

1 **Title:**

2 **Advancing marker-gene-based methods for prokaryote-mediated multifunctional redundancy: exploring**
3 **random and nonrandom extinctions in a watershed**

4

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21

22 **Abstract**

23 Multifunctional redundancy, the extent of loss in multiple ecosystem functions with decreasing biodiversity, stands
24 as a crucial index for evaluating ecosystem resilience to environmental changes. We aimed to refine a
25 marker-gene-based methodology for quantifying multifunctional redundancy in prokaryotic communities. Using
26 PICRUSt2, we predicted KEGG orthologs (KOs) for each Amplicon Sequence Variant (ASV), assessed
27 community-wide KO richness, and validated predictions against experimentally quantified phenotypic
28 multifunctionality. Additionally, we introduced a refined regression on ASV richness–KO richness curves,
29 providing a reliable estimate of the power-law exponent within computational time constraints, serving as the
30 multifunctional redundancy index. Incorporating various non-random extinction scenarios alongside a random one
31 allowed us to quantify estimate variations between scenarios, providing conservative estimates of multifunctional
32 redundancy. Applied to Lake Biwa and four of its inlet rivers, the refined methodology unveiled spatio-temporal
33 variations in multifunctional redundancy. Our analysis demonstrated lower redundancy in Lake Biwa compared to
34 rivers, aiding in prioritizing conservation targets and inferring distinct community assembly processes. Future
35 directions include a deeper exploration of KO composition information for detailed multifunctionality
36 quantification and the refinement of extinction scenarios. This study demonstrates the promising integration of
37 bioinformatic functional prediction and modeling biodiversity loss, offering a valuable tool for effective ecosystem
38 management.

39

40 **Introduction**

41 Microbial diversity is pivotal for sustaining ecosystem functions across both aquatic and terrestrial environments,
42 spanning from local to global scales (1,2). However, this critical role faces challenges posed by the risks of local
43 extinction of microbial species (3,4) and spatial homogenization of microbial community composition (5,6).

44

45 In assessing the resilience of ecosystems to environmental changes and disturbances, the extent of
46 biodiversity loss and its subsequent impact on ecosystem functions are crucial factors (7–10). Functional
47 redundancy, representing the degree of overlap in ecosystem functions among taxonomic units within a community,
48 is a critical concept in addressing the latter concern. Ecosystems characterized by lower functional redundancy are
49 more vulnerable to disturbances, highlighting the necessity for prioritized conservation efforts.

50

51 Functional redundancy is originally defined as the degree of maintaining a single ecosystem function in
52 the face of taxonomic richness loss (11–14). Various methods for quantifying functional redundancy have emerged,
53 largely rooted in the multifunctionality concept (9,15), which addresses the simultaneous assessment of multiple
54 ecosystem functions. In addition to simple indices utilizing taxonomic and functional diversity measures, as
55 comprehensively reviewed by Galland *et al.* 2020, the relationship between taxonomic richness loss and functional
56 richness loss is depicted through simulations of random or non-random taxonomic extinctions (17,18). The shape
57 parameter of the resulting curve, such as the area under the curve, is subsequently employed as an index of
58 functional redundancy (18,19). While some studies focus on the correlation between taxonomic composition
59 dissimilarity and functional dissimilarity to infer the presence and strength of functional redundancy (20), it should
60 be noted that this discussion primarily revolves around the linkage between taxonomic composition and functional
61 composition. This approach is not explicitly designed to predict the loss of multiple functions due to taxonomic
62 richness loss.

63

64 Extending the shape parameter approach to microbial communities, the previous study (21) incorporated
65 the entirety of the functional gene pools within a community as a proxy for genome-based multifunctionality. It
66 introduced the exponent coefficient from power-law fitting on the curve of taxonomic richness and functional gene
67 richness as an index of genomic multifunctional redundancy. In the study, Miki, Yokokawa and Matsui (2014)
68 employed in-silico simulations with habitat-specific pseudo-communities sourced from the Microbial Genomic
69 Database (MBGD, (22)). The study revealed that functional redundancy was considerably lower than indicated by
70 earlier empirical studies (23). This low functional redundancy has since received support from multiple studies
71 using different approaches (2,24,25). However, it is worth noting that quantitative comparison of multifunctional
72 redundancy across various methods poses challenges (26), underscoring potential weaknesses in asserting
73 unanimous support for the observed lowness in functional redundancy.

74

75 Nonetheless, it is crucial to highlight the limitations of the methodology introduced by (21), which we
76 aim to address and improve upon in this study. First of all, this earlier work presents some issues because it

77 concentrated on creating pseudo-communities from the Microbial Genomic Database (MBGD(22)) with a small
78 number of naturally sampled communities. Functional overlap may be underestimated as a result of the deliberate
79 selection of just one strain from each genus to build the pseudo-communities in order to avoid introducing
80 phylogenetic bias. This is particularly significant when compared to natural communities that permit the
81 cooccurrence of closely related taxonomic units, potentially biasing the estimation towards lower functional
82 redundancy. Moreover, when applying the proposed method to marker-gene compositions from natural samples,
83 bioinformatic tool such as PICRUSt or tax4fun (27,28) was not used to connect marker-gene information to whole
84 genome prediction (21). Instead, the phylogenetically closest strain with whole genome information registered in
85 the MBGD database was manually selected, which could introduce biases. Another limitation arises from the
86 consideration of only random extinction, which might underestimate the impact compared to realistic non-random
87 sequences of extinction. Lastly, while Ruhl *et al.* 2022 employed a similar power-law fitting approach with direct
88 measurements of functional genes through both amplicon sequences and metagenomics, it revealed limitations in
89 fitting the power-law curve. This cautions against generalizing findings based on power-law fitting in the context of
90 assessing functional redundancy.

91

92 Due to the persisting cost-ineffectiveness of metagenomic sequencings and the complexity of taxonomic
93 assignment algorism to metagenomes (30), methods that rely only on marker-gene (16S rRNA gene) data are still
94 important for assessing multifunctional redundancy. In light of this, our study aimed to improve the methodology
95 originally proposed by Miki, Yokokawa and Matsui (2014). Subsequently, we applied this refined, and
96 cost-efficient method to analyze natural samples, with the objective of quantifying variations in functional
97 redundancy across different locations within a single watershed, encompassing both spatial and temporal scales.

98

99 In our study, we focused on two key aspects to enhance the existing methodology. Firstly, for the
100 improvement of functional prediction from marker-gene composition, we utilized PICRUSt2 (31). This approach
101 was indirectly validated through a comparison with metabolic profiling, specifically conducted via the Ecoplate
102 incubation experiment (32). Secondly, our study incorporated both random and non-random extinction scenarios to
103 comprehensively estimate the range of functional redundancy. This consideration allowed us to evaluate the impact
104 of varying extinction patterns on the estimation of functional redundancy. Furthermore, we critically assessed the
105 fitting of power-law regression and improved the regression method to ensure robustness and reliability in our
106 analyses.

107

108 We then applied the refined methodology to various sites within a single watershed, including Lake Biwa
109 and its four inlet rivers in Japan. Multiple samplings were conducted at each site, enabling us to investigate
110 variations in functional redundancy across both spatial and temporal dimensions. Additionally, by pooling all
111 samples and simulating species extinctions as a single metacommunity (33–36), we aimed to elucidate potential
112 mechanisms contributing to the variations in multifunctional redundancy among these five sites in the single
113 watershed. These steps are pivotal in contributing to the ongoing debate about redundancy levels, as discussed in
114 earlier studies (2,14,37–39). Moreover, it plays a crucial role in outlining a strategy for assessing biodiversity

115 impact (40) and determining conservation priorities based on functional redundancy (19).

