

# 1    Estimating recent and historical effective population size 2    of marine and freshwater sticklebacks

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11   Running head: Stickleback Effective Population Size

## 12   **Abstract**

13   Effective population size ( $N_e$ ) is a quantity of central importance in evolutionary biology and  
14   population genetics, but often notoriously challenging to estimate. Analyses of  $N_e$  are further  
15   complicated by the many interpretations of the concept and the alternative approaches to  
16   quantify  $N_e$  utilising different properties of the data. However, the alternative methods are  
17   also informative over different time scales, suggesting that a combination of approaches  
18   should allow piecing together the entire continuum of  $N_e$ , spanning from the recent to more  
19   distant past. To test this in practice, we inferred the  $N_e$  continuum for 45 populations of nine-  
20   spined sticklebacks (*Pungitius pungitius*) using whole-genome data with both LD- and  
21   coalescent-based methods. Our results show that marine populations exhibit the highest  $N_e$   
22   values in contemporary, recent, and historical times, followed by coastal and freshwater  
23   populations. They also demonstrate the impact of both recent and historical gene flow on  $N_e$   
24   estimates and show that simple summary statistics are informative in comprehending the  
25   events in the very recent past and aid in more accurate estimation of  $N_e^C$ , the contemporary  
26    $N_e$ , as well as in reconstruction and interpretation of recent demographic histories. Although  
27   our sample size for large populations is limited, we found that GONE can provide reasonable  
28    $N_e$  estimates. However, due to challenges in detecting subtle genetic drift in large  
29   populations, these estimates may represent the lower bound of  $N_e$ . Finally, we show that  
30   combining GONE and CurrentNe2, both sensitive to population structure, with MSMC2  
31   provides a meaningful interpretation of  $N_e$  dynamics over time.

32

33   Keywords: admixture, demographic history, effective population size, nine-spined  
34   stickleback, *Pungitius*

35

## 36 Introduction

37 By quantifying the magnitude of genetic drift and inbreeding in real-world populations, the  
38 concept of effective population size ( $N_e$ ) has numerous applications in both evolutionary  
39 (Charlesworth, 2009; Charlesworth & Charlesworth, 2010) and conservation biology  
40 (Frankham et al., 2010; Allendorf et al., 2012). In evolutionary biology,  $N_e$  is informative  
41 about the efficacy of selection, mutation, and gene flow as systematic evolutionary forces  
42 (Charlesworth & Charlesworth, 2010). In conservation biology, it serves as a crucial measure  
43 of a population's evolutionary potential and long-term viability (Frankham et al., 2010; Hare  
44 et al., 2011; Waples, 2022a). Various genetic methods have been developed to estimate  $N_e$ ,  
45 ranging from approaches that infer long-term historical  $N_e$  dynamics (reviewed in Beichman  
46 et al., 2018; Nadachowska-Brzyska et al., 2022) to those designed for estimating  
47 contemporary  $N_e$  ( $N_e^C$ ; Waples & Do, 2010). While significant methodological progress has  
48 been made, challenges remain in reconciling different estimators and integrating recent,  
49 historical and contemporary  $N_e$  estimates into a full  $N_e$  continuum, shedding light on how past  
50 demographic events shape present-day genetic diversity (Nadachowska-Brzyska et al.,  
51 2022). Of the distinct temporal scales, estimating  $N_e^C$  for large populations remains  
52 particularly challenging as their genetic drift and inbreeding effects are small and detection of  
53 any signal requires extensive sample sizes (Waples et al., 2016; Marandell et al., 2019;  
54 Waples, 2024).

55 In addition to analyses, interpretation of the  $N_e$  estimates remains complex (Waples 2024).  
56 All methods used to estimate  $N_e$  from genetic data make assumptions, and violation of these  
57 assumptions may lead to errors and biases (Beichman et al., 2018; Nadachowska-Brzyska  
58 et al., 2022, Waples 2024). For instance, many  $N_e$  estimation approaches assume  
59 populations to be closed and panmictic (e.g., Waples & Do, 2010; but see also Palamara  
60 and Pe'er 2013; Santiago et al., 2020; Novo et al., 2023a), while in real life, most populations  
61 are structured and affected by at least some level of migration. The methods based on  
62 sequential Markovian Coalescent (Li & Durbin, 2011; Schiffels & Durbin, 2014) commonly  
63 used to estimate dynamics of the historical  $N_e$  are not immune to these effects, and changes  
64 in gene flow can yield  $N_e$  trajectories that mimic changes in population size (Beichman et al.,  
65 2018). Empirical studies examining how introgression and population structure affect  $N_e$   
66 estimates across different temporal scales are limited (but see Palamara & Pe'er 2013;  
67 Kersten et al., 2023 and Li et al., 2024), highlighting a gap in our understanding of these  
68 dynamics.

69 Marine and freshwater fish populations exhibit distinct genetic variability patterns and  
70 demographic histories and offer a compelling case study for investigating the challenges in  
71  $N_e$  estimation. Freshwater populations generally harbor less genetic variation than marine  
72 populations (e.g., Ward et al., 1994; DeWoody & Avise, 2000; DeFaveri & Merilä, 2015;  
73 Kivikoski et al., 2023), indicating smaller long-term  $N_e$  ( $N_e^{LT}$ ; Ellegren and Galtier, 2016). This  
74 makes intuitive sense due to the limited size and fragmentation of freshwater habitats as  
75 compared to more continuous marine environments. However, estimating contemporary  $N_e$   
76 for large marine populations is particularly difficult (cf. Waples et al., 2016; Marandell et al.,  
77 2019) and rigorous comparisons of marine and freshwater populations are rare (e.g.,  
78 DeFaveri & Merilä, 2015). Recent development in Linkage Disequilibrium (LD) -based  
79 methods (Santiago et al., 2020 & 2024) for the estimation of temporal changes in  $N_e$  might

80 bring a solution for this, and the new methods have been shown to provide robust estimates  
81 even for populations of relatively large  $N_e$  (Santiago et al., 2020 & 2024; Kersten et al., 2023;  
82 Atmore et al., 2022 & 2024; Andrews et al., 2024). However, LD-based approaches have  
83 biases and limitations of their own (Santiago et al., 2020, Novo et al., 2023a & b, Waples  
84 2024, Gargiulo et al., 2024), particularly their assumption of population isolation that is rarely  
85 met in natural populations.

86 The primary goal of this study was to explore the capability and feasibility of using  
87 approaches designed to infer recent and historical changes in  $N_e$  to reconstruct the temporal  
88 dynamics of  $N_e$  of natural populations from different ecological contexts. We focused on  
89 nine-spined stickleback (*Pungitius pungitius*), a small euryhaline teleost fish with a  
90 circumpolar distribution. The species is present in both open marine and landlocked  
91 freshwater habitats and previous studies have found evidence for ancient and potentially  
92 ongoing introgression in some parts of its distribution range (Guo et al., 2019; Feng et al.,  
93 2022; Wang et al., 2023). Using whole-genome sequences ( $n = 888$ ) obtained from Feng et  
94 al. (2022), we estimated both recent (1–100 generations ago) and historical  $N_e$  (hundreds to  
95 thousands of generations ago) across 45 marine and freshwater populations originating from  
96 localities with varying connectivity and environmental conditions, representing open outbred  
97 marine to semi- and fully-isolated inbred freshwater populations.

