

1 Reduced injury risk links sociality to survival in a group-living
2 primate

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20

Abstract

21 Affiliative social relationships and high social status predict longer lifespans in many
22 mammal species, including humans. Yet, the mechanisms by which these components
23 of sociality influence survival are still largely unknown. Using 10 years of data and
24 over 1000 recorded injuries from a free-ranging population of rhesus macaques (*Macaca*
25 *mulatta*), we investigated two possible mechanisms that could underpin the relationship
26 between sociality and survival: sociality (1) reduces injury risk; and/or (2) increases the
27 probability of survival after an injury. We found that sociality can affect an individual's
28 survival by influencing their risk of injury, but had no effect on the probability of injured
29 individuals dying. Individuals with more affiliative partners experienced fewer injuries
30 compared to less socially integrated. Social status was also associated with lower risk
31 of injury, particularly for older high-ranking individuals. These results represent the
32 first demonstration of a link between social integration and fatal injury risk in a group-
33 living species, and are the first to link social status, injury risk and survival outside of
34 humans. Collectively, our results offer insights into a mechanism that can mediate the
35 well-known benefits of sociality on an individual's fitness.

36 Uncovering the means by which sociality influences lifespan is of major interest to evo-
37 lutionary biologists, social scientists and biomedical researchers [1, 2, 3]. Evidence from
38 humans and other animals has provided increasing support for the benefits of affiliative
39 social interactions on survival. The strength of social bonds [4, 5, 6, 7], the number of weak
40 connections [8], the number of associates [9, 10, 11], the number of relatives in a group
41 [12] and the number of indirect connections [11, 13] predict the lifespan of individuals; the
42 general pattern being that those with more social partners are the ones that live longer.
43 Similarly, socioeconomic status in humans and social status in other animals are also robust
44 predictors of mortality risk [14, 15, 12, 5, 16, 17] with lower status individuals suffering a
45 greater risk of death. But precisely how the social environment affects survival is less well
46 understood.

47 One way for sociality to influence survival is by mitigating the costs of contest com-
48 petition over resources. Dominance hierarchies, for instance, are believed to have evolved
49 to reduce direct costs associated with competition for resources [18]. Nevertheless, social
50 hierarchies still usually entail disparities in access to resources, with individuals higher in
51 the hierarchy having priority access to food and mates at the expense of their subordinates
52 [19], who may still need to compete for access. Affiliative partners can also help to re-
53 duce engagement in agonistic encounters by providing access to resources via cooperation
54 and social tolerance [20]. For example, food sharing, cooperative feeding and co-feeding
55 have been described in several mammals, including some species of bats [21], cetaceans
56 [22, 23, 24], monkeys [25, 26] and apes [20]. Affiliative partners can also help to deter phys-
57 ical aggression from conspecifics by providing agonistic support. For instance, affiliative
58 interactions predict the formation of coalitions in male and female African wild dogs (*Ly-
59 caon pictus*) [27], Camargue horses (*Equus caballus*) [28], macaques (*Macaca spp.*) [29, 30]
60 and chimpanzees (*Pan troglodytes*) [31]. Agonistic support has also been widely documented
61 in female-philopatric primate species where related females defend one another [32, 33, 34].
62 If social status or affiliative relationships reduce the chance of aggressive interactions, these
63 components of sociality may directly enhance survival by allowing individuals to avoid costly
64 outcomes, such as injuries.

65 In addition to mitigating the immediate costs of aggressive behaviors, sociality may
66 also enhance survival through buffering mechanisms that influence an individual's health.
67 Differences in access to resources according to social status, for instance, may determine
68 the general body condition and health of individuals. Low social status has been related
69 to higher disease risk [15], higher levels of inflammation [35, 36], reduced healing capacity
70 [37] and overall impaired health in several mammal species, including humans [38, 39, 3].
71 Affiliative partners, on the other hand, can be valuable resources that can contribute to
72 better health by providing access to food [40, 41] and reducing the burden of infections via
73 hygienic behaviors (*i.e.*, grooming) [42, 43]. Better health status for high ranking or socially
74 integrated individuals may translate into higher chances of survival in the face of adversity,
75 for example, by improving the chances of healing following an injury.

76 Yet despite clear hypotheses for the potential mechanisms by which social status and af-
77 filiative relationships influence lifespan, there remains a lack of empirical evidence for these

78 mechanisms affecting survival. For example, several studies have shown associations be-
79 tween individual variation in sociality with markers of health and immunity [44, 45, 35, 36],
80 yet the consequences of such differences in the face of naturally occurring challenges to
81 health, and the downstream impact those differences might have on survival are unknown.
82 Similarly, studies supporting a relationship between sociality and lifespan usually do not
83 have the detailed physiological or health data required to test potential mechanisms con-
84 nnecting the two [1]. To fill this gap, we use a long-term data set containing both survival
85 data and detailed information on injuries in a free-living population of rhesus macaques to
86 test whether sociality mitigates the costs of competition (*i.e.*, injuries) and its consequence
87 on survival.

