

1 A unified model for cell-type resolution genomics from 2 heterogeneous omics data

3 Zeyuan Johnson Chen ^{*,1}, Elior Rahmani ^{†*,2}, Eran Halperin ^{‡2}

4 ^{*}These authors contributed equally

5 ¹Department of Computer Science, University of California, Los Angeles, CA, USA

6 ²Department of Computational Medicine, University of California, Los Angeles, CA, USA

7 **The vast majority of population-scale genomic datasets collected to date consist of “bulk”**
8 samples obtained from heterogeneous tissues, reflecting mixtures of different cell types. In
9 order to facilitate discovery at the cell-type level, there is a pressing need for computational
10 deconvolution methods capable of leveraging the multitude of underutilized bulk profiles al-
11 ready collected across various organisms, tissues, and conditions. Here, we introduce Unico,
12 a unified cross-omics method designed to deconvolve standard 2-dimensional bulk matrices
13 of samples by features into a 3-dimensional tensors representing samples by features by cell
14 types. Unico stands out as the first principled model-based deconvolution method that is the-
15oretically justified for any heterogeneous genomic data. Through deconvolution of bulk gene
16 expression and DNA methylation datasets, we demonstrate that the transferability of Unico
17 across different data modalities translates into superior performance compared to existing
18 approaches. This advancement enhances our capability to conduct powerful large-scale ge-
19 nomic studies at cell-type resolution without the need for cell sorting or single-cell biology.

[†]Corresponding Author: eliorrahmani@mednet.ucla.edu

[‡]Corresponding Author: ehalperin@cs.ucla.edu

20 **1 Introduction**

21 Studying cell-type level genomic variation is critical for unveiling complex biology. Unfortunately,
22 collecting large and well-powered datasets at cell-type resolution for population studies has yet to
23 become common practice. Current single-cell datasets typically consist of data collected from no
24 more than several dozens of individuals due to prohibitive costs, and purifying cell types at scale
25 using flow cytometry is laborious and often impractical, particularly for solid and frozen tissues
26 for which cell isolation is very challenging [1–5].

27 Indeed, most transcriptomic and other genomic data types collected to date have been mea-
28 sured from heterogeneous tissues that consist of multiple cell types, resulting in vast amounts of
29 large heterogeneous “bulk” genomic data (e.g., over two million bulk profiles publicly available
30 on the Gene Expression Omnibus alone [6]). This rationalizes the development of computational
31 methods that can disentangle the convolution of cell-type level signals that compose such bulk
32 profiles. The premise, upon successful implementation, offers a transformative capability to con-
33 duct powerful, large-scale studies at the cell-type level in multiple tissues and under numerous
34 conditions for which large bulk data have already been collected.

35 Here, we propose a method for deconvolving 2-dimensional (2D) bulk data (samples by fea-
36 tures) into its underlying 3-dimensional (3D) tensor (samples by features by cell types). Thus far,
37 deconvolution methods have been tailored to specific data types [7–11]. In contrast, we introduce
38 a unified cross-omics method, Unico, the first principled model-based deconvolution method that
39 is theoretically applicable to any heterogeneous genomic data. As we demonstrate through a com-

40 prehensive analysis of multiple gene expression and DNA methylation datasets, this generalization
41 translates into superior performance over existing approaches and improves our ability to conduct
42 powerful large-scale genomic studies at cell-type resolution.

43 **2 Results**

From bulk genomics to cell-type resolution: decomposition versus deconvolution The study of bulk genomics routinely calls for *decomposition*, wherein an observed bulk data matrix is modeled as the product of two matrices: (i) cell-type proportions (fractions) of the samples in the data and (ii) per-feature cell-type genomic levels (“signatures”; Figure 1a). This amounts to solving a matrix factorization problem. For a given bulk observation x_{ij} of genomic feature j in sample i , virtually all decomposition models share the following assumption:

$$x_{ij} = \sum_{h=1}^k w_{ih} z_{jh} + e_{ij} \quad (1)$$

44 where w_{i1}, \dots, w_{ik} are the proportions of k modeled cell types in sample i , z_{j1}, \dots, z_{jk} are the cell-
45 type level signatures of the genomic feature j in each of the k cell types, and e_{ij} is an error term.

46 Numerous decomposition formulations with various assumptions on the products of the
47 factorization have been proposed for the estimation of cell-type compositions and for learning
48 cell-type signatures using different genomic modalities, including gene expression [12–15], DNA
49 methylation [16–20], copy number aberrations [21, 22], ATAC-Seq [23], and Hi-C data [24]. The
50 rich toolbox of decomposition methods has proven successful for a wide range of applications,
51 such as clustering genes and studying their functional relationships [25, 26], inferring tumor com-

52 position [21, 22], and discovering cancer sub-types [27]. However, these methods allow us to infer
53 only a single profile of cell-type level signatures per feature, which corresponds to the unrealistic
54 assumption that all samples in the data share the same genomic levels at the cell-type level.

Every sample may reflect its own – possibly unique – cell-type level patterns, owing to various factors of inter-individual variation, such as genetic background, environmental exposures, and demographics. A natural adjustment of the decomposition model to reflect such variation yields:

$$x_{ij} = \sum_{h=1}^k w_{ih} z_{ijh} + e_{ij} \quad (2)$$

55 where z_{ijh} now represents the level of feature j in cell-type h , specifically in sample i . Learning
56 z_{ijh} from bulk data is essentially a *deconvolution* problem, wherein we disentangle the mixture of
57 signals in a 2D samples by features bulk data into the unobserved underlying 3D tensor of samples
58 by features by cell types (Figure 1a).

59 Decomposition under Equation (1) can be viewed as a degenerate case of the more general
60 deconvolution problem in Equation (2) [28]. *Deconvolving* the data is thus more desired than
61 merely *decomposing* the data, and the higher resolution of a successful deconvolution is expected
62 to improve cell-type context and discovery in the analysis of bulk genomics. This has been high-
63 lighted and demonstrated by several recent deconvolution methods, including CIBERSORTx [8],
64 MIND [9], bMIND [10], and CODEFACS [11] in the context of transcriptomics and TCA [7] in
65 the context of DNA methylation.