116

117 Materials and methods

118 Sample collection

119 We collected water samples from the surface at the Ie-1 station in the north basin of Lake Biwa (35°12'58"N,
120 135°59'55"E) on July 3, July 30, September 10, and October 17 2019, and four of the inlet rivers of Lake Biwa:
121 Yasu River (35°02'35.5"N, 136°01'10.0"E), Hino River (35°06'09.1"N, 136°04'37.4"E), Echi River (35°11'44.0"N
122 136°10'46.5"E), and Ane River (35°24'45.1"N, 136°16'59.0"E) on July 9, August 7, September 17, and October 15
123 2019, respectively. We filtered 250 mL (for the river samples on July 9) or 500 mL (for all the other samples) with
124 φ 0.22 µm Sterivex™ cartridges (SFGV010RS, Merck Millipore Darmstadt, Germany) filled with 1 g of zirconia
125 beads implemented (φ 0.5 mm, YTZ-0.5; AsOne, see Ushio 2019) for the amplicon sequencing. The variations in
126 filtered water volumes were contingent on water sample conditions, primarily influenced by fine particles that
127 tended to clog the filter pores. Following filtration, each filter cartridge received the addition of 1 mL of RNAlater
128 solution. Additionally, 50 mL of unfiltered water was collected specifically for bacterial direct count and the
129 ecosystem functioning experiment. The samples designated for both amplicon sequencing and ecosystem
130 functioning experiments were promptly transported back to the laboratory. Throughout the transportation period (up
131 to 6 hours), the samples were maintained at 4°C. The filter cartridges were subsequently preserved in a freezer at
132 -20°C until DNA extraction and further processing, while the unfiltered samples were immediately utilized for
133 ecosystem functioning experiments.

134

135 Bacterial direct count

136 Bacteria were enumerated directly under an epifluorescence microscope using the SYBR Green I staining method
137 by Honjo *et al.* (2007) with some modifications. Briefly, 1 mL of the water sample was mixed with 10 µl of
138 200-fold diluted SYBR Green I solution for bacterial staining. After staining for 10 minutes, the bacteria in the
139 water sample were trapped onto a 0.2 µm pore-size polycarbonate membrane filter (Advantec, Japan) and then
140 mounted on glass slides with a drop of immersion oil (Olympus, Japan). We randomly selected ten fields per filter,
141 and in total, more than 300 bacteria were counted using an Olympus BX51 epifluorescence microscope equipped
142 with an oil-immersion objective (UPlanFL 100×/1.30) lens at 100x magnification under blue excitation.

143

144 Ecosystem functioning experiment

145 In this study, we measured the capabilities of processing 31 organic carbon substrates as the index of
146 multifunctionality using the Ecoplate (Biolog, Hayward, CA, USA) (21,32). The Ecoplate is a phenotypic
147 microarray containing triplicate wells for each single carbon substrate and three control wells with no substrate.
148 Each well also contained tetrazolium violet dye, which turned purple when the substrate within the well was
149 catabolized. To conduct this experiment, we inoculated 100 µl of sample water into each well of an Ecoplate and
150 conducted the incubation of all the Ecoplates in a single cooling incubator with 20°C (As One). We incubated the
151 Ecoplates for 14 days to ensure that the color development reached saturation. Color development was measured
152 with an optical density (OD) microplate reader (iMark, Bio-rad) set at 595 nm every day from the 0th day to the

153 14th day.

154 To evaluate ecosystem functioning, the cumulative color development from 1st day to the 14th for each
155 well of the Ecoplate was calculated for assessing the integration of the color density development curve. The
156 resulting integrated value was normalized by dividing it by the integration period (32). Furthermore, to standardize
157 the background turbidity and color development in relation to in situ dissolved organic carbon (DOC), the
158 integrated value of the control well was subtracted from each integrated value of the Ecoplate. After averaging the
159 values among triplicate wells to minimize experimental error, values for 31 different functions were obtained,
160 which were then be used as our EF indices.

161 To estimate the multifunctionality based on 31 EFs derived from the Ecoplate as a proxy of phenotypic
162 ecosystem multifunctionality (MF_P), threshold method (Zavaleta et al. 2010) was applied as

$$MF_{Pj} = \sum_{i=1}^{31} 1 \times [EF_{i,j} > \text{threshold}]$$

163 where $EF_{i,j}$ represents the value for ecosystem function i in a given community of the j^{th} observation. The threshold
164 value corresponds to the 5% of the maximum value among 20 samples (including 7 samples that were excluded for
165 functional gene prediction processes) for each function (Zavaleta et al. 2010).

166

167 DNA extraction, PCR amplification and sequencing

168 We employed the DNeasy Blood & Tissue Kit (QIAGEN, Hilden, Germany) for DNA extraction from filter
169 cartridges, following the protocols outlined by Miya *et al.* (2016) and Ushio (2019). Subsequently, the extracted
170 DNA served as a template for the polymerase chain reaction (PCR) targeting the V4 region (~250 bp) of the 16S
171 rRNA gene, utilizing prokaryotic universal primers (515F by Parada, Needham and Fuhrman 2016 and 806R by
172 (45). To ensure enhanced reproducibility and consistent outcomes, we implemented a two-step PCR approach (46).
173 During both the extraction and PCR processes, we included an extraction negative control and a PCR negative
174 control to monitor potential contamination. Sequencing of the PCR amplicons was conducted on the Illumina
175 Miseq platform, generating 2×300 bp paired-end reads. Additional details on the experimental procedure can be
176 found in the Supplementary Data (Supplementary Methods). The raw sequence data have been deposited in the
177 NCBI Sequence Read Archive under the accession number PRJNA1080231.

178

179 Sequence data processing

180 All pipelines used for processing sequence reads and generating amplicon compositions followed Ushio 2019. In
181 brief, the raw MiSeq data were converted into FASTQ files using the bcl2fastq program provided by Illumina
182 (bcl2fastq v2.18) without demultiplexing, and the FASTQ files were subsequently demultiplexed using Claident
183 v0.2.2018.05.29 (<http://www.claident.org>, (Tanabe and Toju 2013)). Only reads that matched both the Illumina tag
184 and primers were utilized for subsequent bioinformatic processes. The demultiplexed FASTQ files were analyzed
185 using the Amplicon Sequence Variant (ASV) method implemented in DADA2 (v1.26.0) (47). Initially, primer
186 removal was carried out using of the external software cutadapt v2.6 (Martin 2011). Subsequently, sequence quality
187 filtering was performed with the DADA2::filterAndTrim() function, and error rates were determined using the
188 DADA2::learnErrors() function, with the MAX_CONSIST option set to 20. Although the DADA2 algorithm

189 typically processes each sample independently, this default approach tends to remove singletons and doubletons
190 within individual samples, thereby impeding the estimation and standardization of sampling coverage. To address
191 this limitation, we deviated from the default setting and combined all samples for sample inference using dada (...,
192 pool = TRUE)(49). This modification allows the preservation of ASVs that appear once or twice in each sample
193 (i.e., local singletons or doubletons) while eliminating ASVs that appear once or twice only in the pooled samples
194 (i.e., global singletons or doubletons). After the removal of spurious sequences by the DADA2 algorithm,
195 paired-end reads (i.e., overlapping by at least 20 bases) were merged into ASVs. Chimeric sequences were
196 eliminated using the DADA2::removeBimeraDenovo() function. Taxonomic assignments of ASVs were performed
197 using the SILVA database (version 138.1) (Quast et al. 2013). Any ASVs classified as Mitochondria or Chloroplast
198 were subsequently eliminated.

199 The presence of ASVs in the negative controls may suggest potential contamination during the
200 experiments (Supplemental Data). Therefore, the maximum abundance for each ASV observed in these negative
201 controls was calculated and excluded from other samples. Note that if the resulting abundance was greater than
202 zero, we assigned the value as normalized ASV abundance. In cases where the resulting abundance was not greater
203 than zero, we set the normalized ASV abundance to zero.