98 During the analyses, we observed isolated cases of population structure, prompting us to  
99 examine how the violations of panmixia influence the  $N_e$  estimates and impact on associated  
100 demographic interpretations. More systematically, we leveraged data on admixture between  
101 two divergent stickleback lineages (Feng et al. 2022) and investigated how introgression—  
102 present at varying levels across populations— influences the temporal dynamics of  $N_e$ . By  
103 integrating multiple  $N_e$  estimators across populations with distinct histories, our study  
104 provides complementary insights into reconstruction and interpretation of demographic  
105 changes across diverse ecological contexts, as well as the impacts of population structure  
106 and gene flow on  $N_e$  estimation.

107

## 108 Methods

### 109 Data Acquisition

110 The sequence data and admixture proportions used in this study were obtained from a  
111 previous study (Feng et al., 2022) and in accordance with the national legislation of the  
112 countries concerned. In brief, the data used in this study originate from 45 *P. pungitius*  
113 populations covering much of the species distribution area in Eurasia, North America and the  
114 Far East (Table S1 and Fig. S1). Of these populations, 12 were from marine and nine from  
115 coastal freshwater populations with connection (or recent connection) to the sea. Of the true  
116 freshwater populations, eleven were from closed ponds (surface area < 4 ha), ten from  
117 lakes, two from rivers and one from a man-made drainage ditch (Table S1). The admixture  
118 proportions of nine admixed marine and seven admixed freshwater populations (Table S1)  
119 from the Baltic Sea area were derived from Feng et al. (2022).

## 120 Data processing

121 The short-read data were mapped to the 21 linkage groups (LG) of the v7 nine-spined  
122 stickleback reference genome (Kivikoski et al., 2021) using the Burrows-Wheeler Aligner  
123 v.0.7.17 (BWA MEM algorithm; Li, 2013) and its default parameters. Duplicate reads were  
124 marked with samtools v.1.7 (Li et al., 2009) and variant calling was performed with the  
125 Genome Analysis Toolkit (GATK) v.3.6.0 and v.4.0.1.2 (McKenna et al., 2010) following the  
126 GATK Best Practices workflows. In more detail, RealignerTargetCreator and IndelRealigner  
127 (from v.3.6.0) tools were applied to realign reads around indels, HaplotypeCaller was used to  
128 call variants for each individual (parameters set as -stand emit conf 3, -stand call cof 10, -  
129 GQB (10,50), variant index type linear and variant index parameter 128000), and finally  
130 GenotypeGVCFs was used to jointly call the variants for all the samples using its default  
131 parameters. Binary SNPs were extracted with bcftools v.1.7 (Danecek et al., 2021) excluding  
132 sites located within identified repetitive sequences (Varadharajan et al., 2019) and negative  
133 mappability mask regions combining the identified repeats and unmappable regions  
134 (Kivikoski et al., 2021). Sites showing low (<8x) or too high (>25x) average coverage, low  
135 (<20) genotype quality, low (<30) quality score and more than 25% missing data were  
136 filtered out using vcftools v.0.1.5 (Danecek et al., 2011). Data from the known sex  
137 chromosomes (LG12) were removed from further analysis. For details about the subsequent  
138 filtering of the dataset used in different analyses, see Table S2.

## 139 Analyses of linkage disequilibrium and genetic relatedness

140 The magnitude of linkage disequilibrium (LD) and its decay are informative on  $N_e$ , level of  
141 inbreeding, and migration (Flint-Garcia et al., 2003). Hence, we characterised LD patterns in  
142 all populations by estimating the squared correlation coefficient  $r^2$  between each pair of  
143 SNPs with PopLDdecay (Zhang et al., 2019) with its default settings and max distance  
144 between two SNPs were set to 1000Kb. We restricted the analysis to the largest linkage  
145 group LG4 and used LG1 for cross-validation (SNP Set 2 in Table S2). The LD decay curve  
146 was plotted with R (R Core Team, 2020).

147 High levels of LD in a population may indicate (i) small  $N_e$ , (ii) increased inbreeding, and/or  
148 (iii) recent migration/admixture. To distinguish between these, we first estimated the average  
149 inbreeding coefficients ( $F_{IS}$ ) for each population using vcftools --het. We then used  
150 ngsRelate v.2 (Hanghøj et al., 2019) to calculate the  $r_{xy}$ , the pairwise relatedness within  
151 populations (Hedrick & Lacy, 2015). As a measure of temporal gene flow, we examined the  
152 LD decay patterns. Within the same ecotype, populations showing atypical LD decay  
153 patterns were considered as potentially affected by temporal gene flow.

## 154 GONE analyses

155 We reconstructed the recent demographic history using GONE (Santiago et al., 2020). This  
156 method utilises the LD patterns in the data and has been shown to be robust for time spans  
157 of 0-200 generations before present, even when  $N_e$  is relatively large (Santiago et al., 2020).  
158 The analyses were performed using the SNP Set1 (Table S2 & S3) along with a genetic map  
159 lifted-over from the reference genome version 6 (Varadharajan et al., 2019) to version 7  
160 (Kivikoski et al., 2021). According to Santiago et al. (2020), the possible bias from recent

161 gene flow can be mitigated by lowering the recombination fractions threshold ( $hc$ ). Following  
162 that, we repeated the analyses using a  $hc$  of 0.01 and 0.05. For each population, twenty  
163 replicates using the SNP Set1 (Table S2 & S3) were performed with the default settings, and  
164 the geometric mean were applied to summarize the  $N_e$  estimates across replicates. The  
165 replicates are not fully independent due to overlapping SNPs. We used a generation length  
166 of two years (DeFaveri et al., 2014) to scale the time and the  $N_e$  estimates at 1 generation  
167 before present was taken as the estimate of  $N_e^C (N_e^C_{GONE})$ . Following the patterns shown in  
168 Fig. 2F of Santiago et al. (2020), we manually inspected and classified the trajectories as  
169 affected by gene flow or not.

## 170 CurrentNe analyses

171 Although the GONE estimates for the most recent generation can be considered measures  
172 of  $N_e^C$ , a newer method called CurrentNe has been shown to provide less biased estimates  
173 (Santiago et al., 2024). CurrentNe estimates  $N_e^C$  from the LD patterns between SNPs  
174 without requiring information on their location (i.e. genetic map) and accounts for the  
175 species' mating system. The latest version of the method, CurrentNe2 (available at  
176 <https://github.com/esrud/currentne2>), allows for the integration of a genetic map and  
177 modelling of gene flow, improving the accuracy of  $N_e^C$  estimation. When the migration option  
178 (-x) is enabled, CurrentNe2 assumes the target population to be a metapopulation consisting  
179 of two subpopulations of equal size and incorporates migration between them into the  $N_e$   
180 estimation. This approach can yield a more appropriate metapopulation estimate in cases  
181 where gene flow or population substructure are present (A. Caballero, pers. comm.). We  
182 applied CurrentNe2 to estimate the  $N_e^C_{CurrentNe}$  and compared that with  $N_e^C_{GONE}$ . In each  
183 population, autosomal SNPs with missing data were removed from the analyses, and the  
184 average number of full siblings (k) was empirically inferred from the data. The analysis was  
185 performed with and without the migration option, and when population substructure was  
186 present, the estimate allowing for migration was taken. The number of SNPs used in this  
187 analysis for each population is detailed in Table S4.

