, resulting in
50 a superoinferiorly short and mediolaterally flaring pelvis relative to the modern great apes (1, 2).
51 These features are believed to have emerged early in hominin evolution and the alteration in
52 pelvic anatomy allowed for the positioning of the upper body above the lower limb joints and
53 facilitated the maintenance of an upright posture (3). While debate continues about the details of
54 gait mechanics in fossil hominins (1) it is clear that the modern human pelvis is adapted to
55 habitual bipedality, and undergoes a specific pelvic motion during walking that is thought to
56 reduce energetic costs associated with bipedal locomotion (4).

57

58 The suite of adaptations for bipedality includes a reduction of the bi-acetabular distance,
59 minimizing pelvic rotation during bipedal movement and consequently enhancing efficiency (5).
60 This narrowing of the bi-acetabular distance results in a narrower birth canal, and is thought to
61 stand in direct opposition to the birthing of children with significantly larger brains than our
62 evolutionary predecessors (e.g., (6–14)). In the 1960s, this functional and evolutionary conflict
63 was coined the “obstetrical dilemma” by Washburn (13). In the six decades since then, the
64 obstetrical dilemma has been a source of intense debate, and different studies have attempted to
65 examine the validity of the hypothesis through empirical data (6, 14–17). One area of contention
66 centers on the relationship between pelvic shape and walking efficiency or walking speed. Some
67 studies have found there is an association between the two (7, 18), while others have not (19–
68 22). Another point of debate revolves around whether differential birth canal proportions are
69 associated with obstruction during delivery (7–14, 17, 23–26). Recently, there has been growing
70 appreciation for the concept of a multifactorial pelvis, which proposes that the role of pelvic
71 width reduction is not just to enable bipedal locomotion, but also to reduce the risk of pelvic
72 floor disorders. Pelvic width reduction improves the pelvic floor's ability to support the fetus and
73 the inner organs, and to prevent incontinence (7, 27, 28).

74

75 In addition to debates about the association between pelvic morphology and locomotion,
76 childbirth, and pelvic floor function, it has been suggested that in modern humans the obstetrical
77 dilemma has been alleviated through evolution. Washburn's initial hypothesis proposed that
78 relative to the other great apes, humans experience a shorter gestation period. This enables
79 human infants to be born relatively earlier in development than their primate counterparts,
80 consequently limiting the extent of brain growth before birth and ensures that the newborn can
81 successfully traverse the birth canal during delivery. However, this hypothesis has been
82 challenged and updated in recent years, as human gestational length and newborn size have been
83 found to align with or exceed expectations for primates of our size, similar to the other great apes
84 (14, 29–31) (see (6, 7, 32) for alternate usages and historical perspectives on the term “obstetrical
85 dilemma”).

86

87 While different aspects of the dilemma have been tackled over the past few decades,
88 these previous studies suffer from several shortcomings. One issue with many studies –
89 particularly those involving clinical outcomes – is that measurements of pelvic dimensions were
90 collected externally (19, 22), which may not adequately reflect the skeletal constraints imposed,
91 particularly with respect to the birth canal. Another issue is that some earlier studies lack

92 complete information about individual lifetime health records and are unable to distinguish
93 between fine-grained but important details such as elective and emergency C-sections. However,
94 the major challenge contributing to the ongoing debate is the limited sample size in many of
95 these studies, which often only have data on a few hundred individuals (sample sizes and
96 references of previous papers are reported in **Table S1**). In addition, data obtained for each study
97 is often only capable of addressing one facet of the dilemma, as datasets examining childbirth
98 outcomes and pelvic morphology often do not include data about pelvic floor function or
99 walking speed/efficiency for the same individuals.

100
101 Finally, the underlying basis of skeletal evolution in the pelvis is genetic. While
102 functional genomic datasets examining gene expression through development as well as
103 comparative gene expression between the great apes and humans for the pelvis have yielded
104 valuable insights (33–35) study of the direct association between pelvic trait variation and
105 genetics has not yet been carried out. Thus, the genetic basis of pelvic morphology underlying
106 variation in humans or indeed any other vertebrate is largely unknown, precluding analysis of
107 natural selection on pelvic phenotypes directly at the genomic level.

108
109 Here, we applied methods in computer vision to derive a comprehensive set of seven
110 skeletal measurements of the human pelvis from full-body dual-energy X-ray absorptiometry
111 (DXA) images at biobank scale. We performed genome-wide scans on these seven phenotypes to
112 identify loci associated with variation in pelvic proportions (PPs). Using summary statistics from
113 these image-derived phenotypes (IDPs), we linked human PPs through phenotypic and genetic
114 correlation with other biobank phenotypes, with an emphasis on locomotor, pelvic floor and
115 childbirth-related outcomes.

116

117 **Results**

118

119 *A deep learning approach to measure pelvic morphology*

120 To study the genetic basis of the human pelvis, we jointly analyzed DXA and genetic
121 data from 42,284 individuals in the UK Biobank (UKB). Individuals from this dataset are
122 between 40 and 80 years old and reflect adult skeletal morphology. We report baseline
123 information about our analyzed cohort in **Table S2**. Using a previously published deep learning-
124 based image quality control (QC) pipeline (36), we retained only DXA images for the full body
125 which included the entire pelvis, and removed images which contained image artifacts, atypical
126 aspect ratios, and other abnormalities, retaining 39,469 images of high quality. These images
127 were then uniformly cropped and padded to focus on the pelvis for subsequent analysis
128 (**Methods: A deep learning model to identify pelvic landmarks on DXA scans**).

129 After performing image QC, we manually annotated 17 landmarks on 293 randomly
130 selected pelvic images (see **Fig. S2**) to train our model. To assess the accuracy of our manual
131 annotations, we re-annotated 20 images from the initial set of 293 and refined this annotation
132 through model-in-the-loop labeling (**Methods: Image quality control, Fig. 1B and 1C**). Our deep
133 learning model was based on a High-Resolution Network (HRNet) architecture chosen because it
134 maintains a high-resolution representation throughout the model which improves the
135 performance of landmarking for this task on benchmarking tasks. These methods were robustly

136 applied to a similar task of identifying joints on the overall skeleton (36, 37) (**Methods: Image**
137 *quality control*).

138 ***Validation of human pelvic phenotype estimates***

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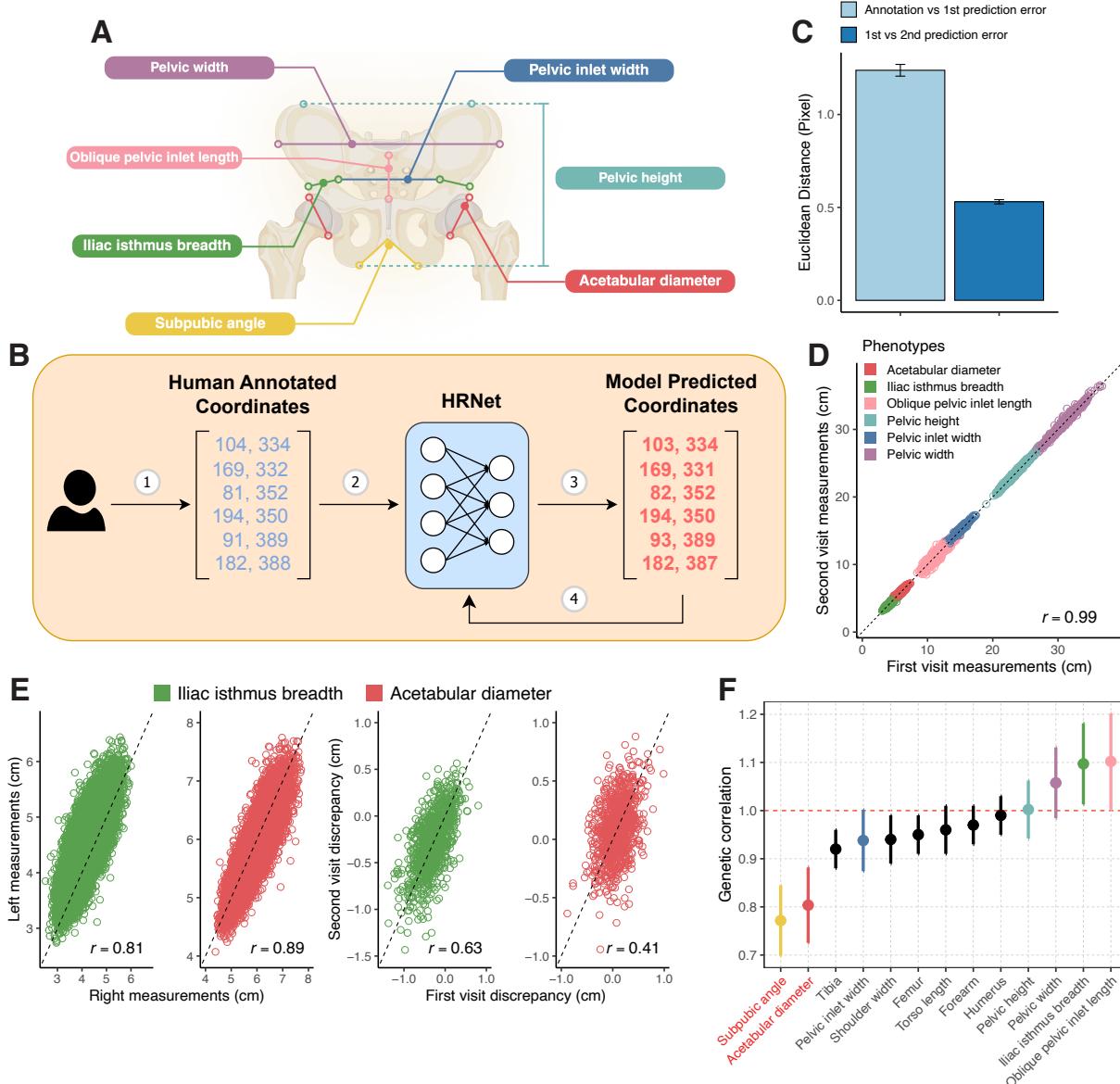
140 After training and validating the deep-learning model on the 297 manually annotated
141 images, we applied this model to predict the 17 landmarks on the rest of the 39,469 full-body
142 DXA images. We then calculated the pixel Euclidean distances between pairs of landmark
143 coordinates to ascertain six length phenotypes: pelvic width, pelvic inlet width, oblique pelvic
144 inlet width, iliac isthmus breadth, pelvic height and acetabular diameter, and one angle
145 phenotype: subpubic angle (**Fig. 1A**). To standardize images with varying aspect ratios, we
146 rescaled pixels into centimeters for each image resolution. This was achieved by regressing the
147 pixel height against the standing height in centimeters, as measured in the UK Biobank
148 assessments (**Methods: Image standardization, Fig. S11**). For all seven pelvic measurements, we
149 excluded individuals exceeding four standard deviations from the mean (**Methods: Removal of**
150 *image outliers, Fig. S11*).

151 Following outlier removal, we validated the accuracy of our measurements on the
152 remaining samples in two ways. First, we calculated the average error between labels in the
153 validation data and model performance: average error was 2 pixels across all 17 landmarks.
154 Second, we analyzed 935 individuals with repeat imaging visits at least two years apart. The
155 correlation of all pelvic length phenotypes between the first and second imaging visits was
156 greater than 0.99 (**Fig. 1D**). This indicates that the phenotype estimations via our deep learning
157 model are both accurate and highly replicable.

158 ***Human pelvic asymmetry is associated with handedness, and is not heritable***

159 Next, we examined the correlation between measurements on the left and the right side of
160 the pelvis. The two phenotypes with measures on each side were iliac isthmus breadth and
161 acetabular diameter. The left-right correlation for iliac isthmus breadth and acetabular diameter
162 were 0.809, and 0.894 respectively (**Fig. 1E**). The average difference between the measurements
163 in the iliac isthmus breadth between the left and right sides was 0.287 cm ($p < 2 \times 10^{-16}$, 95%
164 confidence interval (CI) = 0.294 to 0.280), and for acetabular diameter, it was 0.101 cm ($p < 2 \times$
165 10^{-16} , 95% CI = 0.093 to 0.108). Though these differences were small, we found that they were
166 replicable - left and right discrepancies in individuals across two imaging visits had Pearson
167 correlations of 0.633 and 0.407 for iliac isthmus breadth and acetabular diameter respectively
168 (**Fig. 1E**). This suggests that we can capture a measure of pelvic asymmetry beyond
169 measurement error. On estimating the heritability of this trait using GCTA (38) we found that it
170 was consistent with 0 (h_g^2 for acetabular diameter discrepancy = 0.0131, SE = 0.0149, h_g^2 for
171 iliac isthmus breadth discrepancy = 0.0275, SE = 0.0158). However, we observed a significant
172 association between pelvic asymmetry and handedness - another trait that is also not significantly
173 heritable (left-handed h_g^2 = 0.0104, right-handed h_g^2 = 0.0096 in 150,000 individuals). The
174 genetic correlation between acetabular diameter discrepancy and left-handedness is -0.39, and
175 with right-handedness, it is 0.34. Similarly, the genetic correlation between iliac isthmus breadth
176 discrepancy and left-handedness is 0.15, and with right-handedness, it is 0.11. In addition, we
177 regressed the left and right pelvic phenotype ratio against handedness while controlling for age

178 and sex. Right-handed individuals tended to have larger right acetabular diameters than left-
 179 handed individuals (regression $p = 8.31 \times 10^{-6}$) and larger left iliac isthmus breadth than left-
 180 handed individuals (regression $p = 0.0665$). This suggests that left-right pelvic asymmetry might
 181 be driven by left- or right-side dominance which is itself not heritable, but affect movement
 182 patterns and consequently skeletal development.



183
 184 **Fig. 1. Deep learning-based image quantification and validation** (A) Deep learning-based
 185 image landmark estimation using the HRNet architecture is shown. During this process, 293
 186 training images manually annotated with specific landmarks were used to train the model, which
 187 to perform automatic annotation of landmarks on the rest of images in the dataset from which
 188 pelvic measurements were calculated. (B) Model in the loop training data workflow. The
 189 coordinates from the 293 training images initially annotated by humans were used as a training
 190 set to train a model that was then redeployed on the training data. This helped to remove
 191 variation present in human labelling of the images and refined the training data itself. (C) Model
 192 in the loop training reduces annotation variability. Light blue bar indicates the average Euclidean

193 distances between human annotated landmarks and the model's first prediction on 58 validation
194 set images. The dark blue bar indicates the average Euclidean distance between first and second
195 model prediction on 58 validation set images. **(D)** Correlation of lengths measured from the first
196 and second imaging visits for the same individual. **(E)** The two panels on the left side show the
197 correlation between the left- and right-side measurements of the iliac isthmus breadth and
198 acetabular diameter. The two panels on the right side illustrate the correlation of the left-right
199 discrepancy in the iliac isthmus breadth and acetabular diameter between the first and second
200 imaging visit. **(F)** Genetic correlation between female and male pelvic phenotypes and other
201 skeletal traits including tibia, femur, torso length, forearm, and humerus. The error bars show 1
202 standard error. Heritability greater than 1 is due to small sample size. The two traits shown in red
203 on the x-axis are the only ones that are significantly different from one.