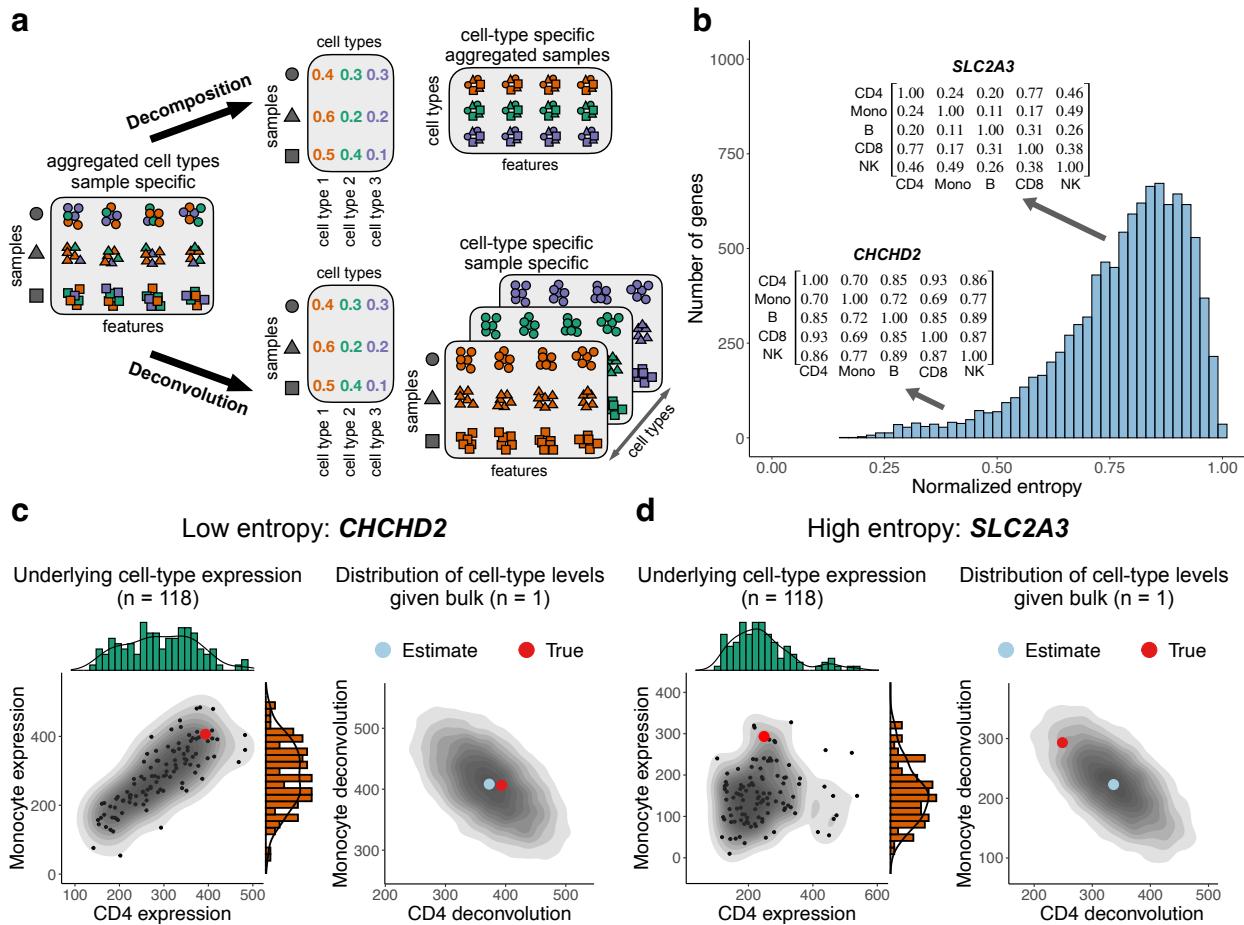


Figure 1: (a) Illustration of decomposition versus deconvolution. (b) The distribution of cell-type covariance structure strength across the top 10,000 most highly expressed genes in scRNAseq from PBMC [29], measured by normalized von Neumann entropy (Methods). (c) The joint distribution of CD4- and monocyte-specific CHCHD2 expression across 118 scRNAseq PBMC samples (left) and the corresponding conditional joint distribution of a Unico deconvolution (right) for one arbitrary individual sample (red circle) given the pseudo-bulk level of the sample. The conditional joint distribution highlights the distance between the true cell-type levels (red circle) and the deconvolution estimate (light blue circle; expectation of the conditional distribution). (d) The same only for the higher-entropy gene SLC2A3.

66 **Unico: A unified cross-omics deconvolution model** Current deconvolution methods can be cate-
67 gorized into two groups: heuristic approaches, including CIBERSORTx [8] and CODEFACS [11],
68 and methods based on the assumption of data following a normal distribution, including TCA [7],
69 MIND [9], and bMIND [10]. The latter group faces limitations rising from the normal distribution
70 assumption, which is known to be invalid at least for transcriptomic data [30–32]. Importantly, the
71 utilization of variance stabilizing transformations, such as log-scaling, would violate the linearity
72 assumption in Equations (1)-(2) and therefore lead to biased estimation [33].

73 We introduce Unico, a deconvolution method for learning cell-type signals from an input
74 of large heterogeneous bulk data and matching cell-type proportions. In practice, the latter is
75 estimated from the input bulk profiles using reference-based decomposition (e.g., [14, 34]), as per-
76 formed by all existing deconvolution methods [7–11]. The primary novelty of Unico stems from
77 taking a model-based approach following Equation (2) while making no distributional assump-
78 tions, which renders it the first principled model-based method that is theoretically justified for
79 analyzing cell type mixtures in any bulk genomic dataset (Methods).

80 A second key component of Unico is the consideration of covariance between cell types.
81 Genomic features may be different yet coordinated across different cell types; for example, tran-
82 scriptional programs can persist through multiple differentiation steps [35, 36]. Indeed, we observe
83 that many genes present a non-trivial correlation structure across their cell-type-specific expression
84 levels, as measured by entropy of the correlation matrix (Figure 1b), with stronger cell-type cor-
85 relations (lower entropy) observed between cell types that are close in the lineage differentiation

86 tree (Supplementary Materials). In the presence of covariance, Unico leverages the information
87 coming from the coordination between cell types for improving deconvolution (Figure 1c,d).

88 **Establishing a new state-of-the-art deconvolution for bulk genomics** We compared Unico to
89 CIBERSORTx, TCA, and bMIND, as well as to a simple baseline approach of naively weighting
90 each bulk profile by the cell-type proportions of the sample. Our evaluation excluded methods that
91 are either not publicly available [11] or require multiple measurements for every sample [9].

92 In order to form a basis for evaluation, we generated pseudo-bulk mixtures using single-cell
93 RNAseq (scRNAseq) data from peripheral blood mononuclear cells (PBMC; n=118 donors) [29]
94 and from lung parenchyma tissues (n=90 donors) [37] (Methods). We first evaluated the per-
95 formance of the different methods in estimating population-level cell-type means, variances, and
96 covariances by establishing gold standard estimates using the known underlying cell-type profiles
97 of the mixtures. Our results yielded Unico, TCA and bMIND as the best performing methods for
98 estimating population-level means and variances (Figure 2a; Supplementary Figures S1). Unico
99 stands out as the leading method for learning cell-type level covariances, showcasing an average
100 correlation improvement of 36.3% over bMIND, the second-best performing method, which also
101 explicitly models cell-type covariance [10] (Figure 2a; Supplementary Figures S1). The ranking of
102 methods remained consistent across different numbers of modeled cell types and various sample
103 sizes (Supplementary Figures S2-S7).

104 We next evaluated how well the 3D tensor estimated by Unico correlates with the true under-
105 lying cell-type expression levels of the pseudo-bulk profiles. Unico consistently outperformed the