204 To standardize the sampling coverage of samples on each sampling date, the coverage-based approach
205 (Chao et al. 2014) was employed. Samples with coverage lower than 90% were excluded from further processing
206 (Table S1). Among the remaining 13 samples, the minimum sample coverage (i.e., 93.32%) was calculated, and this
207 fixed coverage was applied to subsample the ASV abundance 100 times for each sampling date. Subsequently, the
208 average of the 100 subsampled ASV tables was utilized to depict the prokaryotic composition of microbial
209 communities. When estimating the abundance of each ASV (cells/mL), we used the frequency distribution of ASVs
210 (0-1) with the total bacterial count (cells/mL).

211

212 Functional gene prediction

213 To predict functional genes, the representative sequence of ASV, along with a BIOM table that excluded ASVs
214 classified as Mitochondria and Chloroplast, was employed in PICRUSt2 (31). The default PICRUSt2 pipeline
215 utilized the Integrated Microbial Genomes (IMG) databased as week as KEGG database to generate KEGG
216 Ortholog (KO) predictions for each input ASV. Each KO entry denotes an ortholog group associated with a gene
217 product in the KEGG pathway diagram (50). Following the PICRUSt2 guidelines, ASVs with a nearest-sequenced
218 taxon index (NSTI) score above 2 are typically considered as poor alignments with existing reference sequence in
219 the IMG. Thus, these ASVs were removed from both KO compositional information and the averaged-rarefied
220 ASV composition.

221 To estimate genomic multifunctionality (MF_G) of a community, we multiplied the composition vector
222 representing the presence/absence of ASVs, derived from the averaged-rarefied ASV composition, with the matrix
223 indicating the presence/absence of KOs. This multiplication, performed as an inner product, resulted in a
224 community-wise KO list in vector form. Subsequently, we calculated the KO richness for each community, serving
225 as a proxy for MF_G .

226

227 ASV-extinction simulations

228 *Setting for ASV-extinction simulations and fitting of power-law curves*

229 We simulated the ASV richness loss as a proxy of taxonomic diversity (TR) loss and its subsequent impact on the
230 KO richness loss as a proxy of genomic multifunctionality (MF_G) loss within each local community. Our first aim
231 was to improve the power-law regression method of this relationship ($MF_G = cTR^a$) for quantifying microbial
232 community multifunctional redundancy. The exponent a indicates the degree of MF_G changes with TR , while the
233 coefficient c represents the expected MF_G with a single taxonomic unit ($TR = 1$). To achieve this, we first adopted
234 the random extinction scenario, assuming a random order of ASV extinctions within a community. This scenario
235 served as a baseline for evaluating multifunctional redundancy and also facilitated the assessment and refinement of
236 the power-law regression method. Rather than numerically simulating sequences of randomly ordered ASV
237 extinctions, we employed an analytical formula for the TR - MF_G relationship (21). Originally designed for a species
238 accumulation curve and its rarefaction (51), this formula allows rapid quantification of the expected KO richness
239 across a sequence, ranging from a single ASV to the maximum ASV richness (e.g., 1, 2, ..., 1000), resulting in a
240 high-resolution TR - MF_G curve. Additionally, we generated a low-resolution TR - MF_G curve with an arithmetic
241 progression of ASV richness at a consistent interval (e.g., every 5% of the maximum ASV richness). It's important
242 to note that a low-resolution TR - MF_G curve is exclusively suitable for non-random extinction scenarios, as the
243 analytical formula is not applicable to them, and it is impractical to simulate KO richness for the entire sequence
244 from a single ASV to the maximum ASV richness. The comparison between high- and low-resolution TR - MF_G
245 curves was expected to guide the development of a more robust methodology for evaluating multifunctional
246 redundancy.

247 Using linear regression on \log_{10} -transformed data ($\log_{10}(MF_G) = \log_{10}(c) + a\log_{10}(TR)$), the regression
248 line tends to overestimate MF_G levels when TR levels are small. This tendency is also indicated in Fig. S3 of Ruhl
249 *et al.* 2022, where log-transformed data, being less dense with smaller TR levels, leads to overestimation. As our
250 focus is on the initial phases of microbial extinctions and their impacts on ecosystem multifunctionality, the bias
251 introduced by severe extinctions (e.g., 95% reduction of the maximum ASV richness) does not significantly
252 compromise the efficacy of the power-law regression method for assessing multifunctional redundancy. However,
253 developing an appropriate fitting procedure is also crucial to ensure quantitative comparability between high- and
254 low-resolution TR - MF_G curves, corresponding to analytical and numerical methods. To determine the specific
255 ranges where power-law fitting on the low-resolution curve aligns reasonably with the high-resolution curve, we
256 conducted a comparative analysis of slope estimates across three intervals: 1) from a single ASV to the maximum
257 ASV richness, 2) from 10% to 100% of the maximum ASV richness, and 3) from 50% to 100% of the maximum
258 ASV richness.

259

260 *Non-random extinction scenarios*

261 When exploring non-random extinction scenarios, we exclusively utilized low-resolution TR - MF_G curves, as the
262 analytical formula is not applicable for them. The ASV richness range selected for the power-law regression was
263 determined through the analysis of the random-extinction scenario. Our investigation focused on two categories of
264 non-random extinction scenarios (Teichert *et al.* 2017), defining survival probability through either KO

265 presence/absence compositional information (four *function-based* scenarios: F1-F4) or abundance distribution
266 information (four *abundance-based* scenarios: A1-A4).

267 In the function-based scenarios, [F1] the *generalists survival* scenario assumed that survival probability is
268 proportional to KO richness, reflecting a plausible situation where ASVs with greater genetic functional richness
269 within a genome (i.e., generalists) are more resilient to environmental fluctuations (52). Conversely, [F2] the *low*
270 *maintenance survival* scenario assumed that survival probability is proportional to the inverse of KO richness,
271 attributed to the smaller maintenance cost of their genomes. [F3] The *niche uniqueness advantage* scenario
272 assumed that ASVs with greater mean dissimilarities (evaluated by Sørensen dissimilarity measure) in KO
273 composition compared to all other ASVs are more likely to survive. This notion is rooted in the theoretical
274 proposition that ASVs functionally dissimilar to others may indicate less overlapping of niches, allowing them to
275 escape competition (53). Conversely, [F4] the *shared niche survival* scenario assumed that survival probability is
276 proportional to the inverse of the averaged dissimilarities in KO composition. While F4 may not be realized in
277 natural environments, it was included in the scenarios to obtain pessimistic (or cautious) estimates of
278 multifunctional redundancy. The anticipation is that scenario F4 would lead to greater impacts of ASV richness loss
279 on MF_G loss compared to other scenarios. This is because functionally unique ASVs are anticipated to go extinct
280 first in F4, leading to a more pronounced effect on multifunctionality.

281 For the abundance-based scenarios, [A1] the *high mean abundance survival* scenario assumed that
282 survival probability is proportional to the mean abundance across sites and time points, indicating a scenario where
283 rare ASVs go extinct first (54,55). [A2] The *low abundance variation survival* scenario assumed that survival
284 probability is proportional to the inverse of the standard deviation of abundance across sites and time points, based
285 on the rationale that more variable ASVs are more likely to go extinct (56,57). [A3] The high *occurrence richness*
286 *survival* scenario assumed that survival probability is proportional to the number of sites and time points in which
287 the ASVs were detected, reflecting a plausible situation where ASVs widely distributed in space and time are less
288 likely to go extinct (58–60). [A4] The *high occurrence Shannon* survival scenario assumed that survival probability
289 is proportional to the hill number ($q = 1$, i.e., the 1st order effective number, Chao, Chiu and Jost (2014)) of sites
290 and time points in which the ASVs were detected. This modification of scenario A3 considers both the occurrence
291 and the variation in abundance.

