



Risk of hematological malignancies from CT radiation exposure in children, adolescents and young adults

Received: 20 March 2023

Accepted: 29 September 2023

Published online: 09 November 2023

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Over one million European children undergo computed tomography (CT) scans annually. Although moderate- to high-dose ionizing radiation exposure is an established risk factor for hematological malignancies, risks at CT examination dose levels remain uncertain. Here we followed up a multinational cohort (EPI-CT) of 948,174 individuals who underwent CT examinations before age 22 years in nine European countries. Radiation doses to the active bone marrow were estimated on the basis of body part scanned, patient characteristics, time period and inferred CT technical parameters. We found an association between cumulative dose and risk of all hematological malignancies, with an excess relative risk of 1.96 (95% confidence interval 1.10 to 3.12) per 100 mGy (790 cases). Similar estimates were obtained for lymphoid and myeloid malignancies. Results suggest that for every 10,000 children examined today (mean dose 8 mGy), 1–2 persons are expected to develop a hematological malignancy attributable to radiation exposure in the subsequent 12 years. Our results strengthen the body of evidence of increased cancer risk at low radiation doses and highlight the need for continued justification of pediatric CT examinations and optimization of doses.

The use of computed tomography (CT) has grown rapidly in most high-income countries¹ since its introduction² at the beginning of the 1970s. Although the benefits of CT imaging in patient management are undisputed, the potential increased cancer risk³ and relatively high cumulative doses incurred from multiple scans have raised concerns

in the medical and scientific community, leading to a plateauing/reduction in number of pediatric CTs in many countries^{4–6} and a reduction in pediatric doses⁷. A number of alternative modalities, including fast-acquisition magnetic resonance imaging and ultrasound are now replacing CT examinations for specific pediatric

indications⁸. Despite this, up to 7% of all CT procedures in high-income countries are performed on children².

While moderate-dose (≥ 100 mGy) to high-dose (≥ 1 Gy) ionizing radiation exposure is a well-established risk factor for leukemia, in both children and adults^{9,10}, the risk associated with childhood and adolescent low-dose exposure (< 100 mGy), the dose range typically associated with diagnostic CT examinations, is unclear. This is especially concerning given that CT scanning is the largest contributor to the world's average annual effective dose per person from medical radiation sources, in both children and adults^{2,11}.

Several studies estimated the hematological malignancies risk associated with CT scan radiation in children and young adults in large-scale national cohort^{12–18} and case–control studies^{19,20}. Although results of most individual studies^{12,13,17,20} and a recent meta-analysis²¹ suggest an increased risk of leukemia associated with repeated CT examinations, studies were criticized due to low statistical power, inadequate individual dosimetry and potential bias from confounding by indication (when those who undergo CT examinations are at higher risk of cancer than those who do not, due to underlying conditions)²². Current international radiological protection recommendations²³ are, therefore, mainly based on linear extrapolations of risk from the higher doses of the Japanese atomic bomb survivor studies²⁴. These extrapolations, which assume no dose threshold below which the risk of radiation-induced cancer is zero (the linear no threshold model of risk), are controversial^{10,25}.

The EPI-CT study, coordinated by the International Agency for Research on Cancer (IARC), was set up to overcome limitations of previous national studies and improve direct estimates of cancer risk from low-dose radiation exposure from CT scanning in childhood and adolescence. It included 948,174 individuals from nine European countries²⁶. In this Article, we present the EPI-CT analyses of risk of hematological malignancies in relation to radiation exposure from CT examinations in childhood, adolescence and early adulthood.

Results

Descriptive analyses

The analysis included 876,771 individuals, who underwent 1,331,896 CT examinations (mean 1.52, standard deviation (s.d.) 1.46 CT examinations per patient) and were followed up for at least 2 years following their first CT. They contributed 6,863,833 person-years (PYs) of follow-up (Table 1). We identified 790 hematological malignancies (subtype distribution in Supplementary Table 1), including 578 cases of lymphoid malignancies and 203 cases of myeloid malignancies and acute leukemia (AL). Mean follow-up was 7.8 years (6.5 years for cases). Fifty-one percent of the cases were younger than 20 years at diagnosis (ranging from 38% among mature T and natural killer (NK) cell neoplasms to 82% among precursor cell neoplasms), whereas 88.5% (range 76–99%) were younger than 30 years (Table 1 and Supplementary Table 2).

The distribution of age at first scan was skewed towards later ages, with 30% of the cohort (33% of cases) scanned at age 15 years or above (Table 1 and Supplementary Table 2). This distribution varied by outcome. Among lymphoid malignancies, 70.5% of Hodgkin lymphoma (HL) cases were ≥ 10 years at the time of their first CT, compared with 46.5% among non-Hodgkin lymphoma (NHL) cases. Among the latter, 62% of mature T and NK cell neoplasm cases were ≥ 10 years at the time of their first CT compared with 24% precursor cell neoplasm cases (Supplementary Table 2). Among myeloid malignancies and AL cases, the group of myeloproliferative neoplasms (MPNs), myelodysplastic syndrome (MDS) cases and myelodysplastic/myeloproliferative neoplasms (MDS/MPN) also tended to be older at first CT (65% ≥ 10 years).

About 58% of participants were born between 1985 and 1999 (Table 1). Countries contributed heterogeneously to the EPI-CT cohort (Table 1), with the United Kingdom, the Netherlands, Sweden and France representing 35%, 16%, 14% and 12% of individuals in the cohort, respectively (50%, 17%, 14% and 6% of cases).

The distribution of dose to the active bone marrow (referred to as ABM dose or dose throughout the Article) was strongly positively skewed, with most individuals having received low doses (Extended Data Table 1). The mean and median cumulative ABM dose at end of the follow-up were 15.6 mGy and 10.7 mGy (p25–p75: 5.8–18.2 mGy) (Table 1), respectively, in the cohort and 20 mGy and 13.0 mGy (p25–p75: 6.8–23.2 mGy) among cases overall.

As reported in the previous EPI-CT dosimetry paper⁷, the predominant body part scanned was the head, representing, with neck examinations, approximately 81% of all examinations. For this location, the mean ABM dose decreased by about 25% over the study period in newborns aged 0–3 months (from 15 mGy before 1991 to 12 mGy after 2001) but remained constant in adults aged 17.5 years and older (2.6 mGy). Dose reduction over time was greater for examinations of other body regions: for example, for chest CTs by more than 60% in newborns (from 18 to 7 mGy) and approximately 40% in adults (from 8 to 5 mGy).

Risk estimation

Elevated relative risks (RRs) for all hematological malignancies combined were observed across all dose categories ≥ 10 mGy, with a strong dose–response relationship and a RR of 2.66 (95% confidence interval (CI) 1.92 to 3.70) for doses ≥ 50 mGy compared with doses < 5 mGy (reference category) (Table 2). The estimated excess relative risk (ERR) per 100 mGy was 1.96 (95% CI 1.10 to 3.12). Elevated RRs were observed for lymphoid malignancies and for myeloid malignancies and AL separately in most dose categories compared with the reference (Table 2), with risk estimates generally increasing with dose. Continuous risk estimates were very similar for lymphoid malignancies (ERR/100 mGy 2.01, 95% CI 1.02 to 3.42) and myeloid malignancies and AL (ERR/100 mGy 2.02, 95% CI 0.47 to 4.77). The excess absolute risk (EAR) was estimated to be 17.7 per 100,000 PYs per 100 mGy (95% CI 11.6 to 24.0).

The ERR/100 mGy for NHL was 2.51 (95% CI 1.14 to 4.73) and for HL 1.24 (95% CI 0.08 to 3.28). Increasing trends in RRs with dose were seen for all subtypes (Table 3), although the CIs included unity for mature T and NK cell and for precursor cell neoplasms, and for the MPN + MDS + MDS/MPN grouping. An increased RR compared with the reference dose category was seen at doses as low as 10–15 mGy for NHL as a whole and for mature B cell neoplasms, the largest subgroup. A dose-dependent increase in RR was also seen for leukemia excluding chronic lymphocytic leukemia (CLL) in an analysis using previous classification for comparison with published estimates.

Potential confounders of the risk estimates

Removing birth cohort from the model and adjusting for socio-economic status (SES), where available, had little impact on risk estimates (Table 4). Analyses by country (Supplementary Table 3) showed similar numbers of cases of hematological malignancies in the United Kingdom as in the remaining countries combined (394 versus 396). The ERR/100 mGy was about twice as high in the United Kingdom compared with all other countries together, overall (ERR/100 mGy 2.69 versus 1.34), and for lymphoid malignancies and myeloid malignancies and AL separately. Risk estimates varied across countries, particularly for myeloid malignancies and AL, where numbers of cases were low, but estimates were statistically compatible. Analyses removing one country at a time confirmed that only the United Kingdom had a strong influence on the combined risk estimate (Supplementary Tables 3 and 4).