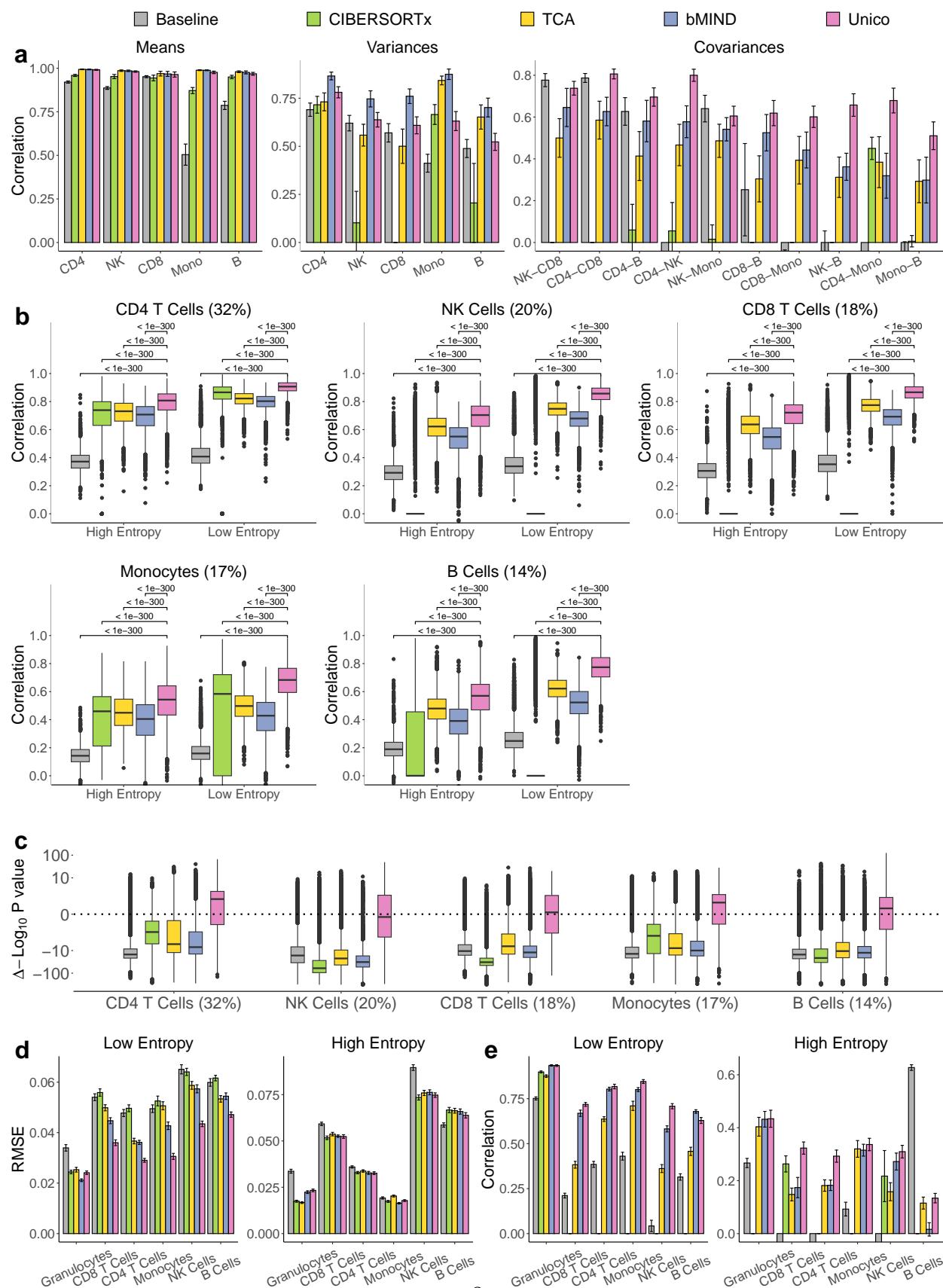


Figure 2: Evaluation of deconvolution methods. (a) Correlation between deconvolution and single-cell based estimates of population-level means, variances, and covariances at the cell-type level across 20 sets of pseudo-bulk mixtures from PBMC scRNAseq profiles of five cell types (500 samples and 600 genes in each set). (b) Evaluation of the concordance between the deconvolution estimates and the known cell-type profiles of the same data in (a). Boxplots reflect the distribution of linear correlation across all genes, and percentages indicate average cell-type abundances. (c) Assessing deconvolution estimates for their information that cannot be explained by pseudo bulk expression. Boxplots reflect the distribution across genes from the same data in (b) of $\Delta \log_{10}(p\text{-value})$, the difference between the log-scaled p-values of the effects of the pseudo bulk expression and deconvolution estimates (higher is better; Methods). (d)-(e) Evaluation of whole-blood DNA methylation deconvolution in terms of RMSE and correlation between estimates and experimentally validated cell-type level methylation across 20 random sets of 1,000 highly variable CpGs. All barplots and error bars in the figure represent means and one standard deviation errors; negative correlations were truncated for visualization purposes, and p-values were calculated using a paired Wilcoxon test.

106 alternative methods across all cell types, providing an average improvement of 17.8% in correla-
107 tion over TCA, the second-best performing method (Figure 2b; Supplementary Figures S1). Unlike
108 Unico, bMIND is a Bayesian method that can perform deconvolution while incorporating prior in-
109 formation on the cell-type level means and covariates. We therefore further compared Unico to
110 bMIND in the presence of informative priors from single-cell data. Remarkably, we found that
111 bMIND could not improve upon Unico even in the unrealistic extreme case where the prior was
112 learned from the true cell-type levels of all samples in the data (Supplementary Figures S8 and S9).

113 As anticipated, the improvement of Unico is more pronounced in genes that exhibit strong
114 cell-type covariance structure (low-entropy genes; average correlation improvement of 20.0%)
115 compared to high-entropy genes (average improvement of 14.9%). This discrepancy highlights
116 the added information Unico gains by modeling the cell-type covariance structure. Importantly,
117 learning a richer model does not come at the cost of significant computational runtime in this case;
118 in fact, Unico is the second fastest deconvolution method (Supplementary Figure S10). The overall
119 ranking of methods remained consistent across different numbers of modeled cell types and various
120 sample sizes (Supplementary Figures S2-S7).

121 Crucially, pseudo-bulk profiles are correlated with their true underlying cell-type levels. We
122 therefore asked whether the 3D tensors estimated by Unico and other methods explain the variation
123 of the true tensor beyond the pseudo-bulk input (Methods). Strikingly, we found that Unico is the
124 only method that learns substantial variation of the true tensor when accounting for the pseudo-bulk
125 profiles, including in lowly abundant cell-types (Figure 2c; Supplementary Figures S1-S7).

126 Lastly, we aimed to confirm the transferability of Unico to other data modalities by decon-
127 volving bulk DNA methylation data. Reinius et al. [38] assayed from the same six individuals both
128 whole-blood methylation and cell-type methylation of six whole-blood cell types. This data col-
129 lection allowed us to establish a ground truth for the cell-type levels composing the whole-blood
130 bulk samples. In order to circumvent the sample size limitation of the Reinius data (n=6), we
131 took a two-step, reference-based approach. Initially, we employed Unico to estimate the model
132 parameters using a separate large whole-blood methylation dataset from a similar population [39].
133 Subsequently, we utilized these parameter estimates in Unico's tensor estimator, which given the
134 model parameters, deconvolves the bulk profile of each individual sample independently of other
135 samples in the data. A similar procedure was adapted for the competing methods (Methods).

136 Unico demonstrated exceptional performance compared to the alternative methods in recon-
137 structing the experimentally known 3D tensor. Considering the top 10,000 most variable methyla-
138 tion CpGs in the data, Unico achieved an average improvement of 8.8% and 8.1% in root median
139 squared error (RMSE) and correlation compared with bMIND, the second best performing method
140 (Figure 2d,e; Supplementary Figures S11 and S12). The ranking of the methods was preserved
141 when considering a set of 10,000 randomly selected CpGs; unsurprisingly, all methods present a
142 noticeable decrease in performance in this case (Supplementary Figures S13-S15).

143 **Detecting cell-type-specific differential expression in heterogeneous tumors** Follicular lym-
144 phoma (FL) is the second most common indolent non-Hodgkin lymphoma (NHL) in the USA and
145 Europe, accounting for nearly 20% of all NHL cases [40]. Previous work using FACS-sorted B

146 cells from FL tumors identified 612 differentially expressed genes in the presence of CREBBP
147 mutation [41]. Here, similarly to previous analysis [8], we asked whether deconvolving bulk FL
148 tumors (n=24, including 14 with CREBBP mutation) [8, 41] would allow us to detect the pre-
149 viously reported effects in B cells from FL tumors. Indeed, B cell expression levels estimated
150 by Unico from bulk FL tumors recapitulate the previously reported down- and up-regulation ef-
151 fects in FL B cells significantly better than alternative deconvolution methods (Figure 3a). More
152 specifically, none of the methods performed significantly better than the others on the up-regulated
153 genes, with the exception of the baseline method, which performed worse than all deconvolution
154 methods. However, Unico performed best on the down-regulated genes, and remarkably, it was the
155 only deconvolution method that performed significantly better than a straightforward bulk analysis
156 (adjusted p-value<0.05; Paired Wilcoxon test).