292

293 *Computational procedures for ASV-extinctions and multifunctional redundancy*

294 To simulate both random and nonrandom extinction scenarios numerically, we utilized the sample() function in R to
295 generate a shuffled sequence of ASV IDs. This sequence was created with uniform (unweighted) sampling
296 probability for random-extinction scenario or weighted sampling probability based on survival probability for
297 non-random extinction scenarios. In simpler terms, the final element of the generated sequence corresponded to the
298 ASV that had gone extinct first while the first element corresponded to the one that had survived until all ASVs had
299 gone extinct.

300

301 Following this sequence, we generated multiple ASV richness levels, covering from a single ASV to
302 100% of the maximum ASV richness, with increments of 5% (1 ASV, 5%, 10%, ... up to 100%), resulting in 21

303 ASV richness levels. We repeated these procedures 100 times by preparing 100 randomly shuffled sequences as
304 100 trajectories of ASV extinctions. From these 100 trajectories, we calculated the mean KO richness at each ASV
305 richness level and fitted the power-law curve using linear regression on \log_{10} -transformed ASV richness and mean
306 KO richness ($\log_{10}(\text{mean KO richness}) = \log_{10}(c) + a\log_{10}(\text{ASV richness})$), where the fitted exponent 'a' served as
307 the index of multifunctional redundancy. A smaller exponent a value indicated higher multifunctional redundancy.
308

309 *Metacommunity assembly scenario*

310 Assuming each local community within the Lake Biwa watershed originated from a shared metacommunity (36),
311 we also employed the $TR\text{-}MF_G$ curve to investigate potential mechanisms influencing variations in
312 multifunctionality across sites and time points. More specifically, we gathered all ASVs from local communities
313 into a common pool as the metacommunity. Subsequently, we systematically increased ASV richness in 5%
314 increments (ranging from 5% to 100% of total ASV richness), including one smaller value (100 ASVs), with
315 following the survival probability defined in the extinction scenarios. Although the processes of extinction and
316 assembly may seem to be opposites, it is essential to note that both the orders of ASV extinction and assembly can
317 be defined by the identical survival probability. As a result, the random and nonrandom extinction scenarios can
318 also be interpreted as the random and nonrandom assembly scenarios, respectively.

319 This iterative process was repeated 200 times under the random assembly scenario, generating $TR\text{-}MF_G$
320 curves. The 95% ranges of KO richness values at each ASV richness level were determined by extracting the 2.5%
321 and 97.5% quantiles from the 200 replications. By overlaying the realized ASV richness and KO richness
322 combinations, we identified local communities that deviated from the 95% range. Furthermore, we created $TR\text{-}MF_G$
323 curves and associated 95% ranges under non-random assembly scenarios (F1-F4 & A1-A4). We examined whether
324 local communities showing deviations from random assembly scenarios aligned with specific non-random
325 assembly scenarios, offering insights into the underlying mechanisms shaping the assembly processes of the focal
326 communities.

327

328 **Computation**

329 All computation processes that are not specified in the sections above were conducted in R (version 4.2.1)
330 (<http://www.r-project.org/>). Specifically, we process sequences with the 'dada2' package (version 1.26.0) (47). We
331 estimated prokaryotic diversity with the 'iNext' package (version 3.0.0)(62). For the linear model analysis, we used
332 lm(), glm(), and step() functions for a simple linear regression, generalized linear model with Poisson distribution,
333 and model selection based on Akaike Information Criterion (AIC). The R notebook as a html file especially for
334 Ecosystem functioning experiment and ASV-extinction simulations and data files for their inputs are available at
335 https://github.com/tksmiki/biwako_redundancy.

336

337 **Result**

338 **Basic information of ASV richness and KO richness**

339 After obtaining the sequencing results from the negative controls, which included 29 and 16 reads and
340 corresponded to 12 and 7 ASVs for the extraction and PCR negative controls, respectively, we normalized the

341 results from other data sets based on these values. Subsequently, we assessed the sampling coverages of the 20 data
342 sets by using their sequence frequency distribution. Among them, 7 data sets exhibited sampling coverages lower
343 than 90%, leading us to exclude them from the estimation of standardized ASV richness and KO richness. The total
344 ASV richness and KO richness for the remaining 13 data sets are represented by 3576 and 7340, respectively, with
345 detailed information provided in **Table S1** and **Figure S1** for each data set. One of the datasets (S11, collected from
346 Lake Biwa on September 10, 2019) initially comprised only 2675 reads before resampling (Table S1). Although
347 this could introduce uncertainties into subsequent analyses, its sampling coverage was sufficiently high at 96%.
348 Additionally, the robustness of the major statistical tests, including those associated with Figure 1 and Figure 2b,
349 persisted even when the results from S11 were excluded.

350

351

352 **Validation and assessment of proposed methods**

353 The generalized linear model (GLM) with a Poisson distribution revealed a significant positive linear relationship
354 between MF_G and MF_P ($\lambda = 0.0004963 * MF_G - 0.1611650$, P value for the coefficient of $MF_G = 1.67e-5$), indicating
355 a positive association between the two variables (**Fig. 1**). Utilizing the step() function to identify the best model
356 explaining the variation in MF_P with potential explanatory variables (ASV richness and MF_G), we observed that
357 only MF_G remained as a significant explanatory variable. This indicated that genomic multifunctionality (MF_G) is a
358 good proxy of phenotypic multifunctionality (MF_P).

359

360 The power-law fitting applied to high- and low-resolution $TR-MF_G$ curves revealed that the estimated
361 intercepts (c) for the entire range of ASV levels (from a single ASV to the maximum ASV) were considerably
362 larger than the expected MF_G with a single ASV (**Fig. 2a** and **Table S2**), indicating a tendency for
363 multifunctionality overestimation when TR is very low. Analyzing the estimates of the exponent (a) across three
364 defined ranges, we found that two intervals, specifically from 10% to 100% and from 50% to 100% of the
365 maximum ASV richness, yielded quantitatively comparable results for a between the high- and low-resolution
366 curves. Considering these findings, we opted to utilize the range from 10% to 100% of the maximum ASV richness
367 for numerical simulations under non-random extinction scenarios, ensuring coverage of a broader interval, provided
368 that the estimates from the low-resolution curves remain comparable to those from the high-resolution curves
369 (within 3% differences, **Table S2**).

370

371 Utilizing estimates obtained from the selected ASV richness ranges (10% - 100% of the maximum ASV
372 richness for each community) under the random extinction scenario, we identified variations in multifunctional
373 redundancy across different sites and sampling dates. Notably, we observed a positive association between these
374 variations and ASV richness (**Fig. 2b**), where a smaller exponent value indicated greater multifunctional
375 redundancy (linear regression, adjusted $R^2 = 0.8838$).

376

377 **Variations in multifunctional redundancy under non-random extinction scenarios**

378 The redundancy exponent (a) exhibited substantial variations across non-random extinction scenarios (**Fig. 3**, **Fig.**

379 4), although the ranking order of the exponent among sites and dates remained generally consistent between
380 scenarios (Fig. 4a). In comparison to redundancy estimated from random-extinction scenarios, F1 and F3
381 demonstrated higher redundancy (i.e., smaller exponent), while F2, F4, and A1 exhibited lower redundancy (i.e.,
382 greater exponent)(Fig. 4b). Notably, scenario F2, representing the *low maintenance survival* scenario, estimated the
383 lowest redundancy, reflected by the highest exponent. The other scenarios did not exhibit a clear trend in
384 redundancy estimation.