204

205 ***Sexual dimorphism in the genetic basis of PPs***

206 The human pelvis plays a critical role in childbirth and is one of the most dimorphic
207 skeletal elements between males and females (39–41). Given the distinct functionalities between
208 male and female pelvis, we examined whether the genetic basis of our seven pelvic phenotypes
209 differed between males and females. To do so, we carried out genetic correlation analysis
210 between a GWAS carried out in males versus females. Functionally similar pelvic phenotypes,
211 such as pelvic height, exhibit similar genetic architectures between males and females, with a
212 genetic correlation of 1.03. In contrast, birth canal-related phenotypes like the subpubic angle
213 showed genetic correlations significantly divergent from 1. This difference in genetic correlation
214 is in striking contrast to virtually all other skeletal traits previously examined such as arm, leg,
215 torso, and shoulder dimensions. These other traits all showed genetic correlations not
216 significantly different from 1 in the same cohort (**Fig. 1F**), suggesting that sex-specific
217 reproductive requirements of the human birth canal are driving genetic differences between sexes
218 for these PP traits.

219 ***GWAS of human PPs***

220

221 We performed GWASs using imputed genotype data in the UKB to identify variants
222 associated with each pelvic phenotype. We applied standard variant and sample QC and focused
223 our analyses on 31,115 individuals of “white British ancestry,” as defined by the UKB genetic
224 assessment, and 7.4 million common biallelic single-nucleotide polymorphisms (SNPs) with
225 minor allele frequency >1%. We used BOLT-LMM (42) to regress variants on each skeletal
226 measure using a linear mixed-model association framework. We included height as a covariate to
227 directly adjust for differences in body size between individuals and focus on skeletal proportions
228 instead of overall length. We also adjusted for body size differences in two other ways: dividing
229 each phenotype by height to generate a skeletal proportion, and including a leave one-
230 chromosome-out polygenic risk score (PRS) for height as a covariate in the GWAS (43). GWAS
231 effect sizes using either height as a covariate or height combined with the one-chromosome-out
232 PRS as a covariate were highly correlated (Pearson correlation = 0.99) (**Methods: Adjusting for**
233 ***height correlation in GWAS by adding height as covariate, Fig. S13***). For downstream analyses,
234 we focused on the results from the GWAS that included height as a covariate. Notably, we show
235 that the effect sizes estimated for our PP phenotypes were uncorrelated with effect sizes from
236 height (**Methods: Adjusting for height correlation in GWAS by adding height as covariate, Fig.**

237 **S14**, average Pearson correlation across all phenotypes = 5.67×10^{-5} , standard deviation =
238 0.0097), suggesting that PPs and height are distinct traits.

239

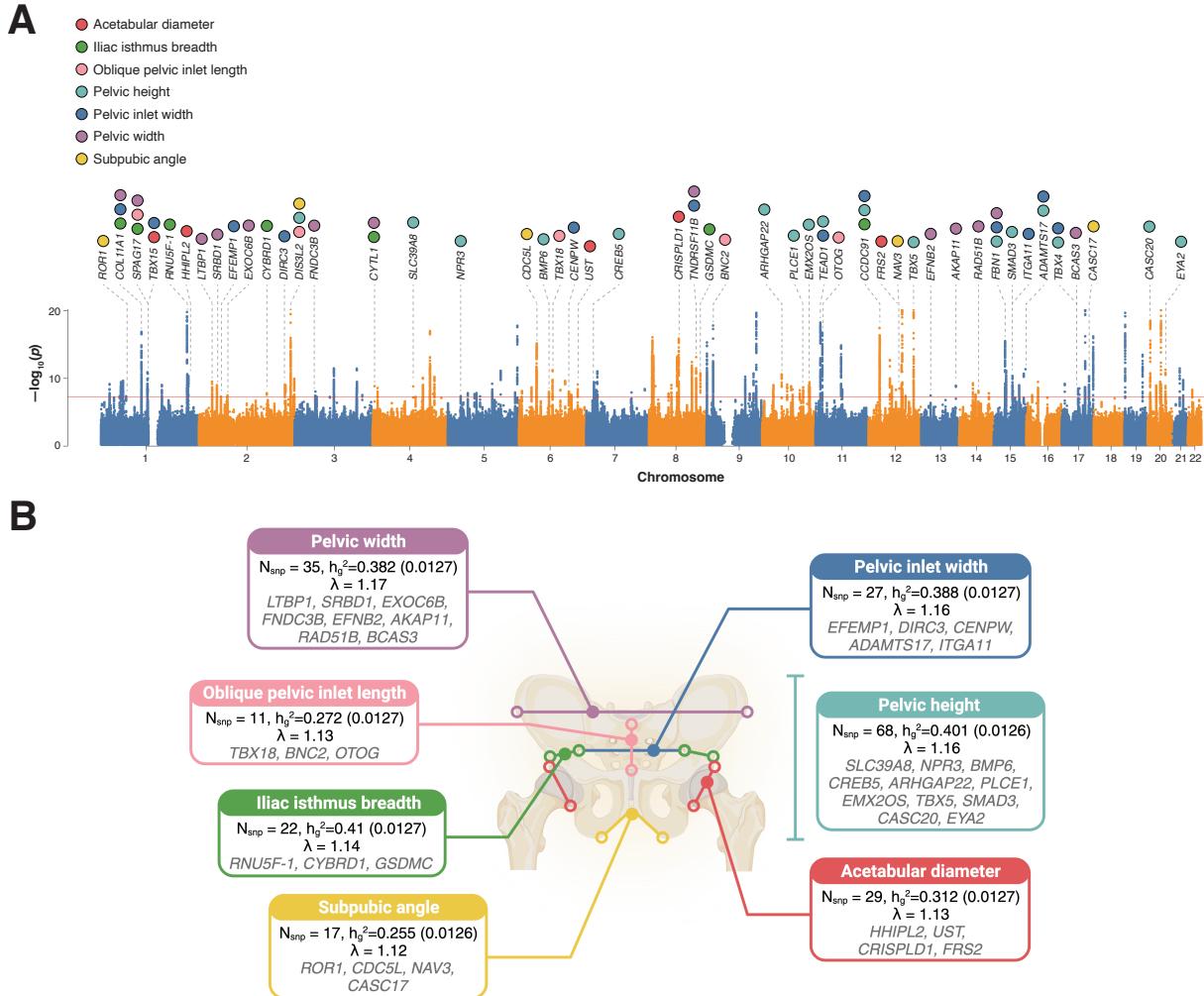
240 After generating summary statistics for each skeletal measure, we estimated SNP
241 heritability using LD Score regression (LDSC) (44) and GCTA-REML (38). All traits were
242 highly heritable, with SNP heritability between 25% and 40% for LDSC and between 17% and
243 50% for GCTA-REML (**Methods: GWAS and Heritability analysis, Fig. 2B, Fig. S15**). Across
244 the six pelvic phenotypes adjusted by height (pelvic width, pelvic height, iliac isthmus breath,
245 acetabular diameter, pelvic inlet width, oblique pelvic inlet length) and subpubic angle, we
246 identified 339 loci at $p < 5 \times 10^{-8}$ and 241 loci at $p < 7.14 \times 10^{-9}$ (Bonferroni correction for
247 seven traits). Of these loci, 179 are independently significant at $p < 5 \times 10^{-8}$ (linkage
248 disequilibrium (r^2) < 0.1) across all seven phenotypes (119 after Bonferroni correction for seven
249 traits at $p < 7.14 \times 10^{-9}$) (**Fig. 2A**).

250

251 *Biological insights from pelvic associations*

252

253 Out of the 179 independent loci identified across GWASs (**Table S11**), 50 loci
254 overlapped a single protein-coding gene within each clumped region (**Fig. 2B**). Notably, of these
255 50 genes, 22 (or 44%) resulted in abnormal skeletal phenotypes when disrupted in mice using the
256 Human-Mouse Disease Connection database (36). Eight genes (*COL11A1*, *NPR3*, *CDC5L*,
257 *TNFRSF11B*, *TBX5*, *FBNI*, *SMAD3*, and *TBX4*) were associated with rare skeletal diseases in
258 humans (**Table S11**). In some cases, genes associated with specific PPs in our GWAS contribute
259 to human pelvic abnormalities. We found that *TBX15* and *TBX4*, two T-box transcription factors,
260 have been associated with differences in pelvic inlet width and pelvic height in model organisms,
261 and mutations in both the *TBX15* and *TBX4* genes lead to pelvic abnormalities such as
262 hypoplasia of the pelvis and small patella syndrome (45, 46). Thus, our GWAS of PPs identifies
263 genes that were previously associated with skeletal developmental biology and Mendelian
264 skeletal phenotypes and demonstrates the potential for future functional and knockout studies.



265
266 **Fig. 2. Genome-wide association results.** (A) Manhattan plot of a GWAS performed across six
267 PPs and subpubic angle; the lowest p for any trait at each SNP is annotated. Loci over the
268 genome-wide significance threshold that are close to only a single gene are annotated. (B)
269 Shown are the total number of genome-wide significant loci per trait, heritability (GCTA-
270 REML), λ (from LDSC), and associated genes of loci that are specific to each skeletal trait
271 (again annotating only loci that map to a region with a protein-coding gene within 1000 kb of
272 each clumped region).

273 **Genetic and phenotypic association of PPs with locomotor phenotypes**

274
275 We examined how PPs were associated with walking pace, and musculoskeletal disorders
276 such as knee, hip, and back osteoarthritis (OA), which are degenerative conditions that arise
277 from lifetime cumulative effects of gait and motion. First, we used logistic regression to examine
278 phenotypic associations between PPs and these phenotypes (Fig. 3A) while controlling for age,
279 sex, weight, height, and other major risk factors for OA (**Methods: Polygenic risk score (PRS)**
280 *association of skeletal phenotypes with musculoskeletal disease*). After correcting for multiple
281 testing at an FDR < 5% across all associations, we found that one standard deviation in two birth
282 canal-related phenotypes was associated with increased self-reported walking pace (oblique

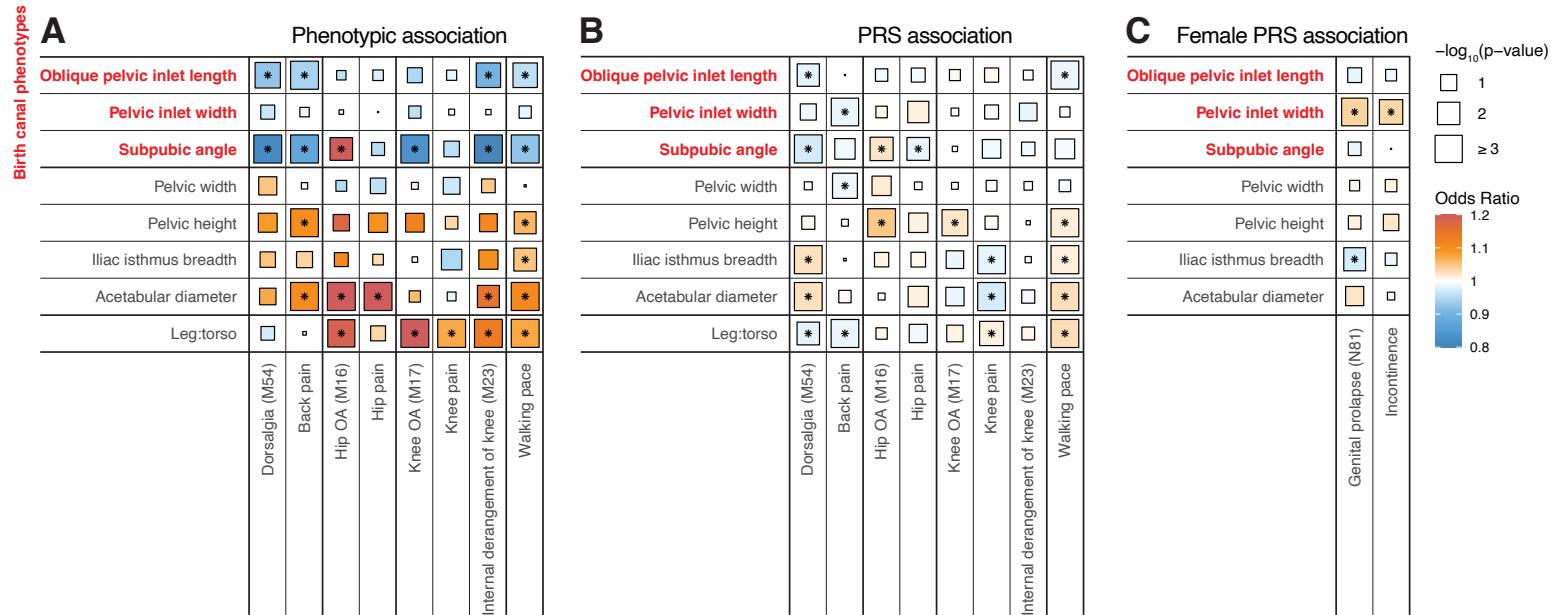
283 pelvic inlet length: $p = 5.3 \times 10^{-3}$, odds ratio (OR) = 0.96; subpubic angle: $p = 4.4 \times 10^{-4}$, OR =
284 0.92) (**Table S15**). As a positive control, we examined another skeletal trait, leg-to-torso length,
285 which we found to be significantly positively associated with walking speed ($p = 2.97 \times 10^{-8}$, OR
286 = 1.08), in line with previous results and with mechanical modeling (6, 7). These results provide
287 empirical evidence that narrower birth canal proportions in humans are associated with increased
288 walking speed (phenotypic association: between oblique pelvic inlet length and walking pace: p
289 = 5.31×10^{-3} , OR = 0.96, between subpubic angle and walking pace: $p = 4.89 \times 10^{-4}$, OR = 0.92).
290 However, examining the associations with OA-related phenotypes we found that having larger
291 birth canal-related phenotypes also decreased the risk of back pain/dorsalgia (phenotypic
292 association: between oblique pelvic inlet length and dorsalgia: $p = 3.45 \times 10^{-3}$, OR = 0.93,
293 between subpubic angle and dorsalgia: $p = 3.28 \times 10^{-7}$, OR = 0.82, between subpubic angle and
294 back pain: $p = 5.16 \times 10^{-7}$, OR = 0.87) (**Fig. 3A, Table S15**). We also found that individuals with
295 larger birth canal phenotypes were also at increased risk of hip osteoarthritis (phenotypic
296 association: between subpubic angle and hip OA: $p = 1.18 \times 10^{-2}$, OR = 1.27) but reduced risk of
297 knee osteoarthritis (phenotypic association: between subpubic angle and knee OA: $p = 9.97 \times 10^{-4}$,
298 OR = 0.83, between subpubic angle and internal derangement of knee: $p = 9.71 \times 10^{-5}$, OR =
299 0.81) (**Fig. 3A, Table S15**).