157 **Unico improves resolution and robustness in epigenetic association studies** We expected that
158 modeling and effectively estimating cell-type covariance will allow Unico to yield better perfor-
159 mance in downstream applications that aim at disentangling signals between cell types. In order to
160 demonstrate this, we evaluated the different deconvolution methods in calling cell-type level differ-
161 ential methylation (DM). While ground truth DM is generally unknown, one can consider the con-
162 sistency of a given method across different datasets as a surrogate for true/false positive/negative
163 rates.

164 We applied each method for testing a set of 177,207 CpGs for cell-type level DM in four
165 large whole-blood methylation datasets (n>590 each) with sex and age information [39, 42, 43].

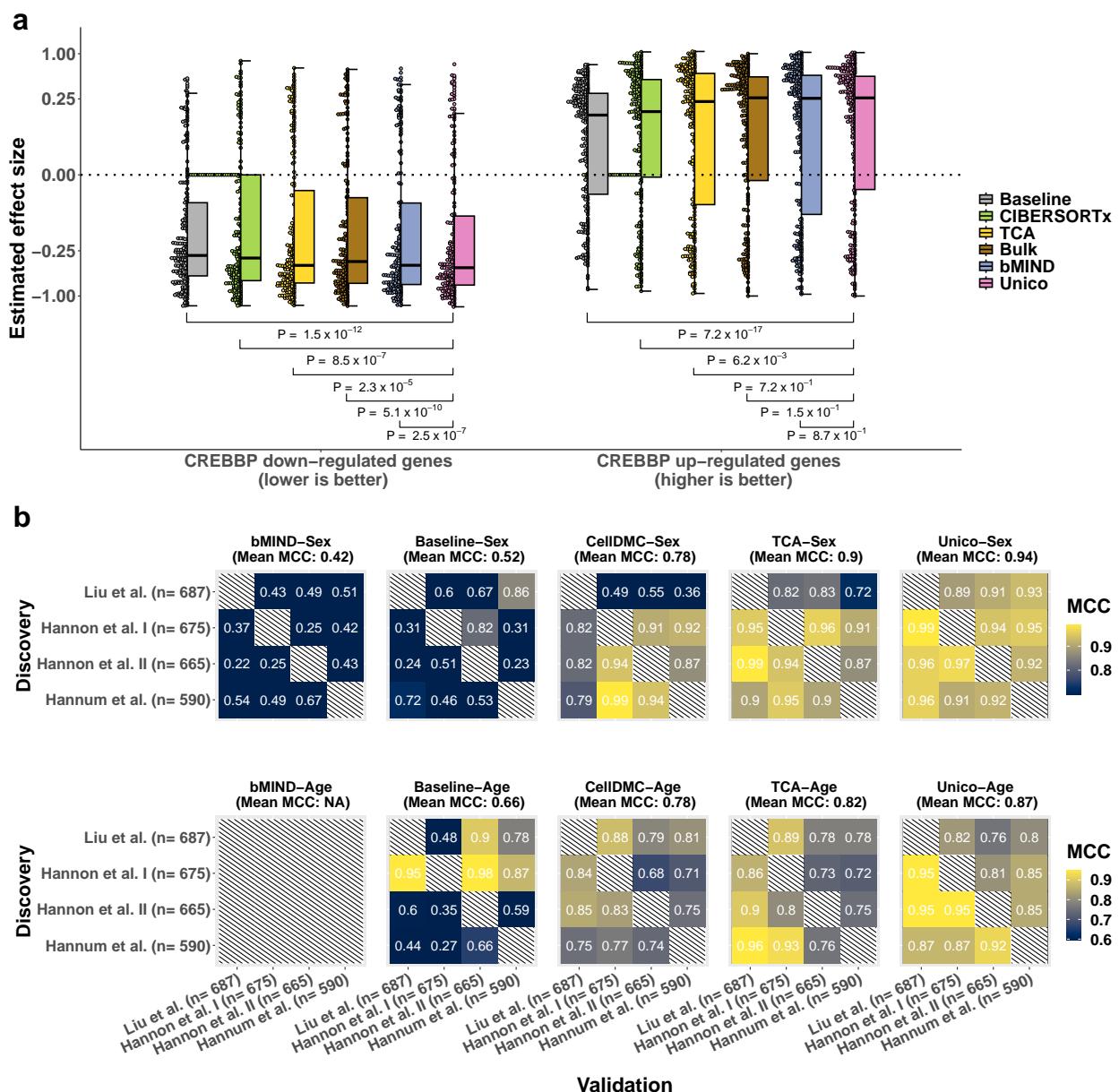


Figure 3: Application of deconvolution to downstream analysis tasks. (a) Deconvolution of bulk FL tumor samples for assessing previously reported CREBBP mutation-related gene expression in B cells. Presented are deconvolution-based B cell effect size distributions for 219 down-regulated and 275 up-regulated genes; comparisons to Unico were calculated using a one-sided paired Wilcoxon test. (b) Consistency in calling cell-type level differential methylation with sex and age across four independent whole-blood DNA methylation datasets. Color gradients represent the Matthews correlation coefficient (MCC) for every possible pairing of two datasets as discovery and validation (Methods). Since bMIND was designed for binary conditions only, it was not evaluated in the age analysis

¹⁶⁶ Specifically, for every possible combination of two out of the four datasets as discovery and val-
¹⁶⁷ idation data, we measured the consistency between datasets using the Matthews correlation coef-
¹⁶⁸ ficient (MCC) [44] (Methods). We excluded from this analysis CIBERSORTx, due to its runtime
¹⁶⁹ (Supplementary Figure S10) and poor performance in deconvolving bulk methylation (Figure 2e;
¹⁷⁰ Supplementary Figures S11-S15). Instead, we considered CellIDMC, a method that was designed
¹⁷¹ specifically for detecting cell-type level DM by evaluating linear effects of interaction terms be-
¹⁷² tween the condition of interest and cell-type proportions [45]. We observe that Unico provides
¹⁷³ the best overall consistency (Figure 3b), and it significantly improves upon TCA, the second best
¹⁷⁴ method (p-value \leq 0.05 for both sex and age; one-sided paired Wilcoxon test). Importantly, the
¹⁷⁵ runtime of Unico was on par with TCA's (Supplementary Figure S10).

¹⁷⁶ The above evaluation disregards a straightforward analysis of the bulk data, which cannot

177 associate DM with specific cell types but rather call CpGs as generally associated with conditions
178 (“tissue-level” analysis). Intuitively, models that provide cell-type resolution are more realistic
179 and are thus expected to improve cross-dataset consistency over a standard tissue-level analysis. In
180 order to verify this intuition, we evaluated a standard linear regression analysis of the bulk data for
181 calling tissue-level DM (Supplementary Figure S16). We observe that cell-type level analysis using
182 any of the deconvolution methods provides a substantial improvement in consistency compared to
183 the bulk analysis. In particular, Unico provides an increase of 107.5% and 40.7% in MCC for
184 sex and age, respectively. Further adapting the different deconvolution methods to call tissue-level
185 DM (Supplementary Methods) yields all methods as better than a standard bulk analysis, with
186 Unico being the top performing method (Supplementary Figure S16) These results demonstrate
187 how carefully modeling the cell-type signals in bulk data improves analysis even if constrained to
188 a tissue-level context.

189 **3 Discussion**

190 We propose Unico, a deconvolution method that is theoretically appropriate for any bulk genomic
191 data type that reflects mixtures of signals across cell types. Here, we demonstrate the utility of
192 Unico for gene expression and DNA methylation, however, our distribution-free treatment sug-
193 gests its applicability to other genomic data types as well. Unico leverages covariance across cell
194 types, and as such, it deconvolves particularly well low-entropy features that exhibit non-trivial
195 correlation structure between cell types. Remarkably, our evaluation, based on two scRNAseq
196 datasets from different tissues and purified methylation data, demonstrates that Unico considerably

197 outperforms state-of-the-art methods in general, even when deconvolving high entropy features.