385

386 Comparison of multifunctionality through community assembly from a shared metacommunity

387 We observed that the majority of communities within the metacommunity can be explained by the random
388 assembly scenario, as indicated by the 95% range. Nevertheless, four local communities (Biwa on 0703, 0730, and
389 0910, and Echi on 1015) deviated from this95% range of random assembly (Fig. 5). One of them (Echi on 1015)
390 could be explained by F1 and F3. Additionally, the deviations of two communities (Biwa on 0703 and 0730,) were
391 explainable by scenario F2, with one of them (Biwa on0703) also aligning with scenario F4 (but at the boundary of
392 the 95% range). One local community (Yasu on 0709) also fell into A2 (Fig. S2). One local community (Biwa on
393 0910) did not fall into any scenarios (random, F1-F4, and A1-A4).

394

395 Discussion

396 Advancements in Methodology for Assessing Multifunctional Redundancy

397 The primary objective of our study was to refine the methodology for assessing multifunctional redundancy,
398 resulting in three key findings with significant implications.

399

400 Firstly, the implementation of PICRUS2 was indirectly validated through a comparison of KO richness
401 predictions with phenotypic multifunctionality assessed by ecoplate incubations (Fig. 1). Unlike the previous study
402 (21), our analysis covered a wide bacterial richness range (from 312 to 2715 ASVs), demonstrating the practical
403 effectiveness of our proposed procedure starting from amplicon sequences. Model selection further revealed that
404 KO richness serves as a superior predictor of phenotypic multifunctionality compared to ASV richness. Although
405 direct metagenomic sequencing for validation is acknowledged, these results highlight the practical efficacy of our
406 proposed approach for obtaining functional information from amplicon sequences.

407

408 Secondly, our detailed analysis of the power-law regression method led to two crucial procedural
409 enhancements. Firstly, we recommended setting the focal range for power-law regression as 10% to 100% of the
410 maximum ASV richness to avoid dependence on the resolution of ASV richness. This adjustment ensures that the
411 estimate is robust across different resolutions. Additionally, we proposed conducting low-resolution simulations
412 (e.g., every 5% of the maximum ASV richness) due to computational efficiency, while confirming that power-law
413 fittings on low-resolution $TR-MF_G$ curves are quantitatively equivalent to those on high-resolution curves. The
414 power-law regression tends to overestimate MF_G when ASV richness is low, limiting its utility for predicting the
415 impacts of severe richness reduction on ecosystem multifunctionality. When predicting such impacts, we
416 recommend directly using the results from extinction simulations instead of relying on power-law fitting estimates.

417

418 Lastly, concerning extinction scenarios, we proposed that random-extinction, generalists survival, and
419 low maintenance survival scenarios represent the minimum requirements to cover the possible range of functional
420 redundancy estimates (**Fig. 4**). The random-extinction scenario acts as the baseline, while the generalists survival
421 and low maintenance survival scenarios provide the highest and lowest redundancy estimates (smallest and greatest
422 exponent values), respectively, among all scenarios. This comprehensive approach ensures a thorough
423 understanding of the potential range of functional redundancy estimates under different extinction scenarios.

424

425 Comparative Analysis of Lake Biwa and Inlet Rivers: Multifunctional Redundancy Dynamics

426 The secondary objective of our study was to apply the refined methodology to multiple systems in the Lake Biwa
427 watershed, elucidating spatio-temporal variability in multifunctional redundancy.

428

429 Three major findings emerge from our analysis: Firstly, Lake Biwa exhibited smaller bacterial richness
430 compared to any of the four inlet rivers, irrespective of the time points (**Fig. 2**). One might argue that the observed
431 lower richness in Lake Biwa could be attributed to a smaller number of sequences or lower bacterial abundance.
432 However, the former is not supported as the sampling coverages were high enough (> 96%, as shown in **Table S1**).
433 The latter is also not supported because ASV richness in rivers with comparable bacterial abundance was much
434 greater, and temporal variations in ASV richness were much smaller than those of bacterial abundance (**Fig. S3**).
435 Instead, this result aligns with the established notion that high richness is likely linked to elevated environmental
436 and spatial heterogeneity, providing multiple niches for diverse microbes to inhabit in rivers (63). Although
437 environmental variables were not directly measured, this hypothesis gains indirect support from the relationship
438 between MF_G and MF_P (Fig 1), indicating higher functional diversity in rivers compared to lakes.

439

440 Secondly, we observed a negative association between bacterial richness and the exponent of
441 multifunctional redundancy, suggesting reduced multifunctional redundancy with a decline in bacterial richness.
442 (**Fig. 2b**). As Lake Biwa exhibited smaller bacterial richness and redundancy compared to any of its inlet rivers,
443 irrespective of the time points (**Fig. 2b**), it highlights a higher conservation priority for Lake Biwa compared to the
444 four inlet rivers, at least concerning prokaryote-mediated ecosystem multifunctionality. Conversely, while it
445 appeared that the Echi River had the lowest redundancy (greatest exponent) on July 9 and October 15, all four
446 rivers displayed significant fluctuations in both bacterial richness and multifunctional redundancy across different
447 time points. This complexity makes it challenging to definitively identify the river most vulnerable to disturbances
448 and losses in bacterial taxonomic richness, affecting ecosystem multifunctionality.

449

450 The challenge in prioritizing ecosystem vulnerability between four rivers (Ane, Echi, Hino, and Yasu)
451 may stem from the non-intensive sampling design (only four time points from each river). Additionally, the lack of
452 clear differences between these rivers could be influenced by random assembly processes, suggested by the
453 metacommunity scenario (**Fig. 5**). Most bacterial communities in these rivers seemed to follow a random-assembly
454 scenario. If true, significant differences in terms of multifunctional redundancy between these four rivers may not

455 exist. To confirm the presence or absence of real differences, more frequent sampling with additional time points
456 and multiple locations within a river is necessary (64). Additionally, expanding the study to encompass a broader
457 selection of inlet rivers is crucial, given the presence of 117 first-class inlet rivers in the Lake Biwa watershed.
458

459 Finally, the results suggest that Lake Biwa was subject to different assembly processes compared to its
460 inlet rivers (**Fig. 5**). While most of the rivers could be explained by random or generalist survival scenarios, Lake
461 Biwa aligns more closely with a low maintenance survival scenario, indicating that generalists are less likely to
462 persist. Such contrasting results may be attributed to the distinct environments of rivers and lakes (65). For example,
463 the running water in the inlet rivers is highly dynamic and fluctuating, which would favor generalists more because
464 they can adapt to various conditions (66–70). In contrast, pelagic water in a large lake exhibits characteristics of a
465 more stable environment. Therefore, the high maintenance cost for multiple genes might be a challenge for species
466 to survive in lakes, resulting in the survival of specialists with low maintenance costs.
467
468

469 Advancing methodology: underexplored aspects and future directions

470 In this section, we highlight three underexplored aspects of our proposed method, each serving as potential avenues
471 for future research.
472