Potential modifiers of the risk estimates

There was no evidence of effect modification by sex, except for myeloid malignancies and AL where the elevated ERR was restricted to women (Table 4). The risk increased with increasing age at exposure, especially for lymphoid malignancies, with estimates in the 5–9 and ≥ 10 years at exposure groups about 2-fold and 3–4-fold those for the < 5 years group, respectively. Risk decreased with time since exposure, with

Table 1 | Characteristics of the cohort

	Hematological malignancies—numbers (%)						Entire cohort—n (%)	PYs of follow-up—n (%)	
	All cases	Lymphoid			All myeloid malignancies and AL	Histiocytic cell			
Overall	790 (100)	578 (73.2)	190 (24.1)	387 (49.0)	203 (25.7)	6 (0.8)	3 (0.4)	876,771 (100)	6,863,833 (100)
Sex									
Male	466 (59.0)	343 (59.3)	117 (61.6)	226 (58.4)	118 (58.1)	3 (50.0)	2 (66.7)	491,426 (56.0)	3,826,559 (55.7)
Female	324 (41.0)	235 (29.7)	73 (38.4)	161 (41.6)	85 (41.9)	3 (50.0)	1 (33.3)	385,345 (44)	3,037,274 (44.3)
Age at first CT (years)									
<1	93 (11.8)	70 (12.1)	12 (6.3)	58 (15)	22 (10.8)	0 (0)	1 (33.3)	100,628 (11.5)	789,500 (11.5)
1 to <5	126 (15.9)	95 (16.4)	13 (6.8)	82 (21.2)	30 (14.8)	1 (16.7)	0 (0)	149,483 (17.0)	1,159,795 (16.9)
5 to <10	132 (16.7)	98 (17.0)	31 (16.3)	67 (17.3)	33 (16.3)	1 (16.7)	0 (0)	168,135 (19.2)	1,308,483 (19.1)
10 to <15	169 (21.4)	123 (21.3)	58 (30.5)	65 (16.8)	42 (20.7)	3 (50.0)	1 (33.3)	190,561 (21.7)	1,525,680 (22.2)
≥15	270 (34.2)	192 (33.2)	76 (40.0)	115 (29.7)	76 (37.4)	1 (16.7)	1 (33.3)	267,964 (30.6)	2,080,375 (30.3)
Years since first CT examination at end of follow-up									
2 to <5	266 (33.7)	197 (34.1)	55 (28.9)	142 (36.7)	64 (31.5)	4 (66.7)	1 (33.3)	215,041 (24.5)	323,031 (4.7)
5 to <10	263 (33.3)	196 (33.9)	71 (37.4)	124 (32.0)	67 (33.0)	0 (0)	0 (0)	305,667 (34.9)	1,624,031 (23.7)
10 to <15	137 (17.3)	99 (17.1)	42 (22.1)	57 (14.7)	35 (17.2)	2 (33.3)	1 (33.3)	188,762 (21.5)	1,938,588 (28.2)
≥15	124 (15.7)	86 (14.9)	22 (11.6)	64 (16.5)	37 (18.2)	0 (0)	1 (33.3)	167,301 (19.1)	2,978,183 (43.4)
Birth cohort									
<1980	162 (20.5)	115 (19.9)	42 (22.1)	73 (18.9)	45 (22.2)	2 (33.3)	0 (0)	65,725 (7.5)	1,169,822 (17.0)
1980 to <1985	143 (18.1)	94 (16.3)	41 (21.6)	53 (13.7)	47 (23.2)	0 (0)	2 (66.7)	84,747 (9.7)	1,101,016 (16.0)
1985 to <1990	144 (18.2)	108 (18.7)	42 (22.1)	65 (16.8)	35 (17.2)	1 (16.7)	0 (0)	152,209 (17.4)	1,434,265 (20.9)
1990 to <1995	148 (18.7)	113 (19.6)	42 (22.1)	71 (18.3)	34 (16.7)	1 (16.7)	0 (0)	189,513 (21.6)	1,303,426 (19.0)
1995 to <2000	107 (13.5)	81 (14)	17 (8.9)	64 (16.5)	24 (11.8)	1 (16.7)	1 (33.3)	163,306 (18.6)	989,004 (14.4)
2000 to <2005	67 (8.5)	55 (9.5)	6 (3.2)	49 (12.7)	11 (5.4)	1 (16.7)	0 (0)	131,115 (15.0)	643,601 (9.4)
≥2005	19 (2.4)	12 (2.1)	0 (0)	12 (3.1)	7 (3.4)	0 (0)	0 (0)	90,156 (10.3)	222,700 (3.2)
Attained age, years									
2 to <20	404 (51.1)	299 (51.7)	75 (39.5)	224 (57.9)	99 (48.8)	5 (83.3)	1 (33.3)	435,894 (49.7)	2,132,791 (31.1)
20 to <30	295 (37.3)	214 (37)	97 (51.1)	116 (30.0)	78 (38.4)	1 (16.7)	2 (66.7)	320,706 (36.6)	2,752,949 (40.1)
30 to <40	86 (10.9)	62 (10.7)	18 (9.5)	44 (11.4)	24 (11.8)	0 (0)	0 (0)	104,767 (11.9)	1,633,680 (23.8)
≥40	5 (0.6)	3 (0.5)	0 (0)	3 (0.8)	2 (1.0)	0 (0)	0 (0)	15,404 (1.8)	344,414 (5.0)
Country									
Belgium	5 (0.6)	3 (0.5)	0 (0)	3 (0.8)	2 (1.0)	0 (0)	0 (0)	9,052 (1.0)	28,131 (0.4)
Denmark	8 (1.0)	7 (1.2)	2 (1.1)	5 (1.3)	0 (0)	1 (16.7)	0 (0)	15,835 (1.8)	68,053 (1.0)
France	47 (5.9)	43 (7.4)	8 (4.2)	35 (9)	4 (2.0)	0 (0)	0 (0)	104,542 (11.9)	453,713 (6.6)
Germany	23 (2.9)	19 (3.3)	2 (1.1)	17 (4.4)	4 (2.0)	0 (0)	0 (0)	39,501 (4.5)	162,615 (2.4)
The Netherlands	137 (17.3)	98 (17)	30 (15.8)	68 (17.6)	37 (18.2)	2 (33.3)	0 (0)	141,294 (16.1)	1,201,627 (17.5)
Norway	48 (6.1)	36 (6.2)	13 (6.8)	23 (5.9)	11 (5.4)	1 (16.7)	0 (0)	70,942 (8.1)	461,963 (6.7)
Spain	21 (2.7)	17 (2.9)	8 (4.2)	9 (2.3)	4 (2.0)	0 (0)	0 (0)	67,031 (7.6)	253,968 (3.7)
Sweden	107 (13.5)	79 (13.7)	27 (14.2)	51 (13.2)	27 (13.3)	1 (16.7)	0 (0)	119,056 (13.6)	1,151,088 (16.8)
United Kingdom	394 (49.9)	276 (47.8)	100 (52.6)	176 (45.5)	114 (56.2)	1 (16.7)	3 (100)	309,518 (35.3)	3,082,675 (44.9)
Mean bone marrow dose (min–max), mGy									
20 (0–286)	20 (0–286)	17 (0–209)	22 (0–286)	19 (0–117)	10 (6–22)	19 (13–25)	15.6 (0–1,684)		

Values are shown as number of participants, PYs and mGy. Histiocytic cell malignancies; Unsp, unspecified; min, minimum; max, maximum. *One case could not be classified as HL or NHL.

risk estimates highest for ABM doses received in the time window '2 to <5 years' and lowest in the time window '≥10 years' before diagnosis. There was, however, no evidence for heterogeneity of risk by time window of exposure, except for myeloid malignancies and AL.

Sensitivity analyses

Lagging doses by 1 year had little effect on the ERR/100 mGy, while a lag of 5 years reduced the risk by slightly less than half for all and for lymphoid malignancies and by two-thirds for myeloid malignancies

Table 2 | RR and 95% CI per cumulative active bone marrow dose category and ERR/100mGy by type of hematological malignancy^a—analyses stratified on sex, birth cohort and country

ABM dose range (mGy)	All hematological malignancies (n=790)			Lymphoid malignancies (n=578)			Myeloid malignancies and AL (n=203)			Leukemia excluding CLL (n=271)		
	#	RR	95% CI	#	RR	95% CI	#	RR	95% CI	#	RR	95% CI
[0,5)	125	1.00		91	1.00		34	1.00		38	1.00	
[5,10)	171	1.10	0.87 1.39	120	1.07	0.81 1.42	47	1.08	0.69 1.71	43	0.79	0.51 1.24
[10,15)	157	1.53	1.20 1.97	123	1.65	1.24 2.20	32	1.16	0.70 1.92	56	1.35	0.87 2.09
[15,25)	165	1.40	1.09 1.80	121	1.41	1.05 1.90	42	1.31	0.80 2.15	66	1.21	0.78 1.89
[25,50)	114	1.87	1.42 2.45	81	1.81	1.32 2.49	32	1.96	1.17 3.29	44	1.61	1.01 2.58
[50+]	58	2.66	1.92 3.70	42	2.64	1.80 3.89	16	2.75	1.47 5.14	24	2.41	1.40 4.17
P for trend		0.02			0.03			0.02			0.02	
	#	ERR/100mGy	95% CI	#	ERR/100mGy	95% CI	#	ERR/100mGy	95% CI	#	ERR/100mGy	95% CI
	790	1.96	1.10 3.12	578	2.01	1.02 3.42	203	2.02	0.47 4.77	271	1.66	0.43 3.74
	#	RR at 100mGy ^b	95% CI	#	RR at 100mGy	95% CI	#	RR at 100mGy	95% CI	#	RR at 100mGy	95% CI
	790	2.96	2.10 4.12	578	3.01	2.02 4.42	203	3.02	1.47 5.77	271	2.66	1.43 4.74

Values are shown in RR, ERR/100mGy and 95% CI. #, number of cases. Statistically significant values are shown in bold. ^aNo analysis of histiocytic and dendritic cell malignancies or of unspecified malignancies were conducted because of the small number of cases (six and three, respectively). ^bNote that the RR at 100 mGy is simply obtained by adding 1 to the ERR/100 mGy.

and AL (Table 5). Using the median of all dose realizations had no major impact on risk estimates. Substantial ERR increases were noted when excluding individuals with the highest cumulative doses (99th, 98th and 95th percentiles). Excluding 5 and 10 years of follow-up after the first CT increased the estimated ERR/100 mGy for all hematological and for lymphoid malignancies but decreased it for myeloid malignancies and AL; the confidence interval of the latter included zero with a 10-year exclusion (two-thirds of cases excluded). Restricting analyses to individuals born after cancer registration was established in their country/region led to a 10–20% reduction in the ERR/100 mGy depending on outcome, while excluding individuals with a CT in a hospital with low CT reporting consistency had little impact on the risk estimates, except for myeloid malignancies and AL (25% decrease). Restricting the follow-up to 2 years after the maximum age at first CT in each country reduced the number of cases from 790 to 491 and duration of follow-up and resulted in lower, but still elevated, risk estimates particularly for lymphoid malignancies. Excluding individuals with no vital status (n = 78,793) slightly reduced risk estimates and increased the width of the CIs, due to the reduction in sample size, for all hematological and lymphoid malignancies, and reduced the myeloid malignancies and AL risk estimates by 31%. Analyses excluding individuals known to have undergone transplantation (United Kingdom only) had little effect on the risk estimate for lymphoid malignancies.