88 We explore two injury-related mechanisms that can link sociality with survival. Specif-
89 ically, we test whether social status and/or affiliative relationships: 1) influence the risk
90 of being injured, and/or 2) alter an individual's survival trajectory after an injury (Fig.
91 1). We did so using 10-years of injury data collected *ad-libitum* together with demographic
92 information from male and female rhesus macaques aged 4-29 years living on Cayo Santiago
93 island, Puerto Rico. Rhesus macaques live in multi-male multi-female despotic societies,
94 where access to resources is highly determined by an individual's position in the dominance
95 hierarchy [46]. Previous studies have shown the benefits of affiliative partners and social
96 status on the survival probability of monkeys in this population [12, 14, 8]. Predators are
97 absent from the island, ensuring injuries are mostly the result of physical aggression be-
98 tween conspecifics. Rhesus macaques are seasonal breeders with a mating season that can
99 last from 3 to 6 months. During these periods both affiliative and agonistic interactions are
100 usually heightened [47, 48] and, thus, important trade-offs between health, reproduction
101 and survival may occur [49, 50].

102 Because our study hinged on the assumption that being injured was detrimental for
103 survival in this population we first tested whether injuries inflicted by conspecifics increased
104 the probability of death in these animals (Fig 1; red arrow). To test if sociality influences the
105 risk of injury (mechanism 1), we asked whether social status and the number of affiliative
106 partners were associated with an individual's injury risk (Fig 1; yellow arrow). Given
107 the protective role of high social status and importance of affiliative partners in deterring
108 aggression [20, 32, 18], we predicted that high status individuals and those with more
109 affiliative partners would have a lower risk of injury. To test if sociality can alter the
110 impact of injuries on survival (mechanism 2), we asked if social status and the number
111 of affiliative partners affected the survival trajectories of injured individuals (Fig 1; green
112 arrow). As both social status and social integration can determine differences in health
113 status that may affect healing rates [39, 37, 51], we predicted that high status animals and
114 those with more affiliative partners would have a lower hazard of death from an injury than
115 low status individuals or those with fewer affiliative partners. Our results demonstrate that
116 sociality plays an important role in mediating the risk of injury, offering one of the few clear
117 mechanistic links between sociality and survival in a non-human mammal to date.

118 Results

119 Effect of injuries on survival

120 To quantify the extent to which injuries affect an individual's survival we used time-
121 dependent mixed effects cox models [52, 53]. Animals that were injured were nearly three
122 times more likely to die in the two months following the injury compared to animals that
123 were not injured (Fig. 2A; Hazard (Hz) = 1.06 ± 0.17 (SEM), $z = 6.58$, $p < 0.01$, injuries
124 ($i = 1041$, deaths ($d = 443$, N injured = 571, N uninjured = 1030), independent of their
125 sex or the reproductive season when the injury occurred. Individuals that were severely
126 injured (*e.g.* broken bones, exposed organs, multiple wounds or wounds in vital areas, see
127 SI Materials and Methods for details) experienced even a higher hazard of death that was
128 dependent on sex (Hz severity*sexM = 1.46 ± 0.72 , $z = 2.02$, $p = 0.04$, $i = 398$, $d = 107$,
129 N severely injured = 295). In males, severe injuries were associated with a higher chances
130 of dying compared to non-severe injuries, while in females, severe and non-severe injuries
131 had similar hazards of death (Fig. 2B).

132 Mechanism #1: Sociality affects the risk of injury

133 Effect of social status on injury risk

134 To test if high status animals were less likely to be injured or severely injured than low status
135 ones, we compared their injury risk separately for males and females using logistic models.
136 Given that observations of social interactions were only available for a subset of our subjects,
137 we used proxies of social status previously used in female (matrilineal rank) [14, 12] and
138 male rhesus macaques (group tenure length) [54, 55, 56] to maximize our statistical power.
139 We found that social status in females had a strong effect on the likelihood of being injured,
140 which was dependent on an individual's age (Odds rankLow*age = 0.3 ± 0.1 , $z = 3.02$, $p <$
141 0.01 , $i = 448$, $N = 827$). Low status females had a higher probability of being injured than
142 high status females, and this probability increased with a female's age (Fig. 3A). Social
143 status had no relationship with the risk of severe injuries in females (Odds = 0.13 ± 0.2 , z
144 = 0.65 , $p = 0.5$, $i = 135$, N severely injured = 114). In males, social status also had a strong
145 effect on the probability of being injured which was dependent on age (Odds status*age =
146 0.1 ± 0.03 , $z = 3.28$, $p < 0.01$, $i = 536$, $N = 748$). In younger males, lower social status
147 was associated with a higher incidence of injuries, while at older ages high status males
148 had higher probability of being injured (Fig. 3B). The same pattern was observed when
149 we focused our analysis on severe injuries (Fig. S2A, Odds status*age = 0.12 ± 0.04 , $z =$
150 2.67 , $p < 0.01$, $i = 245$, N severely injured = 168). Consistent with heightened male-male
151 competition over females [48] and with male harassment of females during the reproductive
152 season [57], we also found that injury-risk increased for both males and females during
153 the mating period compared to outside it, independent of their social status (injury: Odds
154 females = 0.85 ± 0.28 , $z = 3.02$, $p < 0.01$; Odds males = 1.2 ± 0.26 , $z = 4.6$, $p < 0.01$;
155 severe injury : Odds females = 1.04 ± 0.26 , $z = 4$, $p < 0.01$; Odds males = 1.38 ± 0.25 , z
156 = 5.4 , $p < 0.01$).