473 Firstly, our initial expectation that the scenario F4 would lead to lower multifunctional redundancy
474 compared to the scenario F2 was contradicted by the results (**Fig. 4**). While F4 assumes that ASVs with greater
475 functional dissimilarity go extinct earlier, resulting in a greater negative impact on community-wide KO richness
476 than F2, the opposite trend was observed. A simple linear association between the indices for F2 (inverse of KO
477 richness) and F4 (inverse of mean dissimilarity) was not evident (**Fig. S4a**). One potential refinement could involve
478 a more direct method for defining functional uniqueness, such as considering the number of unique KOs not shared
479 with other ASVs. However, our exploration revealed only 127 ASVs with such unique KOs, and the maximum
480 number of unique KOs per ASV was limited to 16 (not shown). This indicates that such an index may not be
481 effective in weighting extinction probability across 3576 ASVs. In the scenario F2, we counted the number of
482 different KOs without considering multiple copies, but an alternative approach could involve considering these
483 copies as well. Future investigations should also explore alternative extinction scenarios capable of yielding lower
484 multifunctional redundancy (greater exponent a) than those from F2 and F4. Conservative estimates of redundancy
485 are crucial for assessing ecosystem vulnerability, particularly when faced with substantial uncertainty. This
486 challenge is inherent because we primarily rely on genomic information, including PICRUSt2 predictions (as in
487 this study) or metagenome-assembled genomes (71,72), as well as abundance distribution information, but have
488 limited access to direct physiological information of unculturable bacteria from environmental samples. Further
489 refinement of extinction scenarios and indices is warranted to enhance the accuracy and reliability of redundancy
490 assessments in the face of such complexities.
491

492 Second, the comparison of redundancy exponents between this study and past studies (21,29) provides

493 insights into the robustness of the proposed method and offers cautions for its application to other systems. The
494 exponent (a) values estimated in Miki, Yokokawa and Matsui (2014) ranged from 0.55 to 0.75, but these may have
495 been overestimated as they were derived from pseudo-communities containing a single strain from each genus. In
496 natural environments, where multiple strains, OTUs, or ASVs from each genus coexist in local communities,
497 functional overlap within communities is likely higher. When applying a power-law regression directly fitting on
498 multiple datasets, our results yielded an exponent (a) of 0.15799 (adjusted $R^2 = 0.9022$) (**Fig. S5**), comparable to
499 the value ($a = 0.1338$) presented in Fig.3d of Ruhl et al. 2022. However, it is important to note that our study used
500 KO richness as a proxy for genomic multifunctionality, while Ruhl et al. used Pfam richness. Additionally, the unit
501 of taxonomic richness (OTUs or ASVs) also plays a crucial role, as highlighted by Ruhl et al. 2022, indicating that
502 the estimated exponent highly depends on both the proxy of genomic multifunctionality and the taxonomic richness
503 unit (Fig. 3 of Ruhl *et al.* 2022).

504

505 The third aspect, while not directly aligned with the main objectives of this study, offers the potential for
506 a more in-depth analysis of the relationship between the predicted KO composition, its richness, and the
507 spatio-temporal distribution of each ASV in the watershed. In line with the exploration of utilizing genomic
508 functional information for predicting bacterial occurrence patterns, as proposed by studies like Barberán *et al.*
509 (2014), we also found a positive yet weak association between KO richness (indicative of functional generalization)
510 and both mean abundance (**Fig. S4b**, linear regression on log10-transformed mean abundance, slope = 4.589e-04 (P
511 < 2.0e-16), adjusted $R^2 = 0.04647$) and the number of occurrences (**Fig. S4c**, linear regression, slope = 0.0020760
512 ($P < 2.0e-16$), adjusted $R^2 = 0.0632$). While KO richness of ASVs is not directly linked to their genome size, these
513 patterns stand in contrast to observations in the surface ocean and lakes, where genome-streamlined groups tend to
514 dominate (74–76). To maintain the simplicity of the multifunctional redundancy assessment procedure, we
515 exclusively utilized KO richness without incorporating specific functional information for each KO. However,
516 introducing functional details of KO composition into the procedure could yield two significant advantages: 1)
517 refining the potential scenarios of bacterial richness loss, thereby reducing uncertainty in the estimated redundancy
518 exponent, and 2) providing more detailed assessments of functional redundancy, such as redundancy within specific
519 functional categories like carbon metabolism and nitrogen metabolism. In the context of this study's primary
520 objectives, maintaining simplicity in the assessment of multifunctional redundancy was prioritized. However, for
521 researchers willing to employ more complex procedures, exploring predictions of growth rate and interspecific
522 interactions through metabolic network-based reverse ecology methods, such as flux balance analysis (FBA), can
523 be a promising avenue. It's worth acknowledging that these tools are still evolving in their development (77,78).

524

525 In light of these underexplored features and the current limitations in the size of amplicon sequence data,
526 future investigations may prioritize refining our methodology, addressing potential complexities, and collecting
527 spatially and temporally high-resolution datasets. The crucial refinement of our methodology is necessary for
528 accurately assessing the vulnerability of ecosystem functions, particularly those mediated by microorganisms, and
529 will contribute to the development of effective strategies for ecosystem management and conservation.

530

531 **Acknowledgements**

532 This study was supported by Center for Ecological Research, Kyoto University, a Joint Usage / Research Center.
533 We appreciate Captain Dr. Y. Goda and Vice-Captain T. Akatsuka of the Research vessel Hasu for field sampling in
534 Lake Biwa, and R. Nakamura and A. Matsuda for helping field samplings in the rivers.

535

536 **Funding**

537 T.M., H.Y., and S.N. were supported by Grant for Environmental Research Projects, Sumitomo Foundation. T.M.
538 and T.Y were supported by JSPS KAKENHI, Grant-in-Aid for Scientific Research (S) (19H05667). T.M. and K.Y
539 were supported by JSPS KAKENHI, Grant-in-Aid for Scientific Research (A) (23H00538). T.M and N.S. were
540 supported by JSPS KAKENHI, Grant-in-Aid for Scientific Research (B) (19H03302). T.M. was also supported by
541 JSPS KAKENHI, Grant-in-Aid for Scientific Research (A) (19H00956) and the Alexander von Humboldt
542 fellowship.

543

544 **Authors' contributions:** WHC and TM leads data analysis, design and write the manuscript, MI leadingly
545 conducted sampling and experiments, KY, KM, and HY helped experimental procedures. TM, KM, TY, HY, and
546 SN developed the whole picture of the project in the lake Biwa watershed and contribute to the improvement of the
547 manuscript.