300 To complement these phenotypic associations, we also analyzed 361,140 UKB
301 participants who had not undergone DXA imaging and were of “white British ancestry” for
302 predictive risk based on PRS derived from our GWAS on PPs for the imaged set of individuals
303 (**Fig. 3B, Table S16**). We generated PRS with Bayesian regression and continuous shrinkage
304 priors (47) using the significantly associated SNPs and ran a logistic regression of the generated
305 risk scores and traits, adjusting for the first 20 principal components of ancestry and imputed sex
306 as well as age, sex, weight and other major risk factors of OA (**Methods: Polygenic risk score**
307 (**PRS**) *association of skeletal phenotypes with musculoskeletal disease*). Our genetic association
308 analysis mirrored our phenotype association analysis and suggests that individuals with smaller
309 birth canal proportions have on average a faster walking pace, but are at the same time more
310 susceptible to back pain and strain, common consequences of bipedal locomotion due to the
311 distribution of weight on just two limbs (genetic association between leg to torso ratio and
312 walking pace: $p = 1.00 \times 10^{-13}$, OR = 1.03, between oblique pelvic inlet length and walking pace:
313 $p = 8.09 \times 10^{-4}$, OR = 0.98, between oblique pelvic inlet length and dorsalgia: $p = 1.31 \times 10^{-2}$,
314 OR = 0.98, between pelvic inlet width and back pain: $p = 1.25 \times 10^{-3}$, OR = 0.98, between
315 subpubic angle and dorsalgia: $p = 1.02 \times 10^{-4}$, OR = 0.97) (**Fig. 3B, Table S15**).

316 ***Genetic and phenotypic association of PPs with pelvic floor function***

317 Next, we combined all incontinence-related phenotypes from the ICD10 record, including
318 stress incontinence (N39.3), other specified urinary incontinence (N39.4), fecal incontinence
319 (R15), and unspecified urinary incontinence (R32), into a single binary phenotype. We
320 conducted a GWAS restricted to female individuals who were imaged and computed a PRS for
321 approximately 200,000 females of “white British ancestry” who were independent from the
322 GWAS set. 18,020 individuals out of the 200,000 individuals had one of these incontinence
323 phenotypes. We then regressed binary incidence of genital prolapse and incontinence against
324 PRS for all female pelvic traits, controlling for the number of live births and age (**Methods:**
325 *Polygenic risk score (PRS) association of skeletal phenotypes with musculoskeletal disease*). Out

326 of the various PPs the only significantly positive association we observed was with the width of
 327 the birth canal (between pelvic inlet with and genital prolapse: $p = 4.3 \times 10^{-4}$, OR = 1.04,
 328 between pelvic inlet with and incontinence: $p = 4.2 \times 10^{-3}$, OR = 1.03) (Fig. 3C, Table S17).
 329 These results offer support for the multifactorial pelvic hypothesis, suggesting that a narrower
 330 birth canal improves pelvic floor function. Pelvic floor function is critical in assisting bladder
 331 and bowel control and evacuation as well as in supporting the fetus during pregnancy - a function
 332 thought to be more critical in upright humans than in quadrupeds (7).



333 **Fig. 3. Association between pelvic traits, pain phenotypes, musculoskeletal diseases as well**
 334 **as walking pace. (A) Phenotypic associations from logistic regression analyses of**
 335 **musculoskeletal disease traits, self-reported pain and walking pace on PPs. (B) PRS associations**
 336 **between musculoskeletal disease traits, walking pace and PPs. (C) PRS associations between**
 337 **pelvic floor disorders and PPs. For (A), (B) and (C), associations that are significant after False**
 338 **Discovery Rate (FDR) correction are annotated with an asterisk (*). ORs for the phenotypic**
 339 **associations and PRS are shown in colors, and the p-values are represented by size. The number**
 340 **notations in parentheses are the ICD-10 codes associated with each disease: M54–Dorsalgia,**
 341 **M16–Coxarthrosis (arthrosis of hip), M17–Gonarthrosis (arthrosis of knee), M23–Internal**
 342 **derangement of knee, N81–Female genital prolapse.**

344 **Genetic association of PPs with childbirth-related outcomes**

345 Finally, we examined outcomes associated with obstructed labor, which is thought to be
 346 more common in humans than any other modern primate species (9). Obstructed labor affects
 347 around sixteen percent of deliveries today and has been a major cause of maternal and fetal death
 348 throughout human history, which suggests it might play a major role in human evolution through
 349 natural selection (48, 49). First, we focused on cesarean sections (C-sections) reported in the UK
 350 Biobank. To avoid confounding effects due to elective C-sections, we focused on emergency C-
 351 sections which are routinely performed in cases of obstruction. We conducted PRS association
 352 analysis and found that narrower birth canals were associated with increased risk of emergency
 353 C-sections ($p = 0.0108$, OR = 0.92). As childbirth-related outcomes were available only for a
 354

355 small portion of individuals in the UK Biobank (<10% of all individuals) we also examined
356 outcomes in the FinnGen's data set for delivery-related traits. We identified a significant genetic
357 correlation between birth canal traits and labor obstructions due to maternal pelvic abnormalities,
358 of which a major component is dilation width (**Methods: Genetic correlation of skeletal**
359 *proportions with pregnancy phenotypes, Fig. 4B, Table S18*). However, we saw no association
360 between obstructed labor due to malpresentation of the fetus and pelvic traits (**Fig. 4B, Table**
361 **S18**). As the position of the fetus can vary independently of the skeletal structure of the pelvis,
362 this childbirth outcome serves as a negative control for this analysis. Combining both types of
363 analysis, our results suggest strong associations between the size of the birth canal and the
364 chance of obstruction during labor.

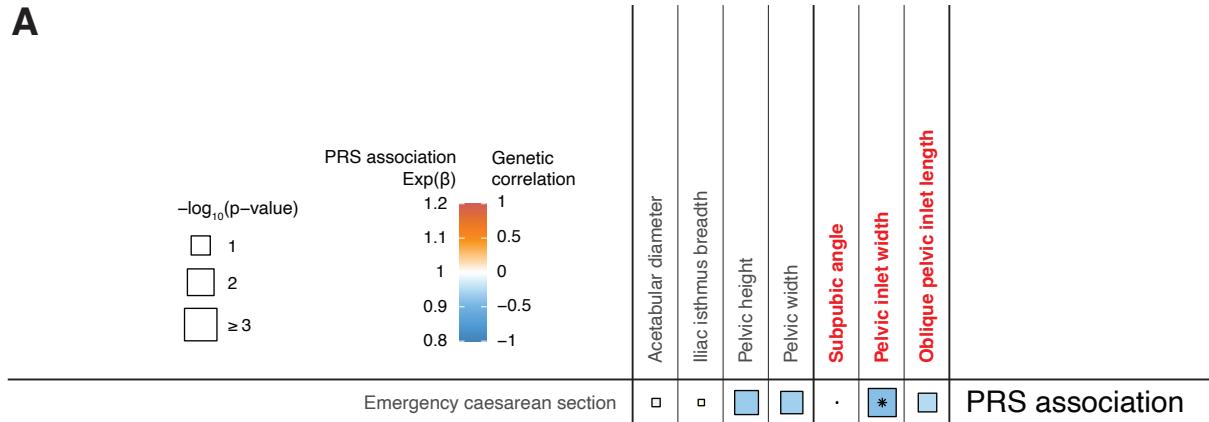
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366 ***Evolutionary escape***

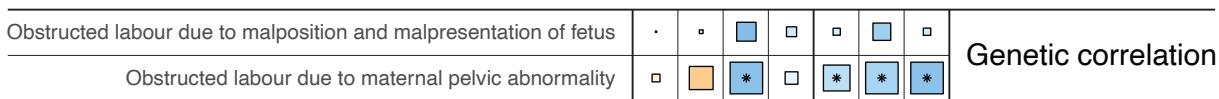
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368 Finally, we investigated associations that might help explain how the obstetrical dilemma
369 may have been alleviated in recent human evolution. First, we examined whether gestation
370 length in humans is shorter than other primates of comparable body size, following from
371 Washburn's proposed that the relatively large-brained human infant must be delivered before its
372 head reaches a volume that cannot pass through the pelvic canal (13, 50, 51). However, we found
373 no association between gestational duration and any PP, including those associated with the birth
374 canal (**Fig. 4C, Table S17**). This result is in line with more recent data on a fairly large dataset of
375 great apes suggesting that human children are not born significantly earlier than those of the
376 other apes (14, 29–31). However, we did see a significant correlation between the proportional
377 width of the birth canal and neonatal birth weight - a proxy for neonatal head size (correlation
378 coefficient ~ 0.7 (52)) (**Fig. 4C, Table S17**). This suggests that natural selection might have led
379 to genetic correlation between pelvic and head proportions, potentially reducing labor
380 obstruction (41, 53).

A



B



C



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Fig. 4. Association between pelvic traits and childbirth-related outcomes. (A) PRS associations between pelvic traits and emergency cesarean section. (B) Genetic correlations between pelvic traits and obstructed labor, including obstructed labor due to malposition and malpresentation of fetus and obstructed labor due to maternal pelvic abnormality. (C) PRS associations between pelvic traits and evolutionary escape variables, including child birth weight and gestation duration. For (A), (B), and (C), associations that are significant after FDR correction are marked with an asterisk. Odds Ratios (ORs) for the PRS associations and genetic correlations are presented in various colors, and the p-values are indicated by size.

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Discussion

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In this study, we used deep learning to understand the genetic basis of skeletal elements that make up human PPs using DXA imaging data in a large population-based biobank. We found sex-specific differences in genetic architecture as well as differences in pelvic symmetry that were associated with handedness. We identified 179 independent genetic loci associated with PPs. We then examined different facets of the obstetrical dilemma, namely the relationship between PPs and locomotor, pelvic floor and childbirth-related outcomes. Lastly, we analyzed possible ways in which evolution and natural selection might have alleviated the dilemma by looking at genetic correlations between gestational period and child birth weight and pelvic proportions.

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In previous work on the obstetrical dilemma, studies have examined locomotor outcomes that are associated with efficiency and energy use rather than speed. Here we did not have access to energetics, but we did have access to outcomes associated with walking speed and OA which

406 relate to gait efficiency accumulated over a lifetime. While self-reported walking pace may not
407 seem an ideal measure of walking speed, several lines of evidence suggest that it is a reliable
408 measure of actual walking pace. First, self-reported walking pace is highly heritable (54).
409 Second, it is associated with muscle strength and declines with BMI and age in line with
410 expectations (55, 56). It is also associated with several disorders that are known to hinder
411 locomotion, including hip osteoarthritis, the leading cause of adult disability in the United States
412 (57–59). Finally, self-reported walking pace has been directly correlated with measured walking
413 pace in a reasonable sample size study and within the Biobank (60, 61). It has also been
414 correlated with mean accelerometer assessed activity (62).
415

416 Our results on locomotion were heavily mixed, with larger birth canal phenotypes related
417 to lower walking speed, reduced risk of back pain and knee OA, but increased risk of hip OA.
418 However, our results provide significant evidence for other facets of the dilemma associated with
419 pelvic floor function and childbirth. Specifically, we show that larger birth canal phenotypes are
420 associated with increased risk of pelvic floor disorders, but at the same time reduced risk of
421 obstruction during labor - two phenotypes that have direct impacts on human evolution due to
422 intense natural selection acting on them. We also investigated several leading hypotheses about
423 how the dilemma could have been alleviated over evolutionary time. Our data does not provide
424 support for the idea that gestational duration has decreased to accommodate birthing large-
425 brained infants - we observed no correlation between any PPs and gestational duration. However,
426 our results indicate that there is a genetic correlation between PPs only related to birth canal
427 width and child head size (which we obtain using birth weight as a proxy phenotype (52)).
428 Across all the skeletal traits we examined, the significantly reduced genetic correlation observed
429 between males and females exclusively for birth canal phenotypes also suggests that sexual
430 dimorphism in these traits may have arisen through natural selection in response to different
431 functional constraints.
432

433 Beyond increasing the sample size by multiple orders of magnitude relative to previous
434 studies that have examined this hypothesis, and presenting high quality measurement of the
435 human pelvic form, our work is also one of the few studies to integrate data from locomotor,
436 childbirth and pelvic floor outcomes all on the same participants. However, a limitation of our
437 study is that we only had individuals aged between 40–80 years old. It has been suggested that
438 age is a source of variation in PPs, and that changes in functional constraints throughout parts of
439 the reproductive lifespan is another means by which the dilemma could be alleviated (63).
440 However, we did not have access to data from individuals from earlier ages to examine this
441 hypothesis.
442

443 Taken together, our work combines, imaging, genetic, health record and survey data on
444 biobank scale data to re-examine a 60 year old theory of human evolution that is standardly
445 taught in textbooks. Our results provide major empirical support for several classical theories
446 of obstetrical dilemma related to locomotion and childbirth, but perhaps for the first time provides
447 evidence for the role of associated factors such as pelvic floor health.
448

449 Materials and Methods

450

451 UKB participants and dataset

452 All analyses were conducted with data from the UKB unless otherwise stated. The UKB
453 is a richly phenotyped, prospective, population-based cohort that recruited 500,000 individuals
454 aged 40–69 in the UK via mailer from 2006 to 2010 (64). In total, we analyzed 487,283
455 participants with genetic data who had not withdrawn consent as of May 4, 2022, out of which
456 42,284 had available DXA imaging data. Access was provided under application number 65439.
457 The baseline participants metadata including age and sex and other variables related to our study
458 are in **Table S2**.