## 188 MSMC2 analyses

189 MSMC2 (Malaspinas et al., 2016) was used to reconstruct the demographic history of the  
190 more distant past. As the method can analyse at most eight haplotypes, we selected and  
191 utilised the four individuals with the highest sequencing coverage from each population. The  
192 input files were generated following Schiffels & Wang (2020), and along with mask files  
193 generated by bamCaller.py, the mappability masks (Kivikoski et al., 2021) were applied.  
194 Estimates were carried out with default settings, and the outputs were processed assuming  
195 mutation rate of  $4.37 \times 10^{-9}$  per site per generation (Zhang et al., 2023) and a generation length  
196 of two years (DeFaveri et al., 2014). To conduct bootstrap estimations, the input data were  
197 chopped into 1 Mb blocks and an artificial 400 Mb long genome was generated by random  
198 sampling with replacement. 20 artificial bootstrap datasets were generated using  
199 Multihetsep\_bootstrap.py from msmc-tools (<https://github.com/stschiff/msmc-tools>) and  
200 analysed with the same settings as the original data. In all analyses, the first two time  
201 segments (which usually are untrustworthy; Schiffels & Durbin, 2014; Sellinger et al., 2021)  
202 were discarded. Point estimates of  $N_e$  at certain times were obtained by interpolation with

203 R's *approx* function and the value at 4000 years before present was used as the estimate of  
204  $N_e^H$ , the historical  $N_e$ , in statistical analyses.

## 205 Long-term $N_e$ estimation

206 The average long-term  $N_e$  ( $N_e^{LT}$ ) were estimated using the formula (Kimura, 1983)  $N_e =$   
207  $\pi/(4\mu)$ , where  $\mu$  is the mutation rate, assumed to be  $4.37 \times 10^{-9}$  per site per generation (Zhang et al.,  
208 2023).  $\pi$  was obtained from folded site frequency spectra (SFS), estimated for each  
209 population directly from the bam data with ANGSD v.0.921 (Korneliussen et al., 2014), using  
210 the R script from Walsh et al., (2022) modified to fit folded SFSs. Sites with more than 70%  
211 heterozygote counts were removed and the mappability masks (Kivikoski et al., 2021) were  
212 applied in data filtering. For details, see Table S2.

213 Table 1. Overview of  $N_e$  estimators in this research.

<b><math>N_e</math> estimator</b>	<b>Method</b>	<b>Definition</b>
$N_e^C_{GONE}$	GONE	Contemporary $N_e$ . Estimates at 1 generation before present.
$N_e^C_{CurrentNe}$	CurrentNe2	Contemporary $N_e$ .
$N_e^H$	MSMC2	Historical $N_e$ . Point estimates at 4000 years before present.
$N_e^{LT}$	Genetic diversity	Long-term $N_e$ . Estimated from nucleotide diversity ( $\pi$ ) using the equation $N_e^{LT} = \pi/(4\mu)$ .

214

## 215 Statistical analyses

216 We assessed the linear relationships among alternative  $N_e$  estimates and admixture  
217 proportions using the Pearson correlation test, implemented in the *cor.test()* function in R (R  
218 Core Team, 2020). The Pearson correlation coefficient ( $r$ ) was calculated to quantify the  
219 strength and direction of these relationships, with  $p$ -values reported to assess statistical  
220 significance (threshold = 0.05). In cases where normality was violated, we conducted  
221 Spearman's rank correlation as an alternative and the Spearman's correlation coefficient  
222 ( $rho$ ) was reported.

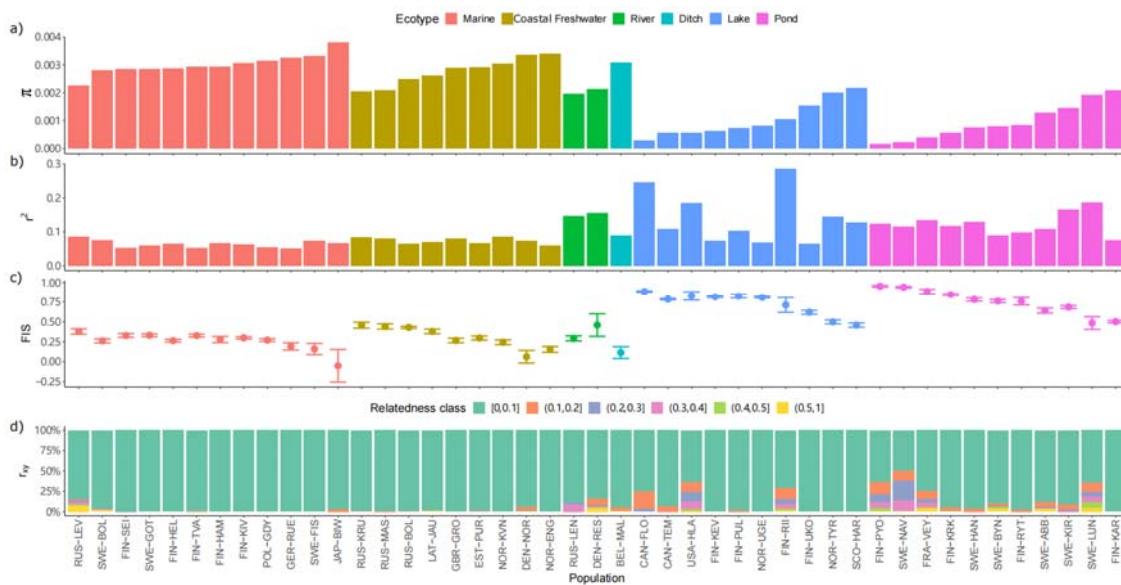
223

## 224 Results

225 Summary statistics show unexpected variation among freshwater  
226 populations

227 We found the genetic diversity to be generally higher in marine and coastal freshwater than  
228 in inland freshwater populations and decreasing together with the connectivity of the habitat  
229 class (Fig. 1a). However, the patterns of LD decay over physical distance showed marked  
230 differences within and between habitat classes (Fig. 1b, Fig. S2). While the decay of LD is

231 faster and shows lower average LD ( $r^2$ ) in marine than in freshwater populations, the latter  
 232 show considerable variation, some being similar to marine populations with low levels of LD  
 233 and others containing very high levels of LD (Fig. 1b). Similarly, we found the marine  
 234 populations to have generally low  $F_{IS}$ , while those for the freshwater populations were highly  
 235 variable (Fig. 1c). Since a few populations showed higher than average levels of within-  
 236 ecotype LD and within-population variation in  $F_{IS}$ , we estimated the pairwise relatedness ( $r_{xy}$ ,  
 237 Hedrick et al. 2015) within each population. The relatedness showed an increase with the  
 238 degree of isolation of the habitat, and unexpectedly, some freshwater samples were more  
 239 closely related than the others ( $r_{xy} > 0.5$  shown in yellow colour in Fig. 1d) within the given  
 240 site (Fig. S3). The initial analysis revealed that the samples from Lake Riikojärvi, Finland,  
 241 show an exceptionally high level of LD and strong patterns of inbreeding in comparison to  
 242 other lake populations, and certain individuals within the population were more closely  
 243 related than others.