Number of CT examinations

An increasing trend in RRs was observed with increasing number of CT examinations (compared with the reference category: one CT examination) both for all hematological and lymphoid malignancies (Supplementary Table 5). In the continuous analyses, risk increased by 43% per examination for hematological malignancies overall, and by 42% and 48%, respectively, for lymphoid and myeloid malignancies and AL.

Discussion

The EPI-CT study is a large-scale multi-center study designed to directly estimate the risk of hematological malignancies associated with ionizing radiation exposure from CT examinations during childhood and young adulthood, aiming to address criticisms of previous studies related to dosimetry, statistical power and potential biases. The size of the study (nearly one million patients) has considerably increased the statistical power compared with previous national studies. EPI-CT also

evaluated risk using the revised World Health Organization (WHO) classification of hematopoietic and lymphoid tissue malignancies^{27,28}. Our results showed a clear dose–response between cumulative ABM dose and risk of hematological malignancies, both lymphoid and myeloid, with increased risk at doses as low as 10–15 mGy for NHL as a whole and for mature B cell neoplasms.

Associations between risk of hematological malignancies and estimated CT radiation dose to the active bone marrow were robust to the different assumptions tested in the sensitivity analyses. Risk estimates decreased by about half but remained increased for all hematological and lymphoid malignancies when doses were lagged by 5 years. Risk estimates increased, rather than decreased, when individuals with the highest 1%, 2% and 5% cumulative doses were excluded from the analyses, suggesting they were not unduly affected by outliers.

Prior publications on subsets of the EPI-CT cohort reported higher leukemia risk estimates for national studies in the United Kingdom¹² and France¹⁶, but much lower estimates for the Dutch¹⁴ study compared with the all-countries EPI-CT risk estimates. When applying the EPI-CT dose estimates to the original UK cohort (exposed before 2002 and with follow-up to 2008), using the same classification of leukemia as in the original publications, the ERR/100 mGy was similar to published estimates, though the dose distribution differed somewhat (Supplementary Table 6). Thus, the difference between the EPI-CT risk estimates and the original UK estimates appears attributable to the expanded cohort and longer follow-up (Supplementary Table 6). EPI-CT leukemia risk estimates for France, using the updated French cohort and follow-up¹⁷, were imprecise due to small numbers of cases in some categories, but compatible with the published French results, even though the dose distribution differed. Differences between the EPI-CT risk estimates and the Dutch data¹⁴ appear to be mainly related to differences in the dose estimates used, as the results of analyses of the Dutch data using the EPI-CT dosimetry were much closer to those of the full EPI-CT study (Supplementary Table 6). The EPI-CT dosimetry used more sophisticated modeling of doses accounting for historical CT practices and uncertainties due to missing data by country and time period. Final absorbed doses to active bone marrow for each CT examination received were estimated by sex, age group at exposure, body part examined, scanner type and technical scan parameters⁷.

Somewhat surprising was the observation of an increased risk of HL in our analysis, particularly in the light of the absence of an

Table 3 | RR and 95% CI per cumulative active bone marrow dose category and ERR/100mGy by type of malignancy—analyses stratified on sex, birth cohort and country

a. Lymphoid malignancies other than HL													
ABM dose range (mGy)	Lymphoid malignancies ^a												
	NHL ^b												
	All NHL (n=387)			Mature B cell (n=204)			Mature T and NK cell (n=29)			Precursor cell (n=140)			
#	RR	95% CI	#	RR	95% CI	#	RR	95% CI	#	RR	95% CI		
[0,5)	53	1.00		32	1.00		7	1.00		13	1.00		
[5,10)	71	1.12	0.78–1.62	48	1.32	0.84–2.09	6	0.86	0.28–2.67	15	0.71	0.33–1.51	
[10,15)	85	1.89	1.32 – 2.72	42	1.87	1.16 – 3.04	3	0.67	0.17–2.74	35	1.71	0.88–3.35	
[15,25)	87	1.57	1.08 – 2.28	36	1.44	0.86–2.41	7	1.58	0.50–4.95	40	1.31	0.66–2.60	
[25,50)	60	2.08	1.40 – 3.10	29	2.18	1.28 – 3.71	4	1.66	0.45–6.06	25	1.63	0.79–3.36	
[50+]	31	3.00	1.87 – 4.81	17	3.63	1.95 – 6.76	2	2.45	0.47–12.71	12	2.10	0.91–4.85	
P for trend		0.038			0.011			0.046			0.133		
	#	ERR/100mGy	95% CI	#	ERR/100mGy	95% CI	#	ERR/100mGy	95% CI	#	ERR/100mGy	95% CI	
		387	2.51	1.14–4.73	204	3.15	1.17 – 6.88	29	2.85	−0.20–20.23	140	1.26	−0.05–4.34
b. HL and myeloid malignancies													
ABM dose range (mGy)	Lymphoid malignancies ^a						Myeloid malignancies ^c						
	HL (n=190)			AML and related precursor neoplasms+ ALMP/ALAL (n=80)			MPN+MDS+MDS/MPN (n=115)						
	#	RR	95% CI	#	RR	95% CI	#	RR	95% CI				
[0,5)	38	1.00		13	1		20	1					
[5,10)	49	1.01	0.65–1.55	15	0.81	0.38–1.74	31	1.28	0.72–2.29				
[10,15)	38	1.32	0.82–2.10	15	1.18	0.54–2.60	17	1.19	0.61–2.33				
[15,25)	34	1.21	0.74–1.98	15	1.01	0.45–2.27	24	1.46	0.76–2.79				
[25,50)	20	1.36	0.77–2.38	15	2.01	0.90–4.47	15	1.75	0.86–3.58				
[50+]	11	2.15	1.08 – 4.30	7	2.61	0.99–6.90	8	2.61	1.10 – 6.20				
P for trend		0.004			0.04				0.01				
	#	ERR/100mGy	95% CI	#	ERR/100mGy	95% CI	#	ERR/100mGy	95% CI				
		190	1.24	0.08–3.28	80	2.39	0.11 – 8.17	115	1.51	−0.15	5.06		

Values are shown in RR, ERR/100mGy and 95% CI. Statistically significant values are shown in bold. ^aOne case of lymphoid malignancy could not be classified as HL or NHL—ICD-O-3, 1st revision code 9820. ^bFourteen NHL cases could not be classified on the basis of cell type—ICD-O-3, 1st revision code 9590. ^cEight cases could not be classified by subgroup: four cases of AL NOS, one acute biphenotypic leukemia and three cases of myeloid leukemia NOS—ICD-O-3, 1st revision codes 9801, 9805 and 9860, respectively. #, number of hematological malignancy cases.

association in the original UK cohort²⁹ and the inconsistent results in older adults in other radiation epidemiology studies³⁰. Applying the EPI-CT doses to the original UK cohort (with follow-up until 2008) resulted in higher RRs for HL in most dose categories compared with the reference category, with little indication of a dose–response relationship, a different dose distribution (with more individuals receiving higher doses) and a higher ERR/100 mGy (1.1 compared with 0.2), with a CI that included zero (Supplementary Table 6). Analysis of the larger UK EPI-CT cohort, with extended follow-up, yielded an increased ERR/100 mGy (1.73, 95% CI 0.09 to 5.46) suggesting that differences between EPI-CT and published results are mainly attributable to differences in the dosimetry and enlarged cohort size with longer follow-up in the EPI-CT study. While the HL results of the categorical analyses of the full EPI-CT cohort using the UK dose categorization do not show a monotonic trend with dose, analyses using a priori EPI-CT cut points, spanning a wider range of doses, showed evidence of a dose–response (Table 3), with an increased RR in the ≥50 mGy dose category (2.15, 95% CI 1.08 to 4.30). Given the relative rarity of HL compared with NHL, with relatively small numbers of cases in most studies, and in light of the increasing HL incidence in young people, our findings based on 190 cases merit further study.

Within EPI-CT, the UK cohort had a strong influence on risk estimates, contributing about 50% of all hematological malignancies and 45% of the PYs of follow-up. Differences in risk estimates between the United Kingdom and the rest of the countries in the study (also seen for brain tumors)³¹ are unexpected in a multinational collaborative study using a common protocol and dose reconstruction approach. One factor that may partly explain this difference may be the adequacy of the assumptions concerning the technical parameters used during pediatric CT examinations in the United Kingdom, particularly in early years, possibly resulting in a systematic underestimation of doses. Hospital-specific protocols were not available for the United Kingdom⁷, and information from Picture Archiving and Communication System (PACS) data was limited and available only for more recent years. Imaging protocols obtained from pre-existing national surveys in Norway and the United Kingdom had to be used to generate probability density functions (PDFs) of machine settings. These may not adequately reflect the local choices regarding technical parameters made in specific hospitals, particularly in earlier years, which could lead to doses substantially higher than anticipated³². Another possible explanation may be related to missing examinations, as the period during which the UK hospitals contributed CT data varied widely between hospitals,

Table 4 | Effects of potential confounders and potential modifiers of the risk estimates

#	All hematological malignancies			#	Lymphoid malignancies			#	Myeloid malignancies			
	ERR/100mGy	95% CI	ERR/100mGy		95% CI	ERR/100mGy	95% CI		ERR/100mGy	95% CI	ERR/100mGy	
Main results ¹	790	1.96	1.10	3.12	578	2.01	1.02	3.42	203	2.02	0.47	4.77
Potential confounders analysis:												
No adjustment for birth cohort												
	790	1.92	1.08	3.05	578	1.99	1.02	3.37	203	1.92	0.43	4.53
a) SES ²												
Unadjusted	210	1.40	0.08	3.83	161	0.99	-0.16	3.37	47	4.22	-0.17	30.6
Adjusted	210	1.44	0.10	3.90	161	1.03	-0.15	3.45	47	4.16	-0.17	29.6
Effect modification analysis:												
a) Sex												
Males	466	1.45	0.55	2.80	343	1.91	0.71	3.85	118	0.65	-0.42	2.89
Females	324	2.82	1.27	5.32	235	2.14	0.71	4.64	85	6.09	1.62	19.1
Het. P value		0.20					0.85				0.03	
b) Age at exposure category (note: one individual can enter in more than one category if they had several CTs)												
<5	219	0.78	0.06	1.78	165	0.74	-0.05	1.93	52	1.12	-0.29	3.60
5 to <10	156	1.81	0.57	3.39	115	1.87	0.48	3.74	40	1.72	-0.71	5.41
10+	466	4.02	2.48	5.99	336	4.25	2.41	6.71	124	3.48	1.05	7.35
Het. P value		0.001					0.002				0.32	
c) Time since exposure (years) (note: one individual can enter in more than one category if they had several CTs)												
2 to <5	303	3.56	1.96	5.57	222	3.09	1.37	5.37	76	4.88	1.66	9.87
5 to <10	291	2.82	1.58	4.33	216	2.90	1.46	4.70	74	2.98	0.66	6.40
10+	260	1.24	0.42	2.29	184	1.46	0.49	2.75	72	0.45	-0.80	2.56
Het. P value		0.07					0.21				0.04	

¹Stratified on sex, birth cohort and country—attained age is used as the underlying time variable. ²Analysis restricted to countries where SES data were available: Belgium, France, the Netherlands and Spain. #, number of cases; Het., heterogeneity.

contrary to the other countries in the study, and a large proportion of cases were diagnosed in adulthood while only CT examinations up to the age of 22 were included in the study.