459

460 Dual-energy X-ray absorptiometry (DXA) imaging

461 The UKB has released DXA imaging data for a total of 50,000 participants as part of
462 bulk data field ID (FID) 20158. The DXA images were collected using an iDXA instrument
463 (GE-Lunar, Madison, WI). A series of 8 images were taken for each patient: two whole body
464 images - one of the skeleton and one of the adipose tissue, the lumbar spine, the lateral spine
465 from L4 to T4, each knee, and each hip. Dual-energy X-ray absorptiometry (DXA) images were
466 downloaded from the UKB bulk data FID 20158. The bulk download resulted in 42,284 zip files,
467 each corresponding to a specific patient identifier otherwise known as each patient's EID, and
468 each file contained several DXA images of the patient as described above. All images were
469 exported and stored as DICOM files which were later converted to high-resolution JPEG files for
470 image analysis and quantification.

471

472 Phenotype and clinical data acquisition

473 Self-reported usual walking pace was obtained from UK Biobank under FID 924, and we
474 combined slow pace and steady average pace to increase the sample size in that category. The
475 binary classification of patient disease phenotypes was obtained from a combination of primary
476 and secondary ICD-10 codes (FID 41270) and the non-cancer self-assessment (FID 20002). Self-
477 assessment codes were translated to three-character ICD-10 codes (Coding 609) and ICD-10
478 codes were truncated to only the initial three characters. Patients received one if a disease code
479 appeared in either self-assessment visit or their hospital records and zero otherwise. The
480 phenotypes related to pelvic floor disorders are derived from ICD-10 codes, specifically
481 including incontinence (stress incontinence ICD-10 code: N39.3, other specified urinary
482 incontinence ICD-10 code: N39.4, fecal incontinence ICD-10 code: R15, and unspecified urinary
483 incontinence ICD-10 code: R32) and genital prolapse (ICD-10 code: N81). As the incontinence
484 phenotypes are very specific and each category of incontinence has only limited data, we
485 combined all incontinence phenotype categories into a single binary phenotype. **Table S12** and
486 **Table S13** contain all ICD-10 and FID codes we used in our analysis.

487

488 Computing infrastructure

489 All analysis was carried out on the Corral and Frontera system of the Texas Advanced
490 Computing Cluster. The deep learning analysis was carried out on NVIDIA Quadro RTX 5000
491 GPUs using the CUDA version 11.1 toolkit.

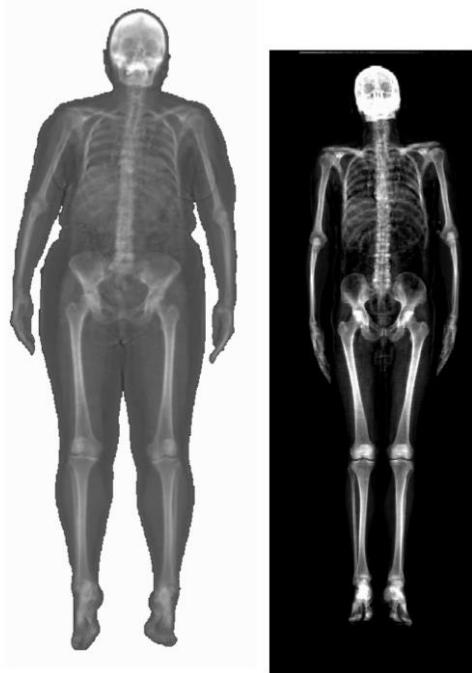
492

493 Image quality control

494 Each individual in the UKBiobank had a DXA image folder containing up to 8 different
495 body parts. In order to check the labels of these body parts that were defined using their file
496 name, we built a convolutional neural network (CNN) to sort the images by body part through
497 the use of a multiclass classification model using a previously published protocol (36). After
498 sorting and removal of images, we were left with 42,228 full skeleton X-rays (**Table S3**). After
499 we determined the final set of full body X-ray images, we performed additional quality control to
500 remove images that were poorly cropped and had other artifacts. To do this we utilized another
501 deep learning classifier also described in (36). Removal of all the cropped images resulted in a
502 total of 39,644 full-body images that we used for analysis (**Table S3**).
503

504 **Image standardization**

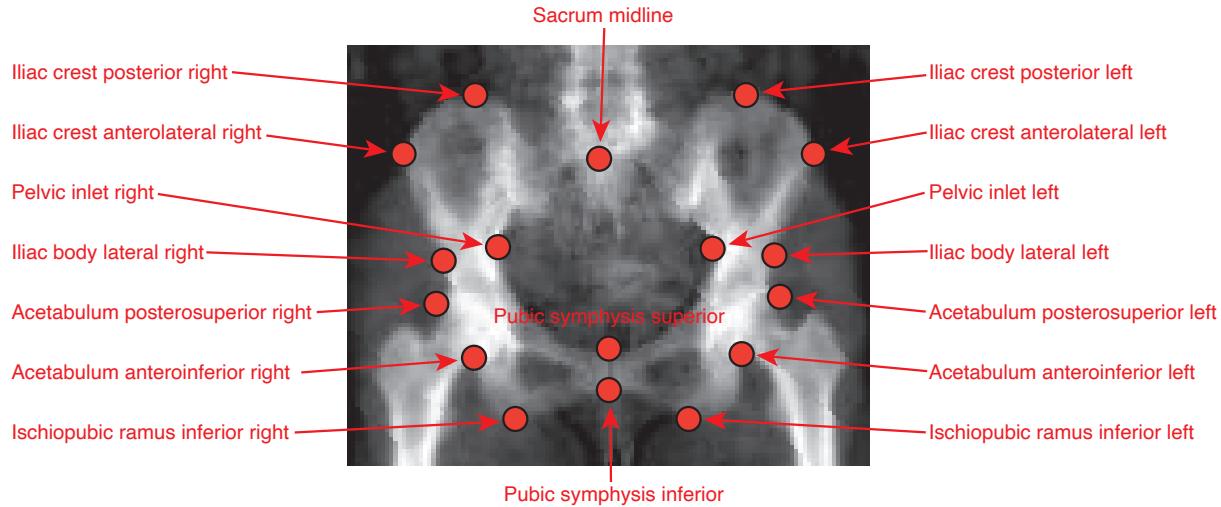
505 From the pool of remaining full-body X-ray images, we discovered that the images varied
506 in both pixel dimension and background. Broadly, the images fell into two main categories: (a)
507 images that were on a black background with sizes between 600-800 by 270 pixels and (b)
508 images on a white background with sizes between 930-945 by 300-370 pixels. The overall
509 distribution of images by pixel ratio and an example of each type of image are shown in **Table**
510 **S7** and **Fig. S1**. To process these images and remove the effects of scaling and resolution change
511 during the deep learning process, we chose to pad all the images to be of consistent size. We
512 removed images that had sizes far out of the normal range and processed each of the two
513 categories of images separately. The black background images were padded equally on all sides
514 of the image to a final resolution size of 816×288 pixels while the white background images
515 were padded in the same fashion to a resolution size of 960×384 pixels. We carried this out by
516 converting each individual DICOM file obtained from the UKB into numpy arrays and added
517 additional rows and columns of black or white pixels as appropriate using standard functions
518 from numpy (65), scipy (66), and skimage (67). These final resolution sizes were chosen based
519 on image size requirements for our deep learning model for landmarking and image
520 quantification. Padding and removing individuals with sizes that did not fit into the two major
521 categories resulted in a final total of 39,469 images - 21,981 images of 816×288 and 17,488
522 images of 960×384 . In our deep learning model for landmarking, we trained using images
523 across both pixel ratios. Despite variations in size and background, the images shared many
524 features, being skeleton X-ray images. Using both pixel ratios enriched the training set,
525 enhancing prediction accuracy.



526
527 **Fig. S1. Types of DXA images acquired from the UKB.** (Left) Image of patient imaged on
528 white background. (Right) Image of patient imaged on black background. Relative sizes of
529 images are true to scale.

530
531 **Manual annotation of human pelvic landmarks**

532 To train our deep learning model, we manually annotated a total of 293 images (with 146
533 images padded to 960×384 pixels on a white background, and 147 images padded to 816×288
534 pixels on a black background). Of these, 239 images were randomly allocated for training and
535 the rest are used for validation. Out of the 293 total images, 10 images were duplicated in each of
536 the image sizes to measure the replicability of our process. We used a single human annotator for
537 all training data and provided an initial dataset of 313 ($293+2 \times 10$ duplicate images) without the
538 annotator's knowledge. We used a standard annotation scheme in computer vision, the Common
539 Objects in Common (COCO) (68) scheme which provides a rubric for joint landmark estimation
540 on the human body. The positions in the pelvis we chose to annotate were the: iliac crest
541 posterior left/right, iliac crest anterolateral left/right, iliac body lateral left/right, pelvic inlet
542 left/right, acetabulum posterosuperior left/right, acetabulum anteroinferior left/right ischiopubic
543 ramus inferior left/right, pubic tubercle, and pubic symphysis inferior. An example of the
544 annotation of one image is shown in **Fig. S2** with landmarks placed at each of the locations listed
545 above.



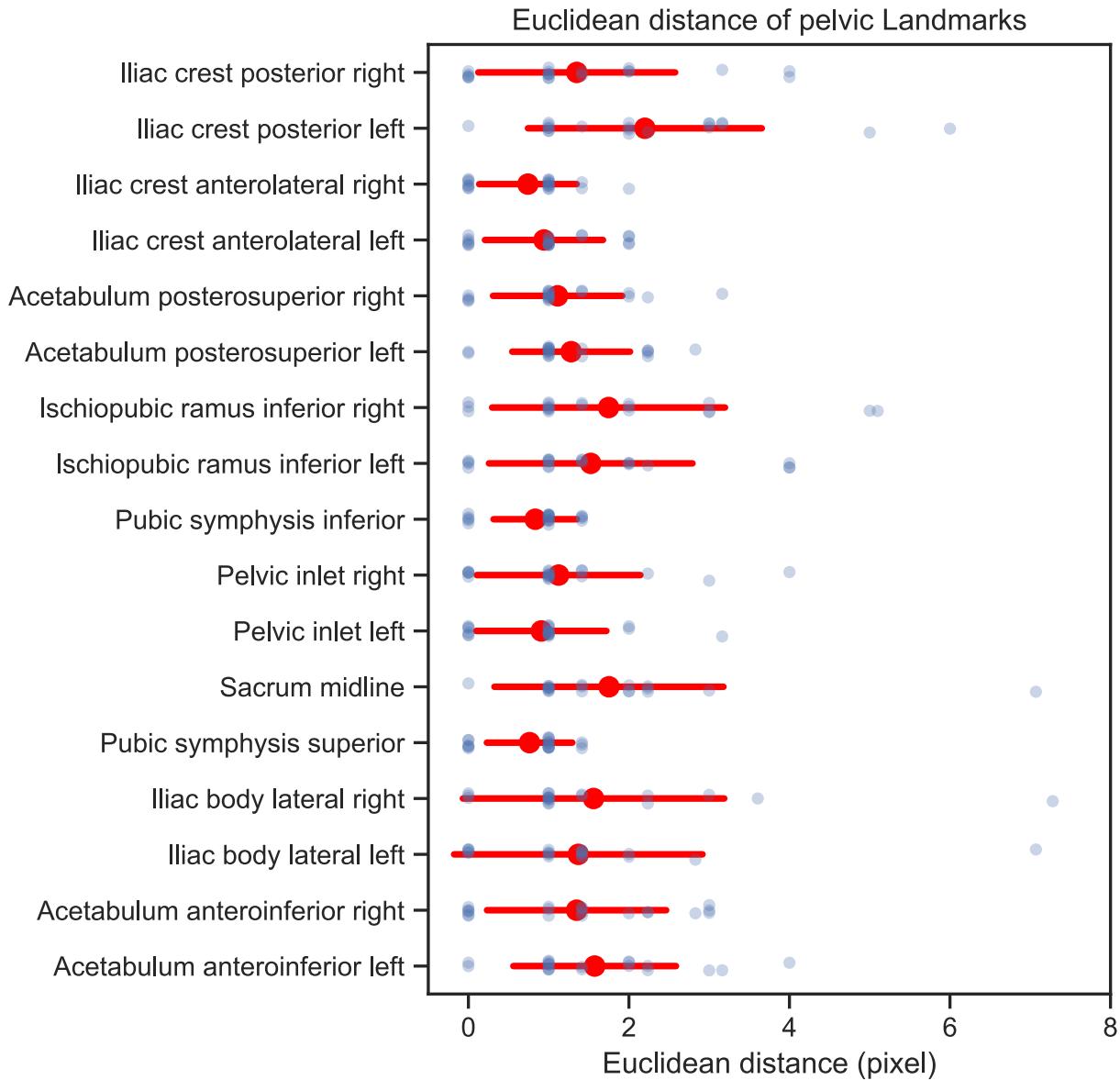
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Fig. S2. The 17 Pelvic landmarks and their corresponding names.

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Replicability Assessment

We measured the replicability of our annotations by taking the Euclidean distance of pixels between the corresponding key points across 10 randomly selected images that were duplicated amongst both of the 816×288 and 960×384 image set without knowledge of whether the image was a duplicate. Our replication analysis of 20 duplicate images was under 8 pixels across the different points that were estimated. Across the body parts, the farthest deviation across annotations was seen in the iliac body lateral, but the mean replicability across 20 images was under 2 pixels for all of the pelvic landmarks (Fig. S3).



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Fig. S3. Annotation Error in Human Pelvic Landmarks. The blue points depict the Euclidean distances of specific landmarks between replicate annotations from 10 images of the 816×288 set and 10 images of the 960×384 set. The red points indicate the mean Euclidean distance for each landmark, while the red error bars represent the standard deviation for these distances.