157 **Effect of affiliative partners on injury risk**

158 To test whether animals with more affiliative partners were less likely to be injured or
159 severely injured than those with fewer affiliative partners we used logistic models. To
160 support robust statistical analyses, we relied on a proxy (*i.e.*, number of female relatives
161 in the group) that has been previously shown to influence survival in this population [12].
162 Female rhesus macaques have a strong bias toward forming partnerships with their maternal
163 kin [58] and this proxy has been positively correlated with network measures of social
164 integration [59]. Males, on the other hand, are the dispersing sex and have few kin in their
165 new groups, and so were excluded from this analysis. We found that the number of close
166 relatives (relatedness coefficient (r) = 0.5, *i.e.*, mother-daughters and full siblings) in a
167 female's group had a weak, but not significant, effect on her probability of being injured
168 (Odds = -0.1 ± 0.05 , $z = -1.84$, $p = 0.06$, $i = 491$, $N = 851$). However, the size of a female's
169 extended family ($r \geq 0.125$, *i.e.*, spanning three generations) was strongly associated with
170 the likelihood of injury, with females experiencing a 13% reduction in the incidence of
171 injuries for every one standard-deviation increase in their number of female relatives (Fig.
172 4A; Odds = -0.14 ± 0.06 , $z = -2.5$, $p = 0.01$, $i = 491$, $N = 851$). The incidence of severe
173 injuries was not affected by the number of close relatives (Odds = -0.06 ± 0.09 , $z = -0.6$,
174 $p = 0.53$, $i = 147$, N severely injured = 123) nor by the size of a female's extended family
175 (Odds = -0.13 ± 0.09 , $z = -1.36$, $p = 0.18$, $i = 147$, N severely injured = 123).

176 **Mechanism #2: Sociality influences the survival of injured animals**

177 **Effect of social status on survival of injured animals**

178 To assess whether social status or affiliative relationship buffer the detrimental effect of in-
179 juries on survival, we used time-dependent mixed effects cox models. We found no evidence
180 of a buffering effect of social status on the survival of injured females (Hz injured*rankLow
181 = -0.45 ± 0.46 , $z = -0.98$, $p = 0.33$, $i = 448$, $d = 103$, $N = 278$) or injured males (Hz in-
182 jured*tenure = 0.00009 ± 0.0002 , $z = 0.47$, $p = 0.64$, $i = 536$, $d = 97$, $N = 272$). Similarly,
183 no buffering effect of social status on survival was observed in severely injured females (Hz
184 injured*rankLow = -0.51 ± 0.92 , $z = -0.55$, $p = 0.58$, $i = 135$, $d = 42$, N severely injured
185 = 114) or males (Hz injured*status = -0.0001 ± 0.0002 , $z = -0.67$, $p = 0.5$, $i = 245$, $d =$
186 57, N severely injured = 168).

187 **Effect of affiliative partners on survival of injured animals**

188 We found no evidence for a relationship between survival after an injury and the number
189 of close relatives a female had available at the time (Hz injured*nkin = -0.22 ± 0.28 , $z =$
190 -0.78 , $p = 0.43$, $i = 491$, $d = 114$, $N = 294$) or current size of her extended family (Hz
191 injured*nkin = 0.03 ± 0.04 , $z = 0.59$, $p = 0.56$, $i = 491$, $d = 114$, $N = 294$). Similarly, the
192 number of affiliative partners did not influence the survival of severely injured females (Hz
193 close kin = -0.82 ± 0.68 , $z = -1.21$, $p = 0.23$; Hz extended family = 0.008 ± 0.09 , $z = -0.1$,
194 $p = 0.92$; $i = 147$, $d = 45$, N severely injured = 123).

195 **Post hoc mediation analysis**

196 Mediation analyses can be used to test the significance of a mediator in the relationship
197 between an independent and a dependent variable and to measure the effect size of that
198 relationship [60]. Although useful, current mediation analysis approaches are unable to
199 estimate effect sizes for data structured in a logistic manner, such as ours, nor are they
200 able to cope with interaction terms in logistic regressions [61]. Given these limitations, we
201 could not use mediation analysis to evaluate if the effect of social status on survival was
202 mediated by injury risk because these results relied on an interaction with individual age
203 (Fig. 3), nor could we use it to estimate the effect size of any of our results. We did,
204 however, use mediation analysis to assess if injury risk significantly mediated some of the
205 effect of affiliative partners on survival. Our mediation analysis confirmed a direct effect
206 of affiliative partners on survival by showing that the size of a female's extended family
207 significantly reduced her hazard of death (direct effect = -0.065 ± 0.02 , $z = -3.19$, $p <$
208 0.01). It also confirmed that this relationship was significantly mediated by the risk of
209 being injured (indirect effect z -score = -2.31 , $p < 0.05$).

210 **Discussion**

211 Taken together, our results suggest that different components of the social environment
212 can modulate the risk of suffering an injury and, therefore, the hazard of death. We found
213 that high social status was associated with a lower injury risk for specific periods of males'
214 and females' lives, and that a female's number of affiliative partners may help to prevent
215 injuries. In contrast to previous research showing that individuals with higher social status
216 had faster healing rates [37], we found that none of the measures of sociality analyzed
217 affected the survival trajectories of injured animals.

218 To our knowledge, this is the first field study to quantify the consequences of injuries
219 on the probability of death in a nonhuman primate. Other studies in wild populations
220 of baboons (*Papio sp.*) and Afro-eurasian monkeys have established the social and demo-
221 graphic predictors of injury risk [62, 37, 63, 64, 65, 66], yet its consequences on survival
222 have yet to be shown. We found sex-differences on the influence of severe injuries in survival
223 that can reflect trade-offs between the energy allocated for reproduction versus immunity
224 [4, 37]. For instance, during the reproductive season the probability of being severely injured
225 was substantially higher for both sexes. During this period, males may be particularly
226 immunocompromised given the high amount of energy and resources required to sustain
227 the effort associated with mating [49, 67], which can impair injury recovery. On the other
228 hand, females usually have higher demand on their immune systems during lactation [4, 68],
229 *i.e.*, outside the reproductive season. Therefore, females may cope better than males with
230 severe injuries during the reproductive season at the expense of being more susceptible to
231 the consequences of injury outside this period.