244

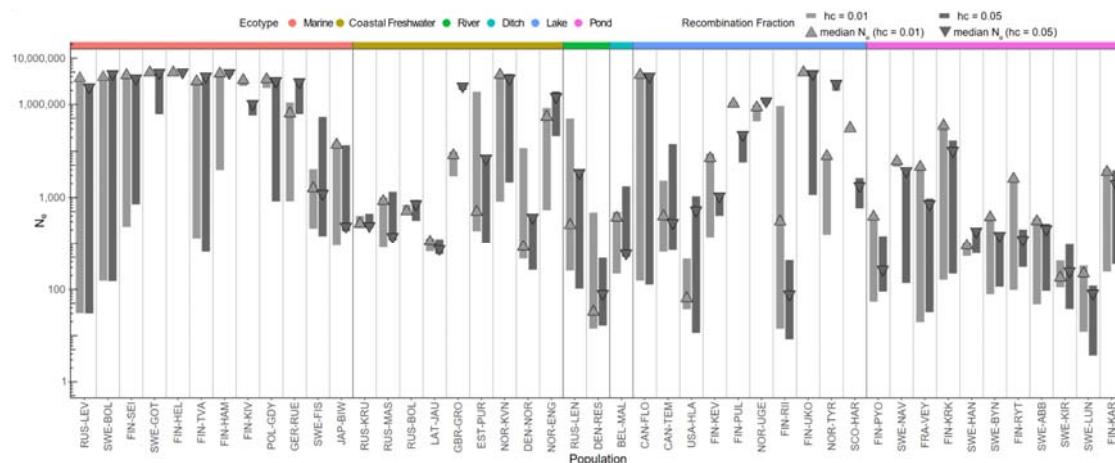
245 Fig. 1. Summary statistics for 45 nine-spined stickleback populations. a) Nucleotide diversity  
 246 ( $\pi$ ) across the autosomal chromosomes. b) LD calculated as the harmonic mean of  $r^2$  for  
 247 SNPs located 100-200 kb apart within LG4. c) Inbreeding coefficients ( $F_{IS}$ ) with standard  
 248 deviations. d) Relatedness ( $r_{xy}$ ) for pairs of individuals within populations with colours  
 249 representing the proportion of pairwise comparisons within a population falling in a specific  
 250 relatedness class. For the LD decay curve for each population, see Fig. S2.

251 Population structure affects inferences of recent demographic history

252 Population structure, bottlenecks and gene flow can distort the LD patterns and bias the  
 253 inferences of recent demographic history (Santiago et al., 2020, Novo et al., 2023a). We  
 254 inferred recent demographic histories using the LD-based methods GONE (Santiago et al.,  
 255 2020) and found sharp declines in  $N_e$  over the last few generations in a few populations (Fig.  
 256 2; Fig. S4). Comparisons of  $N_e$  estimates across different recombination bin cut-offs ( $hc$ )  
 257 showed either inconsistent results or wide ranges of values – the typical symptoms of recent  
 258 gene flow – for populations exhibiting abnormal LD, inbreeding, or relatedness patterns (Fig.  
 259 2; Santiago et al., 2020; Novo et al., 2023a). An extreme case was FIN-RII, an outlier in the

260 LD decay analysis (Fig. 1b, Fig. S2), where  $N_e$  estimates ranged dramatically from 14 to  
261 924,485 (Fig. 2) and varied significantly across different  $hc$  levels (Fig. S4). A lower  $hc$  value  
262 removed the variation in  $N_e$  over the generations in a few populations (e.g. FIN-UKO and  
263 SWE-GOT), but did not completely eliminate it (e.g. FIN-TVA, Fig. S4). Several populations  
264 showed consistent patterns in  $N_e$  estimates with different  $hc$  values. Of the 45 populations,  
265 six showed steady  $N_e$  estimates over time across the two  $hc$  values, indicating their stable  
266 recent history (Fig. S4).

267



268

269 Fig. 2. Variations in recent historical  $N_e$  estimates obtained with GONE. Range of  $N_e$   
270 estimates for each population over 1–50 generations before present, calculated using two  
271 recombination fractions ( $hc = 0.01$  in light gray, and  $hc = 0.05$  in dark gray). Median  $N_e$   
272 values are marked by matching triangles. Populations are ordered as in Fig. 1 and the color  
273 bar on top represents the ecotypes. For the trajectories of each population, see Fig. S4.

## 274 Differences in $N_e^C$ estimates

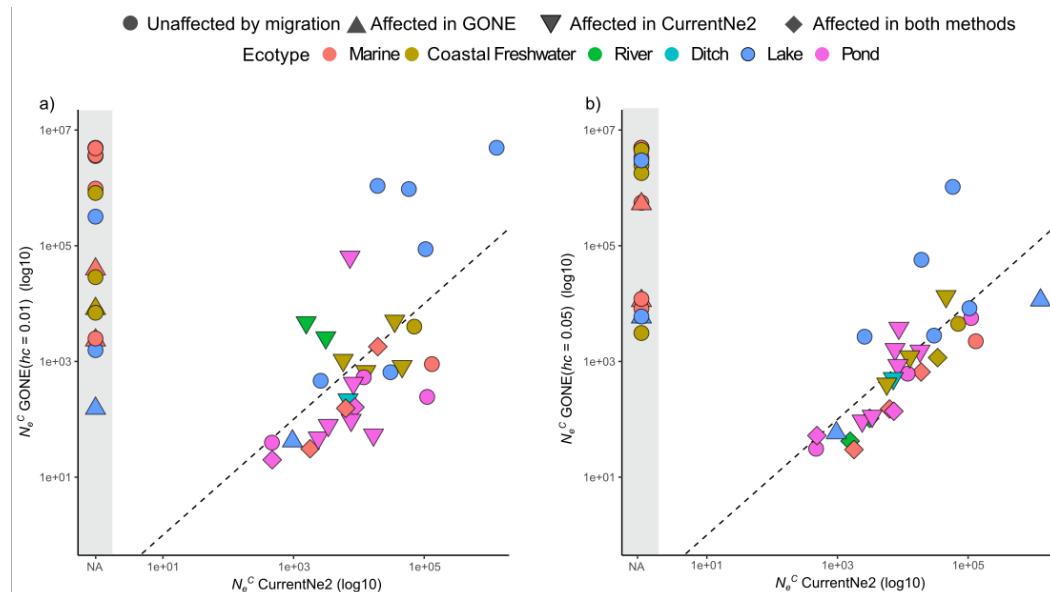
275 We estimated  $N_e^C$  using the LD-based method CurrentNe2 and compared these with the  
276 estimates at 1 generation before present obtained with GONE using different  $hc$  cut-offs.  
277 Overall, the  $N_e^C$  estimates exhibited a similar trend across ecotypes but displayed high  
278 variability within freshwater populations, with pond populations generally having smaller  $N_e^C$ s  
279 than the others. With CurrentNe2, the  $N_e^C$  of eight marine and seven freshwater populations  
280 could not be estimated. Of the 30 remaining populations, 18 were inferred to be affected by  
281 migration. As expected, for the 18 populations affected by migration,  $N_e^C$  estimates were  
282 higher when migration was accounted for. However, migration rates did not show correlation  
283 with the difference ( $r = -0.3$ ,  $p = 0.23$ ,  $n = 18$ ) but were weakly correlated ( $r = -0.41$ ,  $p = 0.09$ ,  
284  $n = 18$ ) with the ratio of  $N_e^C$  estimates between the two models.