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A deep learning model to identify pelvic landmarks on DXA scans

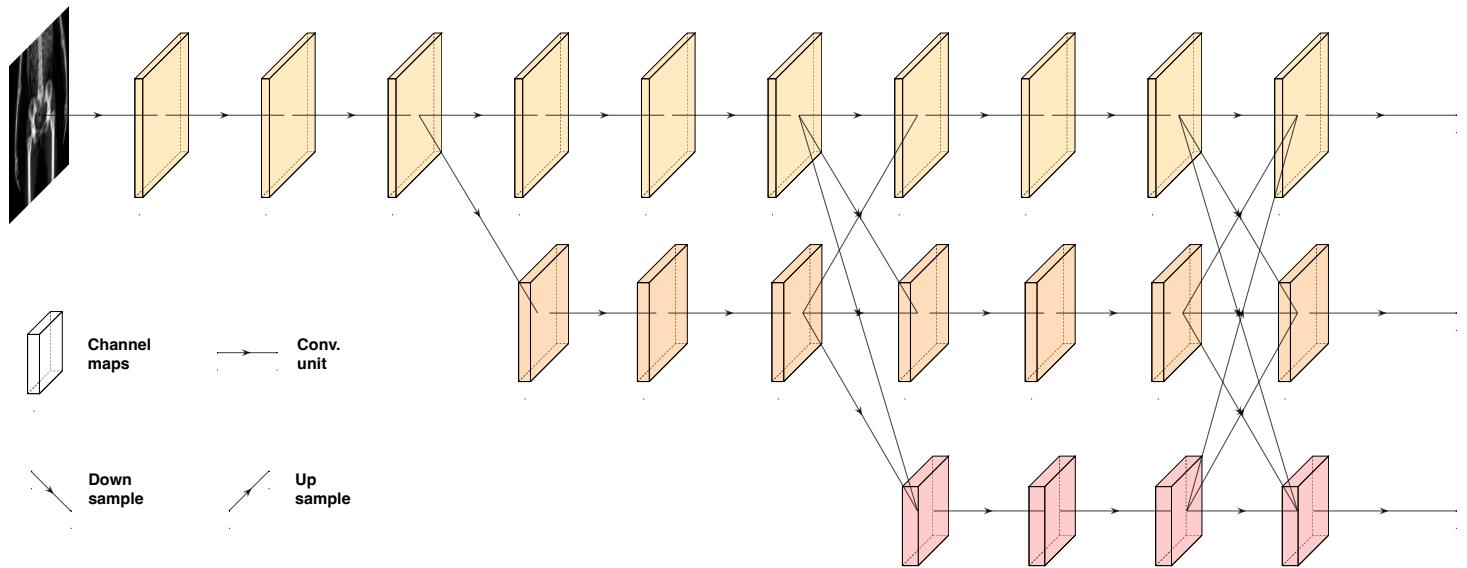
To determine the coordinates of all landmarks across 39,469 images, we initially adjusted the DXA images of both sizes, 816×288 and 960×384 , to a uniform dimension of 256×256 pixels for the pelvis area by applying central cropping and padding. For the upper and lower body sections, the images were resized to 608×608 pixels. Subsequently, 235 images (approximately 80% of the total) were allocated to the training set, with the remaining 58 images set aside for validation. To enhance the training set and to improve the model's ability to generalize, we applied image transformations including rotation and warping.

571 To perform landmarking we used an HRNet architecture based network (69) preserves
572 high-resolution representations throughout its processing, leading to more accurate estimations
573 (refer to **Fig. S3**). Our previous work demonstrated that using a pre-trained HRNet on large
574 human pose estimation tasks, further refined with fine-tuning, results in more precise predictions
575 (36). Consequently, we utilized the pre-trained HRNet architecture (**Fig. S4**), adopting a heatmap
576 size of 64×64 for the 256×256 pelvis images and 152×152 for the 608×608 images of the
577 upper and lower body. The batch sizes were set to 16 for pelvis images and 8 for upper and lower
578 body images.

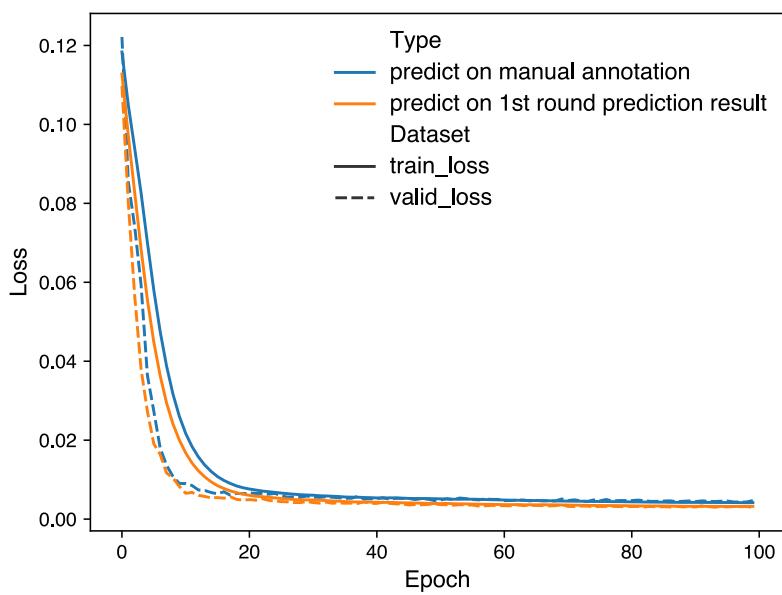
579
580 Initially, our model was trained with 235 manually annotated images. After fine-tuning
581 over 100 epochs using these images and their manually annotated coordinates, we noticed a
582 minimal reduction in loss beyond the 20th epoch (as shown in **Fig. S5**), indicating that 100
583 epochs were adequate for model convergence. The model's performance was then assessed using
584 the remaining 20% of annotated images. According to **Fig. S7**, the mean Euclidean distance error
585 between the human annotation and model prediction was below 2 pixels, similar to the error in
586 human annotation (**Fig. S3**). It is notable that certain landmarks, specifically the ischiopubic
587 ramus inferior (left and right) and the iliac crest posterior (left and right), exhibited larger
588 discrepancies between human annotations and model predictions (**Fig. S6**), a variance also
589 observed in repeated manual annotations (**Fig. S3**), highlighting these landmarks as challenging
590 for human annotators as well.

591
592 We then applied the model to predict landmarks on all remaining images. Initial visual
593 assessments of the model's predictions on original DXA images suggested greater precision
594 compared to human annotations. To explore this observation, we conducted a validation study,
595 selecting a new set of 293 images from both white and black background sets, ensuring they
596 were not part of the initial training set. These images were split into new training and validation
597 sets, maintaining an 80:20 ratio. Training the model anew with these sets and the same
598 architecture and hyperparameters led to faster convergence and slightly lower loss, suggesting
599 the model's predicted coordinates might be less noisy than manual annotations (as indicated in
600 **Fig. S7**). Comparisons between human and model-predicted annotations revealed that the
601 Euclidean distance for all 17 landmarks between the first and second model predictions was
602 significantly smaller than the distance between manual annotations and the first model
603 predictions, reducing the average Euclidean distance to less than one pixel (**Fig. 1C**, **Fig. S6**,
604 **Table S5**).

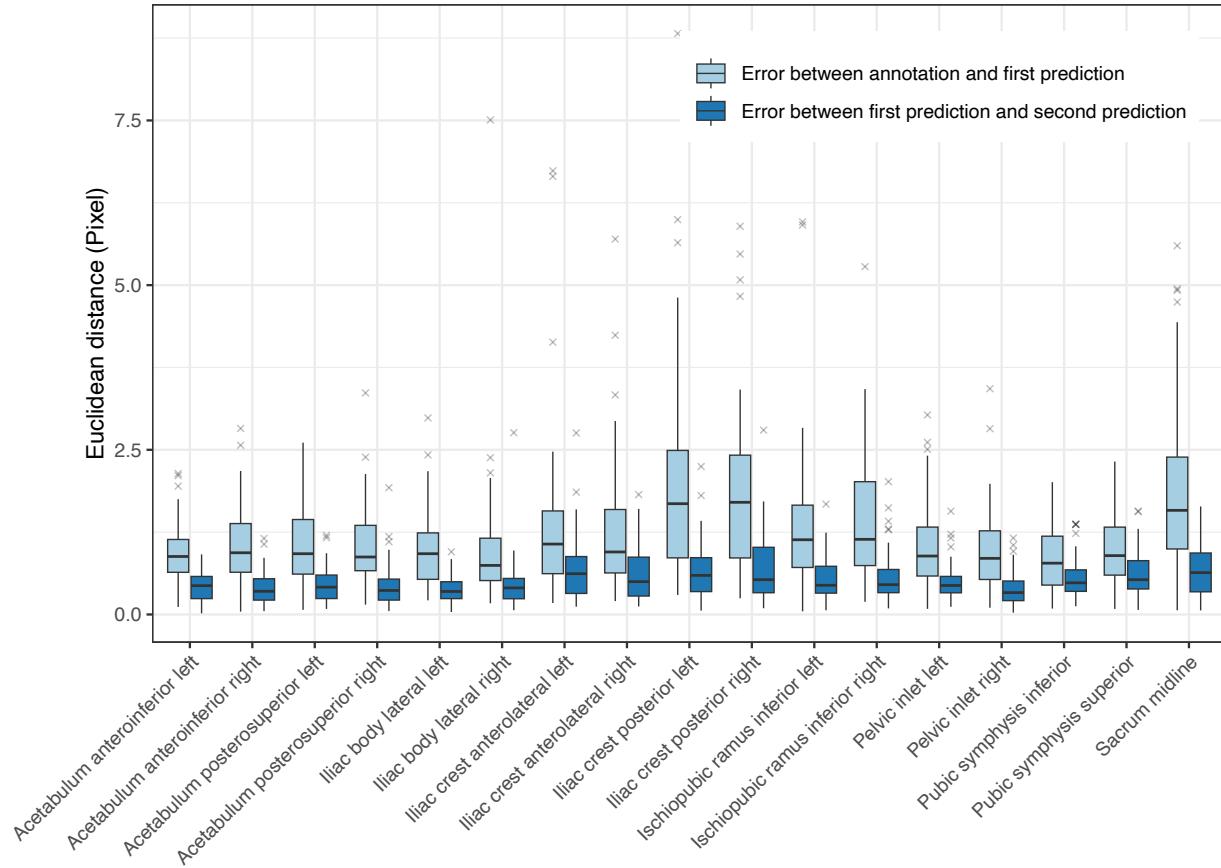
605
606 We further deployed this twice-trained, optimized model on a comprehensive set of
607 39,469 full-body DXA images from the UK Biobank. Additionally, we compared this model's
608 performance to that of a previous study's model (36), particularly assessing the correlation
609 between left and right arm length ratios across two imaging visits (**Fig. S7**). Despite using the
610 same annotated coordinates for model training, our current model showed a significantly higher
611 correlation between imaging visits than our previous model. However, we noticed repeated
612 application of this strategy (outlined in **Fig. 1B**) did not yield significant improvements in model
613 accuracy (**Fig. S8**), suggesting that an additional round of training was useful to reduce the
614 variation in manual annotation to a minimum.



617 **Fig. S4. HRNet deep learning architecture.** High-Resolution Network (HRNet) architecture
618 maintains parallel high to low-resolution subnetworks.

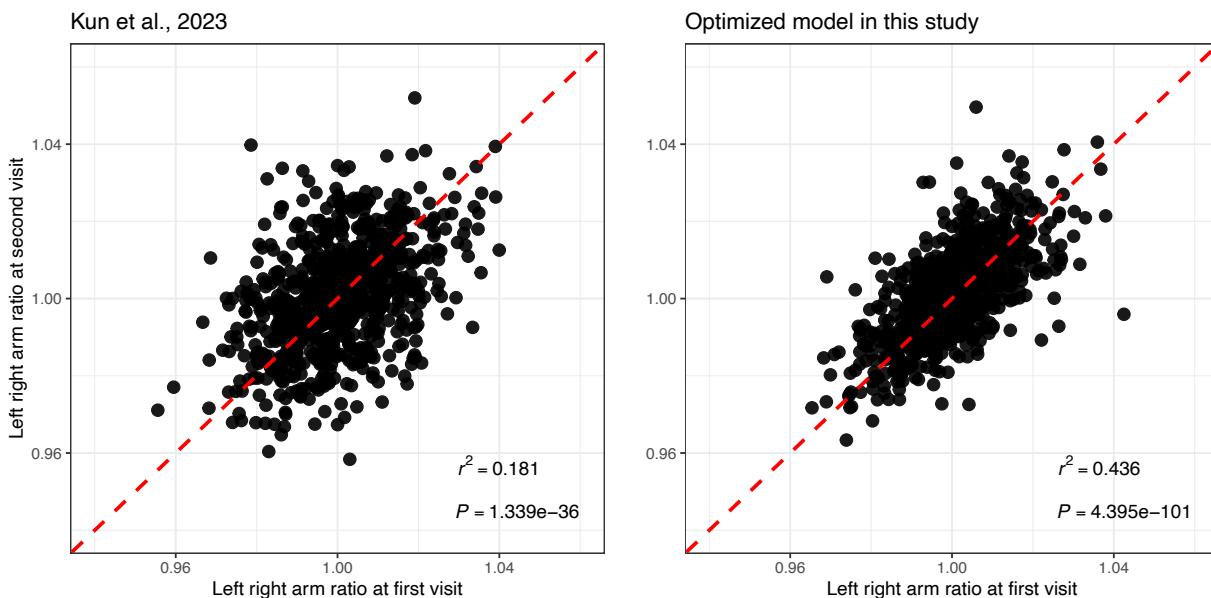


619 **Fig. S5. Training logs of HRNet with two different training sets.**
620



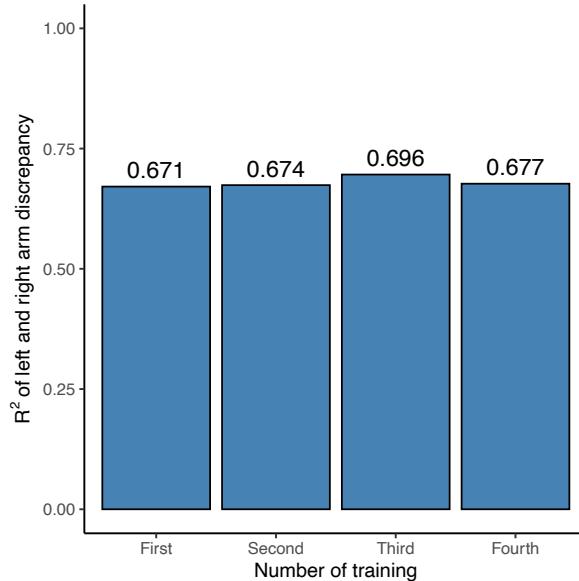
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Fig. S6. Model prediction error. Light blue box plots indicate the Euclidean distance between manually annotated coordinates and HRNet prediction results based on these coordinates across 18 human pelvis landmarks. Dark blue box plots indicate the Euclidean distance between HRNet prediction results based on previous predicted coordinates across 17 human pelvis landmarks.