EPI-CT was designed to address previous methodological criticisms and limitations of similar studies^{12–16}. Reverse causation appears unlikely as risk estimates varied but remained elevated for the major malignancies groupings when greater lags and extended exclusion periods were applied. Neither birth cohort nor SES appeared to confound the associations in the countries where data were available, nor was SES associated with dose in the original UK cohort (A.B.d.G. personal communication).

Despite all efforts, the study presents some limitations. Confounding by indication could not be addressed directly in the full European cohort beyond excluding specific malignancies coded using the International Classification of Diseases for Oncology, 3rd Edition (ICD-O-3) revision 1 as associated with Down syndrome, therapy or organ transplantation, and conducting a sensitivity analysis excluding individuals from the United Kingdom who had undergone organ transplantation. Confounding by indication was, however, evaluated either directly—from a review of medical records—or indirectly—through modeling—in several national EPI-CT cohorts^{15–17,33,34}. These analyses support a low likelihood for confounding by indication for leukemia, though it may be more important for lymphoma, as patients are more likely to have immune deficiencies and may be at higher risk of infectious diseases³⁵. While appropriate adjustment did not modify radiation-related risk in a recent lymphoma case–control study³⁵, the statistical power was low, and the possibility of residual confounding cannot be ruled out.

Information on SES was available in only four of the nine countries (32.3% of the cohort), and the available information covered different SES dimensions, from material deprivation (household income and house value) in the Netherlands to other social determinants of urban vulnerability (including unemployment, unskilled employment and lack of education) in Spain. Adjustment for SES in each country did not materially affect the risk estimates. Residual confounding of the relation between CT radiation dose and risk of hematological malignancies is therefore unlikely to be substantial, particularly since the evidence for an association between different determinants of SES and risk of leukemia (and more generally hematological malignancies) in young people is inconsistent³⁶.

While the EPI-CT dose reconstruction is based on sophisticated modeling of doses and associated shared and unshared uncertainties, uncertainties in individual doses are not negligible (geometric s.d. of the order of 2 on average⁷), particularly in early years, and could not be fully integrated in the risk analyses. These uncertainties are unlikely to be differential between cases and non cases. While the shared uncertainties are expected to have little impact on the continuous linear risk estimates, the unshared uncertainties could lead to underestimation of the risk but would not create a spurious association. Further work is needed to validate retrospective dose estimates and to ensure the systematic prospective collection of appropriate dose quantities and technical parameters in the clinic in real time to improve risk estimates in the future.

Unlike in the atomic bomb survivor study²⁴, the ERR/100 mGy increased with age at exposure and was highest for exposures within 10 years of diagnosis. These findings, also noted for brain cancers³¹,

Table 5 | Results of sensitivity analyses

	All hematological malignancies				Lymphoid malignancies				Myeloid malignancies and AL			
	#	ERR/100mGy	95% CI		#	ERR/100mGy	95% CI		#	ERR/100mGy	95% CI	
Main results	790	1.96	1.10	3.12	578	2.01	1.02	3.42	203	2.02	0.47	4.77
ABM doses												
Doses lagged by 1 year	790	1.99	1.13	3.16	578	2.04	1.05	3.47	203	2.07	0.54	4.80
Doses lagged by 5 years	790	1.06	0.45	1.82	578	1.25	0.53	2.20	203	0.60	-0.43	2.20
Use of median of dose realizations instead of mean	790	2.08	1.09	3.42	578	2.13	0.99	3.75	203	2.17	0.36	5.35
Analyses restricted to individuals with cumulative doses up to:												
99th percentile	777	3.17	1.90	4.91	567	3.02	1.59	5.09	201	3.81	1.43	8.20
98th percentile	761	3.54	2.08	5.55	554	3.27	1.66	5.63	198	4.43	1.67	9.62
95th percentile	710	2.80	1.27	4.94	520	2.79	1.02	5.41	181	2.88	0.34	7.75
Exclusions												
5 years from first CT	524	2.36	1.21	4.05	381	2.54	1.19	4.66	139	1.74	0.05	5.25
10 years from first CT	261	2.67	1.02	5.69	185	3.12	1.13	7.08	72	1.28	-0.62	8.11
Individuals born before the start of cancer registration	490	1.54	0.68	2.80	365	1.59	0.60	3.14	119	1.84	0.20	5.29
Hospitals with low reporting consistency	603	1.93	0.97	3.29	440	2.17	1.00	3.93	158	1.49	0.00	4.31
Follow-up 2 years after country-specific maximum age at exposure	491	1.27	0.49	2.39	356	1.26	0.39	2.63	128	1.61	0.11	6.61
Individuals with no vital status	739	1.73	0.89	2.97	541	1.91	0.91	3.35	189	1.39	0.03	3.85
United Kingdom—main results	NA				276	2.77	1.12	5.55	NA			
Excluding individuals who underwent transplant	NA				256	2.57	0.97	5.31	NA			

#, number of hematological malignancies cases; NA, not applicable.

within EPI-CT, may be an artifact of the generally short follow-up of this cohort (7.8 years on average) and of the heterogeneous distribution of age at exposure, attained age and time since exposure across countries. Further follow-up of this important cohort is needed to increase the statistical power to explore these effects comprehensively.

EPI-CT was conducted to directly estimate risk from CT radiation doses received in childhood, adolescence and young adulthood, avoiding the need for uncertain extrapolations from the atomic bomb survivors and other studies involving higher radiation doses^{10,25}. For comparison, our estimates of the ERR/100 mGy in atomic bomb survivors younger than 20 years at exposure were 0.77 (95% CI 0.31 to 1.2) for leukemia excluding CLL, based on 40 cases, and -0.02 (95% CI -99 to 99) for HL, based on 2 cases. Using the revised WHO classification of lymphoid malignancies³⁷, the ERR/100 mGy for NHL among those with attained age below 35 was 0.88 (95% CI 0.36 to 3.6), based on small numbers of cases (Ritsu Sakata, personal communication). Numbers were too small to derive risk estimates restricted to survivors exposed below age 20 and for NHL subgroups. The risk estimates for atomic bomb survivors were lower than those in our study. Thus, despite the unavoidable differences in dosimetry systems between the two studies, our results suggest that the linear no threshold model does not overestimate risk from pediatric CT radiation. Indeed, our leukemia risk estimate is compatible with those derived in a recent combined analysis of data on individuals exposed before the age of 21 years and ABM dose <100 mGy (ERR/100 mGy 0.84–4.66, depending on leukemia subtype)³⁸.

EPI-CT used the revised WHO classification of lymphoid and myeloid malignancies, which considers cell lineage and different phases of cell differentiation as well as more classical features^{27,28}. To our knowledge, the revised classification has only been used in a re-analysis of

lymphoma incidence in the atomic bomb survivor cohort³⁷. While this classification makes comparisons with previous publications more difficult (we show results for leukemia excluding CLL classification for this purpose), differences in the incidence of different subtypes across populations suggest possible etiological variation, hence possible differences in radiation effects. Indeed, analysis of atomic bomb survivors' data showed a higher radiation risk of precursor cell NHLs than of mature B or T and NK cell NHLs, contrary to our findings. Differences in length of follow-up and attained age between the atomic bomb survivors' and the EPI-CT cohorts make any conclusion difficult but emphasize the need for future radiation epidemiological studies to adopt this revised classification.

The analyses presented here showed consistent associations between CT radiation dose and risk of hematological malignancies as a whole, and of lymphoid and myeloid malignancies and AL, with an ERR/100 mGy around 2. With an average ABM dose of 8 mGy for a typical examination today (the average dose in the cohort in 2012–2014), this translates to about a 16% increased risk (95% CI 8% to 24%) of these rare malignancies per examination. In terms of absolute risk, among 10,000 children who receive such an examination today, we expect about 1.4 cases (95% CI 1 to 2) due to CT radiation during the 12 years after the examination.

In conclusion, this large-scale study was designed to directly evaluate cancer risk from pediatric and young adult CT radiation exposure. The results of this study, in which much effort has gone into considering and accounting for possible biases that could affect the risk estimates, strengthen the findings from previous low-dose studies of a consistent and robust dose-related increased risk of radiation-induced hematological malignancies. The findings highlight the need for raising awareness in the medical community and continued strict application

of radiological protection measures in medical settings through justification and optimization of radiological procedures, particularly in pediatric populations. This includes ensuring doses are kept as low as reasonably achievable (the ALARA principle), while maintaining appropriate image quality for accurate diagnosis, and monitoring delivered doses; ensuring examinations are justified and unproductive exposure is avoided; and ensuring the benefit-to-risk ratio is maximized for all CT examinations³⁹.

Online content

Any methods, additional references, Nature Portfolio reporting summaries, source data, extended data, supplementary information, acknowledgements, peer review information; details of author contributions and competing interests; and statements of data and code availability are available at <https://doi.org/10.1038/s41591-023-02620-0>.