285 Since we previously showed that the GONE analyses of some populations were influenced  
286 by population structure, it is not surprising that the  $N_e^C_{GONE}$  estimates for those populations  
287 were also affected (10 populations at  $hc = 0.01$  and 16 populations at  $hc = 0.05$ , Fig. 3).  
288  $N_e^C_{GONE}$  estimates obtained with  $hc = 0.05$  were generally higher than those obtained with  $hc$   
289 = 0.01, particularly in populations affected by recent gene flow. However, after excluding the

290 affected estimates, the overall trends remained consistent between the two  $hc$  cut-offs ( $r = 0.50$ ,  $p = 0.006$ ,  $n = 29$ ). When migration was accounted for in CurrentNe2, we observed a strong alignment between  $N_e^C_{CurrentNe}$  and  $N_e^C_{GONE}$  estimates ( $hc = 0.01$ ,  $\rho = 0.62$ ,  $p < 0.001$ ,  $n = 30$  and  $hc = 0.05$ ,  $\rho = 0.86$ ,  $p < 0.0001$ ,  $n = 30$ ).

294 Discrepancies emerged when comparing  $N_e^C$  estimates between the methods. For example, 295 freshwater populations from Sweden (Lil-Navartjärn, hereafter SWE-NAV) and Finland 296 (Pyöreälampi, hereafter FIN-PYO) that showed high levels of relatedness also obtained 297 markedly different  $N_e^C$  estimates with the two different methods (Fig. S5). A possible 298 explanation is that their exceptionally low genetic diversity and inbreeding have distorted the 299 LD patterns. Substantial differences in  $N_e^C$  estimates were also observed for the Japanese 300 marine population and several Baltic Sea populations (Fig. S5, Table S5). Interestingly, with 301 both methods, the smallest  $N_e^C$ 's were observed for a pond population from Lund, Sweden 302 ( $N_e^C_{GONE} = 37.33$  at  $hc = 0.01$  and  $31.30$  at  $hc = 0.05$ ,  $N_e^C_{CurrentNe} = 46.63$ ), not in the population 303 with the lowest genetic diversity (FIN-PYO,  $\pi = 0.00015$ ), and the highest  $N_e^C$  was observed 304 in a lake population (Ukonjärvi, Finland), followed by a marine population from Biwase Bay, 305 Japan, for CurrentNe2, and the Gulf of Finland for GONE (Fig. S5). In addition, we found 306 that the confidence intervals for  $N_e^C_{GONE}$  were extremely narrow compared to those for 307  $N_e^C_{CurrentNe}$ , probably due to the overlapping of SNPs in the replicates (Fig. S5).

308



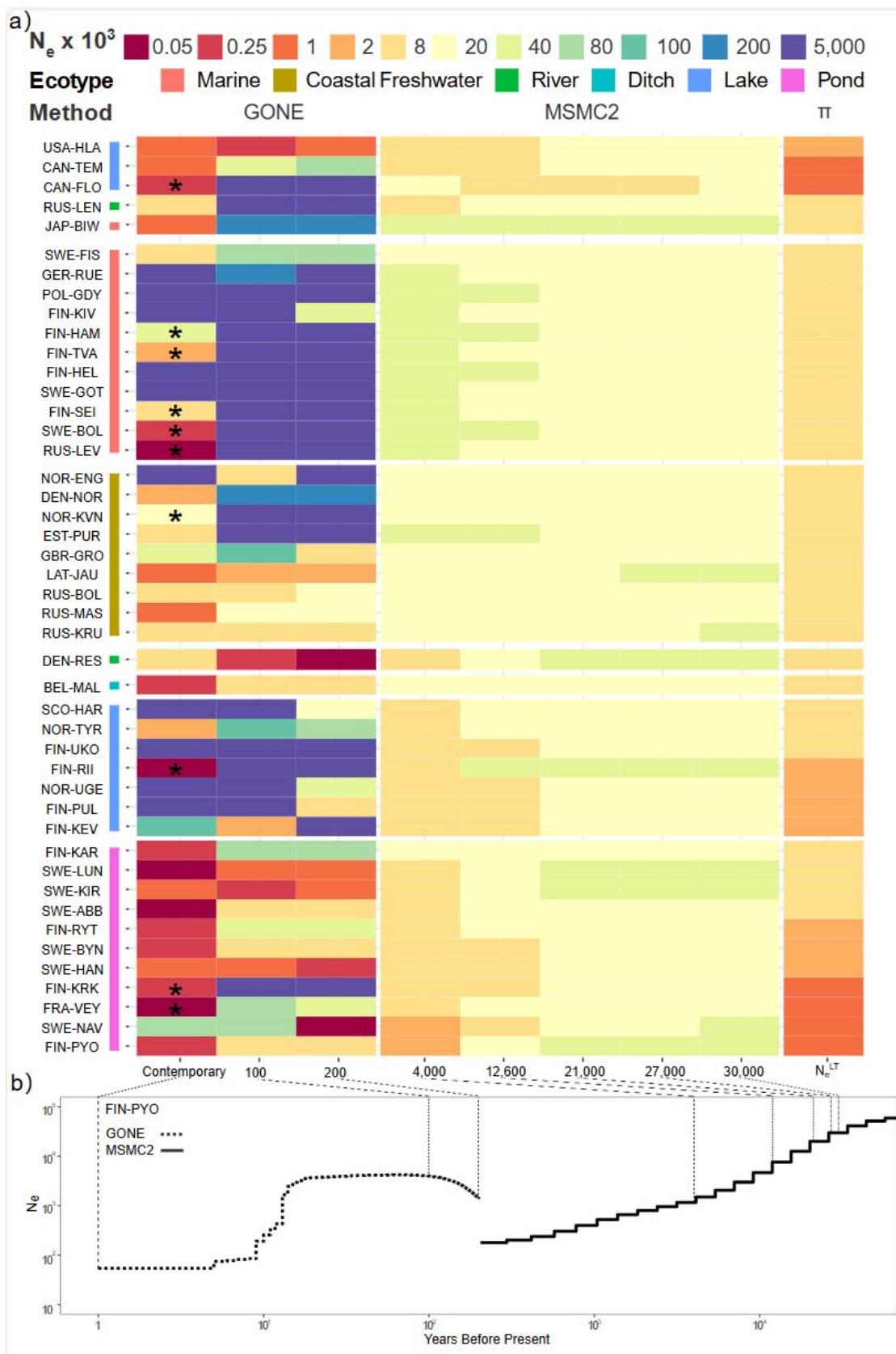
310 Fig. 3. Comparison of estimated contemporary effective population size ( $N_e^C$ ) using GONE 311 and CurrentNe2 for 45 nine-spined stickleback populations. a) and b) show  $N_e^C$  estimates 312 obtained from GONE with  $hc = 0.01$  and  $hc = 0.05$ , respectively. The black dashed line 313 represents the 1:1 slope, and the gray shaded area indicates populations for which  $N_e^C$  314 could not be estimated using CurrentNe2. The shapes denote whether the estimates were 315 affected by migration, while the colors represent the ecotypes.