232 We found support for one of the hypothesized mechanisms linking sociality to survival,
233 whereby sociality reduces an individual's risk of injury. High social status animals were

234 injured less than those of lower status during specific periods of their life, and females
235 with more affiliative partners (i.e., kin) were less likely to be injured than less integrated
236 females and, thus, experienced lower hazard of death. Our results linking social status to
237 reduced risk of injury are consistent with the skewed access to resources in systems with
238 clear linear dominance hierarchies [19]. High status individuals may not need to engage
239 in costly aggression for food or mates, in contrast to low status animals who must gain
240 access through contests. Although we could not test for a mediation effect of injury risk
241 on the relationship between social status and survival, our results suggest that low status
242 individuals experience greater hazard of death as a result of enhanced risk of injury. Our
243 finding that social status did not influence the risk of injury in young females may be
244 because at younger ages females' relative positions in the dominance hierarchy have yet to
245 be fully established [69]. Further, we showed that older high status males were more likely
246 to be injured than older low status males. This finding may reflect heightened aggressive
247 challenges from lower status animals to those higher in the hierarchy as a consequence of
248 a decline in the body condition with age [70] and, thus, the capacity of older high status
249 males to maintain their dominance.

250 Previous studies in matrilineally-structured primate species, in which most of the affil-
251 iative relationships are with female relatives, have shown that females commonly engage in
252 agonistic encounters to support and protect their kin [32, 33, 34], even when confronting
253 higher status individuals [71]. In line with these studies, our results suggest that having
254 more relatives available may provide a numerical advantage to deter physical aggression.
255 Other mechanisms, such as social tolerance when accessing resources [20] could also explain
256 fewer injuries in the presence of more affiliative partners. Interestingly, only the size of a
257 female's extended family, not her number of close relatives, had a significant relationship
258 with risk of injury. This suggests that the number of close relatives in a group (range in our
259 study: 0-5) may not be enough to provide robust agonistic support or access to resources,
260 compared to the size of a female's extended family (range in our study: 0 - 38).

261 We found no support for the second hypothesized mechanism that we explored to link
262 sociality to survival; none of the measures of sociality analyzed influenced an individual's
263 survival trajectory following injury. Despite a vast body of literature supporting differ-
264 ences in health and immunity between individuals of different social status [44, 35, 39],
265 we found no evidence for an effect of social status on the survival trajectories of injured
266 animals. These findings contrast with a previous study on wild baboons where high status
267 males had faster healing rates than lower status males [37]. Although we did not quan-
268 tify differences in healing times, our results suggest that the probability of recovering from
269 an injury was not influenced by an animal's position in the dominance hierarchy. These
270 differences might be explained in part by differences in features of the two study systems.
271 Animals on Cayo Santiago are provisioned with food on a daily basis and access to the nu-
272 trients needed to support immune function might not be as skewed as they are in the wild
273 [45]. Notwithstanding, in both systems high social status has been associated with elevated
274 levels of glucocorticoids and androgens [72, 73, 50], well known immune-suppressors, which
275 suggest that in the Cayo Santiago population, unlike the baboons, the benefits of being of

276 high status may not outweigh the costs in terms of injury recovery.

277 We also found, contrary to our predictions, that the benefits associated with affiliative
278 partners, such as feeding tolerance [74, 75] and social hygienic behaviours [42, 43], seem
279 not to have helped females to cope with the detrimental effect of injuries on their survival.
280 It is possible that social hygienic behavior, such as removal of ectoparasites by grooming,
281 have long-term health benefits but do nothing to enhance the short-term immune response
282 required to heal damaged tissue [76]. Additionally, grooming wounded areas may, in fact, be
283 detrimental for the healing process as it could lead to the removal of protective scabs [43].
284 This could be one reason why females with more affiliative partners, who are presumed
285 to receive more grooming and to have more access to food via social tolerance, did not
286 have improved survival trajectories after an injury. Previous research on this population
287 has shown that the number of close relatives and the size of a female's extended family
288 are associated with increased survival probability [12, 8]. The results of the current study
289 suggest this relationship does not come about because of the reduced risk of death from
290 injury. Further research is needed to elucidate to what extent other mechanisms involving
291 health differences (e.g., disease susceptibility) play a role in the benefits of social partners
292 in the survival of females in this population. Additionally, direct behavioral observations
293 in a large sample of individuals with paired injury data will be required to explore refined
294 ego-networks characteristics and to expand these results to affiliative relationships of males
295 and unrelated females.