References

1. World Bank Country and Lending Groups. *The World Bank* (2023) <https://datahelpdesk.worldbank.org/knowledgebase/articles/906519-world-bank-country-and-lending-groups>
2. UNSCEAR 2020/2021 Report Volume I, Scientific Annex A: 'Medical exposure to ionizing radiation'. *United Nations: Scientific Committee on the Effects of Atomic Radiation*. United Nations, New York (2022) www.unscear.org/unscear/en/publications/2020_2021_1.html
3. Brenner, D. J. & Hall, E. J. Computed tomography—an increasing source of radiation exposure. *N. Engl. J. Med.* **357**, 2277–2284 (2007).
4. Bosch de Basea, M. et al. Trends and patterns in the use of computed tomography in children and young adults in Catalonia—results from the EPI-CT study. *Pediatr. Radiol.* **46**, 119–129 (2016).
5. Pearce, M. S. et al. CT scans in young people in Northern England: trends and patterns 1993–2002. *Pediatr. Radiol.* **41**, 832–838 (2011).
6. Meulepas, J. M. et al. Trends and patterns of computed tomography scan use among children in The Netherlands: 1990–2012. *Eur. Radiol.* **27**, 2426–2433 (2017).
7. Thierry-Chef, I. et al. Dose estimation for the European Epidemiological Study on Pediatric Computed Tomography (EPI-CT). *Radiat. Res.* **196**, 74–99 (2021).
8. Marin, J. R. et al. Trends in use of advanced imaging in pediatric emergency departments, 2009–2018. *JAMA Pediatr.* **174**, e202209 (2020).
9. UNSCEAR 2006 Report. Annex A. Epidemiological studies of radiation and cancer. *United Nations Scientific Committee on the Effects of Atomic Radiation*. United Nations, New York (2006) [https://www.unscear.org/unscear/uploads/documents/publications/UNSCEAR_2006_Annex-A-CORR.pdf](http://www.unscear.org/unscear/uploads/documents/publications/UNSCEAR_2006_Annex-A-CORR.pdf)
10. Health risks from exposure to low levels of ionizing radiation. BEIR, VII Report, phase II. *Committee on the Biological Effects of Ionizing Radiation*, National Academy Press, Washington DC (2006). <http://www.nap.edu/openbook.php?isbn=030909156X>
11. Medical radiation exposure of patients in the United States: recommendations of the National Council on Radiation Protection and Measurements. NCRP <https://ncrponline.org/shop/reports/report-no-184-medical-radiation-exposure-of-patients-in-the-united-states-2019/> (2019).
12. Pearce, M. S. et al. Radiation exposure from CT scans in childhood and subsequent risk of leukaemia and brain tumours: a retrospective cohort study. *Lancet* **380**, 499–505 (2012).
13. Mathews, J. D. et al. Cancer risk in 680,000 people exposed to computed tomography scans in childhood or adolescence: data linkage study of 11 million Australians. *Br. Med. J.* **346**, f2360 (2013).
14. Meulepas, J. M. et al. Radiation exposure from pediatric CT scans and subsequent cancer risk in the Netherlands. *J. Natl Cancer Inst.* **111**, 256–263 (2019).
15. Krille, L. et al. Risk of cancer incidence before the age of 15 years after exposure to ionising radiation from computed tomography: results from a German cohort study. *Radiat. Environ. Biophys.* **54**, 1–12 (2015).
16. Journy, N. et al. Are the studies on cancer risk from CT scans biased by indication? Elements of answer from a large-scale cohort study in France. *Br. J. Cancer* **112**, 185–193 (2015).
17. Foucault, A. et al. Childhood cancer risks estimates following CT scans: an update of the French CT cohort study. *Eur. Radiol.* **32**, 5491–5498 (2022).
18. Krille, L. et al. Erratum to: Risk of cancer incidence before the age of 15 years after exposure to ionising radiation from computed tomography: results from a German cohort study. *Radiat. Environ. Biophys.* **56**, 293–297 (2017).
19. Li, I.-G., Yang, Y.-H., Li, Y.-T. & Tsai, Y.-H. Paediatric computed tomography and subsequent risk of leukaemia, intracranial malignancy and lymphoma: a nationwide population-based cohort study. *Sci. Rep.* **10**, 7759 (2020).
20. Nikkilä, A., Raitanen, J., Lohi, O. & Auvinen, A. Radiation exposure from computerized tomography and risk of childhood leukemia: Finnish register-based case-control study of childhood leukemia (FRECCLE). *Haematologica* **103**, 1873–1880 (2018).
21. Berrington de Gonzalez, A., Pasqual, E. & Veiga, L. Epidemiological studies of CT scans and cancer risk: the state of the science. *Br. J. Radiol.* **94**, 20210471 (2021).
22. Walsh, L., Shore, R., Auvinen, A., Jung, T. & Wakeford, R. Risks from CT scans—what do recent studies tell us? *J. Radiol. Prot.* **34**, E1–E5 (2014).
23. The 2007 Recommendations of the International Commission on Radiological Protection. ICRP Publication 103. *Ann. ICRP* **37** (2–4) <http://www.icrp.org/publication.asp?id=ICRP%20Publication%20103> (2007).
24. Hsu, W.-L. et al. The incidence of leukemia, lymphoma and multiple myeloma among atomic bomb survivors: 1950–2001. *Radiat. Res.* **179**, 361–382 (2013).
25. Tubiana, M. Dose–effect relationship and estimation of the carcinogenic effects of low doses of ionizing radiation: the joint report of the Académie des Sciences (Paris) and of the Académie Nationale de Médecine. *Int. J. Radiat. Oncol. Biol. Phys.* **63**, 317–319 (2005).
26. Bernier, M.-O. et al. Cohort profile: the EPI-CT study: a European pooled epidemiological study to quantify the risk of radiation-induced cancer from paediatric CT. *Int. J. Epidemiol.* **48**, 379–381g (2019).
27. Arber, D. A. et al. The 2016 revision to the World Health Organization classification of myeloid neoplasms and acute leukemia. *Blood* **127**, 2391–2405 (2016).
28. Swerdlow, S. H. et al. The 2016 revision of the World Health Organization classification of lymphoid neoplasms. *Blood* **127**, 2375–2390 (2016).
29. Berrington de Gonzalez, A. et al. No association between radiation dose from pediatric CT scans and risk of subsequent Hodgkin lymphoma. *Cancer Epidemiol. Biomark. Prev.* **26**, 804–806 (2017).
30. Harbron, R. W. & Pasqual, E. Ionising radiation as a risk factor for lymphoma: a review. *J. Radiol. Prot.* **40**, R151–R185 (2020).
31. Hauptmann, M. et al. Brain cancer after radiation exposure from CT examinations of children and young adults: results from the EPI-CT cohort study. *Lancet Oncol.* **24**, 45–53 (2022).
32. Smith-Bindman, R. et al. International variation in radiation dose for computed tomography examinations: prospective cohort study. *Br. Med. J.* **364**, k4931 (2019).

33. Meulepas, J. M. et al. Confounding of the association between radiation exposure from CT scans and risk of leukemia and brain tumors by cancer susceptibility syndromes. *J. Radiol. Prot.* **36**, 953–974 (2016).
34. Berrington de Gonzalez, A. et al. Relationship between paediatric CT scans and subsequent risk of leukaemia and brain tumours: assessment of the impact of underlying conditions. *Br. J. Cancer* **114**, 388–394 (2016).
35. Pasqual, E. et al. Association of ionizing radiation dose from common medical diagnostic procedures and lymphoma risk in the Epilymph case-control study. *PLoS ONE* **15**, e0235658 (2020).
36. Bosch de Basea, M. et al. CT scan exposure in Spanish children and young adults by socioeconomic status: cross-sectional analysis of cohort data. *PLoS ONE* **13**, e0196449 (2018).
37. Fujihara, M. et al. Incidence of lymphoid neoplasms among atomic bomb survivors by histological subtype, 1950 to 1994. *Blood* **139**, 217–227 (2022).
38. Little, M. P. et al. Leukaemia and myeloid malignancy among people exposed to low doses (<100 mSv) of ionising radiation during childhood: a pooled analysis of nine historical cohort studies. *Lancet Haematol.* **5**, e346–e358 (2018).
39. Optimisation of ionising radiation-based medical protocols for diagnostics or therapy. *MEDIRAD* (2022) <http://www.medirad-project.eu/recommendations>

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Methods

Study population

The EPI-CT project set up new cohorts in Belgium, Denmark, the Netherlands, Norway, Spain and Sweden, and included and enlarged existing cohorts in France, Germany and the United Kingdom^{12,15,16}. Detailed methods have been published^{7,26,40}.

The international EPI-CT cohort includes 948,174 individuals who: (1) underwent at least one CT examination in a participating hospital between 1977 and 2014 before the age of 22 years (exact age limit ranging between 10 and 22 years, depending on country²⁶); (2) were residents of geographic areas covered by cancer registries; (3) had no previous history of cancer; and (4) had no cancer diagnosis in the first year following the first CT⁴⁰. In the present analysis we excluded 77,369 individuals with follow-up shorter than 2 years, including 142 individuals with a cancer diagnosis during that period.

The study population was identified through radiology department records of 276 pediatric and general (serving large pediatric patient populations) hospitals. Basic demographic data (including sex, as reported on the clinical history of the patient) and information on each examination was collected for each individual.

Ethics approvals

The study was approved by the ethics committee at IARC (coordinating center) (IARC IEC 12–35), and the appropriate national, regional and hospital ethics committees in participating countries before starting the epidemiological study. This was a record linkage study with no contact with individual patients (and hence no informed consent).

Follow-up

Cohort members were followed up through national and/or regional cancer and mortality registries. Germany and part of France lacked information on mortality. Information on migration status was collected where available: in Denmark, Norway and Sweden. In these countries, only 2.05% of cohort members were known to have emigrated during the study follow-up period.