316

317  $N_e$  continuums reveal distinctive population histories

318 We reconstructed the  $N_e$  continuum spanning the last 30,000 years, capturing demographic  
319 changes from before the Last Glacial Maximum (30 kya) to the present day. Overall, we  
320 found both recent historical and historical  $N_e$  estimates to be generally lower in freshwater  
321 populations compared to marine populations (Fig. 4). However, recent historical  $N_e$   
322 estimates exhibit greater variability in lake and coastal freshwater populations compared to  
323 pond and marine populations (Fig. 4), consistent with the observed variability in levels of LD  
324 and  $F_{IS}$  (Fig. 1), suggesting differences in their recent histories of isolation, recovery from  
325 bottleneck or expansion, and migration.

326 The historical  $N_e$  show that all populations experienced declines until the end of the Last  
327 Glacial Maximum (LGM, approximately 21,000 years ago; Fig. 4, Fig. S6 & S7). After this  
328 period, the historical  $N_e$  of marine populations increased (Fig. 4, Fig. S7), while most  
329 freshwater populations continued to decline (Fig. 4, Fig. S7). Within ecotypes, specific  
330 populations exhibit distinct but predictable deviations from the consensus  $N_e$  history. For  
331 instance, the marine population from Hokkaido, Japan (Fig. S6 & S7), and the freshwater  
332 population from Lake Floating Stone, Canada (Fig. S6 & S7), stand out within their ecotypes,  
333 reflecting their geographic isolation and the regional climatic differences during the Last  
334 Glacial Period. Conversely, some coastal freshwater and pond populations display  $N_e$   
335 trajectories similar to those of European marine populations (Fig. S6 & S7), indicating recent  
336 colonization of the freshwater environment. In general, the demographic histories of  
337 freshwater populations are more variable than those of marine populations (Fig. 4, Fig. S7),  
338 reflecting the differing sizes, connectivities, and regional geographic histories of their  
339 habitats.



341 Fig. 4. Temporal reconstruction of effective population size ( $N_e$ ) for 45 nine-spined  
342 stickleback populations over the past 30,000 years. a) The colors represent  $N_e$  estimates at  
343 key time points with the scale given at the top. The estimates for the last 200 years were  
344 derived with GONE ( $hc = 0.01$ ; results with  $hc = 0.05$  are shown in Fig. S8), while those for  
345 4,000 to 30,000 years ago were obtained with MSMC2. Asterisks indicate estimates  
346 influenced by population structure as inferred by manual inspection following the patterns in  
347 Fig. 2F of Santiago et al. (2020). Long-term  $N_e$  values were calculated from genetic diversity  
348 ( $\pi$ ). Key time points were chosen to reflect significant geographic events: 4,000 years ago  
349 marks the onset of the current Baltic Sea stage, 12,600 years ago the start of the Baltic Ice  
350 Lake stage, and 21,000-27,000 years ago corresponds to the Last Glacial Maximum. The  
351 color next to population labels represents the ecotypes. b) The demographic history of a  
352 representative population, FIN-PYO, over the past 80,000 years. The dashed lines connect  
353 the time points in panels a) and b).

354

### 355 Historical introgression may differently affect alternative $N_e$ estimators

356 The Baltic Sea marine populations were inferred to contain 13-30% of Western Lineage  
357 ancestry (Feng et al., 2022) and such past introgression events should be visible in the  
358 trajectories of historical  $N_e$ . The populations show an increase of  $N_e$  around 10,000 years  
359 ago (Fig. S6 & S7) and similar trajectories are seen in coastal freshwater populations known  
360 to contain moderate levels (12-16%) of admixture. However, the Swedish freshwater  
361 populations with low levels of admixture (3-5%) in a previous study (Feng et al., 2022) show  
362 recent MSMC2 trajectories similar to the consensus pattern of the pond ecotype. This area is  
363 known to have been isolated from the Baltic Sea around 10,000 years ago (Mobley et al.,  
364 2011), suggesting a negligible impact of low degree of admixture on historical  $N_e$  estimates.

365 To more formally test the impact of historical introgression, we assessed the linear  
366 relationships between introgression level and alternative  $N_e$  estimators. In the pond  
367 populations showing small amounts of admixture, the amount of introgressed ancestry  
368 showed no correlation with any  $N_e$  estimate (Table 2). Among the Baltic marine populations,  
369 the admixture proportion was not correlated with  $N_e^H$ , the  $N_e$  estimate for 4000 years before  
370 present, but showed negative correlation ( $r = -0.89$ ,  $n = 5$ ,  $p = 0.04$ ) with  $N_e^C_{GONE}$  ( $hc = 0.01$ ,  
371 Table 2). For the long-term estimates, the effect was opposite and the admixture proportion  
372 correlated positively with  $N_e^{LT}$  (Table 2).

373 Table 2. Pearson product moment correlations between the different  $N_e$  estimates and the  
374 admixture proportions in marine and pond ecotypes. Two to nine marine and seven pond  
375 populations were included.

Method	Ecotype	r	p	N
$N_e^C_{CurrentNe2}$	Marine	NA	NA	2
$N_e^C_{GONE}$ ( $hc=0.01$ )	Marine	-0.89	0.04	5*

$N_e^C_{GONE}(hc=0.05)$	Marine	-0.12	0.85	5*
$N_e^H$	Marine	0.14	0.72	9
$N_e^{LT}$	Marine	0.83	0.006	9
$N_e^C_{CurrentNe2}$	Pond	-0.27	0.56	7
$N_e^C_{GONE}(hc=0.01)$	Pond	-0.14	0.77	7
$N_e^C_{GONE}(hc=0.05)$	Pond	-0.28	0.55	7
$N_e^H$	Pond	0.23	0.63	7
$N_e^{LT}$	Pond	0.41	0.35	7

376 \* Populations that showed drastic decrease in  $N_e$  were excluded from analysis.

377

## 378 Discussion

379 The aim of this study was to reconstruct the temporal dynamics of  $N_e$  in natural populations  
380 by employing alternative methods that leverage different genomic signals. Our results  
381 revealed that all studied populations had experienced glacial contractions and, while marine  
382 populations displayed signs of post-glacial expansion, freshwater populations exhibited  
383 continued declines with some populations showing signatures of recovery from historical  
384 bottlenecks. Across all estimators, the  $N_e$  values were in general highest in marine  
385 populations, followed by coastal freshwater populations, and lowest in pond populations. Our  
386 analysis also revealed a mixed impact of population structure on  $N_e$  estimates. In particular,  
387 the  $N_e^{LT}$  and  $N_e^C$  estimates, based on genetic diversity and obtained with GONE,  
388 respectively, for the admixed marine populations correlate with levels of historical genetic  
389 introgression, whereas no such correlation was detected in other  $N_e$  estimates. We found  
390 that characteristic traces in summary statistics can uncover recent gene flow, a violation of  
391 assumptions potentially distorting LD patterns and complicating the reconstruction of recent  
392 demographic histories, as well as the estimation of  $N_e^C$  with both GONE and CurrentNe2.  
393 While also susceptible to historical introgression, MSMC2 models  $N_e$  across time and  
394 attributes the effects of gene flow to the appropriate historical time periods. We explore  
395 these challenges and their implications for accurate and reliable  $N_e$  estimation in greater  
396 detail below.