198 Finally, Unico has some limitations, and while these limitations are not unique to Unico but
199 are rather common to all the deconvolution methods we evaluated, they may potentially bias and
200 affect the performance of our proposed model. First, given that lowly abundant cell types repre-
201 sent only a small fraction of the variance in bulk data, Unico is expected to perform poorly when
202 attempting to model a large number of cell types. Since heterogeneous tissues often represent
203 mixtures of a large number of cell types and subtypes, the deconvolution of Unico may be biased
204 by unmodeled cell types. Another limitation of Unico pertains to the assumption that cell-type
205 proportions of the input bulk samples are known. Admittedly, this information is rarely available
206 in bulk genomics data, so proportions need to be estimated in practice. While it is commonplace
207 to employ reference-based methods for learning cell-type compositions, using estimates in place
208 of measurements creates yet another source of noise and potential bias. Despite these concerns,
209 we conclude that our comprehensive evaluation of Unico across diverse datasets and data modal-
210 ities provides compelling evidence of its superiority over existing state-of-the-art deconvolution
211 methods.

212 **4 Methods**

Unico: a model for uniform cross-omics deconvolution We denote X_{ij} to be the (tissue-level) bulk gene expression in sample $i \in \{1, \dots, n\}$ of gene $j \in \{1, \dots, m\}$. For simplicity of exposition, we use the notion of gene expression, however, j can represent any other genomic feature that may

vary across cell types. We assume:

$$X_{ij} = w_i^T Z_{ij} + (c_i^{(2)})^T \beta_j + e_{ij} \quad (3)$$

$$E[e_{ij}] = 0, V[e_{ij}] = \tau_j^2 \quad (4)$$

213 The first term in Equation (3) defines X_{ij} as a weighted linear combination of cell-type expression
 214 levels. Specifically, $w_i = (w_{i1}, \dots, w_{ik})$ is a vector of sample-specific cell-type proportions of k
 215 cell types that are assumed to compose the studied tissue and $Z_{ij} = (Z_{ij1}, \dots, Z_{ijk})$ is a vector of
 216 cell-type expression levels of gene j in sample i . The second and third terms in Equation (3) model
 217 systematic and non-systematic variation. Specifically, e_{ij} is an i.i.d. component of variation that
 218 reflects measurement noise, $c_i^{(2)}$ is a p_2 -length vector of known covariate values of sample i that
 219 may be associated with unwanted global effects (i.e., “tissue-level” effects that may affect many
 220 genes and are not cell-type-specific, such as batch effects), and β_j is a vector of the corresponding
 221 gene-specific fixed effect sizes.

We assume that cell-type proportions $\{w_i\}$ are fixed and given. In practice, these can be estimated using a reference-based approach (e.g., [14, 34]), as suggested by other deconvolution methods [7–11]). In contrast to a standard decomposition problem, which assumes shared cell-type expression levels across all samples, the unknown $\{Z_{ij}\}$ components are modeled as random variables; this is emphasized by the use of upper-case notation. Specifically, for Z_{ijh} , the gene expression in sample i of gene j and cell type $h \in \{1, \dots, k\}$, we assume:

$$Z_{ijh} = \mu_{jh} + (c_i^{(1)})^T \gamma_{jh} + \epsilon_{ijh} \quad (5)$$

$$E[\epsilon_{ijh}] = 0, V[\epsilon_{ijh}] = \sigma_{jh}^2 \quad (6)$$

222 where μ_{jh} is the mean level, specific to gene j and cell type h , ϵ_{ihj} is an i.i.d. noise term with mean
 223 zero and variance σ_{jh}^2 that may be specific to gene j and cell type h , $c_i^{(1)}$ is a p_1 -length vector of
 224 known covariate values of sample i that may present cell-type-specific effects, and γ_{jh} is a vector
 225 of corresponding fixed effect sizes.

Lastly, we further model cell-type covariance. Concretely, we model the covariance of a given gene j across cell types h, q and denote:

$$\sigma_{jh,jq} \equiv \text{Cov}[Z_{ijh}, Z_{ijq}], \quad \sigma_{jh,jh} \equiv \sigma_{jh}^2 \quad (7)$$

226 The Unico model makes no assumptions on the distribution of the components of variation
 227 in Equations (3) and (5), which makes it naturally applicable to all heterogeneous tissue-level
 228 omics that can be represented as linear combinations of cell-type level signals. Finally, Unico can
 229 be viewed as a generalization of the TCA model and as a frequentist alternative for the bMIND
 230 model. See Supplementary Methods for details.

Estimating the underlying 3D tensor with Unico. Given a single realization x_{ij} of the bulk level coming from X_{ij} , we wish to learn z_{ij} , the realization of the cell-type-specific expression levels Z_{ij} of the corresponding sample i and gene j . Our goal is hence to compose a 3D tensor (samples by genes by cell types) based on the 2D input matrix. We address this problem by setting the estimator of z_{ij} to be the expected value of the conditional distribution $Z_{ij}|X_{ij}$:

$$\hat{z}_{ij} = \text{E}[Z_{ij}|\theta_j, w_i, X_{ij} = x_{ij}] \quad (8)$$

where θ_j is the set of parameters that are specific to gene j , that is,

$$\theta_j = \{\mu_{jh}\}_h \cup \{\gamma_{jh}\}_h \cup \{\beta_j\} \cup \{\sigma_{jh,jq}\}_{h,q} \quad (9)$$

231 The following theorem provides an analytical solution for the estimator \hat{z}_{ij} under the Unico model
 232 in Equations (3)-(7).

Theorem 1 (The Unico 3D tensor estimator) *The solution for the estimator stated in Equation (8)*

under the Unico model is given by:

$$\hat{z}_{ij} = E[Z_{ij}|\theta_j] + (Sum((w_i w_i^T) \odot \Sigma_j) + \tau_j^2)^{-1} \Sigma_j w_i \left(x_{ij} - w_i^T (\mu_j + (c_i^{(1)})^T \gamma_j) \right) - (c_i^{(2)})^T \beta_j$$

233 where $\gamma_j = (\gamma_{j1}, \dots, \gamma_{jk}) \in \mathbb{R}^{p_1 \times k}$ is a matrix composed of the vectors $\{\gamma_{jh}\}$, $\Sigma_j \in \mathbb{R}^{k \times k}$ is the
 234 cell-type covariance matrix of gene j , the \odot operator is the Hadamard product of two matrices,
 235 and the $Sum(\cdot)$ operator is a summation across all entries of a matrix.

236 Proof is given in the Supplementary Methods.

237 Theorem 1 provides an analytical solution for the 3D tensor given the cell-type proportions
 238 $\{w_i\}$ and model parameters θ_j . As mentioned above, in practice, cell-type proportions are es-
 239 timated using decomposition methods, and as we later describe, the model parameters can be
 240 estimated from the observed bulk data and the estimated cell-type proportions.

241 Unico essentially defines the estimator \hat{z}_{ij} as the expected value of the conditional distribu-
 242 tion $Z_{ij}|X_{ij} = x_{ij}$, which was previously suggested in TCA [7]. However, Under the richer Unico

243 model this conditional distribution becomes more informative owing to the correlation structure
244 between cell types. Intuitively, learning cell-type levels that better capture cell-type covariance
245 will enhance our capacity to assign deconvolution signals accurately to the respective cell types in
246 downstream analysis.

247 A-priori one may wonder whether modeling cell-type covariance is necessary for a decon-
248 volution method to recapitulate the true cell-type covariance in the data. Put differently, one could
249 expect an accurate deconvolution method to capture cell-type covariance regardless of an explicit
250 modeling of the covariance. However, our empirical results suggest that such modeling is valuable
251 for accurate deconvolution, and the following theorem provides intuition into why modeling the
252 covariance is indeed desired in order to achieve accurate deconvolution. Besides Unico, TCA [7]
253 is the only existing deconvolution method that offers an analytical estimator for the 3D tensor.
254 Hence, the following exclusively focuses on Unico and TCA, as the theoretical analysis for other
255 methods remains unclear.