Outcome definitions

Diagnoses were coded using the International Classification of Diseases for Oncology, 3rd Edition (ICD-O-3) 1st revision (2013). Only cases with behavior code 3 (malignant) were included⁴¹. Given changes in classification of hematological malignancies according to cell lineage and maturation^{27,28}, the analyses were conducted using the revised WHO classification of lymphoid and myeloid malignancies^{27,28}, focusing on the following groupings, types and subtypes (morphology codes in Supplementary Table 1):

- All hematological malignancies, excluding those coded as related to therapy or predisposing syndromes as they are unlikely to be related to CT exposure^{16,33};
- All lymphoid malignancies and subgroups of HL, NHL and lymphoid malignancy subtypes (mature B cell, mature T and NK cell, and precursor cell);
- All myeloid malignancies and AL and subgroups of:
 - Acute myeloid leukemia (AML) and related malignancies together with AL of mixed phenotype and ambiguous lineage (ALMP/ALAL);
 - MPN, MDS, together with MDS/MPN–MPN + MDS + MDS/MPN.

For comparison with previous studies, analysis of leukemia, excluding CLL, was also conducted.

Confounding factors

Information on socioeconomic status (SES) was collected, based on nationally available data sources, in the following countries using the information, for individuals from the following countries, representing 32.3% of the EPI-CT cohort:

- Belgium: SES derived from the healthcare reimbursement classification based on the annual income of the household (two categories: lower or normal);
- France: SES based on Townsend deprivation scores, obtained from linkage of residential postal code (five quintiles) with census data;
- the Netherlands: SES derived from average household income and house value for six-digit postal codes (average population, 40 persons) of cohort members' residential addresses from Statistics Netherlands;
- Spain: SES based on the Synthetic index of urban vulnerability generated according to the socio-economic characteristics of the census tract that included the area of residence (five quintiles).

No information was available regarding the indication or reasons of the CT examinations.

Organ dose estimates

The organ dose estimation methodology is described elsewhere⁶. Briefly, it was based on a multi-level approach integrating CT imaging information from hospital questionnaires, national reports, scientific publications, expert opinion together with CT parameter values obtained directly from the PACS from 23% of 276 participating hospitals. Doses were estimated using the National Cancer Institute Dosimetry System for CT⁴². Uncertainty associated with missing parameters, for example, in earlier periods when PACS did not exist, was characterized by a range of possible, realistic values for each missing parameter using the aforementioned sources of information and PDFs defined by age group, sex, body region scanned, machine type representative of technology evolution inferred from questionnaires and time period. For each CT examination, a set of 200 dose realizations was derived where, in each iteration, different values of the parameters were sampled from the PDFs, maintaining proper correlations between parameters.

Our main analyses were based on dose to the active (red) bone marrow (ABM), as commonly used in analyses of hematological malignancies in radiation epidemiology, and the arithmetic mean of all dose realizations for each CT examination. The cumulative dose for each participant was obtained by summing the dose (mean of all realizations) from all examinations the participant received.

Statistical analysis

Descriptive analyses included the distribution of cases and cohort members by sex, country, age-at-exposure, attained age and time since exposure.

Dose–response analyses were conducted for all outcomes listed above by modeling the RR as $1 + \beta Z$, where Z is the cumulative dose and β is the ERR per unit dose. The model was fit with proportional hazards regression using the custom-developed R module rERR: Excess Relative Risk Models R package version 0.1 (ref. 43). Exact age of the individuals was used as the underlying time variable, and all models were stratified by sex, country and birth cohort (1960–1979, 1980–1984, 1985–1989, 1990–1994, 1995–1999, 2000–2004 and 2005–2012). We also fit an EAR model using the PEANUTS module of the EPICURE software (version 2.00.02) to estimate the absolute excess number of hematological malignancies per 10,000 PYs and per dose Z . We used this model to predict the number of cases that would be expected in the European population from CT scanning today as the difference between the total number of cases expected under the fitted model at a typical dose level and the ‘background’ number of cases expected in the absence of radiation exposure.

Analyses used cumulative dose as a continuous variable (in mGy), as well as a categorical variable, with cut points defined on the basis of the cohort dose distribution: 0.0004 to <5, 5 to <10, 10 to <15, 15 to <25, 25 to <50, and 50–1,684 mGy). Due to the skewness of the dose distribution, 95% likelihood-based CIs were used in the continuous analyses. For the categorical analyses, we used Wald-based CI. Trends

in RRs by level of dose were tested by fitting the categorical variables as a continuous ordinal variable.

Follow-up started 2 years after the first CT scan (to minimize reverse causation potential) or when complete cancer registration was available in the country/region, whichever was later. Exit date was defined as the earliest of dates of any cancer diagnosis, death, emigration (where available) or end of follow-up in the country/region.

To account for a minimal latency between radiation exposure and malignancy, doses were lagged by 2 years. As EPI-CT is a record linkage study, no information about confounding factors other than birth cohort, sex and country/region was systematically available. The effect of country was assessed in country-specific analyses, and removing one country at a time, and SES effect in analyses restricted to countries with available SES. Effect modification by age at exposure (<5, 5 to <10, and ≥10 years), time since exposure (2 to <5, 5 to <10, and ≥10), sex and birth cohort was tested by including an interaction term with dose in the linear dose model. The statistical significance of model parameters was tested using the likelihood ratio test.

Supplementary and sensitivity analyses were performed to test the findings' robustness. Regarding doses, analyses included: lagging doses by 1 and 5 years (instead of 2), using the median of all dose realizations instead of the mean, and excluding individuals with the highest cumulative doses (above the 99th, 98th and 95th percentiles of the cumulative dose distribution). Additional analyses were conducted excluding: the first 5 and 10 years of follow-up, individuals born before the start of cancer registration in their respective country/region, and hospitals with low reporting consistency (≥1 consecutive years without reporting CT examinations), as well as excluding individuals from the United Kingdom known to have undergone organ transplantation (transplant data were available only for this country) from the lymphoid malignancies analysis as they are at increased risk of post-transplant lymphoproliferative disorders. We also terminated follow-up 2 years above the age limit for inclusion of scans in each country (as doses received later in life were not collected within the project) and excluded the subcohorts lacking mortality follow-up.

We repeated analyses using the number of CT examinations instead of ABM dose.

To allow comparison of our estimates with those of the atomic bomb survivor study, we conducted analyses of leukemia and HL risk in that study using publicly available grouped incidence data¹⁹, restricted to the population and follow-up most relevant for EPI-CT: less than 20 years old at time of bombing, with attained age less than 35 years. These analyses, adjusted on attained age, sex, birth cohort and city, were conducted using the AMFIT module of EPICURE.

Reporting summary

Further information on research design is available in the Nature Portfolio Reporting Summary linked to this article.

Data availability

The data collected and generated in the study are not freely available because of ethical and data protection constraints. The pseudonymized data analysis file for this manuscript is stored at ISGlobal and cannot be shared. Proposals for possible collaborations in further analyses of these data should be addressed to E.C. (elisabeth.cardis@isglobal.org) and will be reviewed by the EPI-CT steering committee. Scientific collaborations will require a written agreement with all involved parties. Requests are normally processed within 1 month. Agreed analysis will be carried out internally by EPI-CT study members, following the agreed scientific collaboration and under the supervision of the proposing researcher. Note that the Data Transfer Agreements (DTAs) ruling the provision of data for the international EPI-CT analyses are time limited and IARC and ISGlobal will be under obligation to destroy the data from individual cohorts when the DTAs expire. Data from these cohorts will be held only by the original data provider, as long as the national data protection legislation permits.

Code availability

The software used to fit ERR models (the rERR R package) is freely available at <https://rdrr.io/cran/rERR/>. All EAR models were performed using the PEANUTS module of the EPICURE software (version 2.00.02) commercially available at <https://risksciences.com/epicure/>. The EAR code applied is available at https://github.com/Mbb2022-23/EPI_CT_EAR. The ERR code applied is available at https://github.com/radiationISGlobal/EPI_CT_Scripts.

References

40. Bosch de Basea, M. et al. EPI-CT: design, challenges and epidemiological methods of an international study on cancer risk after paediatric and young adult CT. *J. Radiol. Prot.* **35**, 611–628 (2015).
41. International Classification of Diseases for Oncology, 3rd Edition (ICD-O-3). World Health Organisation, Geneva (2013) <http://www.who.int/classifications/icd/adaptations/oncology/en/>
42. Lee, C., Kim, K. P., Bolch, W. E., Moroz, B. E. & Folio, L. NCICT: a computational solution to estimate organ doses for pediatric and adult patients undergoing CT scans. *J. Radiol. Prot.* **35**, 891–909 (2015).
43. Badia Roca, F. B. & Moriña Soler, D. rERR: excess relative risk models (2018). <https://rdrr.io/cran/rERR/>

Acknowledgements

The authors are grateful to the radiologists, clinicians, physicists and administrators of participating hospitals and the national/regional cancer, mortality and cause of death registries that provided exposure and radiological data for the study. The authors gratefully acknowledge scientific and technical assistance provided by: Belgium: J. Geens and H. Bosmans from the participating Belgian hospitals and to the Belgian Cancer Registry for providing cancer data. Denmark: the radiological departments participating in organizing data collection. IARC: C. Chassin for administrative assistance in the overall study coordination. France: the radiologists, physicists and administrators working in the participating hospitals who took so much of their time to provide us with the necessary radiology and clinical data: D. Loisel, B. Ory, D. Weil (CHU Angers), J.-M. Garcier, J. Guerssen, S. Mangin (CHU Clermont-Ferrand), S. Baron, C. Gaborit, D. Sirinelli (CHU Tours), J.-M. Chave, E. Chirpaz, O. Fels (CHU La Réunion), N. Boutry, G. Potier (CHU Lille), D. Defez, Perrot, M. Teisseire (CHU Lyon), B. Bourlière, P. Petit (CHU Marseille), M. Saguintaa (CHU Montpellier), F. Collignon, M.-A. Galloy, E. Pozza, E. Schmitt (CHU Nancy), B. Dupas, T. Lefrançois, M. Salaud (CHU Nantes), J.-F. Chateil, C. Barat, C. Bertini, M. Hajjar (CHU Bordeaux), M.-A. Perrier, H. Daubert, L. Froment (CHU Rouen), S. Dupont, L. Molinier, J. Vial (CHU Toulouse), H. Ducou Le Pointe, A. Bouette, P. Chambert (CHU Armand Trousseau—Paris), N. Boddaert (CHU Necker-Enfants-Malades—Paris), E. Dion (CHU Louis Mourier—Colombes), J. Costa (CHU Robert Debré—Paris), G. Khalifa (CHU Saint-Vincent de Paul—Paris), J. Betout (APHP Siège), D. Musset (CHU Antoine Béclère—Clamart), C. Adamsbaum, S. Franchi, D. Pariente (CHU Bicêtre) and N. Sellier (CHU Jean Verdier—Bondy). We also warmly thank S. Ben Salha, L. Faure and B. Lacour (Registre National des Cancers de l'Enfant) for their valuable help in providing data about cancer diagnoses. Germany: the radiologists, physicists and administrators working in the participating hospitals who took so much of their time to provide us with the necessary radiology and clinical data: T. Albrecht (Vivantes, Klinikum Neukölln), M. Asmussen (Städtisches Klinikum Karlsruhe), J. Barkhausen (Universitätsklinikum Schleswig-Holstein), J. D. Berthold (Medizinische Hochschule Hannover), A. Chavan (Klinikum Oldenburg), C. Claussen (Universitätsklinikum Tübingen), M. Forsting (Universitätsklinikum Essen), K. Jablonka (Klinikum Bremen-Mitte), M. Langer (Universitätsklinikum Freiburg), M. Laniado (Universitätsklinikum Carl Gustav Carus Dresden), J. Lotz