397 Summary statistics can reveal hidden within-population aberrations

398 The effective population size  $N_e$  has many definitions (Husemann et al., 2016; Waples,  
399 2022) but they all basically return to the Wright-Fisher model (Fisher, 1931; Wright, 1931)  
400 and aim to determine the size of an idealised population behaving genetically in a similar  
401 fashion to the target population. However, population subdivision and gene flow are common  
402 in natural populations (Patton et al., 2019) and such violations of model assumptions impact

403 the inferences of  $N_e$  and complicate their interpretation (reviewed in Loog, 2020; Marchi et  
404 al., 2021; Waples 2024). In non-model species, limited sample sizes and uncertainty about  
405 population conditions pose additional challenges.

406 Our findings demonstrate that simple summary statistics can be informative in elucidating  
407 demographic histories and reduce misinterpretation risks. For example, the Riikojärvi lake  
408 population (FIN-RII) was an outlier, exhibiting the highest level of LD, increased relatedness,  
409 and highly variable  $N_e$  estimates in GONE analyses. These patterns are inconsistent with the  
410 population's nucleotide diversity and the LD levels observed in other populations from the  
411 same geographic region and ecotype. We hypothesize that recent hybridization between  
412 native fish and genetically distinct migrants may have elevated its LD. Given the isolation of  
413 the lake, natural migration seems improbable, implying that these migrants were introduced  
414 with fish stocking that has taken place at the lake. Consequently, both  $N_e^C$  estimation and  
415 recent demographic reconstructions for this population may be unreliable due to the  
416 distortion of LD patterns by recent gene flow.

#### 417 Impact of population structure and historical gene flow on $N_e$ 418 estimations

419 The relationship between population's census size and effective size is complex (Frankham  
420 1995; Palstra & Ruzzante, 2008; Waples et al., 2013). The  $N_e$  estimates in this study should  
421 be interpreted as indicators of relative demographic changes in each population, and, as  
422 such, be indicative of the influence of various demographic processes (e.g., population  
423 bottlenecks and inbreeding) and climatic changes. Since the alternative  $N_e$  estimators are  
424 based on different definitions and utilize distinct features of the data, they probably should  
425 not be expected to be fully congruent and none of them should be considered as definitive  
426 values of  $N_e$ .

427 Interestingly, we found that gene flow drives divergent responses in  $N_e$  estimators: LD-based  
428 methods underestimated  $N_e$  when migration is not adequately accounted for (Saura et al.,  
429 2021; Santiago et al., 2024; Gargiulo et al., 2024), whereas coalescent-based methods  
430 tended to overestimate  $N_e$  due to the increased genetic diversity introduced by gene flow.  
431 This contrast underscores the differences in methodology and sensitivity to demographic  
432 changes in different timescales: LD-based methods, which assess  $N_e$  based on the variance  
433 in reproductive success within recent generations ( $N_eV_k$ , Novo et al., 2023a&b, Waples  
434 2024), are sensitive to immediate demographic fluctuations, while coalescent-based  
435 methods focus on historical genetic drift and mutation over longer timescales (Hudson, 1990;  
436 Li and Durbin, 2011). The distinct temporal scales of the methods emphasize the complexity  
437 of interpreting  $N_e$  in structured and admixed populations (Kersten et al., 2023) and highlight  
438 the importance of integrating multiple  $N_e$  estimators to obtain more reliable demographic  
439 inferences.

440 Although GONE is known to be sensitive to population structure (Novo et al., 2023a; Kersten  
441 et al., 2023), our results suggest that the method is generally robust across very different-  
442 sized populations. The trajectories for marine populations are largely consistent, suggesting  
443 that the estimation of recent demographic histories, previously limited to small populations  
444 (DeFaveri & Merilä, 2015; Marandel et al., 2019; Nadachowska-Brzyska et al., 2022), may  
445 be applied to them as well. However, we also detected deviations and unstable estimates in

446 some populations, particularly those with fluctuating connectivity or showing evidence of  
447 population substructure. Such violations of the expected closed population history are known  
448 to bias recent historical  $N_e$  estimates (1–50 generations before present), often causing  
449 sharply decreased  $N_e$  estimates (Santiago et al., 2020, Novo et al., 2023a; Kersten et al.,  
450 2023; Atmore et al., 2022 & 2024). Limiting recombination between SNPs partially reduced  
451 the variation in  $N_e$  estimates across generations (Fig. S4; Santiago et al., 2020) but may also  
452 lead to a loss of informative SNPs, complicating the reconstruction of recent demographic  
453 history (Santiago et al., 2020 & 2023; Novo et al., 2023a) and affecting the estimation of  $N_e^C$   
454 with GONE (Gargiulo et al., 2024).

455 Consistent with previous observations (Ryman et al., 2019, Novo et al., 2023a&b), our  
456 findings indicate that  $N_e^C$ 's tend to be underestimated in structured populations. By  
457 examining different  $hc$  cut-offs in GONE, we found that estimates at  $hc = 0.01$  were  
458 generally lower than those at  $hc = 0.05$  but correlated with them, suggesting that reducing  
459 the  $hc$  cut-off might not effectively mitigate the impacts of population structure. On the other  
460 hand, lowering the  $hc$  cut-off reduced the number of informative SNPs and amplified the  
461 effects of pseudoreplication, narrowing the confidence intervals (Waples et al., 2022).  
462 Furthermore, we found that several populations that GONE inferred unaffected by gene flow  
463 showed signals of population structure in CurrentNe2 analyses. The affected populations  
464 tended to have lower  $N_e^C$  estimates when migration was not accounted for, with the effect  
465 being more pronounced at lower migration rates. Moreover, substantial differences in  $N_e^C$   
466 estimates between the two methods were observed in a few freshwater populations with  
467 extremely low genetic diversity. We postulate that the very low genetic variation in these  
468 populations causes even minor deviations in the data to appear more pronounced and  
469 exaggerate the estimates of changes in drift and LD patterns. Such discrepancies between  
470 the methods highlight the importance of careful evaluation of the population structure when  
471 estimating  $N_e^C$  and emphasize the need for caution in interpreting the estimates, particularly  
472 when the recent conditions of the population are uncertain.