548

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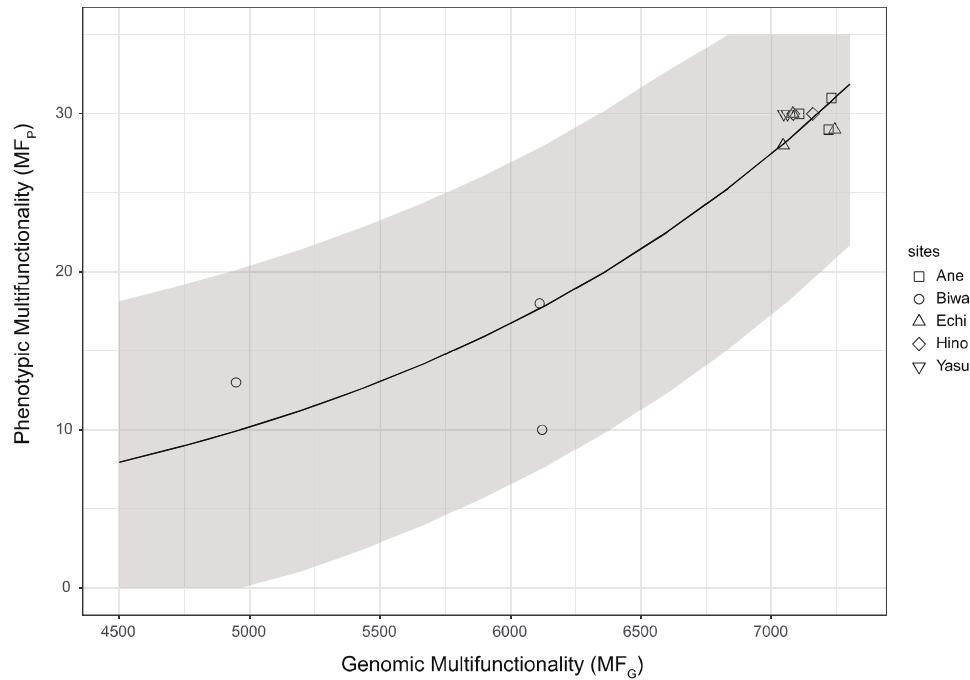
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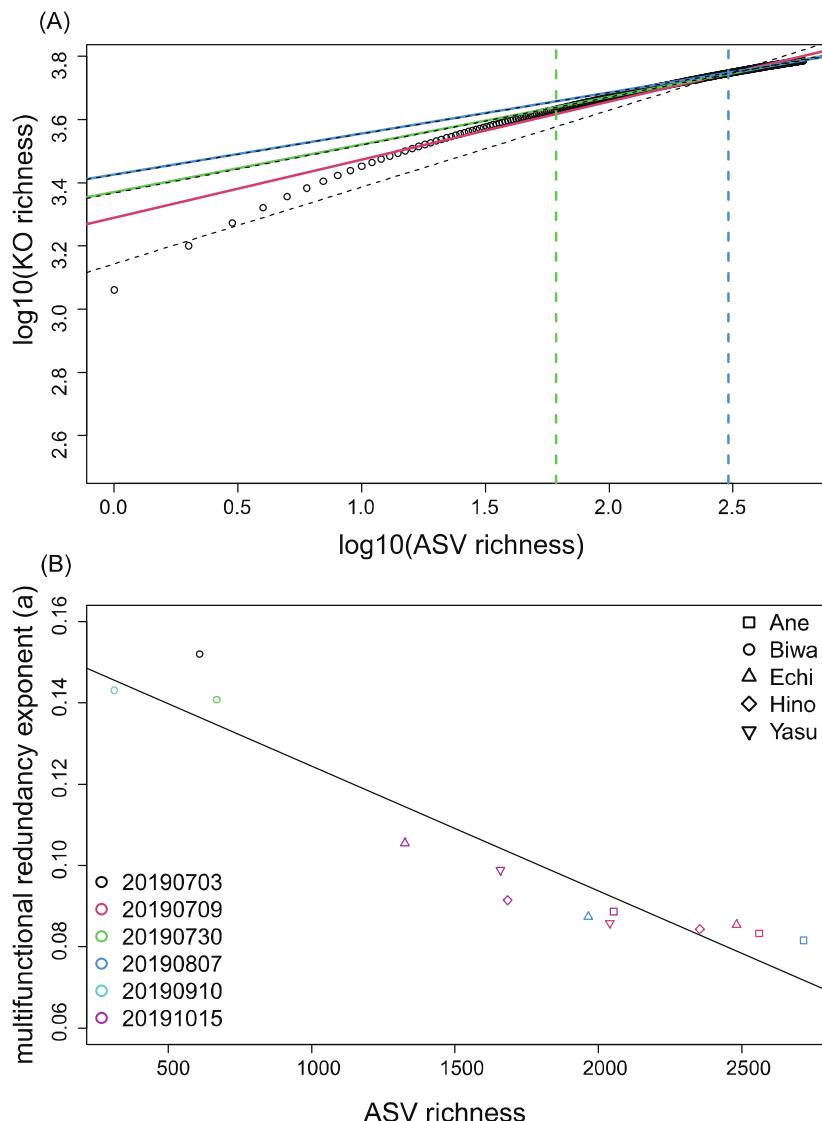
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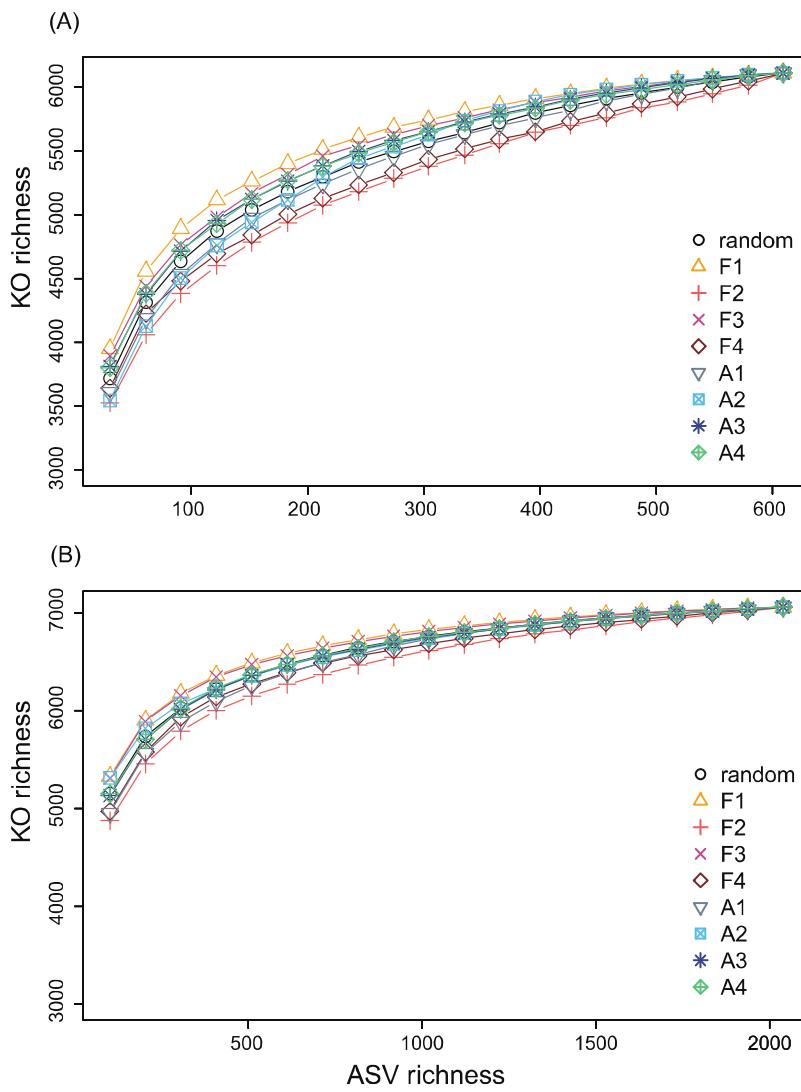


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725 **Figure 1:** Positive association between MF_G and MF_P . GLM with Poisson distribution demonstrated the
726 statistically positive association between genomic multifunctionality index (MF_G) and phenotypic
727 multifunctionality index (MF_P) evaluated by Ecoplate incubation experiment. The line and shaded region represent
728 the regression line and $\pm 2\sigma$ ranges, respectively.