(Universitätsmedizin Göttingen), H. J. Mentzel (Universitätsklinikum Jena), P. Mildenberger, A. Queißer-Wahrendorf and G. Staatz (University Medical Center Mainz), O. Rompel (Universitätsklinikum Erlangen), J. Schlick (Klinikum Nürnberg Süd), K. Schneider and M. Seidenbusch (Klinikum der Universität München, Dr. von Haunersches Kinderspital), M. Schumacher (Universitätsklinik Freiburg), B. Spors (Charité-Universitätsmedizin Berlin), T. Vogl (Klinikum der Johann Wolfgang Goethe-Universität Frankfurt/Main), J. Wagner (Vivantes, Klinikum im Friedrichshain) and G. Weisser (Universitätsklinikum Mannheim). We also warmly thank H. Zeeb, S. Dreger (BIPS) and C. Spix (Deutsches Kinderkrebsregister) for their valuable help. The Netherlands: the staff members, clinicians and boards of directors of all hospitals who provided data for our study (AMC Amsterdam, Amphia Ziekenhuis, Albert Schweitzer Ziekenhuis, Bethesda Ziekenhuis, Canisius-Wilhelmina Ziekenhuis, Diaconessenhuis Leiden, Isala Diaconessenhuis Meppel, Diaconessenhuis Utrecht, Elkerliek Ziekenhuis, Erasmus MC, Flevoziekenhuis, Groene Hart Ziekenhuis, Medisch Centrum Haaglanden, HagaZiekenhuis, Ikazia Ziekenhuis, Isala Ziekenhuis, Jeroen Bosch Ziekenhuis, Kennemer Gasthuis, Streekziekenhuis Koningin Beatrix, Leids Universitair Medisch Centrum, Medisch Centrum Alkmaar, Medisch Centrum Leeuwaarden, Meander Medisch Centrum, Medisch Spectrum Twente, Onze Lieve Vrouwe Gasthuis, Radboudumc, Rijnland Ziekenhuis, Rijnstate, Rivas Beatrix Ziekenhuis, Saxenburgh groep Ro“pcke-Zweers Ziekenhuis, Sint Franciscus Gasthuis, Slotervaartziekenhuis, St. Antonius Ziekenhuis, St. Elisabeth Ziekenhuis, Sint Maartenskliniek, Maastricht UMC, UMC Groningen, UMC Utrecht, Het Van Weel-Bethesda Ziekenhuis, VUMC Amsterdam, Ziekenhuisgroep Twente and ZorgSaam Zeeuws-Vlaanderen). Norway: staff members, clinicians and boards of directors of all hospitals who provided data for our study. We especially thank M. Gårseth, B. Kothe-Næss (Helse Nord-Trøndelag HF), S. Tveiten (Sørlandet sykehus HF) and E. Meen (Cancer Registry Norway) for their valuable contributions. The radiologists, physicists and administrators working in the participating hospitals who took so much of their time to provide us with the necessary radiology and clinical data: G. Andersen, J. Gunnar Andersen, F.-H. Andersen, A. Aslaksen, T. Bakkelund, G. Brandseth, C. De Lange, A. Erikson, K. Fredriksen, G.-E. Gustavsson, S. Hanssen, L. Heiberg, B. Hjelmstad, K. Holen, B. Erik Johansson, N. Kaldahl Wold, M. Alexander Olsen, T. Rehn Holm-Johnsen, H. Roterud, K. Roth, B. Åse Rue Gotaas, Å. Sætevik, E. Marie Sager, Y. Skar, H. Jørgen Smith, E. Søvik and L. Thomassen. UNN Tromsø, UNN Harstad, UNN Narvik, Helgelandssykehuset Mo i Rana, Helgelandssykehuset Mosjøen, Helgelandssykehuset Sandnessjøen, Kristiansund sjukehus, Sykehuset Levanger, Sykehuset Namsos, St. Olavs Hospital Trondheim, Orkdal sjukehus, Molde sjukehus, Volda sjukehus, Ålesund sjukehus, Sykehuset Innlandet, OUS Radiumhospitalet, OUS Rikshospitalet, OUS Ullevål sykehus, Haukeland universitetssjukehus, Voss sjukehus, Kysthospitalet i Hagevik, Sykehuset Telemark, SSHF Kristiansand, SSHF Arendal, SSHF Flekkefjord and Stavanger universitetssjukehus. Spain: F. Badia at ISGlobal; L. Donoso, T. Fonoll, S. Pedraza, L. Riera, A. Capdevila, I. Barber Martínez de la Torre, M. Pardina, J. Muchart, J. Palmero, J. Vilar, L. Martí-Bonmatí, R. García, P. Tallón, Berna, Á. Chans, M. Castañeda, J. Esparza, M. de Blas, A. Gamarra, D. Grande Icaran, J. Lafuente, C. Serrano, G. Gómez Mardones, M. López Nieto, M. Parrón Pajares and M. Luisa Llorente at the radiology departments of the participating Spanish hospitals; additional doctors and IT specialists who contributed to the study and the members of the national and autonomic cancer, mortality, hospital discharge (Conjunto Mínimo Básico de Datos - CMBD) and hospital registries of Spain. Thanks for all the efforts and contribution to bring forward this study. Sweden: the staff members, clinicians and boards of directors of all hospitals in the regions who provided data for our study (Region Skåne, Västra Götaland, Östergötland and Stockholm-Gotland). United Kingdom:

the radiologists, radiographers and medical physicists at participating hospitals for providing data. We also thank M. Pearce, R. Hardy, K. Kirton, J. Salotti, C.-L. Chapple and E. Slack (Newcastle University) and K. McHugh (Great Ormond Street Hospital). Authors who are identified as personnel of the IARC, WHO, are alone responsible for the views expressed in this Article and do not necessarily represent the decisions, policy, or views of the IARC, WHO. M.S.P. and R.H. are affiliated with the National Institute for Health Research Health Protection Research Unit (NIHR HPRU) in Chemical and Radiation Threats and Hazards at Newcastle University in partnership with Public Health England (PHE). The views expressed are those of the authors and not necessarily those of the NHS, the NIHR, the Department of Health or Public Health England. This report makes use of data obtained from the Radiation Effects Research Foundation (RERF), Hiroshima and Nagasaki, Japan and cites risk estimates calculated by RERF. RERF is a private, nonprofit foundation funded by the Japanese Ministry of Health, Labour and Welfare and the US Department of Energy, the latter through the National Academy of Sciences. The conclusions in this report are those of the authors and do not necessarily reflect the scientific judgment of RERF or its funding agencies. This work was partly supported by the European Community's Seventh Framework Programme (FP7/20011-2017) (grant number 269912 - EPI-CT: Epidemiological study to quantify risks for paediatric computerised tomography and to optimise doses) (A.K., E.C., M.H., M.-O.B., A.J., H.O., H.E., C.J., M.B., M.K. and K.K.). In Spain, this study was partially supported by grants (E.C. and M.B.B.) from the Instituto de Salud Carlos III-ISCIII from the Spanish Government (reference: PI16/00120) cofunded by FEDER funds/European Regional Development Fund (ERDF)—a way to build Europe. Additionally, complementary Spanish funding was received from the Consejo de Seguridad Nuclear (E.C.) and M.B.B. was the recipient of a fellowship of the Centro de Investigación Biomédica en Red de Epidemiología y Salud Pública (CIBERESP) for a short stay abroad at Newcastle University. ISGlobal also acknowledges support from the grant CEX2018-000806-S funded by MCIN/AEI/10.13039/501100011033, from the Generalitat de Catalunya through the CERCA Program and from the Secretariat of Universities and Research of the Department of Business and Knowledge of the Generalitat of Catalonia through AGAUR (the Catalan Agency for Management of University and Research Grants) (Project 2017 SGR 1487) to EC. The International Agency for Research in Cancer (IARC) received complementary funding from The Ministry of Health, Labour and Welfare of Japan (grant agreement number 2012-02-21-01) (A.K. and J.S.). In France, complementary funding was obtained from the association 'La Ligue contre le Cancer' (grant number PRE09/MOB) and from the French Institute of cancer (INCa, grant number 2011-1-PL-SHS-01-IRSN-1, grant number SHS-ESP-2019-025) (M.-O.B.). In Germany, complementary funding was provided by the German Federal Ministry of Education and Research (grant numbers 02NUK016A, 02NUK016B and 02NUK016CX) (M.B.) for the German KICT study. In the Netherlands, Worldwide Cancer Research, formerly known as the Association for International Cancer Research, provided partial funding (Grant 12-1155) (M.H.). C.R. was supported by a personal grant for Junior group leaders from the Dutch Cancer Society (Grant UVA2021-5517). The original UK cohort study was funded by the UK Department of Health and the US National Cancer Institute, and further funding for the study has been obtained from Cancer Research UK. In Norway, it was funded by the Norwegian Research Council through the EURATOM program, project no. 209096/E40 (K.K.). Denmark received complementary funding from the Danish Cancer Society (C.J.).