296 In sum, our study provides evidence for a mechanism linking sociality to lifespan. Growing
297 literature has supported a strong relationship between the social environment and sur-
298 vival in many mammal species [3], but the ultimate function of some components of sociality,
299 such as social relationships, remain unclear [77]. Although sociality has been demonstrated
300 to enhance health and immunity [44, 35, 45], here we showed that these benefits did not
301 translate to an improved ability to cope with the risk of death from injuries. Instead, we
302 found that sociality plays an important role in preventing individuals from suffering injuries
303 that would likely lead to death. Given how rare injuries are in this population, we do not
304 expect that this is the only mechanism linking sociality to survival. Other mechanisms may
305 include sociality-mediated differences in components of health related to disease suscepti-
306 bility. In wild animal populations, social partners may also help with predator detection
307 [78], predator mobbing [79], finding food sources [80], thermoregulation [81], among other
308 possibilities. Nevertheless, here we provide rare empirical evidence for an ultimate function
309 of social relationships, showing one mechanism by which high status and socially integrated
310 individuals live longer. Demonstrating the relative importance of different mechanisms
311 linking sociality and survival will be challenging but a crucial goal of future research. Our
312 study provides insight into the essential role that long-term datasets that combine both
313 demographic and health data will play in meeting this challenge.

314 Materials and Methods

315 Subjects

316 We studied a population of free-ranging rhesus macaques on the island of Cayo Santiago
317 in Puerto Rico. The island is home to a population of \sim 1800 individuals living in 6-10
318 mixed sex naturally formed social groups. The field station is managed by the Caribbean
319 Primate Research Center (CPRC), who monitor the population daily, and maintain the
320 long-term (>75 years) demographic database including data on births, deaths, social group
321 membership for all animals and a genetic parentage database for animals born after 1992
322 [82]. Animals have *ad-libitum* access to food and water, the island is predator-free and
323 there is no regular medical intervention for sick or wounded individuals. We focused on all
324 subadult and adult females and males between 4 and 29 years of age that were alive between
325 the years 2010 and 2020, a period for which records on injuries exist (see below for details
326 on how injury data was collected). In this study we included data on 571 injured individuals
327 (294 females, 277 males) and 1030 uninjured individuals (557 females, 473 males). From
328 these animals, 342 (85 injured, 258 uninjured) were removed from the population by the
329 CPRC for population control purposes [83]. For all individuals, birth dates were known
330 within a few days. Removal dates were known for all removed individuals. Dispersal from
331 the island almost never occurs, therefore death dates were also known within a precision of
332 a few days.

333 Observation of injuries

334 From 2010 to 2020 CPRC staff collected *ad-libitum* observations on the incidence and re-
335 covery of injuries, during the daily monitoring of social groups for demographic purposes.
336 Monkeys were individually recognized based on their identity tattoos located on their chest
337 and leg. Whenever a staff member noticed a wounded animal or an animal displaying signs
338 of injury (*e.g.* bleeding, limping), they recorded the animal ID, type of injury and addi-
339 tional details on the general state of the animal (*e.g.* by evidence of weight loss or poor
340 physical condition). If there was a visible wound, observers additionally recorded the area
341 of the body affected, if it was a recent or old wound based on the presence of scars, and
342 whenever possible, an estimate of its size. Observers updated the records every time they
343 encountered the injured animal during their daily census routine with an average update
344 time for an injured individual across the 10 years of 42.17 days. In total, 1137 injury events
345 were observed with an average of 107.6 ± 63.5 per year. Here, we included all the records
346 of injuries that were considered non-ambiguous (*i.e.*, those with visible damage to the skin)
347 including bites, scratches, cuts and abrasions along with other clearly observable injuries
348 such as fractures and exposed organs. Our final sample consisted of 1041 injuries collected
349 from September 2010 to April 2020. We classified these injuries based on their degree of
350 severity, where severe injuries were those involving broken bones, exposed organs, multiple
351 wounds and any wound in vital areas, including head, neck, abdomen or genitalia ($n =$
352 398). All other injuries were considered non-severe ($n = 643$).

353 **Measures of sociality**

354 We used proxies of social status (dominance rank) in our analyses. Observations of agonistic
355 interactions between pairs of animals- from which dominance rank is often computed- were
356 only available for a subset of subjects (194 unique individuals injured in 292 injury events,
357 485 uninjured individuals). To maximize statistical power, we decided to use the complete
358 dataset and to use known proxies of social status instead; group tenure in males [54, 55, 56]
359 and matrilineal rank in females [14, 12]. Male rhesus macaques reach dominance through
360 queuing [84]; those that have been in a group for longer are usually high-ranking [54]. We
361 determined tenure length using information on monthly social group membership. Group
362 tenure length was computed as the time (in days) a male has been observed in his current
363 group at the date of interest (current date minus date of dispersal). If a male had not yet
364 dispersed and remained in his natal group, we computed group tenure since their birth date.
365 If a male died or was removed from the population before the end of the period of interest,
366 we computed group tenure up to that point. We established tenure length for all the males
367 in our dataset ($n = 750$, n injuries = 550). However, 67 of those males had periods where
368 they were observed living outside a social group (*i.e.*, they were “extra-group”). These
369 specific periods when group tenure could not be computed were dealt differently depending
370 on the analysis in question and we discuss this on a case-by-case basis below.

371 Female rhesus macaques are philopatric and form maternally inherited stable linear
372 dominance hierarchies whereby daughters occupy a rank just below their mothers [85].
373 Members of a same matriline tend to be adjacent to one another in the hierarchy, thus
374 the rank of an entire matriline can be used as a proxy for individual rank in social groups
375 containing more than one matriline [14]. We determined matrilineal rank using known social
376 status based on pairwise agonistic interactions from females in our dataset. We identified
377 only one matriline per group as ‘high-ranking’- the one containing the alpha female - while
378 all the others in the group were classed as ‘low-ranking’. Females in groups with a single
379 matriline were disregarded as rank is a relative measure and females from groups with a
380 single matriline are all of the same rank. We established matrilineal rank for 827 females
381 (407 high ranking, 420 low ranking, n injuries = 448).