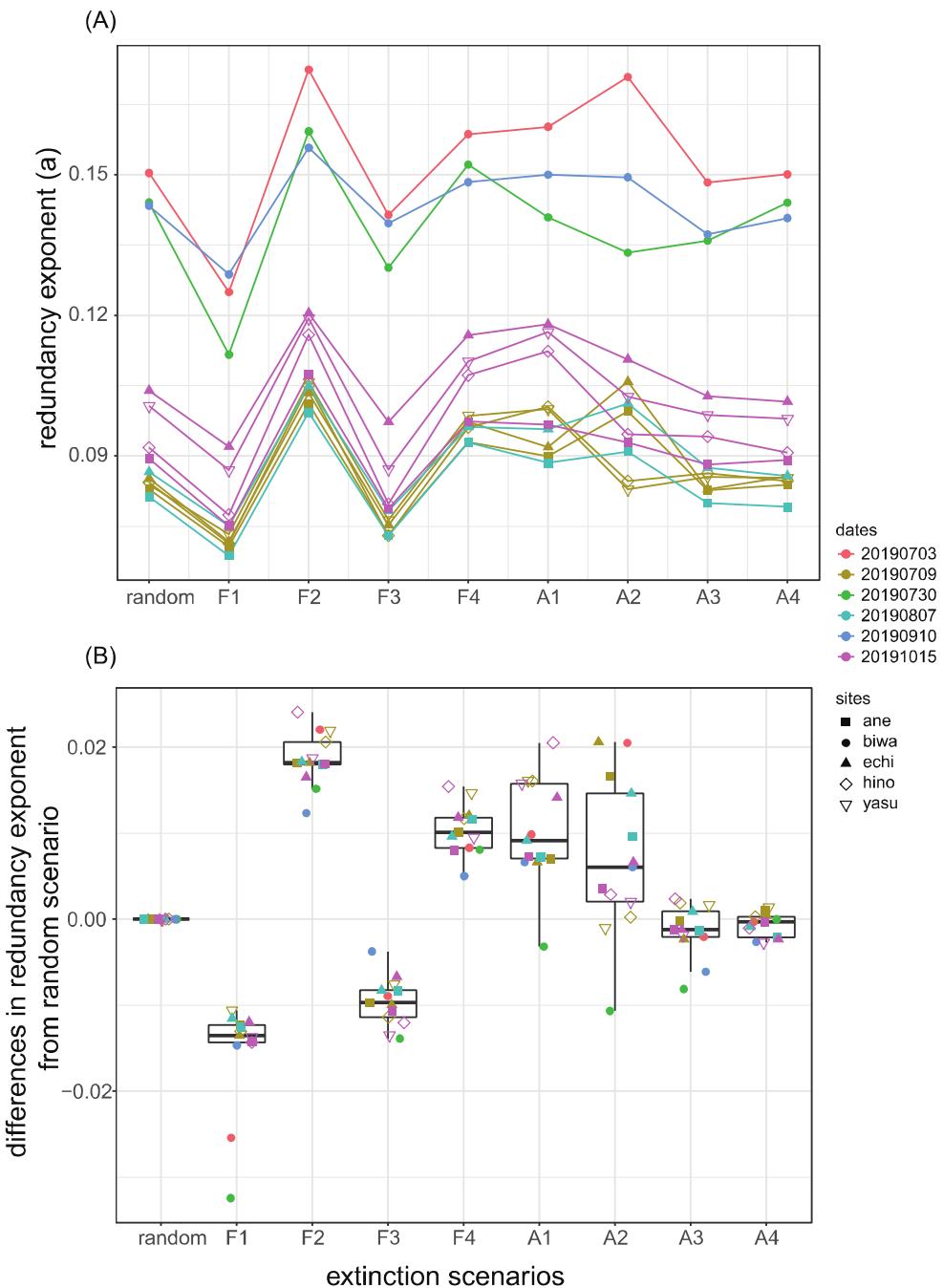
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730 **Figure 2:** (A) Example of $TR\text{-}MF_G$ curve and its power-law regression from Lake Biwa on July 03. The points
 731 represent the analytical estimated expected KO richness for every different level of ASV richness, i.e., the
 732 high-resolution $TR\text{-}MF_G$ curve. The red, light green, and light blue solids lines (or the dashed lines) represent the
 733 power-law regression on the high-resolution curve (or the low-resolution curve) from 1 ASV to 100% maximum
 734 ASV richness, 10% to 100% maximum ASV richness, and 50% to 100% maximum ASV richness, respectively.
 735 The low-resolution curve is just a subset of the high-resolution curve. Except for the fitting from 1 ASV to 100%
 736 maximum ASV richness, low-resolution regressions showed quantitatively comparable results with high-resolution
 737 regressions. The vertical light green and light blue dashed lines represent the 10% and 50 % maximum ASV
 738 richness levels, respectively. (B) Relationship between ASV richness and multifunctional redundancy exponent
 739 from the random-extinction scenario. The negative association between ASV richness and multifunctional
 740 redundancy exponent indicates the positive association between ASV richness and the magnitude of multifunctional
 741 redundancy.



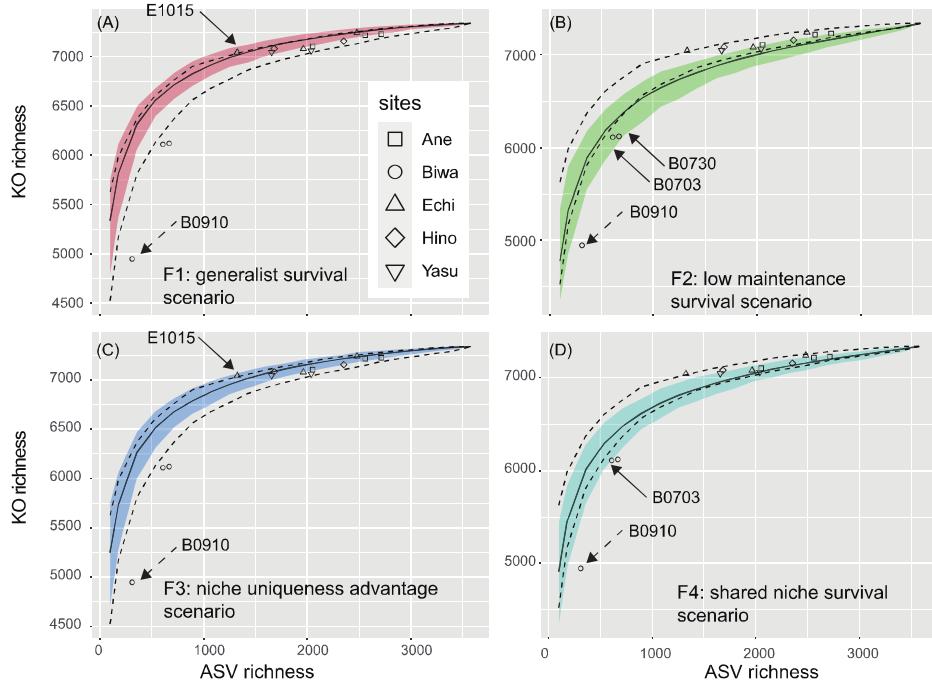
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743 **Figure 3:** Examples of TR-MFG curves from all scenarios. (A) The dependence of KO richness reduction along
744 with different ASV extinction scenarios for the community collected on July 03 from Lake Biwa. (B) The
745 dependence of KO richness reduction along with different ASV extinction scenarios for the community collected
746 on July 09 from Yasu river.



747

748 **Figure 4:** Variations in redundancy exponents between scenarios. Summary of the dependency of the redundancy
749 exponent (a) on extinction scenarios, sites, and time points. (a) The direct comparison of the redundancy exponent;
750 the greater exponent value represents lower multifunctional redundancy. (b) The differences between the exponent
751 values between those from the random extinction scenario and those from eight non-random extinction scenarios.
752 The positive (or negative) values represent the lower (or higher) multifunctional redundancy under the focal
753 non-random extinction scenario than that under the random-extinction scenario.



755 **Figure 5: Deviation from random-assembly from the metacommunity for function-based scenarios (F1-F4).**

756 The variations in ASV richness and KO richness between 13 communities and their relationship with the
757 random-assembly (i.e. random-extinction) scenarios' 95 % confidence interval from 200 repeated simulations (the
758 range between the dashed lines) and with the 95% confidence interval (color-shaded region) of one of the four
759 function-based nonrandom-extinction scenarios (F1-F4). Communities indicated by the solid arrows (i.e., B0703,
760 B0730, and E1015) correspond to the communities from Lake Biwa on July 3, July 30, and from Echi river on
761 October 15, respectively. These communities cannot be explained by random assembly processes but can be
762 attributed to a part of the non-random assembly scenarios (F1-F4). Community indicated by dash arrow (i.e.,
763 B0910, Community from Lake Biwa on September 10) cannot be explained by any scenarios (random, F1-F4, and
764 A1-A4).