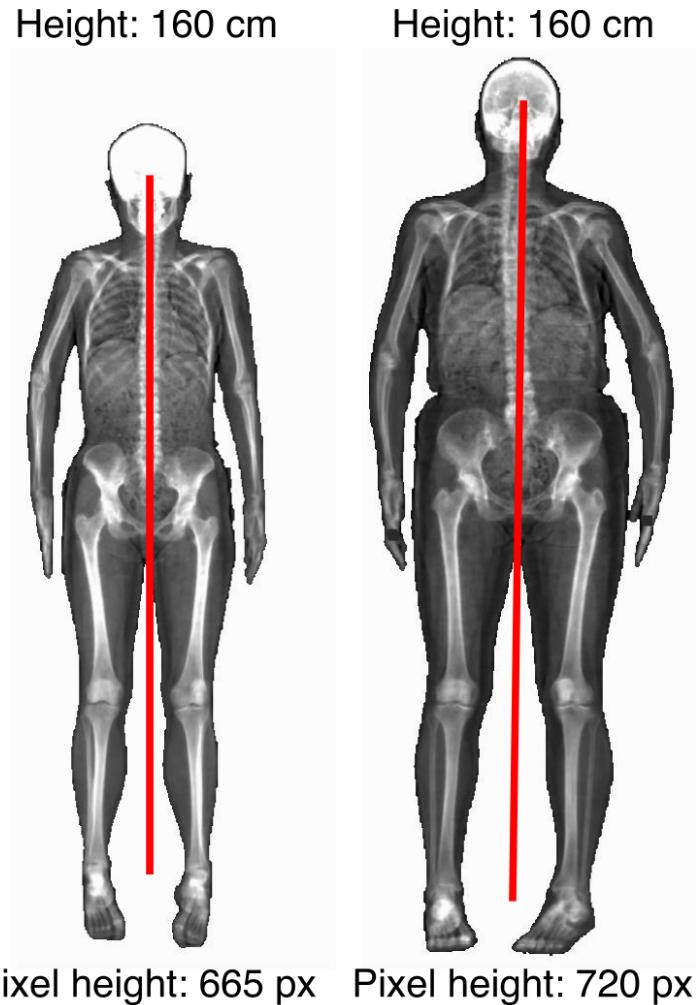
626

627 **Fig. S7. Model Comparison.** The left panel displays the scatter plot of the left-to-right arm ratio
628 from two imaging visits using HRNet, sourced from (36). The right panel shows the scatter plot
629 of the same ratio from two imaging visits but using our optimized HRNet model.



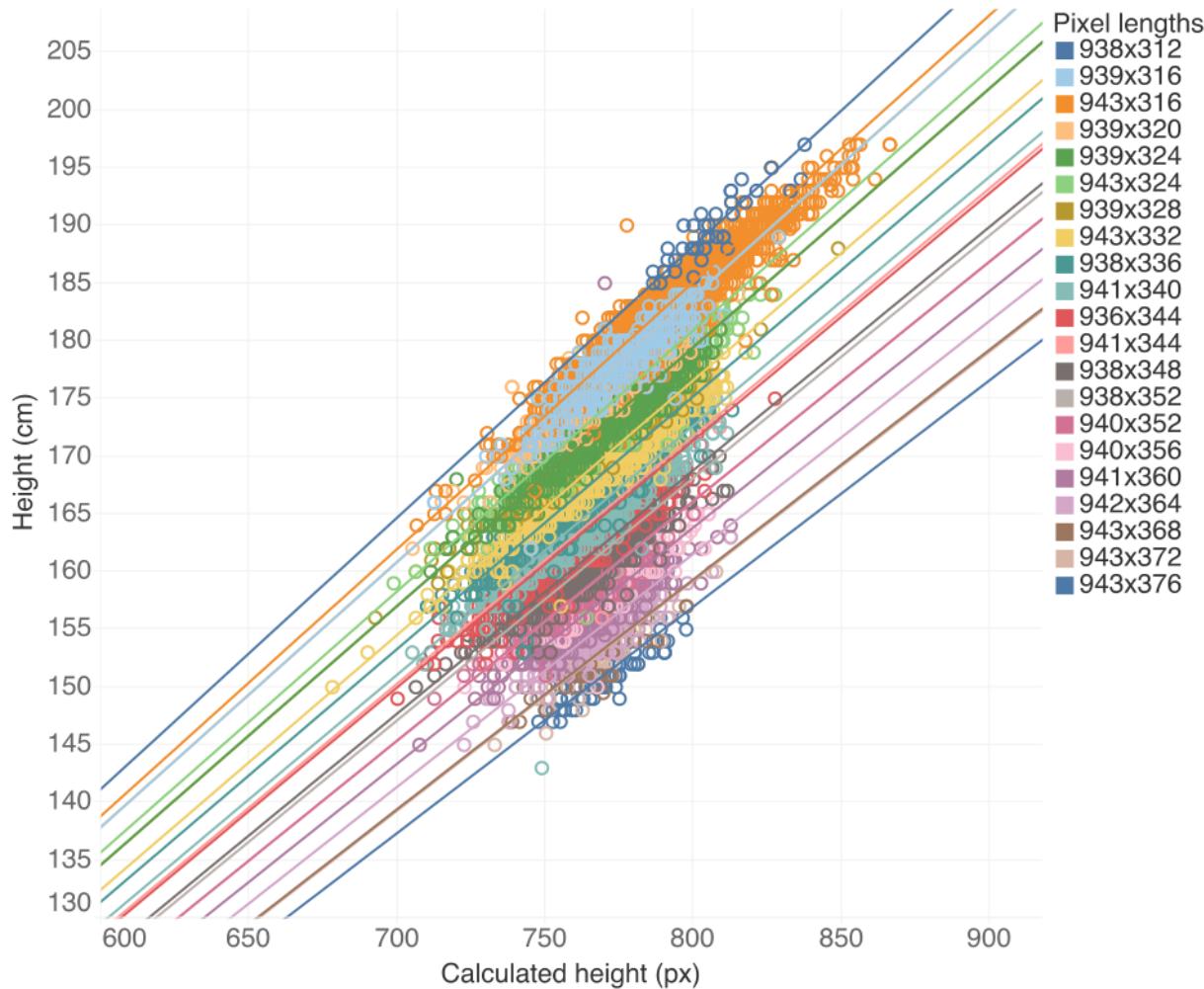
630
631 **Fig. S8. Model prediction accuracy across multiple repeat training.**

632 A major issue in combining our analysis across input pixel ratios was that these pixel
633 ratios represented different resolution scalings, perhaps due to distances that the scanner was
634 held above the patient (**Fig. S9**). That is, in one image a pixel could represent 0.44 cm and in
635 another 0.46 cm. To control for this scaling issue and to standardize the images, we chose to
636 regress height measured directly on our image using the midpoint of the eyes and the midpoint of
637 the two ankle landmarks that could be taken across all image pixel ratios and overall height (FID
638 50) computed externally from the UKB (**Fig. S10, Table S7**). While the height measure we
639 utilized did not include the forehead, it was a relative measure that we used to obtain a scaling
640 factor for each image pixel ratio that we could for normalization. Measurement error of
641 individuals either in our image-based height measure or as reported in the UKB is not expected
642 to affect our conversion from pixels to cms as we are regressing over many individuals.
643 Importantly, we validated this regression and normalization using duplicate individuals taken by
644 different scanners, imaging modes and technicians (**Fig. 1D, 1E**).



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646
647 **Fig. S9. DXA images from the UKB that have undergone different image scaling.** Example
648 of two individuals who were measured to be the same height in the FID 50 in the UKB (overall
649 height) but pixel-based measurements of one image were considerably smaller than the other due
650 to image scaling/resolution differences.

651



652
653 **Fig. S10. A linear regression of image-measured height against UKB-measured height.** For
654 each image pixel-ratio, we regressed height measured in the UKB with height we calculated in
655 pixels from the DXA scan. This provided a conversion from pixels to cm that we used as a
656 normalization factor to correct for differences in resolution.

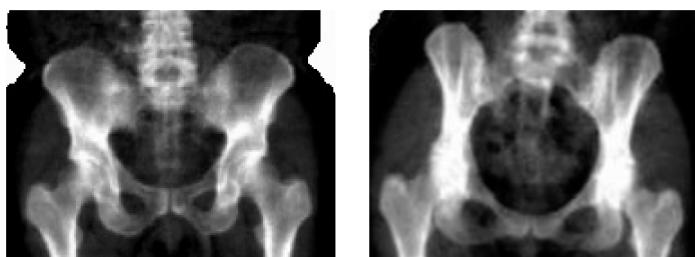
657
658 **Obtaining skeletal element length measures**
659 From each of the 17 landmarks, we generated a total of one angle measure and seven
660 skeletal length measures (Fig. 1) in pixels which we converted to centimeters using coefficients
661 from the regressions with height. We averaged iliac isthmus breadth and acetabular diameter
662 across the left and right sides of the pelvis for all analyses. For all measurements, the phenotype
663 values are shown in Table S8, and the mean and standard deviations are shown in Table S9.

664
665 **Participant data quality control**
666 For all of the following analyses, we filtered the participants with correctly labeled full
667 body DXA images (FID 20158 and 12254) to just Caucasian individuals (FID 22006) from the
668 white British population as determined by genetic PCA (FID 21000). We removed individuals
669 whose reported sex (FID 31) did not match genetic sex (FID 22001), had evidence of aneuploidy
670 on the sex chromosomes (FID 222019), were outliers of heterozygosity or genotype missingness

671 rates as determined by UKB quality control of sample processing and preparation of DNA for
672 genotyping (FID 22027), had individual missingness rates of more than 2% (FID 22005), or
673 more than nine third-degree relatives or any of unknown kinship (FID 22021). We also removed
674 individuals whose standing height (FID50) and weight (FID21002) didn't recorded in their
675 imaging visit. In total 30,370 individuals remained.
676

677 **Removal of image outliers**

678 We removed individuals who were more than 4 standard deviations from the mean for
679 any imaging-derived phenotype from the analysis. In total 31,115 individuals remained (**Table**
680 **S4**). Examination of these outliers by comparing left and right symmetry as well as comparison
681 of other body proportions revealed a heterogeneous set of issues that were associated with the
682 poor prediction by our deep learning model. In some cases, individuals had a limb, or another
683 body part amputated. Some poorly classified images were individuals who had had major hip or
684 knee replacement surgery or had various implants that were causing incorrect model landmark
685 prediction. Another class of outlier images was those that were too poor in quality for any
686 landmarking of any of the points on the image or had abnormal pelvis shape perhaps due to a
687 mendelian disorder (**Fig. S11**).
688



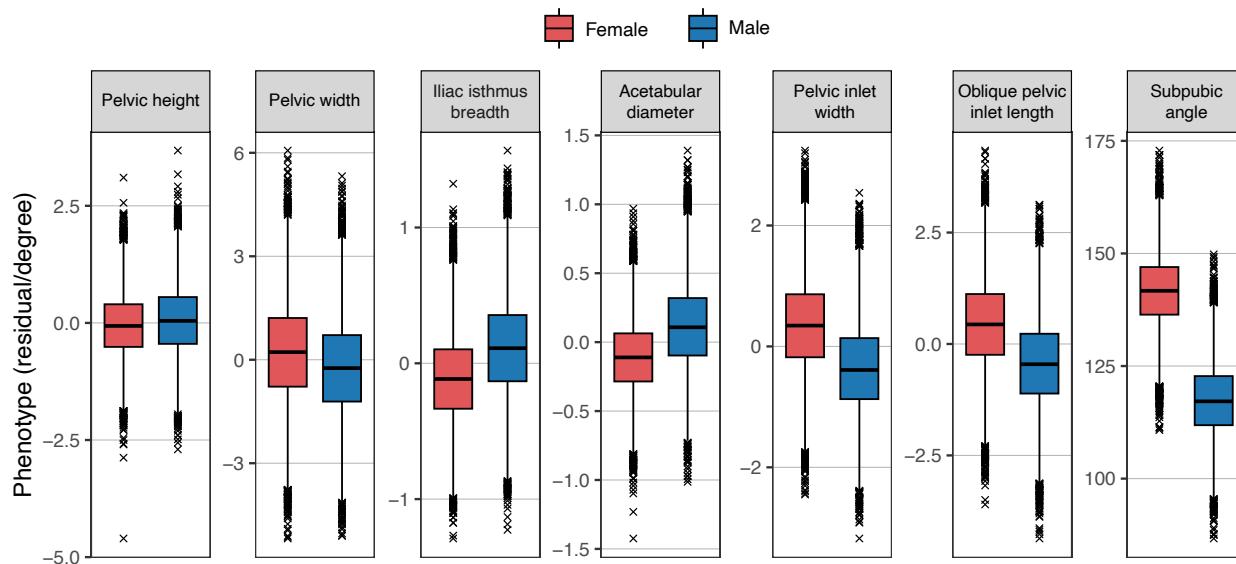
689
690 **Fig. S11. Examples of individuals with normal and abnormal pelvic morphology.** Left panel
691 - normal morphology, right panel - pelvis with highly atypical morphology.

692 693 **Correlations of pelvic left and right discrepancy with handedness**

694 Upon examining **Fig. 1E**, we observed significant differences between the left and right-
695 side measurements for iliac isthmus breadth and acetabular diameter. Given that the discrepancy
696 between these left and right measurements consistently appeared in two separate imaging visits,
697 this is unlikely to be due to random error/noise in model prediction. To investigate whether this
698 left-right discrepancy is associated with handedness, we calculated the ratio between the left and
699 right-side measurements. We then conducted t-test analyses on the ratio of each phenotype to
700 examine differences in these ratios based on handedness (as indicated by FID1707). These
701 analyses were restricted to white British patients.
702

703 **Phenotypic comparison between males and females**

704 For each length phenotype, including pelvic height, pelvic width, iliac isthmus breadth,
705 acetabular diameter, pelvic inlet width, and oblique pelvic inlet length, we regressed out standing
706 height and compared the residuals obtained from the regression analysis between males and
707 females. All six length phenotypes and one angle phenotype showed significant differences
708 between males and females in a t-test (**Fig. S12**).



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Fig. S12. Phenotypic comparison between males and females morphology.