256 **Theorem 2 (Improved capacity to reduce covariance bias)** *Assume for simplicity $\forall h : \mu_{jh} =$*
257 $0, \sigma_{jh}^2 = 1, \tau_j = 0$, *and no covariates for some feature j under Equations (3)-(7). If $n \rightarrow \infty$ then*
258 *(i) the cell-type covariances of the 3D tensor estimated by TCA are fixed and do not depend on*
259 *feature j , and (ii) the cell-type covariances of the 3D tensor estimated by Unico are a function of*
260 *the cell-type covariance of feature j .*

261 Proof is given in the Supplementary Methods.

262 **Optimization** We estimate the parameters of the model by following concepts from the General-
263 ized Method of Moments (GMM) [46]. The GMM framework allows us to learn the parameters of
264 a model by iteratively solving equations (moment conditions) that match population moments (or,
265 more generally, a function of population moments) with their corresponding data-derived sample
266 moments. We tailor the optimization to the Unico model to form asymptotically consistent estima-
267 tors as in classical GMMs [46], while introducing practical considerations and constraints that are
268 essential for finite data. The full details about the optimization and implementation of Unico are
269 provided in the Supplementary Methods.

270 **Implementation of Unico and practical considerations** We implemented Unico in R. In order
271 to stabilize the parameter estimation, in practice, we consider non-negativity constraints when
272 estimating the means and a small L_2 penalty when estimating the variances and covariances in
273 the model. The latter alleviates the risk of multicollinearity and therefore inaccurate estimation
274 owing to the high correlation between the proportions of different cell types. Additionally, when
275 estimating the parameters of a given feature, we disregard samples with values that diverge from the
276 mean by more than two standard deviations. This measure prevents extreme and non-representative
277 data points from dominating the solution.

278 We optimize the Unico model iteratively. At the end of each iteration, we update the weights,
279 which can then be used for weighting the samples in the following iteration (Supplementary Meth-
280 ods). At a given iteration, we learn the means using the constrained least squares solver `pcls`
281 from the `mgcv` R package, and we learn the variances and covariances using the COBYLA al-

282 gorithm [47] as implemented in the `nloptr` R package [48]. Empirically, we found that Unico
283 works well using as few as two iterations (i.e., updating the weights once) for estimating the means
284 and using three iterations for estimating the variances and covariances (data not shown).

285 **PBMC and lung scRNAseq data** We obtained the PBMC scRNAseq dataset from a COVID-19
286 study by Stephenson et al. [29]. We arbitrarily selected only one sample for donors with multiple
287 measurements, which resulted in a total of 118 samples for the analysis. After excluding cells with
288 high percentage of hemoglobin ($\geq 1\%$) or mitochondria ($\geq 5\%$), and low percentage of ribosomal
289 content ($\leq 1\%$), in addition to requiring a minimal and maximal number of unique expressed genes
290 ($\geq 500, \leq 2500$) and total UMI counts ($\geq 2000, \leq 15000$), 499,336 cells remained for the analysis.

291 In addition, we used scRNAseq from the data collection presented by Sikkema et al. [37] as part of
292 a study for integrating multiple datasets collected from the human respiratory system. We focused
293 on the lung parenchyma samples (n=90) that composed most of the carefully annotated group of
294 samples in the original study (defined by the authors as the “core reference” group). Employing
295 the same data filtering criteria as for the PBMC data resulted in a total of 296,227 cells for the
296 analysis. For both the PBMC and lung datasets we used the cell-type annotations provided by the
297 authors and applied a counts per million (CPM) normalization.

298 **Gene expression data with follicular lymphoma** We used a preprocessed microarray bulk FL
299 data (n=302) by Newman et al. [8]. In total, out of the 302 samples available, 14 were confirmed
300 to have the CREBBP mutation and 10 samples were confirmed to exhibit a wild-type allele. The
301 CREBBP status for 12 of these samples was collected by Green et al. [41] and the remaining 12

302 samples by Newman et al. [8]; the CREBBP status of all 24 samples was made available in the
303 supplementary files of Newman et al. For defining a ground truth list of differentially expressed
304 genes with CREBBP mutation in FL B cells, we considered the set of 334 up-regulated and 279
305 down-regulated genes that were previously reported in a study with sorted B cells from FL tu-
306 mors [41]. Intersecting these sets with the genes available in the bulk FL data left us with 275 and
307 219 up- and down-regulated genes for evaluation.

308 **Whole-blood DNA methylation datasets** We used a total of five beta-normalized DNA methyla-
309 tion datasets that were collected using the Illumina 450K methylation array. For the methylation
310 deconvolution analysis, we obtained data from Reinius et al. [38], including whole-blood (n=6)
311 and matching cell-sorted methylation data from the same individuals (granulocytes, monocytes,
312 NK, B, CD4 T, and CD8 T cells). For the cell-type level differential methylation (DM) analysis,
313 we considered whole-blood datasets from liu et al. (n=687) [42], Hannum et al. (n=590; samples
314 with missing smoking status were excluded) [39], and two datasets from Hannon et al. (n=675,
315 n=665) [43]. In all datasets, we removed CpGs with non-autosomal, polymorphic, and cross-
316 reactive probes [49], and we excluded low variance CpGs (variance<0.001). This left us with
317 153,155, 144,743, 134,250, and 95,360 CpGs for the Liu, Hannum, and the two Hannon datasets,
318 respectively. For the Reinius dataset, we considered CpGs at the intersection between the Reinius
319 data and a preprocessed version of the Hannum dataset (restricted to samples with European an-
320 cestry; 93,086 CpGs). Lastly, cell-type proportions were estimated for all whole-blood datasets
321 using EpiDISH, a reference-based methylation decomposition method [50].

322 **Implementation and application of competing deconvolution and cell-type association meth-**

323 **ods** We ran all CIBERSORTx [8] related codes under a docker container version 1.0 encapsulating

324 both the “High Resolution” mode (for estimating cell-type level profiles) and the “Fractions” mode

325 (for estimating cell-type proportions) with default parameters and authentication token granted by

326 the CIBERSORTx team upon request. CIBERSORTx evaluates the maximum value in a bulk input

327 and automatically assumes the data have been log-normalized if the maximum is less than 50. This

328 choice is reasonable for transcriptomic data, for which CIBERSORTx was designed, however, it is

329 not justified for beta-normalized methylation levels that are restricted to the interval $[0, 1]$. We thus

330 scaled the methylation beta values by a factor of 10,000 prior to the application of CIBERSORTx

331 and rescaled the results back to original scale.

332 We installed the TCA [7] R CRAN package version v1.2.1 deposited on CRAN and evalu-

333 ated its performance under default parameters. We fitted the model using the function `tca` and

334 performed deconvolution using the `tensor` function. For the cell-type level DM analysis, both

335 the joint (tissue-level) and marginal (cell-type level) statistical tests were automatically evaluated

336 as part of the model parameter learning step in the `tca` function.

337 bMIND [10] is available via the MIND R CRAN package version 0.3.3. We obtained the cell-

338 type specific profiles and the estimated model parameters with the function `bMIND` and performed

339 association testing with the function `ttest`. bMIND evaluates the maximum value in the bulk

340 input and automatically log transforms the data if the maximum is larger than 50. We therefore

341 scaled the bulk expression profile (and the single-cell derived prior) by the inverse of the largest

342 detected value before applying bMIND, and then rescaled the output back to the original scale. This
343 approach ensured consistency and comparability across all deconvolution methods. Specifically,
344 allowing the default log transformation of the data would have violated the assumption that bulk
345 levels represent linear combinations of cell-type levels.