382 To confirm that group tenure and matrilineal rank were appropriate proxies for social
383 status we looked at the correlation between dominance rank computed from animals with
384 known social status based on agonistic interactions and our proxies. The correlation be-
385 tween group tenure and dominance rank- measured as the percentage of same-sex animals
386 outranked in the group [86] - was moderate and significant (Fig. S1A; *Pearson’s r* = 0.62,
387 $p < 0.01$). Matrilineal rank and categorical dominance rank were strongly correlated (high-
388 ranking: $\geq 80\%$ outranked, low ranking: $\leq 79\%$ outranked [87]) based on Cramer’s V
389 coefficient (Fig. S1B; *Cramer’s V* = 0.39, chi-square = 159.42, $p < 0.01$), which measures
390 the association between two categorical variables [88].

391 As above, we only had data on affiliative interactions for a subset of our subjects.
392 Therefore, to maximize our sample size we followed a previous study [12] and used the
393 number of female relatives (4 years and older) that were present in a female’s social group
394 as a proxy for social capital. Female rhesus macaques preferentially interact with their

395 female kin compared to non-kin individuals [58], thus those with greater number of relatives
396 are expected to have more opportunities for social support. We limited this approach to
397 females as males, being the dispersing sex, often have very few close kin in their new groups,
398 and might not be able to recognise unfamiliar kin [89]. Using the Cayo genetic pedigree
399 database we computed the number of close kin ($r = 0.5$) and extended family ($r \geq 0.125$)
400 for all injured and uninjured females in our dataset ($n = 851$, n injuries = 491). We
401 decided to test these two levels of relatedness as the first represents the strongest kin-bias
402 (*i.e.*, mother-daughter or full sisters) and the second the lowest threshold for kin bias in
403 affiliative interactions for rhesus macaques [90].

404 **Statistical approach**

405 For all of the statistical analyses we defined a two-month time window (hereafter, bimonthly
406 interval) as the period from which the injury status could transition from injured to not
407 injured based on the average update time for an injured animal (*i.e.*, average time between
408 two consecutive records) and the computed average healing time. Thus, all variables were
409 evaluated on a bimonthly basis (*i.e.*, each row in the dataset represents a two-month inter-
410 val). For each of the questions we ran two models, one that included injury status based on
411 all injuries (model 1) and other that included injury status for severe injuries only (model
412 2).

413 **Effect of injuries on survival**

414 To establish the effect of injuries on survival we used time-dependent Cox proportional
415 hazard (PH) models [52]. For the analyses we used the whole dataset ($n = 1061$), including
416 injured and uninjured animals from both sexes. Animals that were removed from the
417 population or that were still alive at the end of the study period were censored. The
418 predictor of interest was the injury status (*i.e.*, all injuries or severe injuries) along with
419 other relevant variables that may influence survival probability, such as reproductive season
420 (*i.e.*, mating vs no-mating) and sex. Age was accounted for implicitly in the models.
421 Additionally, we included random effects for the specific bimonthly interval within the
422 study period to control for potential mortality sources at the population level and individual
423 identity to account for repeated measures. To determine the bimonthly interval we divided
424 the whole study period (10 years) in intervals of two months- ranging from 1 to 58 - where
425 1 represents the first two months since September 2010. We tested for interaction effects
426 among our predictors and only retained them if statistically significant to avoid issues of
427 overfitting.

428 **Mechanism #1: Sociality affects the risk of injury**

429 To assess the effect of social status and the number of affiliative partners on the risk of
430 injuries, we used generalized linear mixed models with binomial distribution (logit models).
431 In all the models we asked whether our measures of sociality influenced the probability of

432 being injured in a given bimonthly interval. To test if high status animals were less likely
433 to be injured compared to low status ones, we ran the analyses separately for each sex (n
434 females = 827, n males = 750). For males, social status was estimated from group tenure
435 computed up to the end of each bimonthly interval. Bimonthly intervals where males were
436 extra-group and so group tenure could not be computed, were excluded. For females, we
437 used matrilineal rank, which remains constant across the lifespan and, thus, remained the
438 same in every interval. To test if animals with more affiliative partners were less likely
439 to be injured compared to animals with social partners we used only females (n = 851),
440 fitting separate models for the two thresholds of relatedness (close kin and extended family).
441 The number of relatives present in a group was computed for each bimonthly interval. We
442 modelled injury status as a function of social status or number of affiliative partners, while
443 controlling for age and reproductive season. As group tenure and age could be correlated,
444 we checked for collinearity between these predictors using the variance inflation factor (vif),
445 but no correlation was found (vif = 1.01). Random effects were included for individual ID -
446 to account for repeated measures - and for the specific bimonthly interval within the study
447 period. We z-scored continuous variables to help convergence and tested interaction terms
448 among all our predictors, which were retained if significant.