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Genetic data quality control

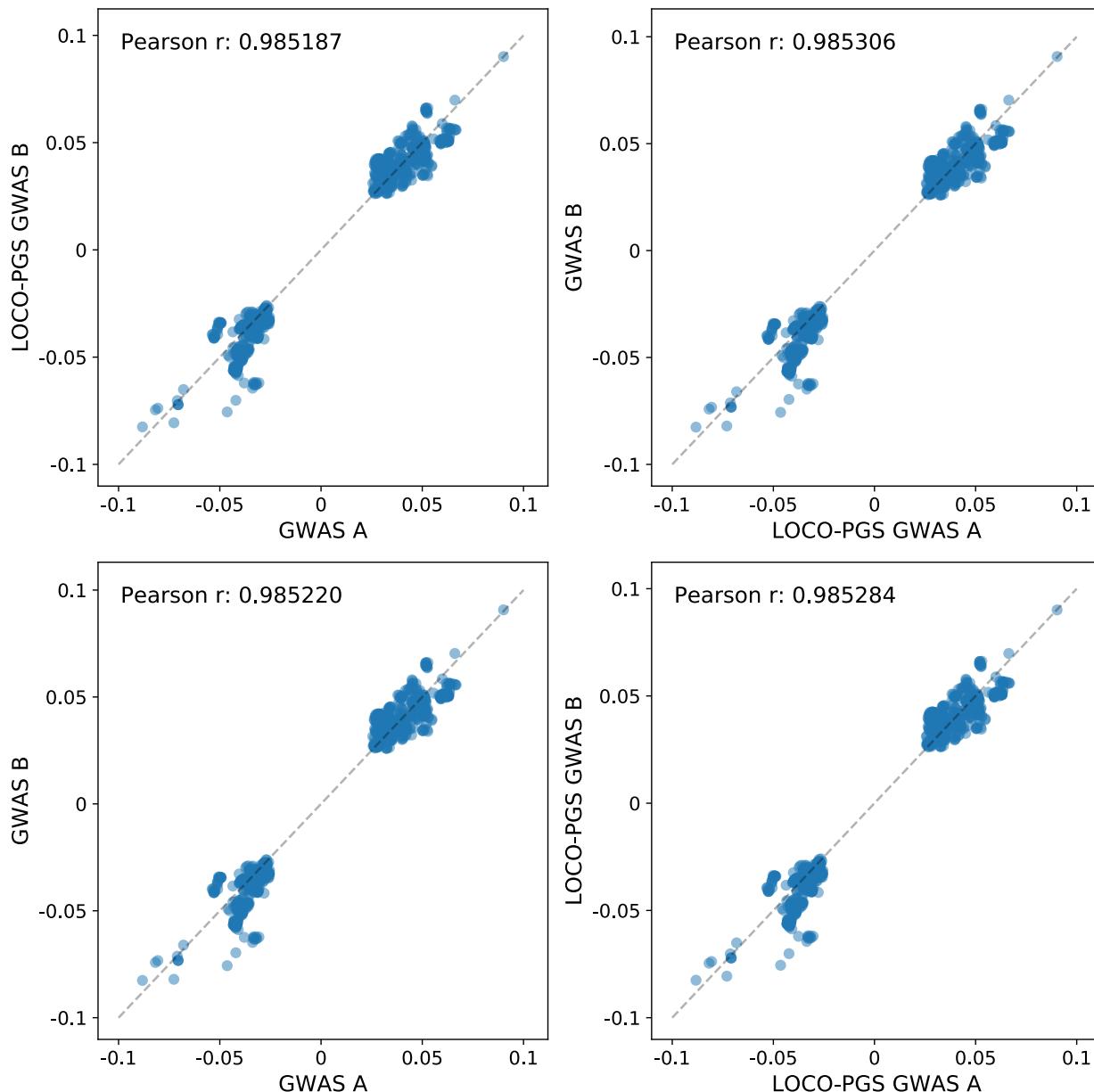
713 Imputed genetic data for 487,253 individuals was downloaded from UKB for
714 chromosomes 1 through 22 (FID 22828) then filtered to the quality-controlled subset using
715 PLINK2 (70). All duplicate single nucleotide polymorphisms (SNPs) were excluded (--rm-dup
716 'exclude-all') and restricted to only biallelic sites (--snps-only 'just-acgt') with a maximum of 2
717 alleles (--max-alleles 2), a minor allele frequency of 1% (--maf 0.01), and genotype missingness
718 no more than 2% (--maxMissingPerSnp 0.02). In total 8,638,168 SNPs remained in the imputed
719 dataset. Non-imputed genetic data (genotype calls, FID 22418) did not contain duplicate or
720 multiallelic SNPs but were filtered to the quality-controlled subset; 652,408 SNPs remained.
721

722

Adjusting for height correlation in GWAS by adding height as covariate

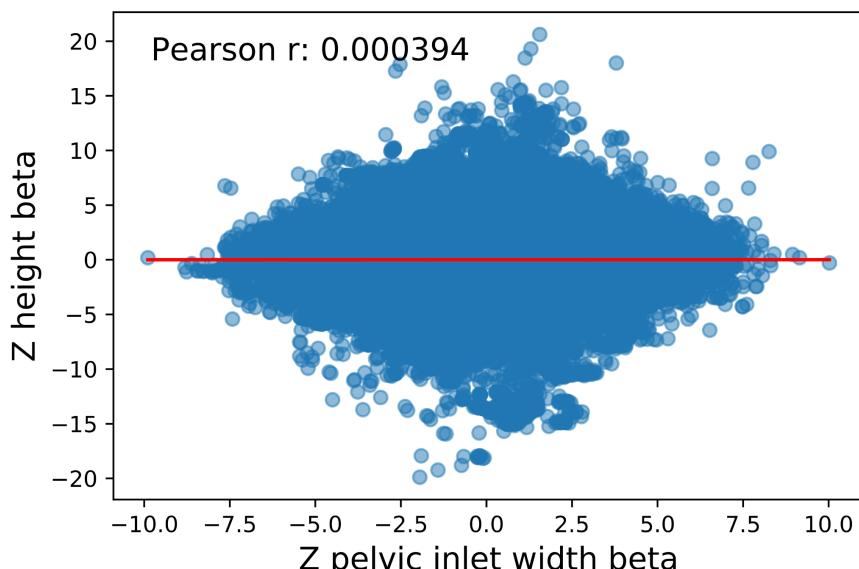
723 A major issue in carrying out GWAS for phenotypes such as ours where we would like to
724 control for height is the potential for confounding due to the adjustment. McCaw et al. highlight
725 the pitfalls in GWAS of ratio traits and describe ways to reduce this confounding (43). Following
726 their pipeline, we carried out GWAS, adjusting not only for height but also for leave-one-
727 chromosome-out (LOCO) polygenic scores (PGS) of height. Briefly, we first performed GWAS
728 on approximately 370k white British individuals without imaging data using PLINK (70).
729 Second, we estimated LOCO-PGS for each individual with imaging data for each chromosome
730 using Bayesian regression with continuous shrinkage priors (47), employing associated single
731 nucleotide polymorphisms from HapMap3. In total, each individual received 22 LOCO-PGS one
732 for each chromosome. Third, we randomly split 30k imaged individuals into two groups, with
733 each group containing around 15k individuals, and conducted two GWAS models for each
734 chromosome. These models adjusted for the first 20 principal components, age, sex, height, and,
735 optionally, the corresponding LOCO-PGS as covariates. Finally, we concatenated all
736 chromosomes' GWAS summary statistics. In examining the effect sizes across all manners of
737 carrying out the analyses, we found that results ($p < 5 \times 10^{-4}$) adjusted for and not adjusted for
738 LOCO-PGS were very similar in the split-sample study (Fig. S13), suggesting that the genetic

739 component of collider bias is minimal. In **Fig. S12** we present an image of the correlation of
740 effect sizes for only a single trait but the results of all traits ranged in correlation between 0.959
741 to 0.989. To provide additional confirmation of reduced confounding with the adjustment for
742 height, we compared the effect size correlation between snps at $p < 5 \times 10^{-4}$ for height and a
743 specific trait (without adjusting for LOCO-PGS) and observed that effect sizes estimated for
744 height and the SNP effect size of a specific trait were completely uncorrelated (**Fig. S14**). This
745 further indicates that the results of the GWAS we conduct for particular pelvic proportions are
746 largely independent of height.



747
748 **Fig. S13. Two GWAS model comparison.** To compare with or without adjusting for LOCO-
749 PGS we compared the effect size correlation between two separate samples with the same or
750 different GWAS models with SNP $p < 5 \times 10^{-4}$. Here we only show one randomly picked trait
751 birth canal width, but we observed similar signals for all traits.

752



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Fig. S14. SNP effect size correlation between height and birth canal width.

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756 **GWAS and Heritability analysis**

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GWAS was performed with BOLT-LMM (47). LD Score v1.0.1 was used to compute linkage disequilibrium regression scores per chromosome with a window size of 1 cM (44). PLINK2 --indep-pairwise with a window size of 100 kb, a step size of 1, and an r^2 threshold of 0.6 was used to create a list of 986,812 SNPs used as random effects in BOLT-LMM. Covariates were the first 20 genetic principal components provided by UKB (FID 22009), sex (FID 31), age (FID 21003), age-squared, sex multiplied by age, sex multiplied by age-squared, and standing height (FID 50). In addition, the DXA scanner's serial number and the software version used to process images were combined into one covariate, resulting in 5 factor levels.

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SNPs in each resulting GWAS were clumped in PLINK using --clump with a significance threshold of 5.0×10^{-8} , a secondary significance threshold of 1.0×10^{-4} for clumped SNPs, an r^2 threshold of 0.1, and a window of 1 Mb. SNPs were assigned to genes with --clumpverbose --clump-range glist-hg19 downloaded from PLINK gene range lists (71). The genomic inflation factor of each phenotype was assessed in R version 4.2.1 as the ratio of the median of the observed chi-squared distribution (an output of BOLT-LMM --verbose) to the expected median of the chi-squared distribution with one degree of freedom.

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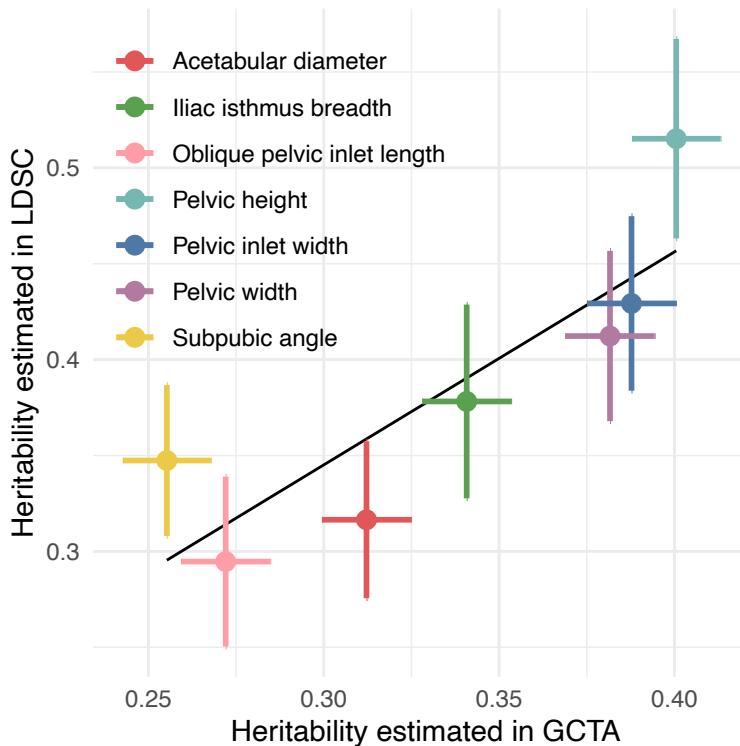
We created the genetic relationship matrix for our quality-controlled subset but without any related individuals and a minor allele frequency of 0.01, then ran GCTA for each phenotype pair with the first ten genetic principal components provided by UKB (FID 22009).

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The heritability of each phenotype was assessed with European HapMap3 SNPs using GCTA (38) with the same covariates as GWAS, excluding age-squared and sex by age-squared. We also estimated heritability using LDSC (44) and found similar heritabilities (20-50%) (Fig. S15, Table S10).

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Fig. S15. Heritability estimated in GCTA and LDSC.

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Sex-specific analysis

We performed a GWAS independently in males and females using the same process and covariates we used in the combined GWAS analysis in the previous section. Subsequently, we used LDSC to carry out genetic correlation analysis between GWAS conducted in males and females. As depicted in **Fig. 1F**, outer pelvic morphology, such as pelvic height, had genetic correlation consistent with 1. However, phenotypes related to the birth canal, such as pelvic inlet width, oblique pelvic inlet length, and subpubic angle, exhibit differences significantly different from 1. This aligns well with previous studies, underscoring the functional importance in females to accommodate childbirth.

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To determine if any sex-specific loci were present in our pelvic phenotypes, we also carried out additional GWAS in PLINK involving a Sex-Genotype interaction for each SP on our original population of 31,115 individuals to determine loci with sex-specific effects. Across all the traits that we examined we did not find evidence for interaction at any locus which would signify sex-specificity. However, we note that this lack of evidence could possibly be due to reduced power for detecting interaction effects at this sample size. We also report the summary statistics for this GWAS with interactions along with the other GWAS that we performed in the Supplementary Data.