473 We found that past gene flow had mixed effects on different  $N_e$  estimates. In Baltic Sea  
474 populations, we observed a positive correlation between  $N_e^{LT}$  and admixture proportions,  
475 suggesting that historical introgression has significantly shaped the genetic landscape of  
476 these populations. In contrast, historical  $N_e$  estimates sampled at 4000 years before present  
477 (ybp), did not show such a correlation. This likely reflects the fact that the ages of the  $N_e^H$   
478 point-estimates are much younger than the recent secondary contacts and the admixture  
479 can correctly be taken into account in the more distant time segments. Coalescent-based  
480 methods can theoretically accommodate for gene flow and population substructure and are  
481 considered relatively robust to low levels of admixture (Beichman et al., 2017; but see  
482 Palamara & Pe'er 2013). However,  $N_e$  estimates may still be inflated if introgressed genetic  
483 variation is not adequately accounted for, as indicated by recent studies on archaic hominins  
484 (Li et al., 2024). In the Baltic Sea marine populations, the complex admixture history is mixed  
485 with post-glacial population expansion (Feng et al., 2022), and disentangling the effects of  
486 these processes on  $N_e$  estimates is not straightforward. Utilizing alternative approaches like  
487 fastsimcol2 (Excoffier et al., 2021) and dadi (Gutenkunst et al., 2009) might provide a path to  
488 a deeper understanding of the history of Baltic Sea populations.

489

490  $N_e$  estimators and demographic history of the nine-spined  
491 sticklebacks

492 While none of the  $N_e$  estimates is immune to the impact of population structure and gene  
493 flow, they remain valuable for understanding population histories and dynamics. For  
494 example, the Japanese marine population from a tide pool connected to Biwase Bay, which  
495 experienced inter-species introgression in the distant past (Yamasaki et al., 2020), exhibited  
496 the highest  $N_e^{LT}$ , a median  $N_e^H$ , but a lower  $N_e^C$  compared to Baltic Sea populations. Despite  
497 their apparent relative incongruence, the results are consistent with the highly complex  
498 history of the population:  $N_e^{LT}$  reflects the mixed species ancestry with no apparent historical  
499 bottleneck and high levels of current genetic diversity,  $N_e^H$  its marine origin and limited  
500 impact during the Last Glacial Period (LGP),  $N_e^C$  its very recent past in a shallow-water tide  
501 pool thereby reduced connectivity to its adjacent marine population and increased genetic  
502 drift in the contemporary population. Although anecdotal, the case well represents the  
503 differences behind the alternative estimates of  $N_e$ , their timescales, differences in sensitivity  
504 to demographic changes and associated biological meanings.

505 The demographic history of nine-spined stickleback populations reflects the significant  
506 impact of Pleistocene climatic oscillations. Glaciations and subsequent deglaciations are  
507 associated with bottlenecks and expansions, which have shaped the genetic diversity and  
508 adaptive potential of contemporary populations (reviewed in Hewitt, 2004). Consistent with  
509 other European species (e.g., Backström et al., 2013; Liu et al., 2016), our results show that  
510 European nine-spined sticklebacks experienced population contractions during glacial  
511 periods. After the LGP (~11,000 years ago), marine and freshwater ecotypes began to  
512 diverge. The timing of bottlenecks in various freshwater populations reflects their  
513 colonization history and the formation of new habitats as glacial ice sheets retreated. For  
514 instance, pond populations close to the Gulf of Bothnia display bottlenecks 5,000–10,000  
515 years ago, consistent with the historical coastal line in the area (Mobley et al., 2011). In  
516 contrast, the White Sea pond populations appear to have resulted from more recent  
517 colonizations (Ziuganov & Zotin, 1995), with a bottleneck dated to approximately 500 years  
518 ago (Figs. S3 and S4). Additionally, two river populations from the Baltic Sea coast (from  
519 Estonia and Latvia) share similar proportions of Western Lineage ancestry with nearby  
520 marine populations (Feng et al., 2022), suggesting their establishment occurred after the  
521 secondary contact between the two lineages. Such fine details highlight the fundamental  
522 differences among the studied freshwater populations, particularly their ages and their origin  
523 from ancestral populations representing different pools of standing genetic variation. The  
524 latter is a crucial resource for local adaptation (Barrett & Schlüter, 2008). Differences in  
525 access to the pool of standing variation due to gene flow or historical demography can either  
526 constrain or enhance local adaptation (Fang et al., 2021; Kemppainen et al., 2021). In this  
527 respect, our results provide a foundation for further investigation into the role of ancestral  
528 polymorphism in the local adaptation of nine-spined sticklebacks.

529 Feasibility of applying GONE in large populations

530 In our analyses, estimating the  $N_e^C$  with CurrentNe2 was particularly problematic for certain  
531 large populations (e.g. FIN-HEL,  $n = 22$ ), likely due to the rapid decay of LD and the weak  
532 genetic drift signals available for analysis (Gargiulo et al., 2024; Waples, 2025). The latter is

533 inversely proportional to  $N_e$  and is quickly diminished by sampling noise as  $N_e$  increases  
534 (Wang et al., 2016; Waples, 2016). Accurate  $N_e$  estimates with LD-based methods require  
535 sample sizes comparable to  $N_e$  to reduce signal-to-noise ratios to acceptable levels (Hill  
536 1981; Marandel et al., 2019), which is impractical for large populations as this often  
537 translates into sampling and sequencing hundreds to thousands of individuals—an  
538 unfeasible scale in most cases. In contrast, GONE showed a clear advantage by providing  
539 reasonable  $N_e$  estimates even with relatively small sample sizes. Furthermore, results from  
540 GONE were well aligned with  $N_e^C$ 's inferred from CurrentNe2 for both large and small  
541 populations when recent gene flow was adequately accounted for. This capability is  
542 particularly valuable for estimating  $N_e$  of large populations. However, due to the inherent  
543 challenges of detecting very subtle drift signals with limited samples, the  $N_e^C$  estimates from  
544 GONE might be best viewed to represent the lower bounds of the  $N_e$ .

## 545 **Conclusions**

546 Using whole-genome data from 45 populations of nine-spined sticklebacks, we demonstrate  
547 the utility and limitations of a set of genomic methods for estimating  $N_e$  over time. Our study  
548 highlights the usefulness of integrating population summary statistics for studying  
549 demographic histories and reveals the complex impact of population structure on different  $N_e$   
550 estimates. Despite their biases, the estimates remain valuable for gaining a comprehensive  
551 understanding of population histories and dynamics. By comparing the results from a  
552 method specifically designed to estimate  $N_e^C$ , we demonstrated the potential of using GONE  
553 for the estimation of  $N_e^C$  in large populations, which is often a challenge due to their low  
554 levels of genetic drift. Overall, our findings underscore the importance of employing a  
555 combination of methods to account for both historical and recent demographic processes,  
556 providing a more holistic view of population histories and resilience.

## 557 **Authors' contributions**

558 J.M. started the project. X.F., J.M. and A.L. devised the research idea. X.F. performed the  
559 analyses with the help of A.L. X.F. and A.L. wrote the first draft and all authors participated in  
560 the writing of the final manuscript.

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571

## 572 **Conflict of interest**

573 The authors declare no competing interests.

574

575 **Data availability statement**

576 The whole-genome re-sequencing data have been published previously by Feng et al.  
577 (2022) and all the raw sequence data for this study can be accessed through European  
578 Nucleotide Archive (ENA) (<https://www.ebi.ac.uk/ena>) under accession code PRJEB39599.  
579 Other relevant data (e.g., input and output files) are available from the Zenodo Open  
580 Repository:<https://zenodo.org/record/14999855>.

581

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