449 **Mechanism #2: Sociality influences the survival of injured animals**

450 To examine the effect of sociality (social status and number of affiliative partners on the
451 survival of injured animals we used time-dependent cox ph models. As before, we tested for
452 an effect of social status on survival in separate models for males and females and examined
453 only females to test the effect of affiliative partners on survival post-injury. In all the models
454 the predictor of interest was specified by an interaction term between injury status and the
455 sociality measure. Variables were evaluated on a bimonthly basis with a time-dependent
456 covariate for reproductive season. Random effects were included for individual ID and
457 bimonthly interval. We additionally included a time-dependent fixed effect for group size to
458 control for its potential effect on the number of kin available and on survival [2]. As some
459 bimonthly intervals had missing information for group tenure, we ran two models for males;
460 a complete case analysis and a model using mean-matching multiple imputation with 20
461 iterations to fill the missing data [91, 92], yet the estimates were identical between both
462 procedures. Given that the main predictor was an interaction term, we did not attempt to
463 fit other interactions.

464 **Post hoc mediation analysis**

465 To further confirm our findings that sociality significantly influences survival by reducing
466 risk of injury we ran a mediation analysis. Given limitations to use mediation analyses
467 with different type of models (logistic and cox), we translate our cox model to predict
468 survival into a logistic regression, where the outcome represents if the animal was still alive
469 (0) or death (1) as a function of injury status on each bimonthly interval. Unlike cox
470 models, logistic regressions can not handle individuals for which the outcome is unknown

471 (i.e., censored), therefore for those individuals the last bimonthly interval in the study was
472 not considered. Different methods for testing mediation using logistic models have been
473 proposed. However, to date there are still no robust methods to quantify the effect size
474 or to consider interaction terms [61]. Given this limitation, we were only able to test the
475 significance of the mediation effect of injury risk on the relationship between the number
476 of affiliative partners ($r \geq 0.125$) and survival. We ran first a model where the number of
477 affiliative partners and covariates predicts the injury risk (injuries ~ sociality + covariates).
478 From this model, we extracted the estimate and standard error for affiliative partners.
479 Then, we ran a second model where both sociality and injury risk predict survival (survival
480 ~ sociality + injuries + covariates), and extract the estimate and standard error for injury
481 risk. Finally, we computed the standardized element (z-score) following Iacobucci [93]. We
482 determined significance by contrasting the z-score against a standard normal distribution,
483 thus an absolute value greater than 1.96 represents a statistically significant mediation
484 effect.

485 **Ethics**

486 This research complied with protocols approved by the Institutional Animal Care and Use
487 Committee (IACUC) of the University of Puerto Rico (protocol no. A6850108) and by the
488 University of Exeter School of Psychology's Ethics Committee. The Caribbean Primate
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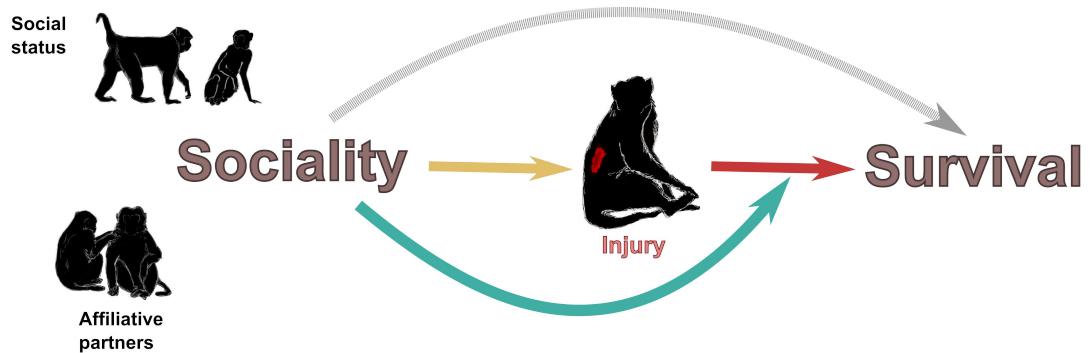


Figure 1: Injury-related mechanisms by which components of sociality (social status, affiliative partners) can influence survival. A direct effect of sociality on survival (gray arrow) has been well established in mammals [4, 5, 7, 10, 11], including studies in the Cayo Santiago population [12, 8]. We explore mechanisms related to injury by which the relationship between sociality and survival might come about. According to the first mechanism, sociality influences the risk of injury (yellow arrow) and, therefore, survival (red arrow). According to the second mechanism (green arrow), sociality affects the survival trajectories of injured individuals.