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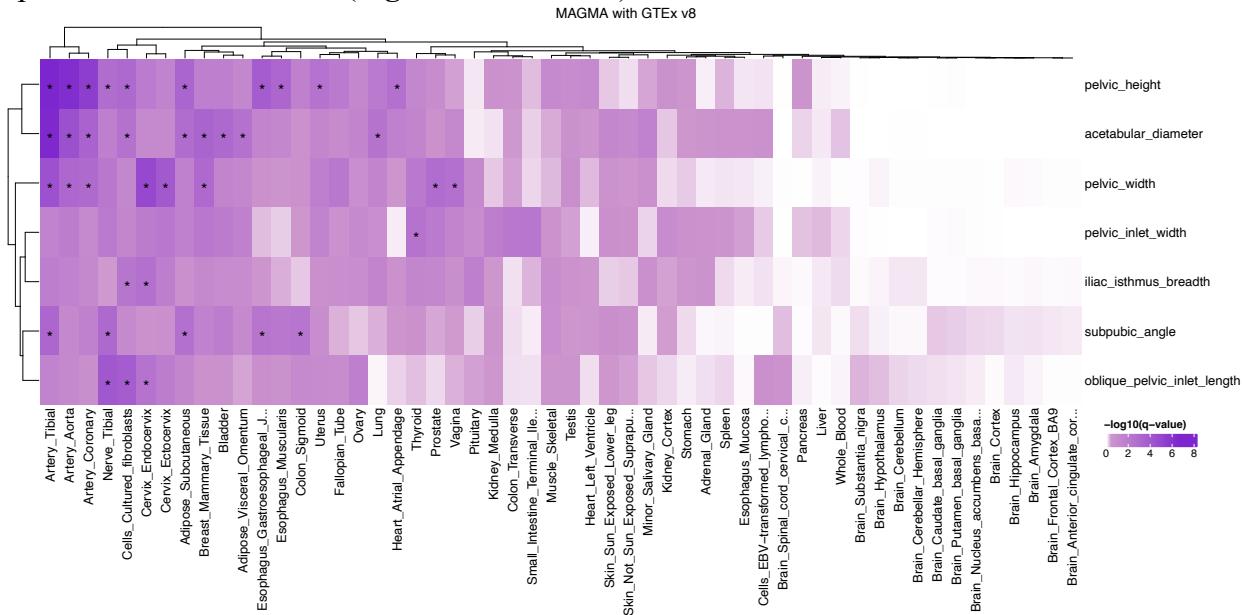
Clumping and identification of genes associated with loci

To obtain a set of independent SNPs associated with each PP phenotype, we first performed clumping analysis for each phenotype using plink and assigned SNPs to genes with --clump-verbose --clump-range glist-hg19 with an r^2 window of 0.1 and a 1 Mb threshold of

809 physical distance for clumping. We downloaded gene ranges from plink for hg19 (72).
810 Following clumping, we looked at a subset of 7 phenotypes and combined the significant SNPs
811 across the chosen phenotypes resulting in 339 unique SNPs.
812

813 **Functional mapping and gene enrichment analysis**

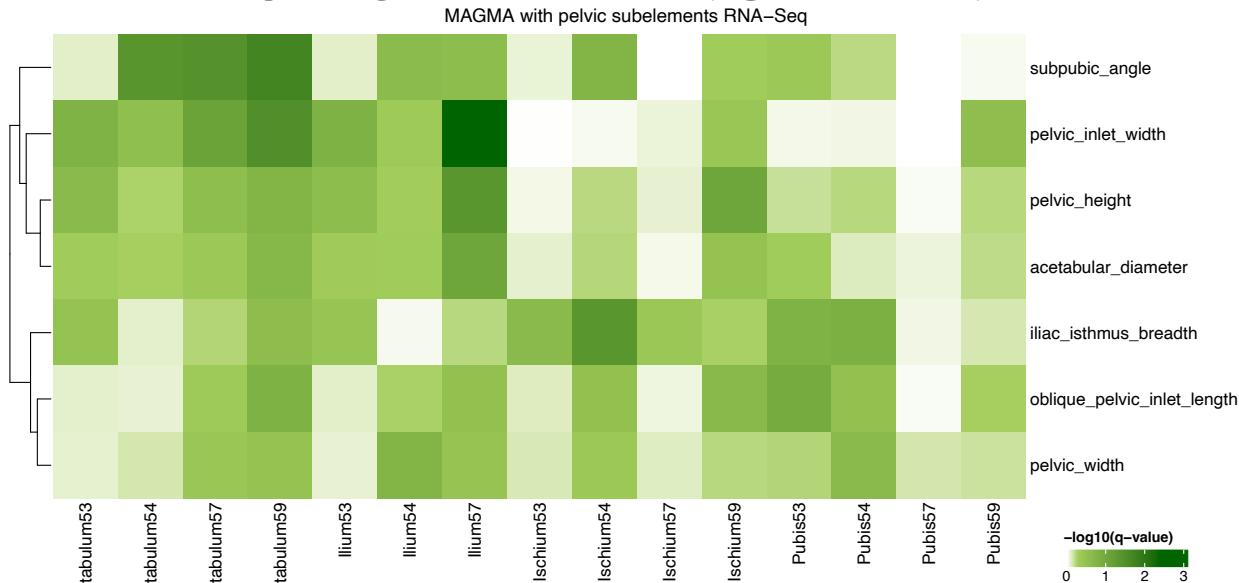
814 We ran FUMA (73) without any predefined lead SNPs on a sample size of 31,115
815 individuals. GENE2FUNC was run with all types of genes selected as background genes using
816 Ensembl v92 with GTEx v8 gene expression data sets and we set window sizes 10 kb for both
817 upstream and downstream (**Fig. S16, Table S19**).



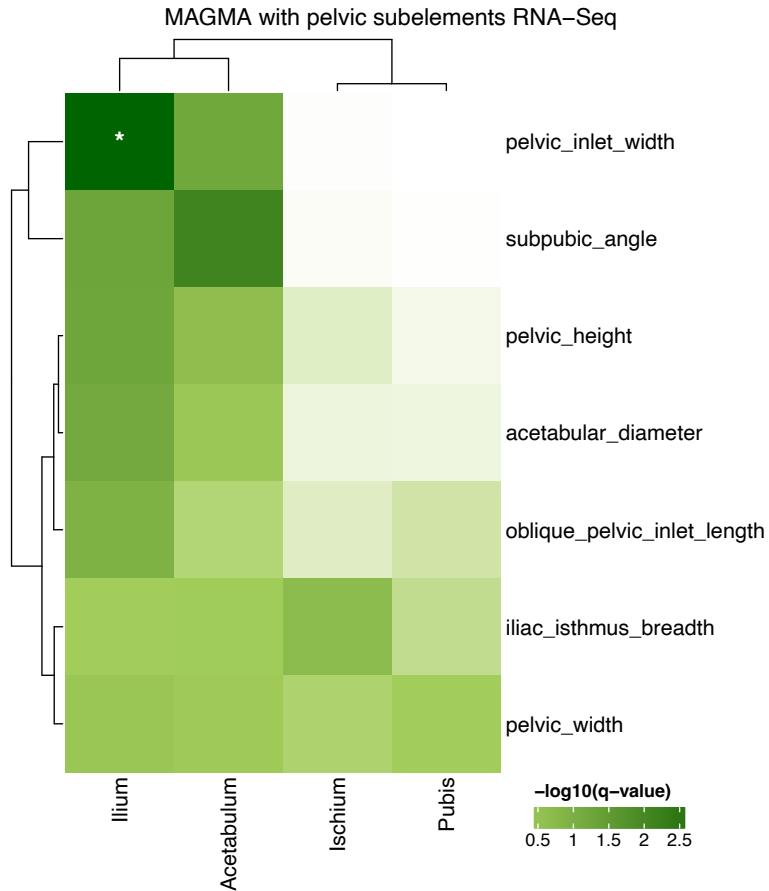
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819 **Fig. S16. MAGMA gene property analysis with GTEx 8 and image-derived phenotypes**
820 **GWAS**

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822 **Transcriptome analysis**
823 To explore the genetic underpinnings of pelvis-related phenotypes in relation to specific
824 developmental stages of the human pelvis, we sought enrichment of genes associated with our
825 GWAS results for pelvis-related phenotypes in gene expression data across four distinct
826 developmental stages of the human pelvis during the embryonic period, as detailed in (33). Our
827 primary objective was to discern which developmental stage (E53, E54, E57, or E59) might be
828 linked to changes in pelvis shape. We downloaded RNA-Seq data for human embryonic pelvises
829 at different developmental stages from the GEO data repository (GSE165930). Subsequently, we
830 converted gene names to Ensembl gene IDs using the biomaRt package (version 2.52.0) in R. To
831 compute the relative gene expression level for a specific subelement at a particular
832 developmental stage, we subtracted the average expression from other stages for that specific
833 subelement and from other subelements across different stages. Following this, we conducted a
834 MAGMA gene property analysis to assess enrichment between genes expressed during specific
835 developmental stages and our phenotypes. However, our analysis did not reveal any significant
836 enrichment for any developmental stage in our GWAS after adjusting for multiple comparisons
837 using FDR correction for both the number of traits and developmental stages (**Fig. S17, Table**
838 **S20**). In a subsequent approach, we combined data from different developmental stages to

839 investigate potential associations between pelvis-related phenotypes and specific pelvis
840 subelements. We determined the relative expression of specific subelements by subtracting the
841 average expression of other subelements. Another round of MAGMA gene property analysis
842 revealed a significant effect between the Ilium and pelvic inlet width, as well as between the
843 Acetabulum and subpubic angle, after FDR correction (Fig. S18, Table S21).



844
845 **Fig. S17. MAGMA gene property analysis with pelvis subelements in different**
846 **developmental stages ATAC-Seq and image-derived phenotypes GWAS**



847
848 **Fig. S18. MAGMA gene property analysis with pelvis subelements ATAC-Seq and image-
849 derived phenotypes GWAS**

850
851 **Phenotypic association of skeletal phenotypes with musculoskeletal disease**
852 To examine correlations between our pelvis phenotypes with musculoskeletal disease,
853 musculoskeletal or connective tissue diseases related to the hip, knee, and back we obtained data
854 from UKB Chapter XIII (FID 41270) ICD-10 codes as well as self-reported pain phenotypes
855 (FID 6159) for the hip, knee and back. We then regressed the binary outcome of disease or
856 reported pain against pelvis phenotypes controlling for clinically relevant covariates that are
857 known to affect OA (74) including age, sex, diet, BMI, and other factors. A full list of variables
858 we controlled for are reported in **Table S14**. After running the regressions, we used Bonferroni
859 correction for significance at the level of the total number of disease/pain traits multiplied by the
860 total number of skeletal phenotypes.

861
862 **Polygenic risk score (PRS) association of skeletal phenotypes with musculoskeletal disease**
863 This analysis only utilized the ~370,000 white British individuals who were not included
864 in our imaging dataset for which GWAS was conducted. We generated PRS for each of the
865 generated traits with Bayesian regression and continuous shrinkage priors (47) using the
866 associated single nucleotide polymorphisms from HapMap3. We ran a logistic or linear
867 regression of the PRS on traits across all individuals, adjusting for weight, household income,
868 non-insulin-dependent diabetes mellitus (ICD-10 code: E11), depressive episode (ICD-10 code:

869 F32), recurrent depressive disorder (ICD-10 code: F33), chronic ischaemic heart disease (ICD-10
870 code: I25), smoking status (FID 20116), and sleep duration (FID 1160). For female PRS
871 regression we also adjusted for the number of live births (FID 2734).

872

873 **Genetic correlation of skeletal proportions with pregnancy phenotypes**

874 We utilized cross-trait LD score regression

875 (<https://github.com/bulik/ldsc/wiki/Heritability-and-Genetic-Correlation>) for estimating genetic
876 correlations between each of our pelvis-related phenotypes and case-control pregnancy
877 phenotypes from the Finngen (https://www.finngen.fi/en/access_results) by using GWAS
878 summary statistics.

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881 **Supplementary Tables**

882

883 **Table S1** - Previous studies have attempted to test the obstetrical dilemma hypothesis.
884 This table contains the papers involved in the debate on the obstetrical dilemma.

885 **Table S2** - GWAS population summary

886 This table contains summary data on the population subset used in our GWAS from the UKB.

887 **Table S3** - Initial deep learning QC

888 This table contains the number of patients removed from each QC step before landmark
889 estimation.

890

891 **Table S4** - Image filtering

892 This table contains the number of patients excluded at each step of quality control following
893 landmark estimation.

894 **Table S5** - Human annotation vs model prediction

895 This table contains the error between human annotation and the first model prediction, as well as
896 the error between the first model prediction and the second model prediction.

897 **Table S6** - Image pixel data

898 This table contains the number of full-body skeletal DXA images for each pixel aspect ratio in
899 the UKB.

900 **Table S7** - Image scaling coefficient

901 This table contains the scaling factor estimated from the regression analysis, which is used to
902 convert pixels to centimeters.

903 **Table S8** - 7 Pelvic phenotype values across 39413 individuals

904 This table contains a list of all generated IDPs.

905 **Table S9** - Pelvic phenotypes summary

906 This table contains the basic statistics of IDPs.

907

908 **Table S10** - GCTA and LDSC heritability estimation

909 This table contains the heritability for each IDP as determined by GCTA.

910 **Table S11** - Clumped independent SNPs and corresponding genes

911 This table contains output from PLINK --clump ranges command including lead SNP, p-value,
912 the number of kilobases in each clump, gene mapping for each clump range as well as whether
913 the single clump range genes are related to known mouse phenotypes and rare human disease.

914

915 **Table S12** - ICD10 Codes

916 This table contains all ICD10 codes used in our analyses.

917 **Table S13** - UKB phenotypes FID

918 This table contains the FID of each UKB trait used in our analyses.

919 **Table S14** - Association analysis covariates

920 This table contains the list of covariates used in our regression analyses and the FID from the
921 UKB

922

923 **Table S15** - Phenotypic association results

924 This table contains the results from the phenotypic association analysis.

925

926 **Table S16** - PRS association results

927 This table contains the results from the PRS association analysis.

928

929 **Table S17** - Female PRS association results

930 This table contains the results from the female PRS association analysis.

931 **Table S18** - Female genetic correlation results

932 This table contains the results from the female genetic correlation analysis.

933 **Table S19** - MAGMA with GTEx v8

934 This table contains the results from the MAGMA analysis with gene expression data from GTEx
935 v8.

936 **Table S20** - MAGMA analysis across pelvic subelements across different time points

937 This table contains the results from the MAGMA analysis with gene expression data from
938 different parts of the pelvis across different development time points.

939 **Table S21** - MAGMA analysis across pelvic subelements

940 This table contains the results from the MAGMA analysis with gene expression data from
941 different parts of the pelvis.

942

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950

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952 E.K. performed analysis. L.X., E.K., D.P., and J.W. performed data preprocessing. M.B. and T.S.
953 provided comments and helped supervise the work.

954

955 **Competing interests:** The authors declare no competing interests.

956

957 **Data and materials availability:** Code used for performing the deep learning–based key point
958 identification and quality control of the DXA data is available at
959 <https://github.com/xliaoyi/Human-pelvic-form>. Our GWAS summary statistics are available at
960 <https://utexas.box.com/s/w1n8oz61sb7km2yotqyyg8em2td7b2wb>. Individual-level information
961 of skeletal lengths has been reported back to the UKB and will be available via the Access
962 Management System.

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