346 Throughout this work, we also evaluated a baseline approach in our analysis and evaluation
347 by simply considering the product of the observed bulk data and the cell-type proportions as cell-
348 type level estimates. That is, we estimated z_{ijh} , the cell-type level of sample i , gene j , and cell
349 type h as $z_{ijh}^{\text{Baseline}} = x_{ij} \cdot w_{ih}$. Finally, we applied CellDMC [45] for DM using the implementation
350 in the Bioconductor R package EpiDISH, version 2.10.0.

351 **Deconvolving mixtures of gene expression profiles and estimating cell-type level moments** We
352 used both the PBMC and lung scRNAseq datasets for generating pseudo-bulk mixtures. Briefly,
353 for creating a new pseudo-bulk sample, we first drew (with replacement) all cell-type level profiles
354 of one randomly selected sample. The cell-type profiles of each individual sample were defined
355 as normalized pseudo-bulk counts at the cell-type level. We then drew (with replacement) the
356 cell-type proportions of one randomly selected sample in the data (total number of cells coming
357 from each cell type, normalized to sum up to 1). Eventually, these were used as the weights for a
358 weighted linear combination of the cell-type level profiles to create one pseudo-bulk sample.

359 In the PBMC analysis we considered either five major cell-type groups (monocytes, NK, B,
360 CD4 T, and CD8 T cells) or seven cell types by further stratifying B cells into canonical B cells and
361 plasma cells and monocytes into CD16 and CD14 monocytes. In the analysis with lung cells we

362 considered either four major cell-type groups (endothelial, stromal, immune, and epithelial cells)
363 or six cell types by further stratifying immune cells into myeloid and lymphoid compartments and
364 epithelial cells into airway and alveolar epithelium cells. Our evaluation was restricted for the top
365 10,000 most highly expressed genes in the data. See Supplementary Methods for more details.

366 The pseudo-bulk mixtures, along with the corresponding mixing proportions, were provided
367 as the input for all deconvolution methods to learn 3D tensors. We assessed these tensors for their
368 accuracy by comparing them against the known cell-type profiles. Particularly, for a given cell
369 type and a given gene, we evaluated the correlation between the true cell-type expression levels of
370 the pseudo-bulk samples and their deconvolution-based estimates.

371 We obtained estimates of population-level cell-type moments from the data (means, vari-
372 ances, and covariances per gene) directly from the output of the deconvolution methods. For
373 methods which do not explicitly output such estimates (e.g., no method except for bMIND and
374 Unico outputs covariance estimates), we used the estimated tensor for calculating these moments.
375 To evaluate the accuracy of the estimated moments, we established gold standard estimates based
376 on the cell-type profiles underlying the pseudo-bulk mixtures. In order to mitigate the potential
377 influence of outliers, we considered only samples within 2 standard deviations from the mean for
378 the moments estimation of a given gene.

379 Finally, we used multiple linear regression for evaluating whether an estimated 3D tensor of
380 a given deconvolution method captures variation of the true tensor beyond its correlation with the
381 deconvolution input (i.e., pseudo-bulk and cell-type proportions). In more detail, for every gene

382 and cell type, we fitted a regression model for the known cell-type expression levels as the depen-
383 dent variable using several independent variables, including the pseudo-bulk levels of the gene,
384 the cell-type proportions, and the cell-type tensor estimates. This allowed us to quantify to what
385 extent the deconvolution-based estimates provide information beyond the bulk data. Specifically,
386 we defined $\Delta \log_{10}(p\text{-value})$ as the difference between the log-scaled (basis 10) t-test derived p-
387 values of the pseudo-bulk variable and the estimated cell-type levels in the regression. Of note, we
388 defined the p-values to be 1 in cases where cell-type levels were estimated to have no variation. In
389 order to mitigate potential biases due to heavy-tailed distributions of expression levels, we log1p-
390 transformed expression levels and considered only samples within 2 standard deviations from the
391 mean.

392 **Deconvolving the Reinius whole-blood DNA methylation data** Unlike our deconvolution of
393 gene expression mixtures, the size of the Reinius data (n=6) does not allow for drawing reliable
394 conclusions through a straightforward evaluation. Particularly, Unico, as well as current decon-
395 volution methods, are designed to operate on large bulk data. We circumvented this limitation by
396 taking a two-step reference-based procedure. First, we learned the parameters of the Unico model
397 from the larger Hannum whole-blood methylation data [39]. Acknowledging that population struc-
398 ture affects methylation [51], we focused solely on Caucasian individuals from the Hannum data
399 (n=426), anticipating that they would adequately represent the Swedish individuals in the Reinius
400 study. Then, we plugged these parameter estimates into Unico's 3D tensor estimator together with
401 the Reinius bulk profiles and their cell-type proportion estimates. We performed the same proce-
402 dure for TCA, however, CIBERSORTx and bMIND, which do not provide an analytical estimator

403 of the tensor, required a different strategy. In order to inform the deconvolution of CIBERSORTx
404 and bMIND with the same additional information, we applied these methods to the concatenation
405 of the Reinius and Hannum datasets and extracted the cell-type level estimates for the Reinius
406 samples.

407 Benchmarking methods based on the Reinius data presents a second challenge: determining
408 a proper way to evaluate their performance given that data from only six individuals is available
409 for the analysis. We tackle this limitation by collapsing methylation levels in the estimated tensor
410 along both the CpGs and samples axes. That is, for every cell type, we evaluated how correlated
411 is the vector of all methylation estimates of the cell type (i.e., by pooling estimates across all
412 CpGs and samples) with the experimentally measured ground truth levels from purified cells. This
413 yielded a single correlation score per cell type. Importantly, when stacking CpGs for evaluation, a
414 deconvolution that only correctly estimates relative means and scales of CpGs but performs poorly
415 in terms of per-CpG correlation (i.e., across samples) may achieve spuriously high correlation
416 levels. We addressed this by removing from every CpG its cell-type level mean methylation level.

417 Since beta-normalized methylation levels are bounded to the range [0,1], unlike in the de-
418 convolution of relative expression levels, we further evaluated the divergence of the estimated 3D
419 tensors from the true cell-type levels in absolute terms. Specifically, we evaluated the root median
420 square error (RMSE) between the true and each estimated 3D tensor; we expected that a median
421 metric in place of a standard mean square error would improve robustness to outliers. Similarly to
422 the evaluation of correlation, we calculated an RMSE value per cell type after collapsing methyla-

423 tion levels in the tensors along both the CpGs and samples axes.

424 Finally, our benchmarking focused either on randomly selected CpGs or on a set of highly
425 variable CpGs based on the Reinius data. For defining the latter, we ranked the CpGs in the
426 intersection of the Reinius and Hannum datasets (93,086 CpGs) by the sum of their variances in
427 the different cell types using the sorted methylation Reinius data and chose the top 10,000 CpGs
428 with the largest values.

429 **Calculating robust linear correlation** All the correlation values reported throughout our analysis
430 and evaluation were calculated using a robust linear correlation metric in place of the standard
431 Pearson correlation. Specifically, we used the function `cov.rob` from the MASS R package [52],
432 which performs an approximate search for a subset of the observations to exclude such that a
433 Gaussian confidence ellipsoid is minimized in volume. Effectively, this procedure trims outliers
434 that may otherwise dramatically bias correlation levels. In particular, if either input vector has an
435 interquartile range (IQR) of 0, `cov.rob` defines the correlation as 0. Throughout the paper, we
436 set the fraction of outliers to exclude to 5% of the data points.