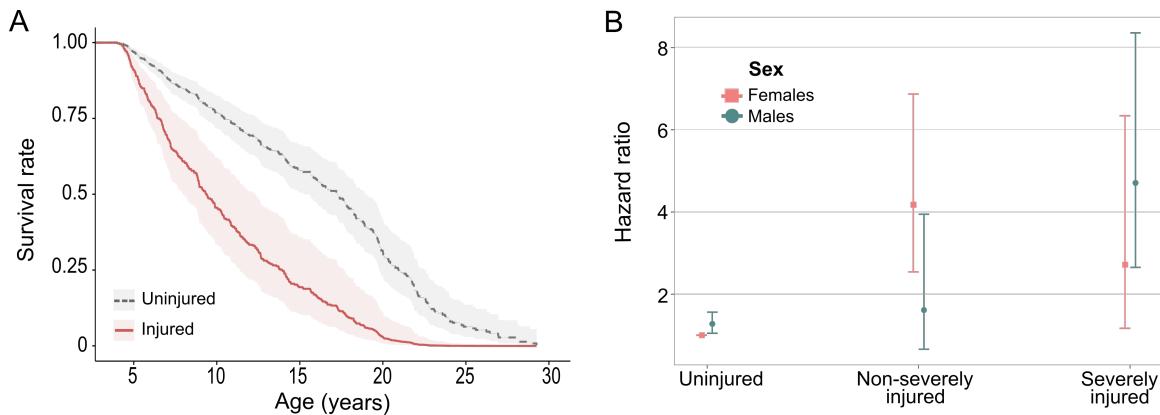


Figure 2: Effect of injuries on survival. **A)** Survival curves adjusted for covariates for injured and uninjured individuals. Injured individuals (red solid line, $n = 571$, 294 females, 277 males) had near a 3-fold increase in the probability of dying compared to uninjured animals (gray dashed line, $n = 1030$, 557 females, 473 males) ($\text{Hz} = 1.06 \pm 0.17$, $z = 6.58$, $p < 0.01$, injuries (i) = 1041, deaths (d) = 443). Curves represent males during the mating season, but those for females were similar. Shaded areas represent standard errors. **B)** Hazard ratios of death for females and males as a function of the severity of injuries. Severe injuries increased the hazard of death relative to non-severe injuries in males (green circles, n uninjured = 473, n non-severely injured = 189, n severely injured = 251), but not in females (Pink squares, n uninjured = 557, n non-severely injured = 232, n severely injured = 147) ($\text{Hz severity*sexM} = 1.46 \pm 0.72$, $z = 2.02$, $p = 0.04$, $i = 398$, $d = 107$).

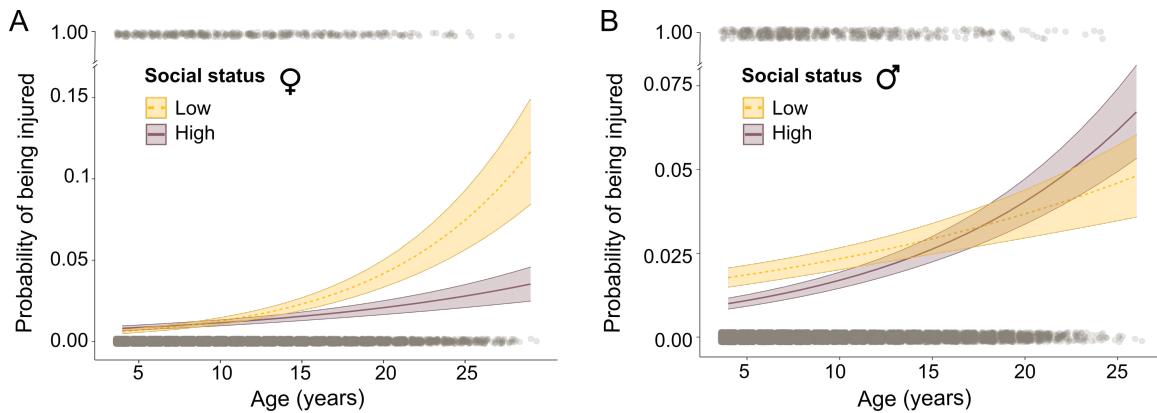


Figure 3: Predicted injury risk in relation to social status. **A)** Injury risk for females as a function of social status and age. Low status females (yellow dashed line, $n = 420$, 237 injuries) had higher chances of being injured than high status females (purple solid line, $n = 407$, 211 injuries), with increasing probabilities for older females (Odds rankLow*age = 0.3 ± 0.1 , $z = 3.02$, $p < 0.01$). **B)** Injury risk for males as a function of social status and age. For visualization, social status was categorized by selecting the 20th (273 days of tenure) and 80th (2029 days of tenure) percentiles depicting low status (yellow dashed line) and high status (purple solid line), respectively ($n = 748$, 536 injuries). Younger males from low status had higher injury risk than high status young males, yet the opposite occurred at later ages (Odds tenure*age = 0.1 ± 0.03 , $z = 3.28$, $p < 0.01$). In both plots, shaded areas represent standard errors and gray dots the raw data used in the models (top: injured, bottom: uninjured).

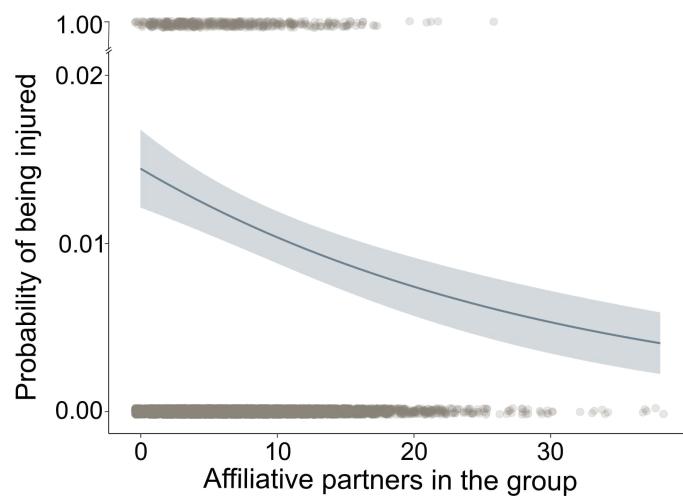


Figure 4: Predicted injury risk as a function of the number affiliative partners. X-axis represents the number of female relatives (extended family, $r \geq 0.125$) present in a female's group ($n = 851$, injuries (i) = 491). Females with more relatives had lower chances of suffering from an injury compared to females with fewer relatives (Odds = -0.14 ± 0.06 , $z = -2.5$, $p = 0.01$, $i = 491$). Shaded areas represent standard errors and gray dots the raw data used in the models (top: injured, bottom: uninjured).