Author contributions

E.C., A.K., M.H. and I.T.-C. were responsible for the study design. A.K. was responsible for the overall coordination of the study, E.C. for the

epidemiological methodology and M.P. for the coordination of the field work. I.T.-C. and S.L.S. developed the exposure reconstruction strategy. I.T.-C., G.F., S.L.S., J.D., T.S.I., L.L.C., H.O., A.J., J.F., C.M., F.M., R.H. and C.L. participated in development and validation of the exposure reconstruction approach. M.B.d.B., M.H., C.R., E.C., M.-O.B., M.B., J.D., H.E., C.J., M.K., K.K., N.J., J.F., J.M.M., L.L.C., R.P., T.S.I., A.N., A.J., A.B.d.G. and R.W.H. were responsible for patient accrual and obtaining and processing the data, including data on exposure. E.C., M.H. and G.B. wrote the statistical analysis plan. M.B.d.B.G., E.C., J.F., A.P., M.M. and A.K. had full access to, and verified, the data. E.C., A.P., J.F., M.B.d.B.G. and M.H. analysed the data and produced the results and figures. M.B.d.B.G. and E.C. wrote and edited the paper. All authors had access to and interpreted the data, edited the manuscript draft and approved the final version of the manuscript. All authors had final responsibility for the decision to submit for publication.

Competing interests

The authors declare no competing interests.

Additional information

Extended data is available for this paper at
<https://doi.org/10.1038/s41591-023-02620-0>.

Supplementary information The online version contains supplementary material available at
<https://doi.org/10.1038/s41591-023-02620-0>.

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Peer review information *Nature Medicine* thanks Jacqueline Vo, Oleg Belyakov and the other, anonymous, reviewer(s) for their contribution to the peer review of this work. Primary Handling Editor: Ming Yang, in collaboration with the *Nature Medicine* team.

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Extended Data Table 1 | Number of CT examinations and cumulative dose to the active bone marrow per patient, for all individuals and by country

Country	Number of individuals	Total number of CT exams	Mean number of CT exams per individual	Estimated cumulative ABM dose per patient (mGy) at the end of follow-up			
				Mean	Median	Min	Max
Overall	876 771	1 331 896	1.52	15.5	10.7	0.0	1684
Belgium	9 052	12 895	1.42	14.9	11.2	0.0	431
Denmark	15 835	29 837	1.88	17.2	10.7	0.0	630
France	104 542	153 258	1.47	13.7	11.2	0.0	670
Germany	39 501	61 280	1.55	23.1	16.5	0.0	708
Netherlands	141 294	209 235	1.48	16.1	10.8	0.0	728
Norway	70 942	129 463	1.82	14.6	9.4	0.0	731
Spain	67 031	96 970	1.45	13.1	9.3	0.0	1088
Sweden	119 056	181 405	1.52	16.3	11.9	0.0	593
UK	309 518	457 553	1.48	15.3	9.7	0.0	1684

CT: computed tomography; ABM: active bone marrow; Min: minimum; Max: maximum; UK: The United Kingdom

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Codes developed are available in GitHub https://github.com/Mbb2022-23/EPI_CT_EAR, https://github.com/radiationISGlobal/EPI_CT_Scripts

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Elisabeth Cardis (elisabeth.cardis@isglobal.org) and will be reviewed by the EPI-CT steering committee. Scientific collaborations will require a written agreement with all involved parties. Requests are normally processed within 1 month. Agreed analysis will be carried out internally by EPI-CT study members, following the agreed scientific collaboration and under the supervision of the proposing researcher. Note that the Data Transfer Agreements (DTA) ruling the provision of data for the international EPI-CT analyses are time limited and IARC and ISGlobal will be under obligation to destroy the data from individual cohorts when the DTAs expire. Data from these cohorts will be held only by the original data provider, as long as the national data protection legislation permits.

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Reporting on sex and gender

All analyses include both females and males and are adjusted for sex through stratification. We also conducted an analysis to evaluate whether the risk might differ between males and females. There was no evidence for a difference of risk between sexes overall though the risk of myeloid malignancies and AL appeared to be higher in women than men.
No data was available on gender in this study

Reporting on race, ethnicity, or other socially relevant groupings

Race and ethnicity information was not available in this study. Information on socio-economic status (SES) was collected, based on nationally available data sources, in the following countries using the information, available for study subjects from the following countries, representing 32.3% of the EPI-CT cohort:

- Belgium: SES derived from the healthcare reimbursement classification based on the annual income of the household (2 categories: lower or normal);
- France: SES based on Townsend deprivation scores, obtained from linkage of residential postal code (5 quintiles) with census data;
- the Netherlands: SES derived from average household income and house value for six-digit postal codes (average population, 40 persons) of cohort members' residential addresses from Statistics Netherlands;
- Spain: SES based on the Synthetic index of urban vulnerability generated according to the socioeconomic characteristics of the census tract that included the area of residence (5 quintiles).

This variable was used to evaluate whether the relation between radiation dose from CT scans and risk of haematological malignancies might be confounded by SES and to adjust for this if this is the case. The reason SES might be a confounder is that on one side, SES and urban vulnerability or deprivation could be related to the likelihood of undergoing a CT scan (because of trauma for example) and, on the other side SES has been suggested to be related to the risk of developing leukaemia though the evidence is not conclusive.

Potential confounding was evaluated by conducting analyses of risk restricted to the four countries where SES information was available including and excluding SES as a covariate in the model and checking whether inclusion of SES modified risk estimates by at least 10%.

Population characteristics

The study population includes 948,174 subjects (males and females aged 0 to over 50 years old) who: 1) underwent at least one CT examination in a participating hospital between 1977 and 2014 before the age of 22 years, 2) had no previous history of cancer, and 3) had no cancer diagnosis in the two years following the 1st CT

Recruitment

The study population was identified through radiology department records of 276 paediatric and general (serving large paediatric patient populations) hospitals in the study regions. The population includes all patients in these services who meet the criteria described above under Population Characteristics.

This is an entirely record based study and hence is not subject to participation bias.

A few hospitals did not provide data for the entire study period thus possibly leading to underestimation of doses for some patients (see discussion).

The hospitals do not cover all of the hospitals in the study regions hence there is the potential for missing CT examinations. However the participating hospitals are expected to cover the vast majority of the CT scans in this population in the study region since the study was based on the specialised paediatric hospitals and the large general hospitals with large paediatric populations.

Ethics oversight

IARC Ethics Committee (IARC IEC 12–35) and the appropriate national, regional and hospital ethics committees in participating countries (the study included participation of 276 radiology departments from 9 countries, in addition to the various PI institutions and national cancer and population registries)

Note that full information on the approval of the study protocol must also be provided in the manuscript.

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Sample size

Statistical power was evaluated in a feasibility study in the EC funded CHILD-MED-RAD project, before the launch of the EPI-CT study. It was estimated based on the expected number of subjects that could be included in each country, the expected duration of follow-up and risk estimates of radiation induced cancer risk by age at exposure and time from the latest follow-up of the atomic bomb survivors study. Calculations were made using the US NIH "Power software" <http://dceg.cancer.gov/tools/design/power>) based on the following assumptions. Distributions of numbers of scans and body part scanned are based on data collected to date in the previously published UK study (Pearce et al 2012).

The specific assumptions for leukaemia analyses were as follows:

- 7% non-exposed as doses to red bone marrow from head scans appear similar (about 5 mGy) to those to red bone marrow in the spine from chest and abdominal CTs – hence only extremity scans would be assumed to be non-exposed
- 5 mGy to the red bone marrow per scan
- RRs of the order of 1.75, 2, and 2.5 associated with 10 scans (i.e. 50 mGy) in the first 10 years of follow-up
- an average incidence rate of 5 per 100 000 per year (based on Spanish figures in this age range)
- an average follow-up time of about 11 years
- and hence the probability of developing a disease would be 5.6 per 10 000 for each person
- about 2% of the paediatric population undergoing CT every year.

For a RR of 2.5 associated with 10 CT scans (ie 50 mGy) in the first 10 years of follow-up, we expected 80% power with a cohort of 500 000 patients. For a RR of 1.75 we needed a cohort of 1.2 Million patients to reach 80% power.

Data exclusions

We excluded from follow-up the first 2 years after the first CT examination to minimize reverse causation potential as well as the years when complete cancer registration was not available in the subject's country/region

We also excluded from analyses of haematological malignancies those coded as related to therapy or predisposing syndromes as they are unlikely to be related to CT exposure

Replication

This is a very large scale epidemiological study and no replication was logistically feasible.

However, to ensure the validity of the results, analyses were conducted by country/group of countries to ensure that no single country drove the results and conclusion of the study.

We also conducted numerous sensitivity analyses to address potential biases which could affect the interpretation of the results.

We compared our results to those of other similar studies and to those of other studies of low to moderate doses of ionising radiation in childhood and adolescence (see discussion)

Randomization

Not applicable as this was an observational study

Blinding

The investigators who extracted the information from the radiological records and who estimated the radiation doses were blinded to the cancer and mortality status of the study subjects.

Reporting for specific materials, systems and methods

We require information from authors about some types of materials, experimental systems and methods used in many studies. Here, indicate whether each material, system or method listed is relevant to your study. If you are not sure if a list item applies to your research, read the appropriate section before selecting a response.

Materials & experimental systems

n/a	Involved in the study
<input checked="" type="checkbox"/>	<input type="checkbox"/> Antibodies
<input checked="" type="checkbox"/>	<input type="checkbox"/> Eukaryotic cell lines
<input checked="" type="checkbox"/>	<input type="checkbox"/> Palaeontology and archaeology
<input checked="" type="checkbox"/>	<input type="checkbox"/> Animals and other organisms
<input checked="" type="checkbox"/>	<input type="checkbox"/> Clinical data
<input checked="" type="checkbox"/>	<input type="checkbox"/> Dual use research of concern
<input checked="" type="checkbox"/>	<input type="checkbox"/> Plants

Methods

n/a	Involved in the study
<input checked="" type="checkbox"/>	<input type="checkbox"/> ChIP-seq
<input checked="" type="checkbox"/>	<input type="checkbox"/> Flow cytometry
<input checked="" type="checkbox"/>	<input type="checkbox"/> MRI-based neuroimaging