437 **Calculating von Neumann entropy** We quantify the amount of signal coming from the covari-
438 ance structure of a given gene by the von Neumann entropy [53]. For a given gene, the von
439 Neumann entropy is defined as the entropy applied to the eigenvalues of the normalized cell-type
440 covariance matrix of the gene (i.e., a $k \times k$ matrix of correlations between cell types). High entropy
441 corresponds to cases where no substantial cell-type covariance structure exists, and low entropy in-
442 dicates strong positive or negative correlations between cell types. Throughout our evaluation of

443 the deconvolution results we grouped genes into high- and low-entropy groups. This classification
444 was based on ranking the genes by their entropy and assigning genes with above-median (below-
445 median) entropy to the high (low) entropy group. Lastly, the normalized von Neumann entropy
446 presented in figure 1b simply refers to von Neumann entropy values scaled to the range [0,1].
447 Since the von Neumann entropy is bounded by a number that depends on the number of cell types
448 k , this normalization enables us to evaluate and visualize the distribution of entropy across genes
449 using covariance matrices of different sizes.

450 **Deconvolving bulk profiles from follicular lymphoma tumors** For every deconvolution method,
451 we first estimated the 3D tensor of the bulk FL dataset (n=302) while considering only the sets of
452 275 and 219 genes that were previously reported as up- and down-regulated with the CREBBP
453 mutation. We provided each method with cell-type proportions estimated using CIBERSORTx
454 (“Fractions” mode) with the LM22 signature matrix [54], while collapsing the estimated propor-
455 tions into 4 categories: B cells, CD4 T cells, CD8 T cells, and “remaining”.

456 A straightforward evaluation would include calculating for every method log-fold changes
457 (LFCs) with the CREBBP mutation based on the estimated B cell expression levels. This would
458 allow assessing the concordance between the LFCs and the previously-reported direction of the
459 differentially expressed genes. However, the group of CREBBP-mutated tumors presents an el-
460 evated B cell composition, which is expected to lead to an overly-optimistic performance on the
461 set of up-regulated genes in cases of deconvolution estimates that are biased by cell composition
462 (Supplementary Figure S17). Most notably, since the baseline method estimates B cell expression

463 levels by naively multiplying bulk levels by B cells proportions, the baseline estimates are ex-
464 pected to be artificially higher for samples with higher B cells composition. The baseline method
465 therefore consistently estimates higher B cell expression levels for the CREBBP-mutated tumors,
466 regardless of whether the genes are truly down- or up-regulated. Consequently, genes that are truly
467 up-regulated in CREBBP tumors are expected to present strong LFCs under the baseline given the
468 combination of both real and artificial up-regulation effects.

469 In order to account for the B cell composition bias, we used multiple linear regression to test
470 whether the estimated tensors capture the mutation effects beyond the effect of B cell composition.
471 In more detail, for every gene, we fitted a regression model for the estimated B cell expression lev-
472 els as the dependant variable using the B cell composition and the mutation status as independent
473 variables. We performed the same procedure while using the bulk expression levels as the depen-
474 dent variable to evaluate a standard analysis of bulk expression. In order to allow a comparable
475 evaluation of the estimated mutation effect sizes across the different methods and to alleviate the
476 potential effect of outliers, we standardized the log1p-scaled B cell expression estimates of every
477 gene. For methods that do not constrain non-negativity in their estimated tensor, for every gene and
478 cell type, we shifted the distribution of the estimates by subtracting the minimum value detected,
479 which enforced non-negatively prior to the log1p transformation. The effect size of a gene that was
480 estimated to have a constant B cell expression level across all samples was set to 0.

481 **Cell-type level epigenetic association studies with sex and age** We performed statistical testing
482 for calling DM using Unico, TCA [7], bMIND [10], and CellDMC [45] (Supplementary Methods).

483 As a baseline model, we evaluated the linear effects of the conditions on the tensor estimates of
484 our baseline deconvolution. Concretely, for a given CpG and cell type, we fitted a linear regression
485 model with the baseline-estimated cell-type level methylation as the dependent variable and the
486 condition (and covariates) as the independent variable. This allowed us to calculate t-statistics and
487 derive p-values for the cell-type level effects of the conditions under a baseline deconvolution.

488 Our analysis included cell-type level covariates ($\{c_i^{(1)}\}$ under the Unico notations) and tissue-
489 level covariates ($\{c_i^{(2)}\}$ under the Unico notations). For cell-type covariates, we considered age
490 and sex in the analysis of all four whole-blood methylation datasets (Liu et al. [42], Hannum et
491 al. [39], and two cohorts by Hannon et al. [43]). In addition, we accounted for rheumatoid arthritis
492 and smoking status in the Liu data, schizophrenia status in the Hannon data, and ethnicity and
493 smoking status in the Hannum data. Across all datasets, smoking status was classified into three
494 major categories: never, past, and current smoker. For tissue-level covariates, we considered sur-
495rogates of technical variability. In more detail, for each methylation dataset, prior to filtering any
496 CpG, we took a previously suggested approach [7, 28] of estimating factors of technical variation
497 by calculating the top 20 principal components (PCs) of the 10,000 least variable CpGs of each
498 methylation array. We expected these PCs to capture only global technical variation and no bio-
499 logical variation due to the use of CpGs with nearly constant variance. In addition to these PCs,
500 we further accounted for plate information, which was available for the Hannum data. All the
501 benchmarked methods were designed to account for cell-type and tissue-level covariates, except
502 for the baseline model. For the latter, we simply included the full set of covariates as independent
503 variables in the linear regression model.

504 The inter-individual distribution of array-probed methylation levels is approximately nor-
505 mally distributed for most CpGs. For that reason, TCA and CellDMC, which were designed for
506 methylation data, assume the data is normally distributed; bMIND assumes normality as well, even
507 though it was not designed for methylation. We therefore similarly applied statistical testing under
508 a normality assumption when evaluating Unico on calling DM (Supplementary Methods). No-
509 tably, this assumption is not required given that the Unico framework is generally distribution-free
510 and allows us to derive asymptotic p-values (Supplementary Methods). Indeed, we empirically ob-
511 serve that asymptotically-derived p-values are highly correlated with their parametric counterparts,
512 while also being calibrated under the null (Supplementary Figure S18-S21).

513 For any given ordered pair of datasets (discovery and validation), we considered the CpGs at
514 the intersection of the two datasets. True positives (TPs) were defined as CpGs that are (i) genome-
515 wide significant in the discovery dataset under a Bonferroni-corrected threshold and (ii) significant
516 in the validation dataset, under a Bonferroni-corrected threshold adjusting for the number of sig-
517 nificant hits identified in the discovery data. CpGs that only satisfied condition (i) and either failed
518 to satisfy condition (ii) or demonstrated inconsistent direction of their estimated effect size were
519 considered as false negatives (FNs). CpGs with $p\text{-value} > 0.95$ in the discovery dataset were con-
520 sidered as negative controls for the evaluation of false positives (FPs) and true negatives (TNs).
521 That is, negative controls with significant (non-significant) p-values under a Bonferroni-corrected
522 threshold adjusting for the number of negative controls in the validation data were counted as FPs
523 (TNs).

524 Finally, as a metric of consistency across datasets, we calculated the MCC per method for
525 every pair of discovery and validation datasets. We favored MCC over the widely-used F1 score
526 since the former incorporates true negatives, which makes it a better choice for assessing model
527 performance on imbalanced class distributions [44]. Yet, for completeness, we further considered
528 the F1 score as the consistency metric, which revealed qualitatively similar results (Supplementary
529 Figure S22 and S23).

530 **Data availability**

531 The bulk FL data is available from GEO (accession number GSE127462). The whole-blood
532 methylation data with matching sorted cells, as well as the whole-blood methylation datasets
533 used for cell-type level DM analysis are available from GEO (accessions GSE35069, GSE42861
534 GSE40279, GSE80417, GSE84727). The PBMC scRNASeq dataset was downloaded from EMBL-
535 EBI (accession E-MTAB-10026), and the lung scRNASeq is available on cellxgene [55] as the
536 integrated Human Lung Cell Atlas.

537 **Code availability**

538 Unico is available as an R package under the GPL-3 license license at: <https://github.com/cozygene/>
539 Unico.

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